



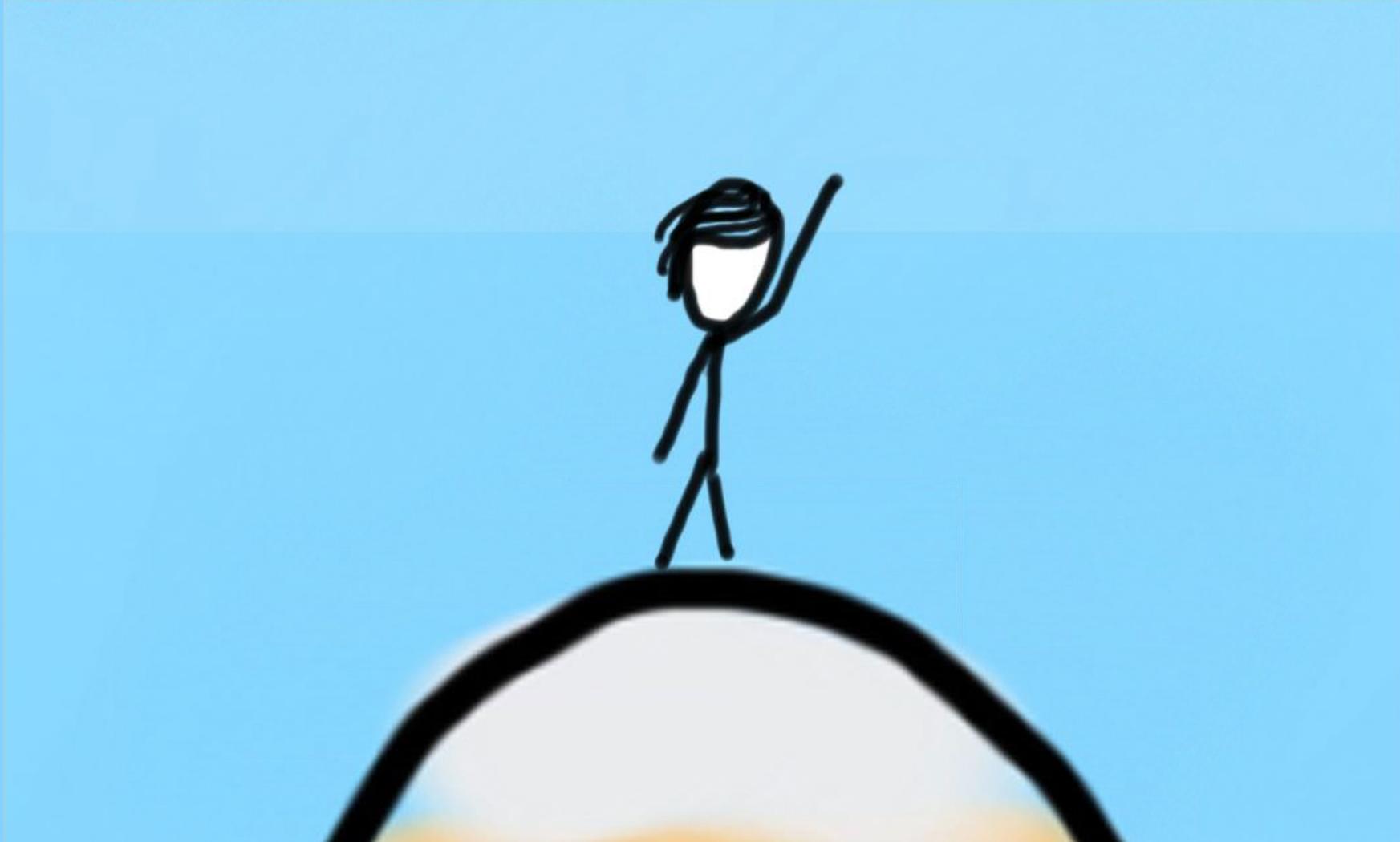
# Tender Moments

Chris Piech

CS109, Stanford University

# Learning Goals

1. Use New Random Variables!
2. Calculate Expectation of a Random Variable



# Quick Announcements

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Python review session 2 – after class today !! (will be recorded)

Pset 2 goes out today !! Due Jan 23<sup>rd</sup> (next Friday)

With MLK holiday – we made Pset 2 a bit shorter than normal so that you don't have to work over the holiday if you don't want to.

That said – get started early 😊

No lecture on Monday



# Review



A **random variable** is a number which takes on values probabilistically.



A discrete random variable is fully described by a **probability mass function**.

Let  $Y$  be a random variable



$Y$

For example  $Y$  is the number of heads in 5 coin flips

Let  $Y$  be a random variable



$$Y = 2$$

\*note: here equals means `==` in coding

It is an event when  
 $Y$  takes on a value

For example  $Y$  is the number of heads in 5 coin flips

Let  $Y$  be a random variable

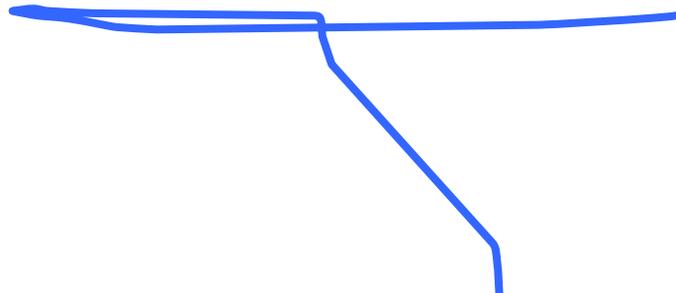


$$Y < 3$$

It is an event when  
you ask any comparison question

For example  $Y$  is the number of heads in 5 coin flips

If this is a number


$$P(Y = 2)$$


Then this is a probability  
(between 0 and 1)

For example  $Y$  is the number of heads in 5 coin flips

If this is a variable

$$P(Y = k)$$

Then this is a function

For example  $Y$  is the number of heads in 5 coin flips

This is a function

$$P(Y = k)$$

The diagram illustrates a function. A blue arrow points from the expression  $k = 5$  to the variable  $k$  in the probability expression  $P(Y = k)$ . A second blue arrow points from the expression  $P(Y = k)$  to the numerical value  $0.03125$ .

$$k = 5 \qquad 0.03125$$

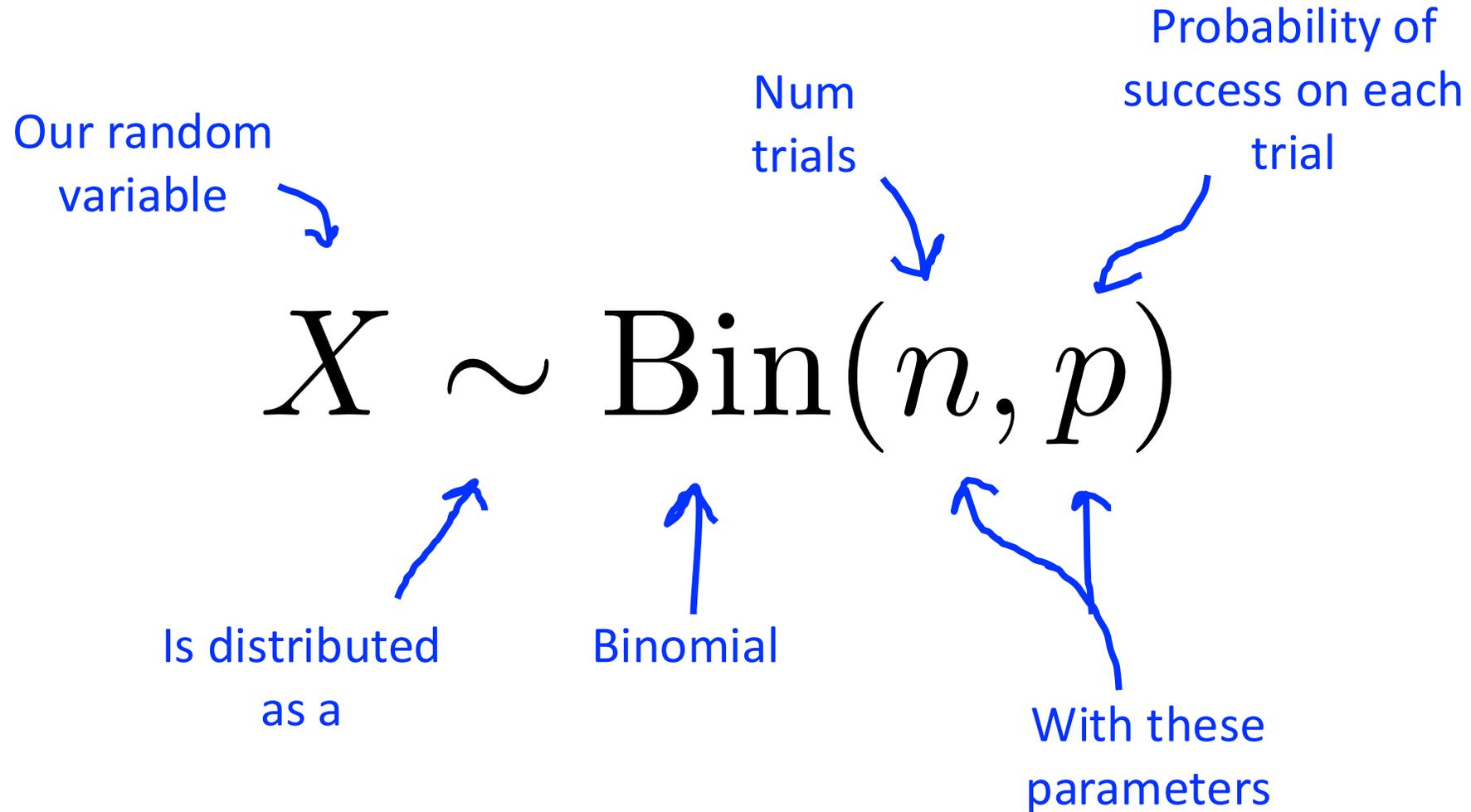
For example  $Y$  is the number of heads in 5 coin flips

Random Variables are a big deal, because they allow other people to give you a PMF (and other helpful equations)

# Classics



# Declare a Random Variable to be Binomial



Exactly  $k$  heads in  $n$  coin flips. Probability of exactly  $k$  heads:

(H, H, H, H, T, T, T, T, T, T)  
(H, H, H, T, H, T, T, T, T, T)  
(H, H, H, T, T, H, T, T, T, T)  
(H, H, H, T, T, T, H, T, T, T)  
(H, H, H, T, T, T, T, H, T, T)  
(H, H, H, T, T, T, T, T, H, T)  
(H, H, H, T, T, T, T, T, T, H)  
(H, H, T, H, H, T, T, T, T, T)  
(H, H, T, H, T, H, T, T, T, T)  
(H, H, T, H, T, T, H, T, T, T)  
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(H, H, T, T, H, T, T, H, T, T)  
(H, H, T, T, H, T, T, T, H, T)  
(H, H, T, T, H, T, T, T, T, H)

$$\binom{n}{k} p^k (1 - p)^{n-k}$$



# Automatically Know the PMF

Probability Mass Function for a  
Binomial

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

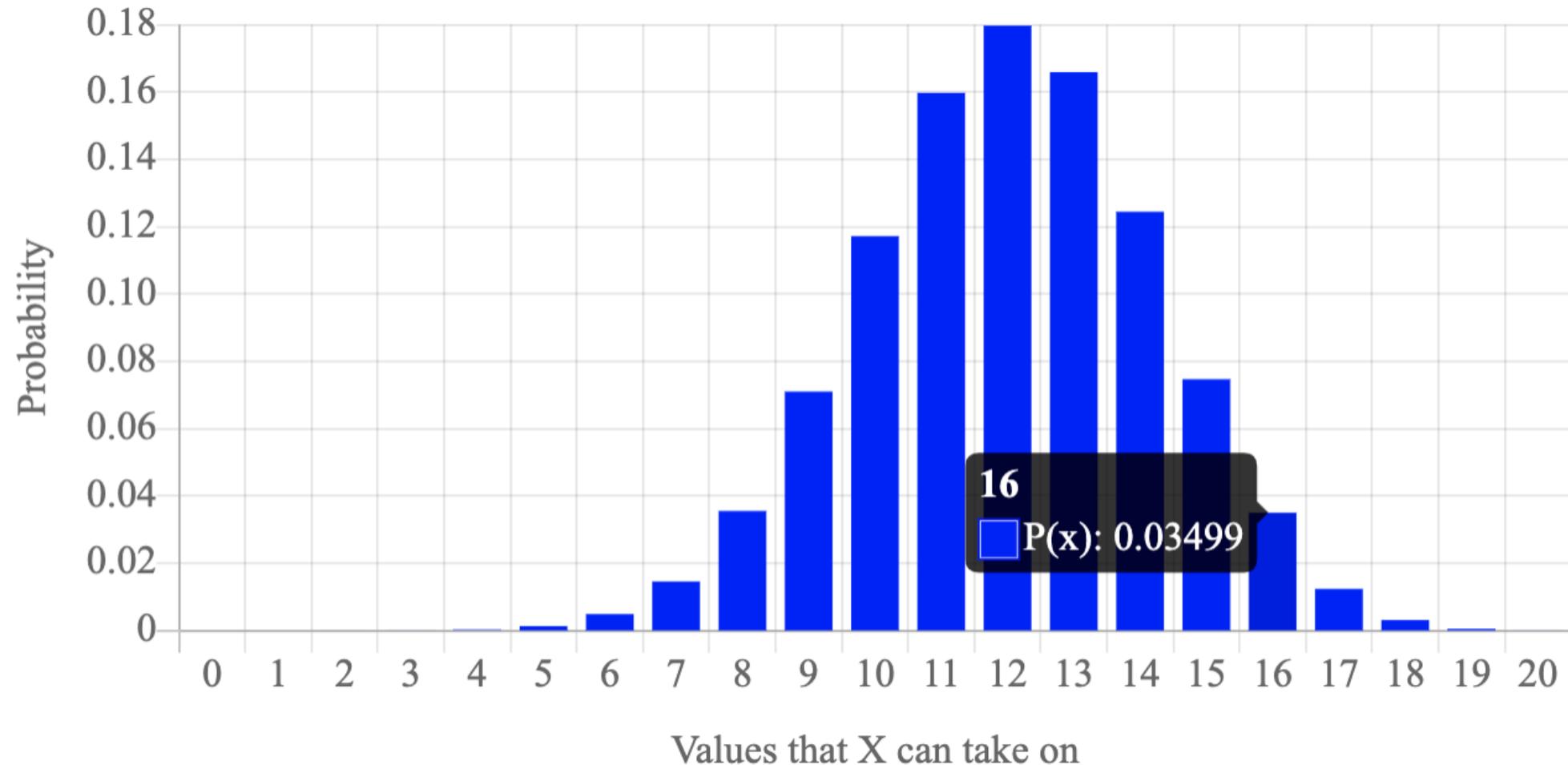
↑  
Probability that our  
variable takes on the  
value  $k$

↑  
\* This is also called the  
binomial term

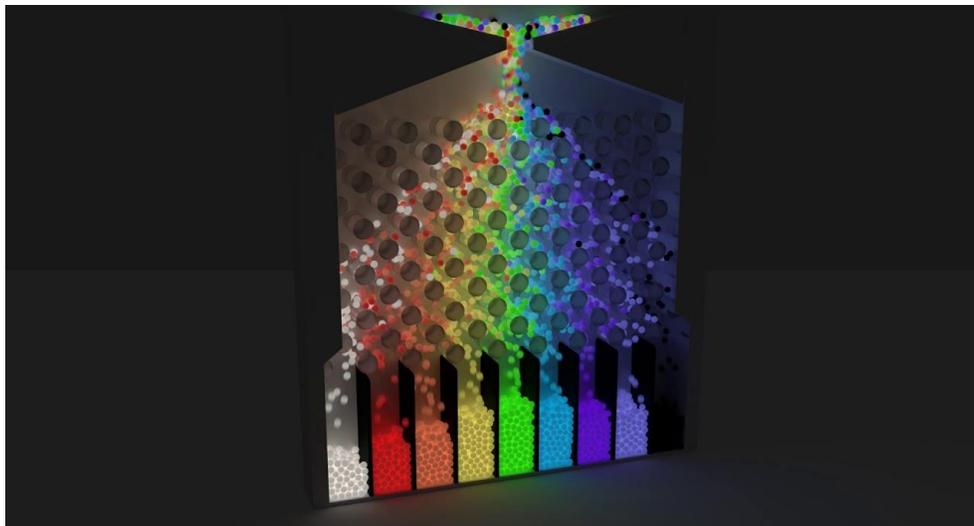
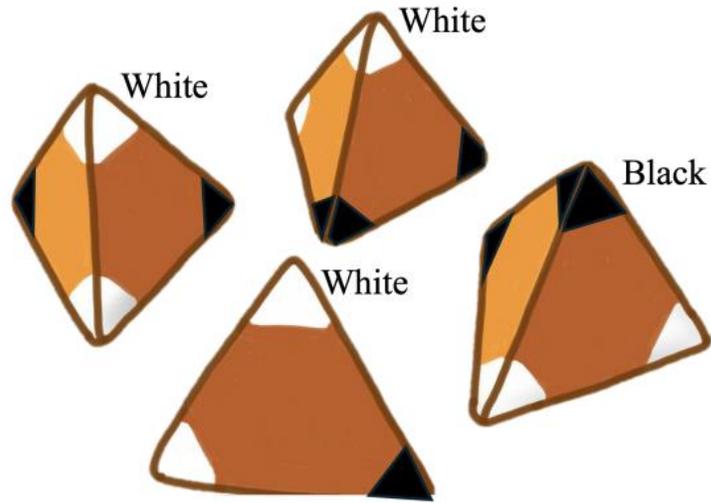


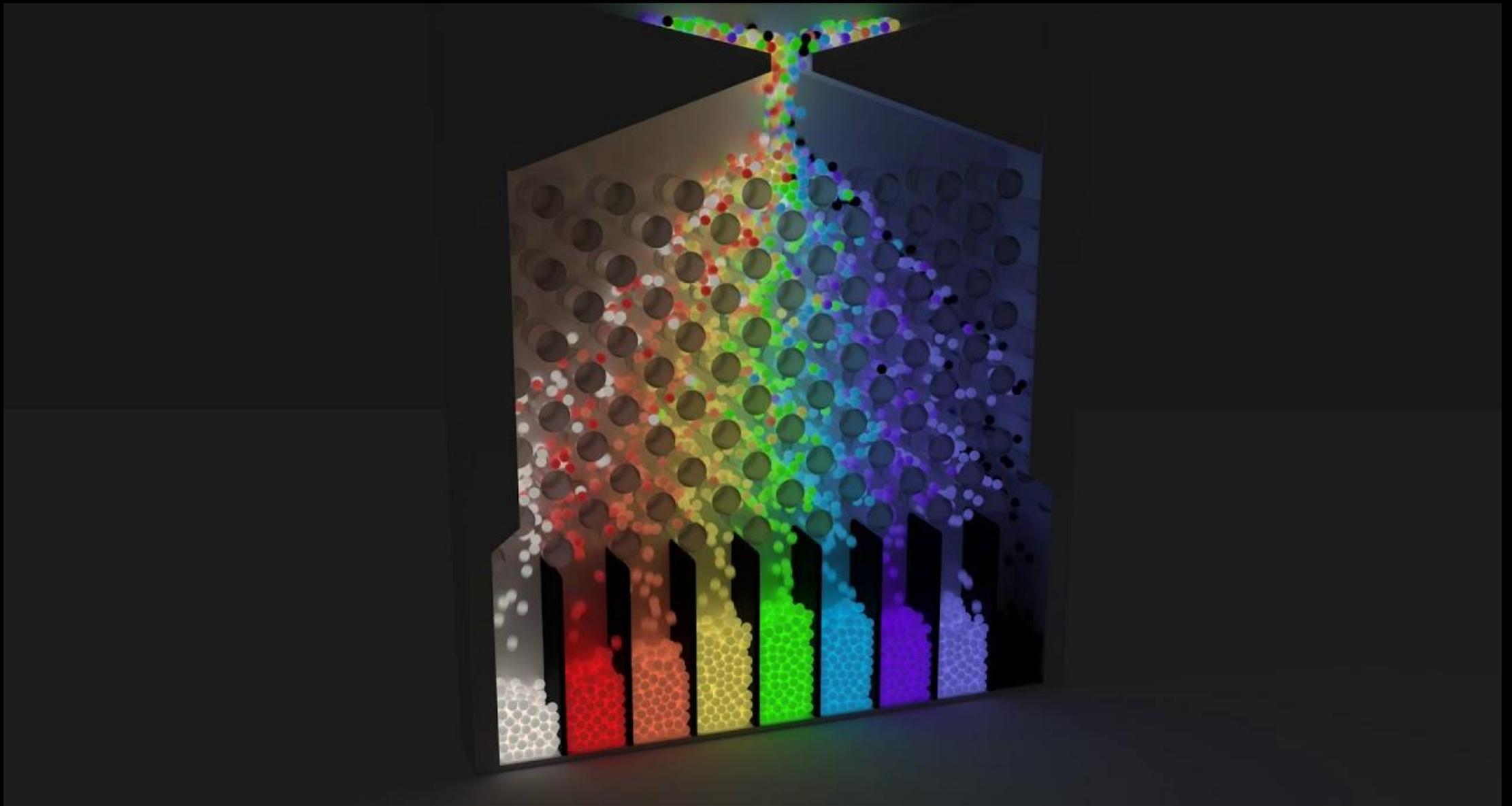
# The PMF as a Graph: $X \sim \text{Bin}(n = 20, p = 0.6)$

Parameter  $n$ :  Parameter  $p$ :



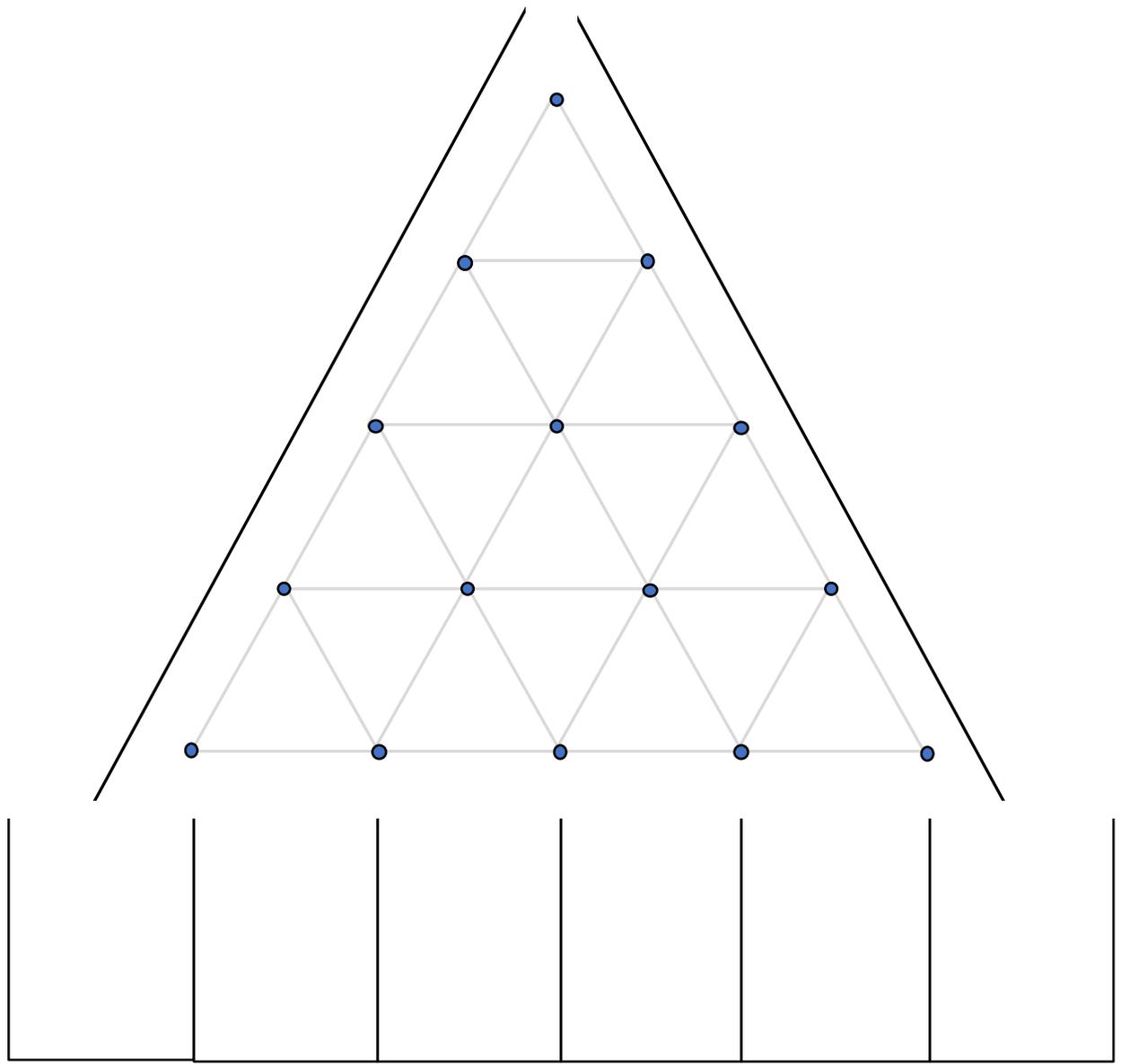
# Many Stories fit the Binomial



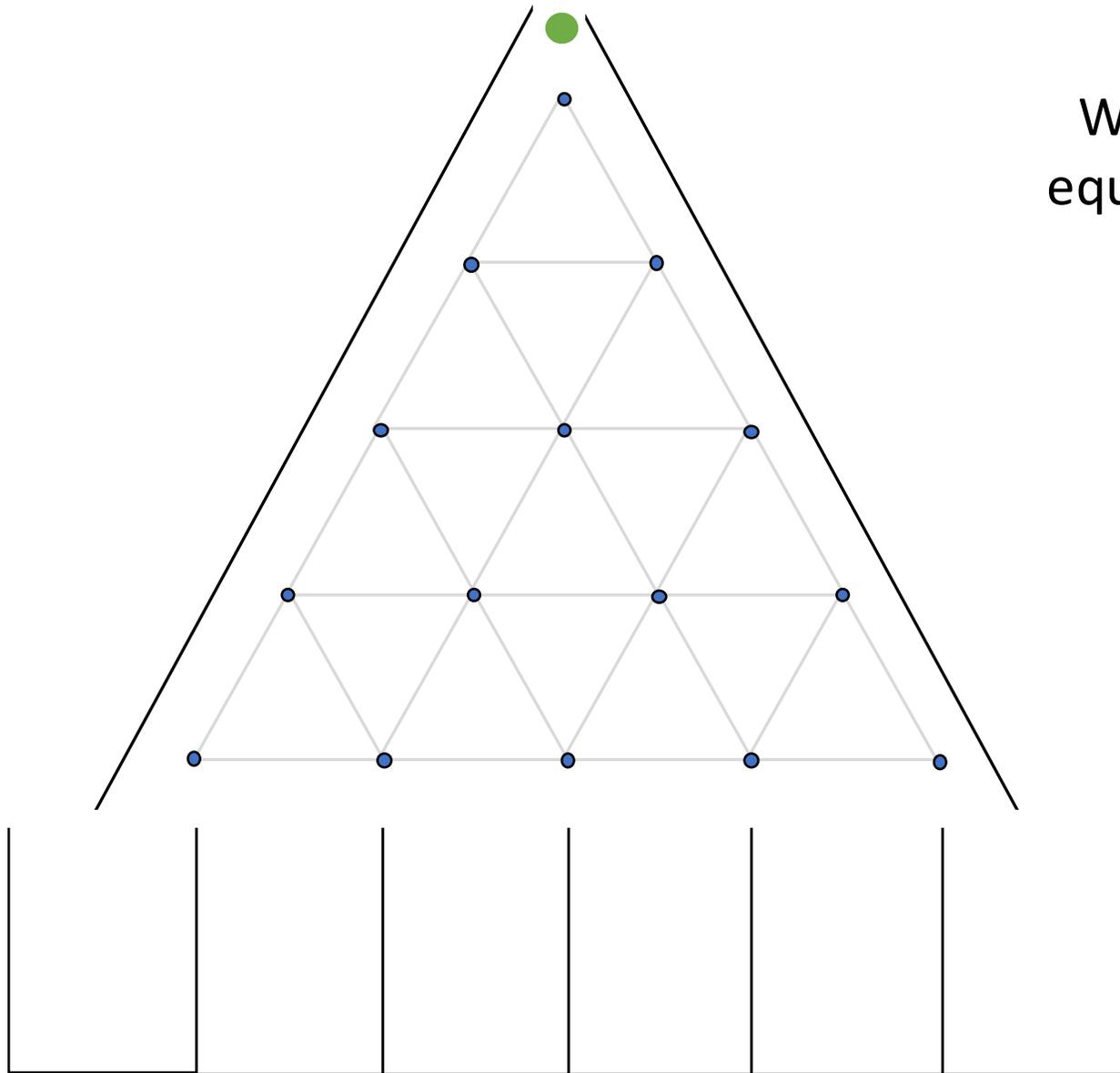


Galton Board Time!

# Galton Board Fun

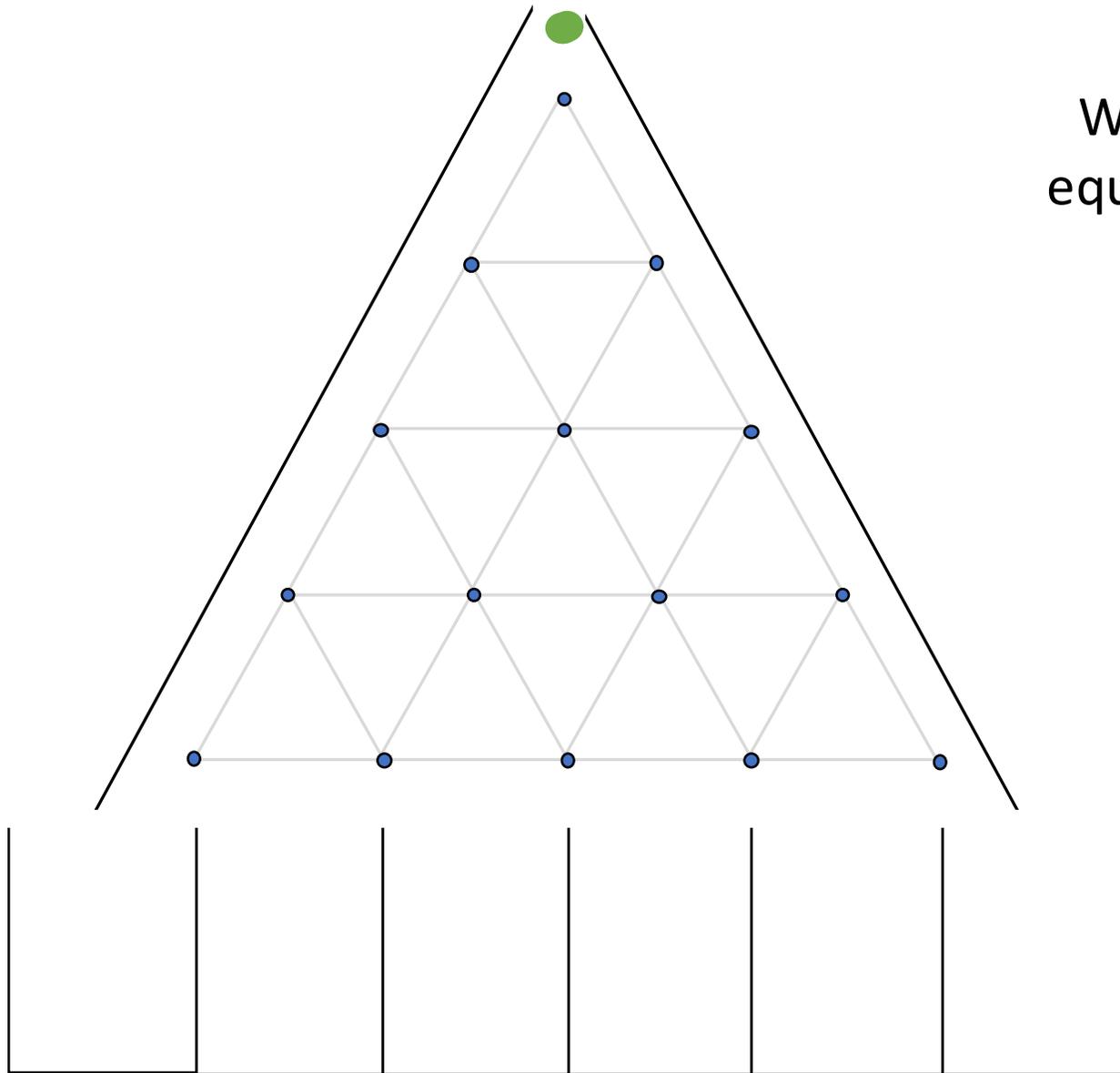


# Galton Board Fun



When a marble hits a pin, it has equal chance of going left or right.

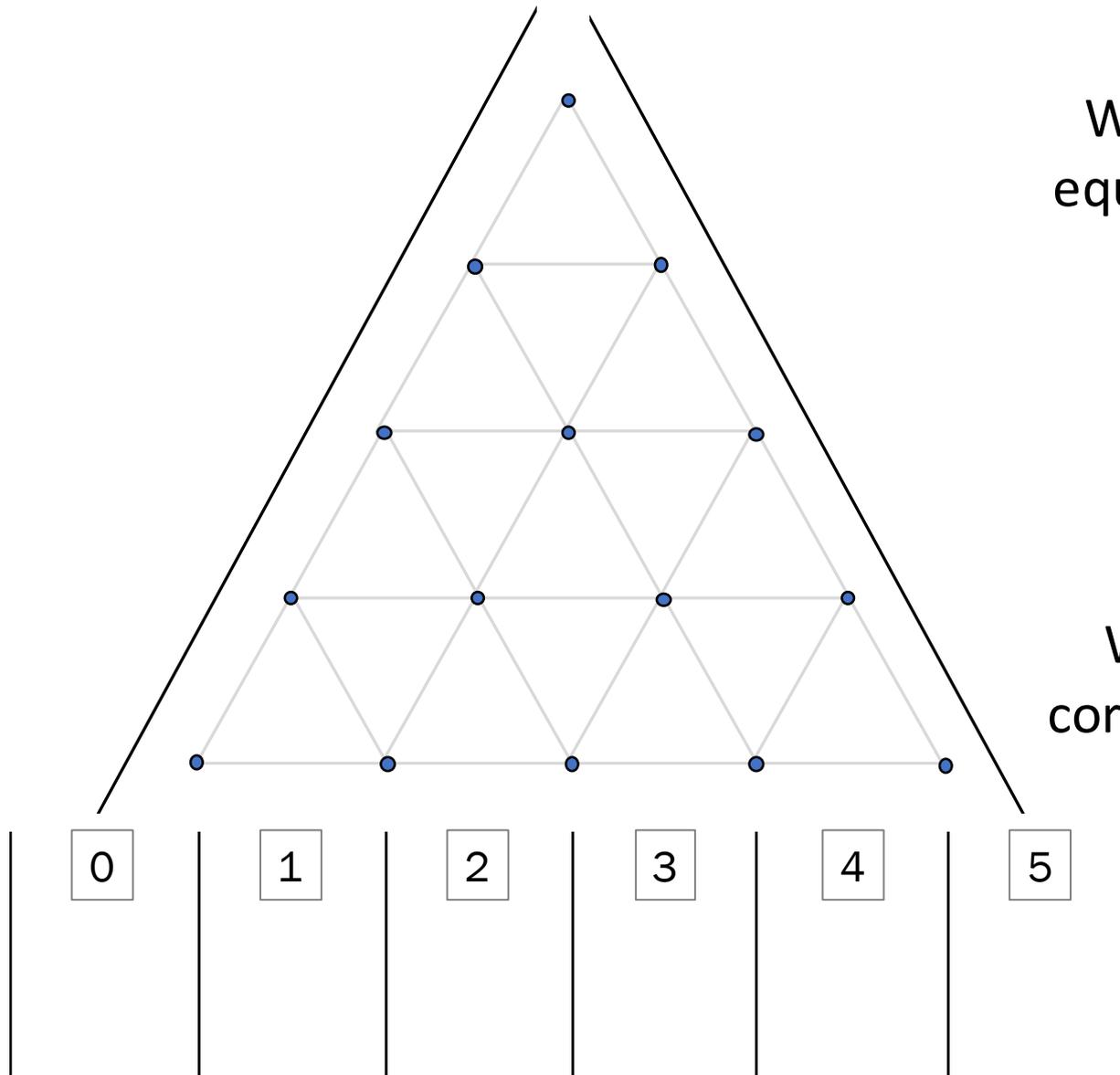
# Galton Board Fun



When a marble hits a pin, it has equal chance of going left or right.

Each pin represents an independent event.

# Galton Board Fun

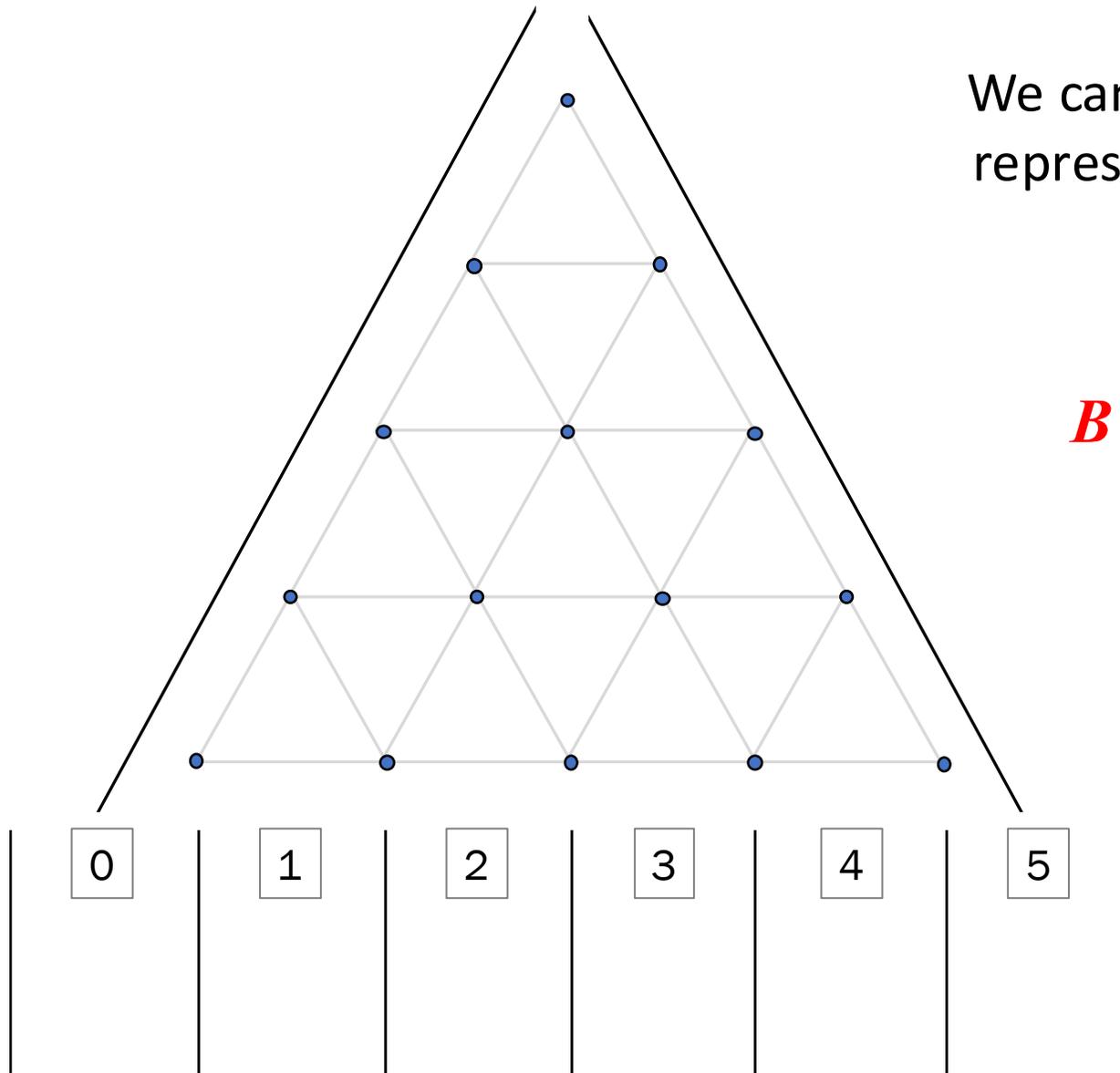


When a marble hits a pin, it has equal chance of going left or right.

Each pin represents an independent event.

Which bucket a marble lands in corresponds to the number of times the marble went right.

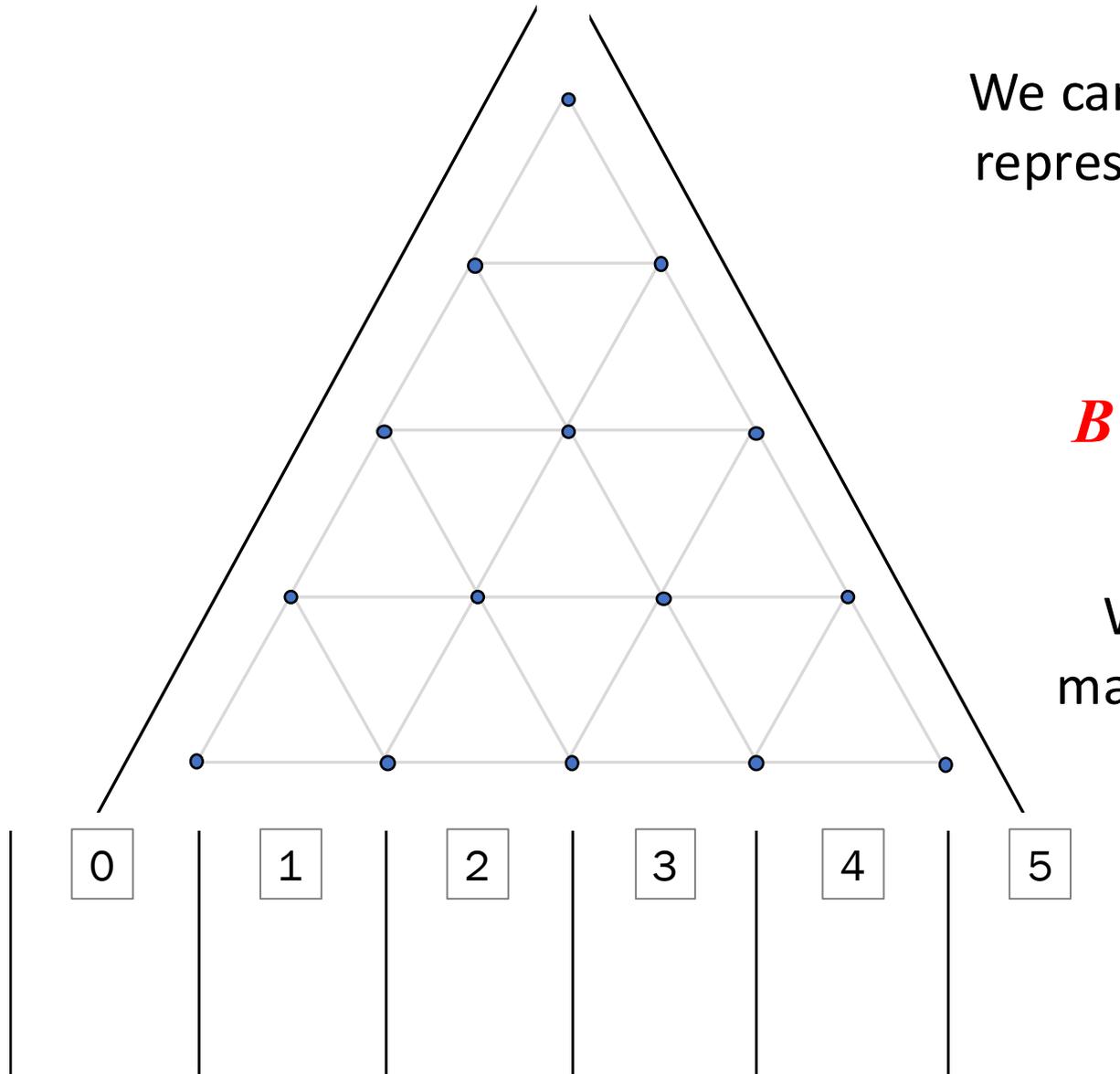
# Galton Board Fun



We can define a random variable ( $B$ ) representing which bucket a marble lands in.

$$B \sim \text{Bin}(n = \text{levels}, p = 0.5)$$

# Galton Board Fun

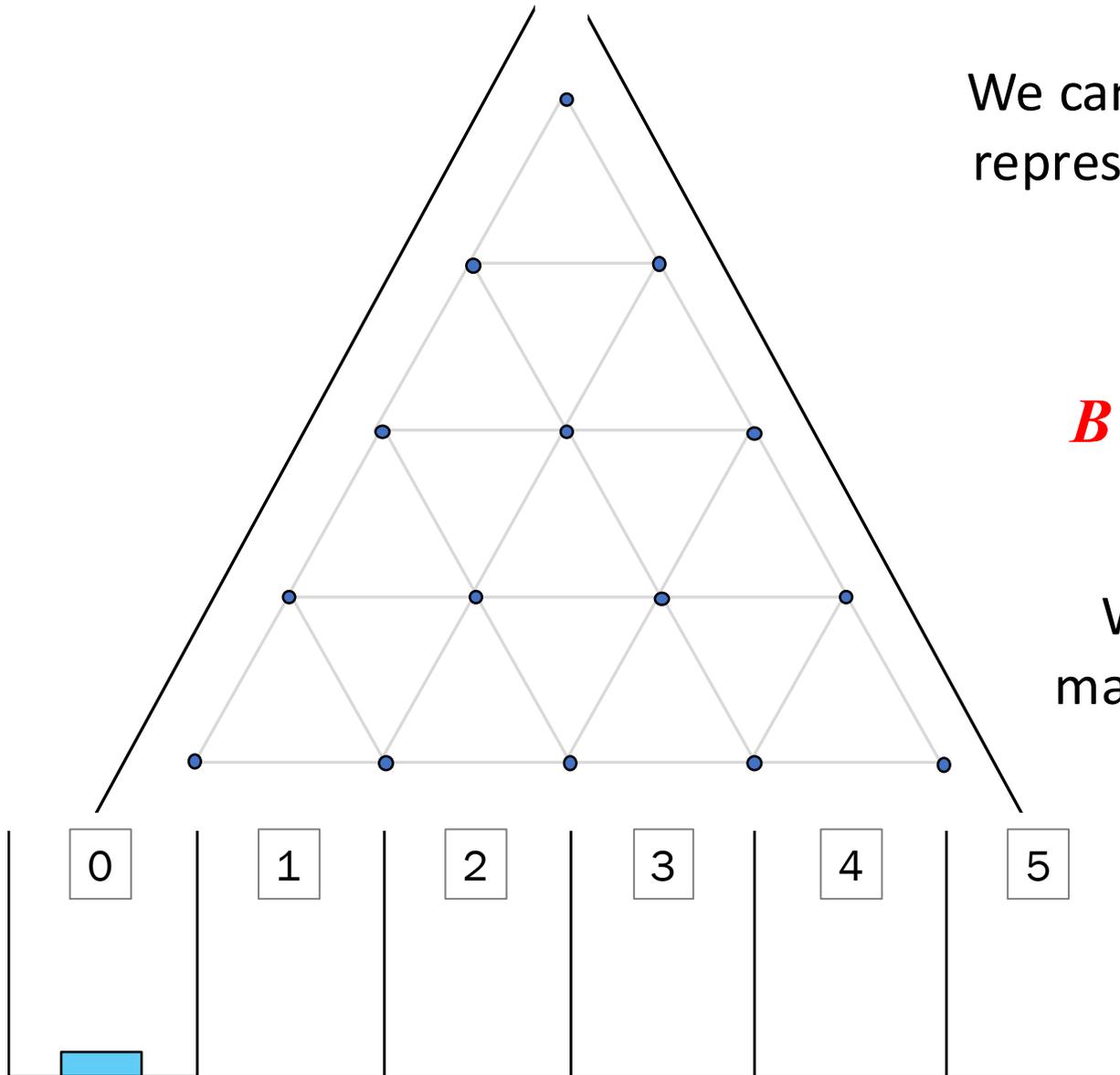


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What is the probability of a marble landing in each bucket?

# Galton Board Fun



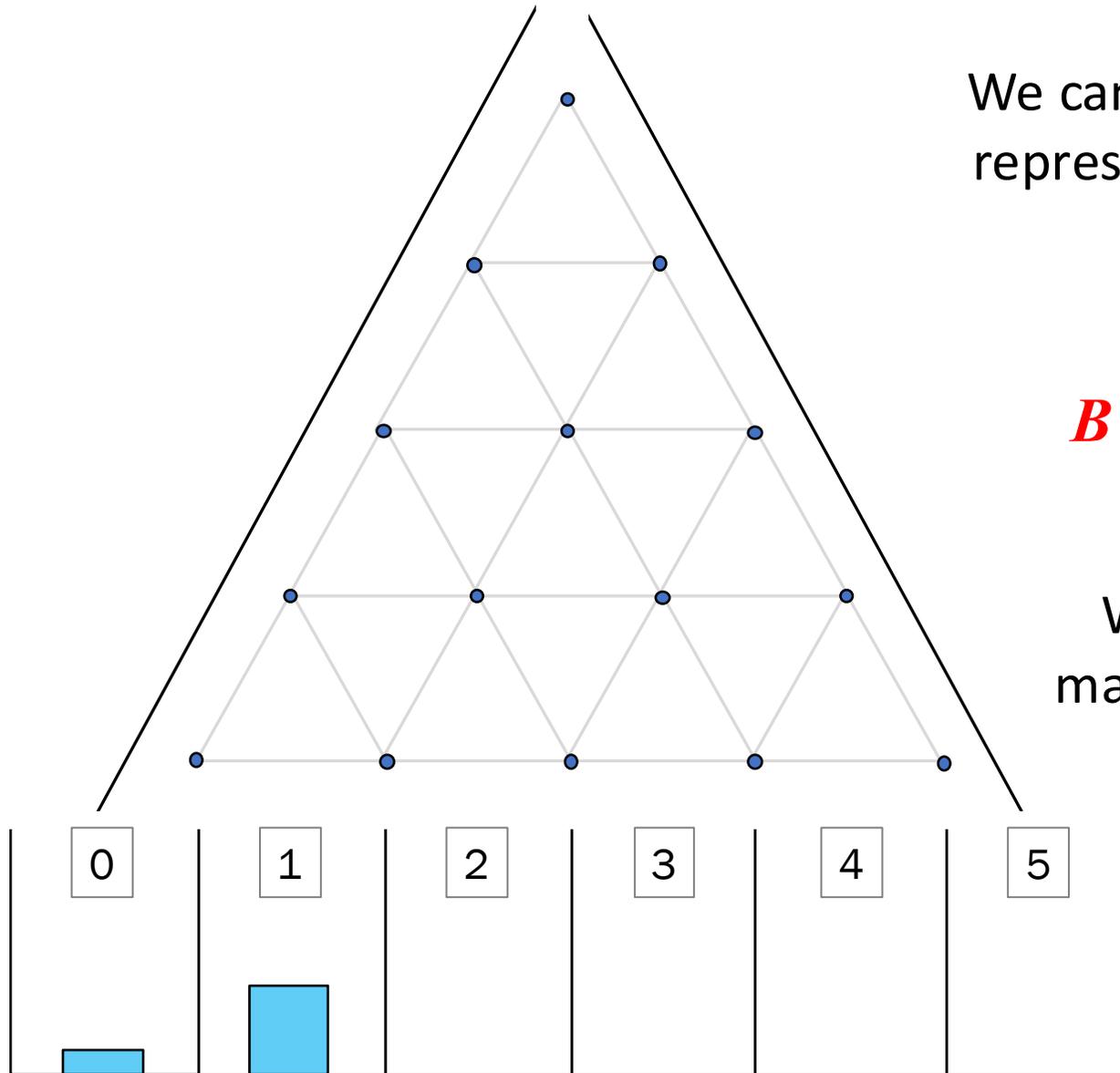
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$$B \sim \text{Bin}(n = \text{levels}, p = 0.5)$$

What is the probability of a marble landing in each bucket?

$$P(B = 0) = \binom{5}{0} \frac{1}{2}^5 \approx 0.03$$

# Galton Board Fun



We can define a random variable ( $B$ ) representing which bucket a marble lands in.

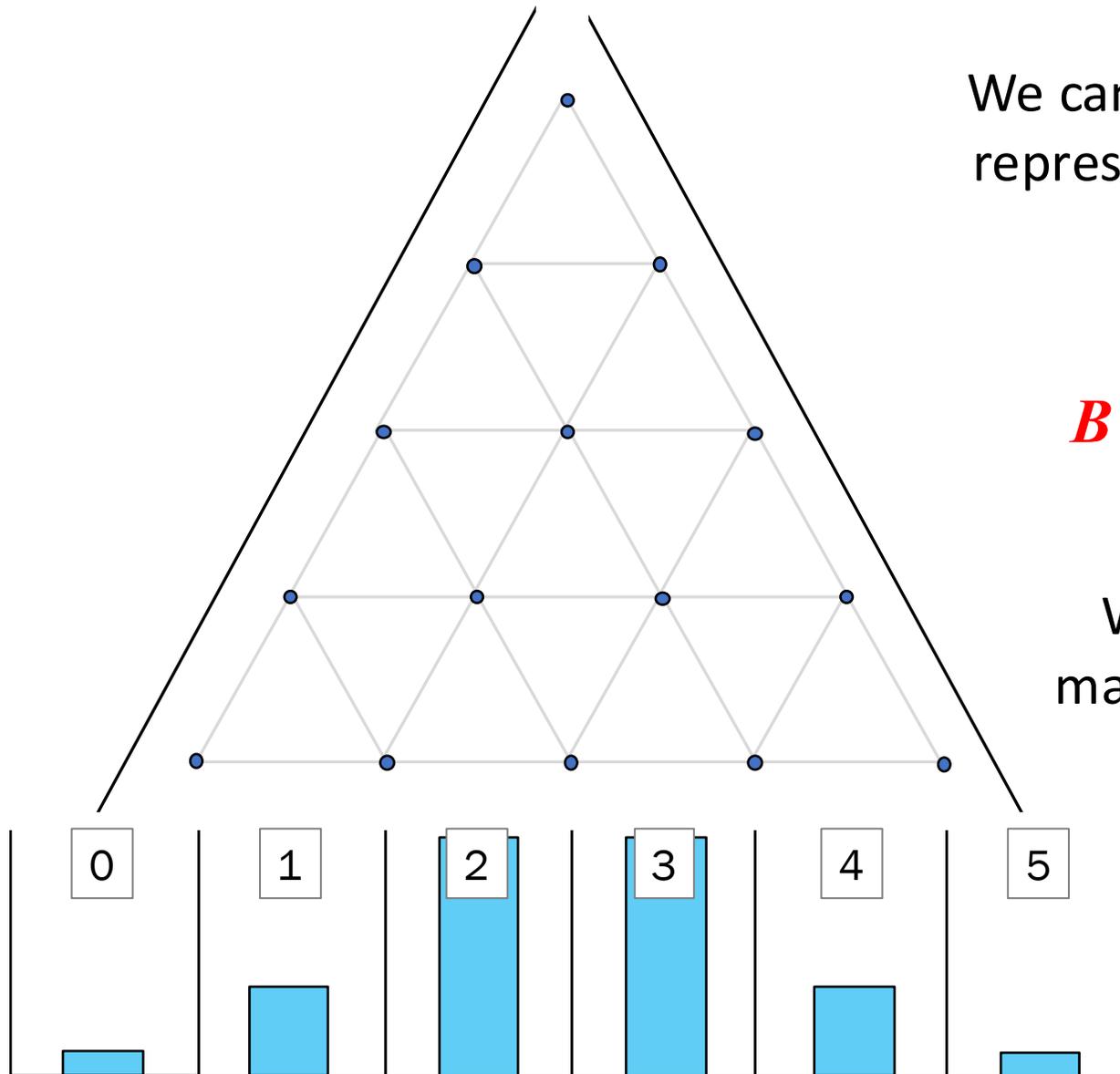
$$B \sim \text{Bin}(n = \text{levels}, p = 0.5)$$

What is the probability of a marble landing in each bucket?

$$P(B = 0) = \binom{5}{0} \frac{1}{2}^5 \approx 0.03$$

$$P(B = 1) = \binom{5}{1} \frac{1}{2}^5 \approx 0.16$$

# Galton Board Fun



We can define a random variable ( $B$ ) representing which bucket a marble lands in.

$$B \sim \text{Bin}(n = \text{levels}, p = 0.5)$$

What is the probability of a marble landing in each bucket?

This is the PMF of the binomial!

End Review

# Where are We in CS109?

---

You are here



Core Probability

$X_2$

Random  
Variables



Probabilistic  
Models



Uncertainty  
Theory



Machine  
Learning



# Classic Random Variables (with PMFs)

---

$$X \sim \text{Bern}(p)$$

Successes in one trial

$$X \sim \text{Geo}(p)$$

Trials until one success

$$Y \sim \text{Bin}(n, p)$$

Successes in  $n$  trials

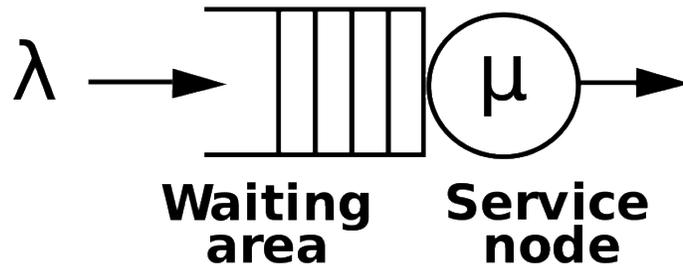
$$Y \sim \text{NegBin}(r, p)$$

Trials until  $r$  success



# Goal: Be Able to Use a New Random Variable

You are learning about servers...



You read about the MD1 queue...

You find a paper that says the length of a server “busy period” is distributed as a Borel with parameter  $\mu = 0.2$  ...

**Borel distribution**

Borel distribution	
Parameters	$\mu \in [0, 1]$
Support	$n \in \{1, 2, 3, \dots\}$
pmf	$\frac{e^{-\mu} (\mu^n)}{n!}$
Mean	$\frac{1}{1 - \mu}$
Variance	$\frac{\mu}{(1 - \mu)^3}$

**Definition** [ edit ]

A discrete random variable  $X$  is said to have a Borel distribution<sup>[1][2]</sup> with parameter  $\mu \in [0, 1]$  if the probability mass function of  $X$  is given by

$$P_{\mu}(n) = \Pr(X = n) = \frac{e^{-\mu} (\mu^n)}{n!}$$

for  $n = 1, 2, 3, \dots$

**Derivation and branching process interpretation** [ edit ]



# Geometric Random Variable

---

$X$  is **Geometric** Random Variable:  $X \sim \text{Geo}(p)$

- $X$  is number of independent trials until first success
- $p$  is probability of success on each trial
- Assumes  $p$  does not change
- $X$  takes on values  $1, 2, 3, \dots$ , with probability:

$$P(X = n) = (1 - p)^{n-1}p$$



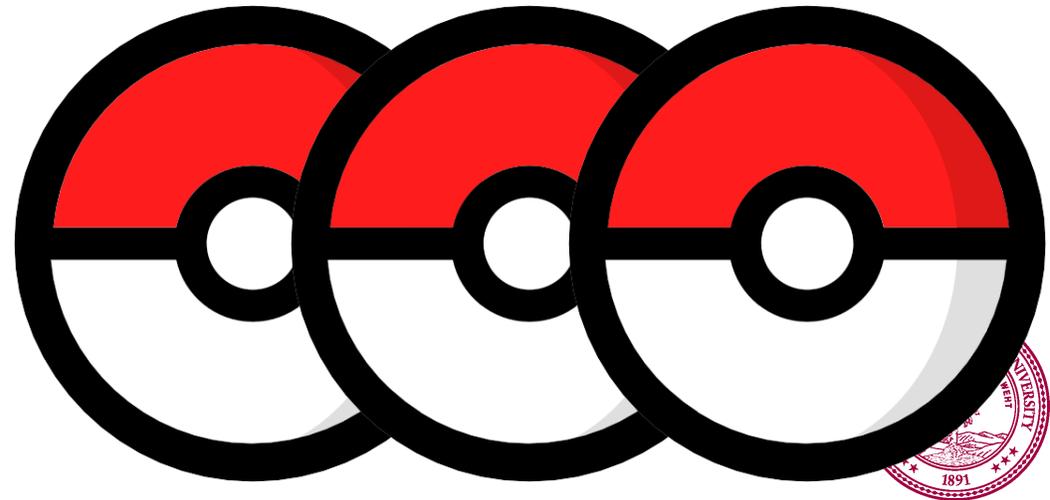
# Negative Binomial Random Variable

---

$X$  is **Negative Binomial** RV:  $X \sim \text{NegBin}(r, p)$

- $X$  is number of independent trials until  $r$  successes
- $p$  is probability of success on each trial
- Assumes  $p$  does not change.
- $X$  takes on value  $n$  with probability:

$$P(X = n) = \binom{n-1}{r-1} p^r (1-p)^{n-r}, \text{ where } n = r, r+1, \dots$$



# Classic Random Variables (with PMFs)

---

$$X \sim \text{Bern}(p)$$

$$X \sim \text{Geo}(p)$$

Trials until one success

$$P(X = x) = (1 - p)^{x-1}p$$

$$Y \sim \text{Bin}(n, p)$$

Successes in  $n$  trials

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$$

$$Y \sim \text{NegBin}(r, p)$$

Trials until  $r$  success

$$P(X = x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}$$





# Recipe For Solving Problems:

1. Recognize a classic random variable type

2. Define a random variable to be that type, with parameters

3. Profit off the PMF



$$X \sim \text{Bin}(n, p)$$



# Dating at Stanford

---

Each person you date has a 0.2 probability of being someone you spend your life with. What is the probability you need to date more than 5 people? **Your meta goal: what steps would you take to answer this question?**



# Equity in the Courts

---

## Berghuis v. Smith

*If a group is underrepresented in a jury pool, how do you tell?*

Justice Breyer [Stanford Alum] opened the questioning by invoking the binomial theorem. He hypothesized a scenario involving **“an urn with a thousand balls, and sixty are yellow, and nine hundred forty are navy-blue, and then you select them at random... twelve at a time.”** According to Justice Breyer and the binomial theorem, if the purple balls were under represented jurors then **“you would expect... something like a third to a half of juries would have at least one yellow ball”** on them.

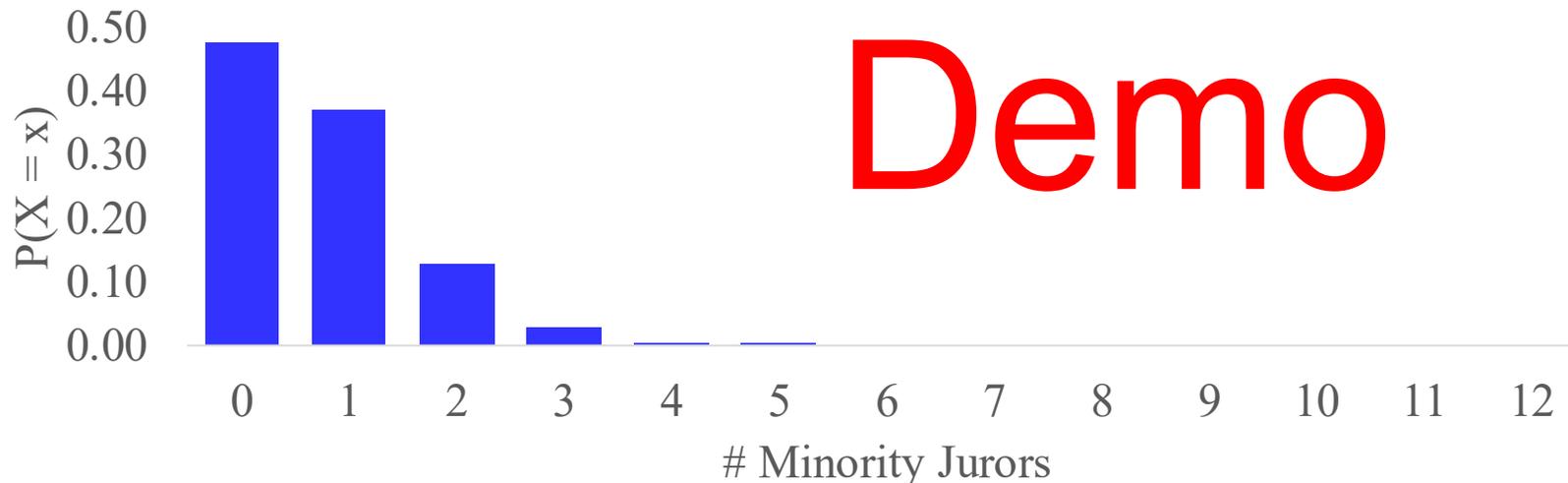


# Equity in the Courts

Approximation using Binomial distribution

- Assume  $P(\text{blue ball})$  constant for every draw =  $60/1000$
- $X = \#$  blue balls drawn.  $X \sim \text{Bin}(12, 60/1000 = 0.06)$
- $P(X \geq 1) = 1 - P(X = 0) \approx 1 - 0.4759 = 0.5240$

*In Breyer's description, should actually expect just over half of juries to have at least one non-white person on them*



# Bitcoin Mining



SHA-256 Hash( , )

Data  
*Fixed*

Salt  
*Choice*

Number that looks like random bits

You “mine a bitcoin” if, for given data  $D$ , you find a salt number  $N$  such that Hash( $D$ ,  $N$ ) produces a string that starts with  $g$  zeroes.



You “mine a bitcoin” if, for given data  $D$ , you find a number  $N$  such that  $\text{Hash}(D, N)$  produces a string that starts with  $g$  zeroes.

---

(a) What is the probability that Hash outputs a bit string which starts with  $g$  zeroes (in other words you mine a bitcoin)?

Let  $X$  be the number of zeros in the first  $g$  bits.  $X \sim \text{Bin}(n = g, p = 0.5)$

$$P(X = g) = \binom{g}{g} \frac{1^g}{2^g} = \frac{1}{2^g} \quad \text{Call this answer } p_a$$

(b) What is the probability that you will need under 100 attempts to mine 2 bit coins?

Let  $Y$  be the number of tries until you mine 2 bitcoins.  $Y \sim \text{NegBin}(r = 2, p = p_a)$

$$\begin{aligned} P(Y < 100) &= \sum_{x=2}^{99} P(Y = x) \\ &= \sum_{x=2}^{99} \binom{x-1}{r-1} p^r (1-p)^{x-r} \end{aligned}$$



# Classic Random Variables (with PMFs)

---

$$X \sim \text{Bern}(p)$$

$$X \sim \text{Geo}(p)$$

Trials until one success

$$P(X = x) = (1 - p)^{x-1}p$$

$$Y \sim \text{Bin}(n, p)$$

Successes in  $n$  trials

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$$

$$Y \sim \text{NegBin}(r, p)$$

Trials until  $r$  success

$$P(X = x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}$$



# Can Jacob Bernoulli Have a Variable Named After Him?



*Here yee. I want to have a random variable named after myself. Huzzah.*

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Yes - the Bernoulli random variable:  $X \sim \text{Bern}(p)$

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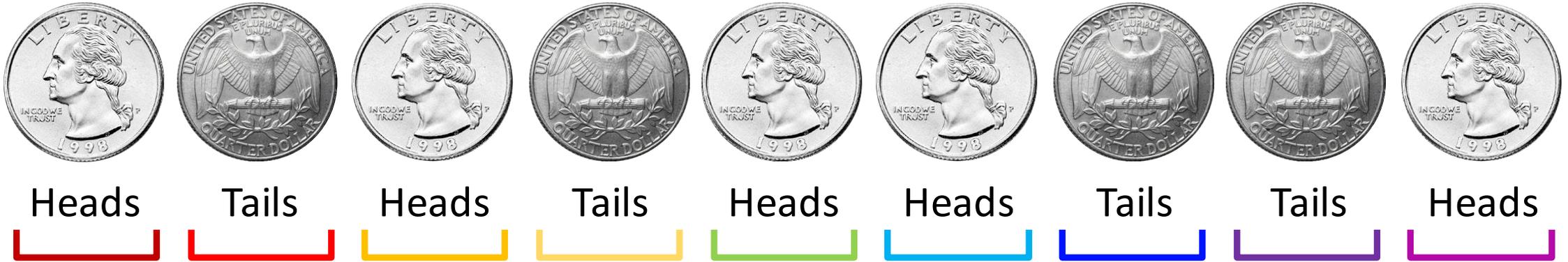
Yes - the Bernoulli random variable:  $X \sim \text{Bern}(p)$

- The Bernoulli is an **indicator** random variable (value is either 0 or 1).
- $P(X = 1) = p$
- $P(X = 0) = 1 - p$
- Examples: a single coin flip, one ad click, any binary event

*(this is the whole PMF)*

# Random Variable Sums

## The Binomial



# Random Variable Sums

The Binomial

...is a sum of Bernoulli random variables



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The Binomial

...is a sum of Bernoulli random variables



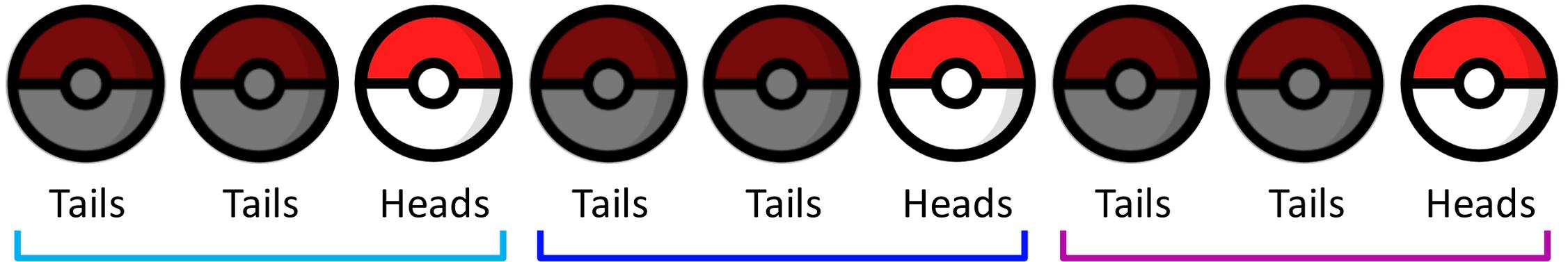
Let  $X_1 \sim \text{Bern}(p = 1/2)$  and  $X_2 \sim \text{Bern}(p = 1/2)$ .

$$Y \sim \text{Bin}(n = 2, p = 1/2)$$

$$Y = X_1 + X_2$$

# Random Variable Sums

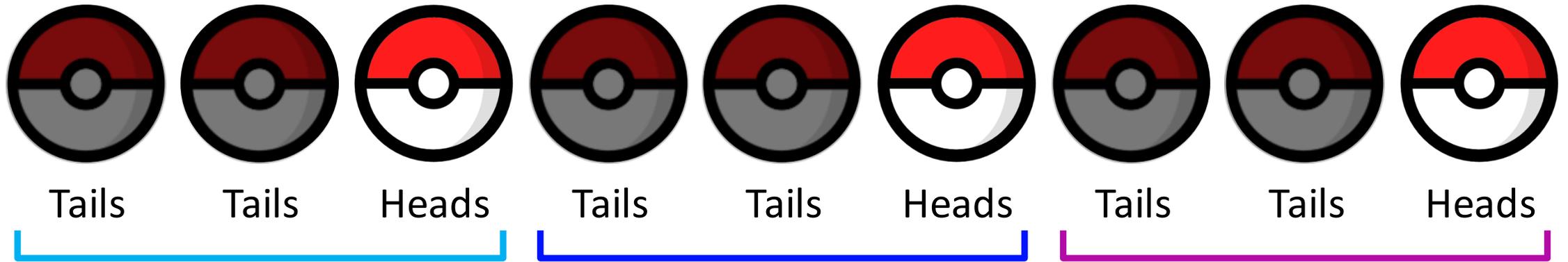
## The Negative Binomial



# Random Variable Sums

The Negative Binomial

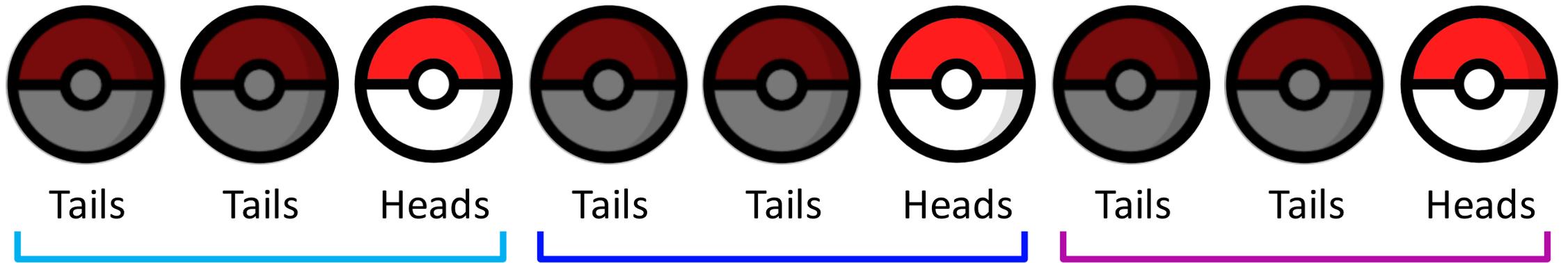
...is a sum of Geometric random variables



# Random Variable Sums

The Negative Binomial

...is a sum of Geometric random variables



Let  $X_1 \sim \text{Geo}(p = 1/3)$ ,  $X_2 \sim \text{Geo}(p = 1/3)$ , and  $X_3 \sim \text{Geo}(p = 1/3)$ .

$Y \sim \text{NegBin}(r = 3, p = 1/3)$

$$Y = \underline{X_1} + \underline{X_2} + \underline{X_3}$$

# Classic Random Variables (with PMFs)

$$X \sim \text{Bern}(p)$$

Successes in one trial

$$P(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \end{cases}$$

$$X \sim \text{Geo}(p)$$

Trials until one success

$$P(X = x) = (1 - p)^{x-1}p$$

$$Y \sim \text{Bin}(n, p)$$

Successes in  $n$  trials

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$$

$$Y \sim \text{NegBin}(r, p)$$

Trials until  $r$  success

$$P(X = x) = \binom{x-1}{r-1} p^r (1 - p)^{x-r}$$

[Short Pedagogical Pause]

# Time for some tender moments\*

\*Moments: numbers that summarize different aspects of a random variable

Expectation

# Expected Value

$$E[X] = \sum_x x \cdot P(X = x)$$

The value  $x$  is indicated by a blue arrow pointing to the  $x$  in the summand.

The probability of that value  $P(X = x)$  is indicated by a blue arrow pointing to the probability term.

Loop over all values  $x$  that  $X$  can take on



# Expected Value

---

Expected value answers the question:

*What is the average value we could expect some random variable to be?*

Also called: **Mean**, *Expectation*, **Weighted Average**, **Center of Mass**,  
*1<sup>st</sup> Moment*

$$E[X] = \sum_x x \cdot P(X = x)$$



# Example: Expected Value of Dice Roll

---

Let  $X$  be the result of rolling a 6-sided dice.

$$P(X = x) = \frac{1}{6} \text{ for } x \in \{1, 2, 3, 4, 5, 6\}$$

What is the expectation of  $X$ ?



$$E[X] = \sum_x x \cdot P(X = x)$$

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What is the expectation of  $X$ ?

$$\begin{aligned} E[X] &= \sum_{x=1}^6 x \cdot P(X = x) \\ &= 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6} \\ &= 3.5 \end{aligned}$$



$$E[X] = \sum_x x \cdot P(X = x)$$

## Example: Expected Value of Dice Roll

Let  $X$  be the result of rolling a 6-sided dice.

$$P(X = x) = \frac{1}{6} \text{ for } x \in \{1, 2, 3, 4, 5, 6\}$$

What is the expectation of  $X$ ?

$$E[X] = \sum_{x=1}^6 x \cdot P(X = x)$$

$$= 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + 3 \cdot \frac{1}{6} + 4 \cdot \frac{1}{6} + 5 \cdot \frac{1}{6} + 6 \cdot \frac{1}{6}$$

$$= 3.5$$

$E[X]$  is not always an actual possible outcome for  $X$



# Lying With Statistics



Imagine a university has 3 classes, with 5, 10, and 150 students in each class. We randomly choose a **class** with equal probability.

Let  $X$  be the chosen class's size. What is  $E[X]$ ?



# Lying With Statistics



Imagine a university has 3 classes, with 5, 10, and 150 students in each class. We randomly choose a **class** with equal probability.

Let  $X$  be the chosen class's size. What is  $E[X]$ ?

$$P(X = 5) = 1/3$$

$$P(X = 10) = 1/3$$

$$P(X = 150) = 1/3$$

$$E[X] = \sum_{x \in \{5, 10, 150\}} x \cdot P(X = x)$$

$$= 5 \cdot \frac{1}{3} + 10 \cdot \frac{1}{3} + 150 \cdot \frac{1}{3}$$

$$= 55$$



# Lying With Statistics



Imagine a university has 3 classes, with 5, 10, and 150 students in each class. We randomly choose a **student** with equal probability.

Let  $X$  be the chosen student's class size. What is  $E[X]$ ?



# Lying With Statistics



Imagine a university has 3 classes, with 5, 10, and 150 students in each class. We randomly choose a **student** with equal probability.

Let  $X$  be the chosen student's class size. What is  $E[X]$ ?

$$P(X = 5) = 5/165$$

$$P(X = 10) = 10/165$$

$$P(X = 150) = 150/165$$

$$E[X] = \sum_{x \in \{5, 10, 150\}} x \cdot P(X = x)$$

$$= 5 \cdot \frac{5}{165} + 10 \cdot \frac{10}{165} + 150 \cdot \frac{150}{165}$$

$$= 137$$



# Expectation from Data

List called data

$X$
3
2
6
10
1
1
5
4
...

$$E[X] = \sum_x x \cdot P(X = x)$$

$$\approx \sum_x x \cdot \frac{\text{count}(X = x)}{N}$$

Length of data



$$\approx \frac{1}{N} \sum_x x \cdot \text{count}(X = x)$$

$$\approx \frac{1}{N} \sum_{v \in \text{data}} v$$



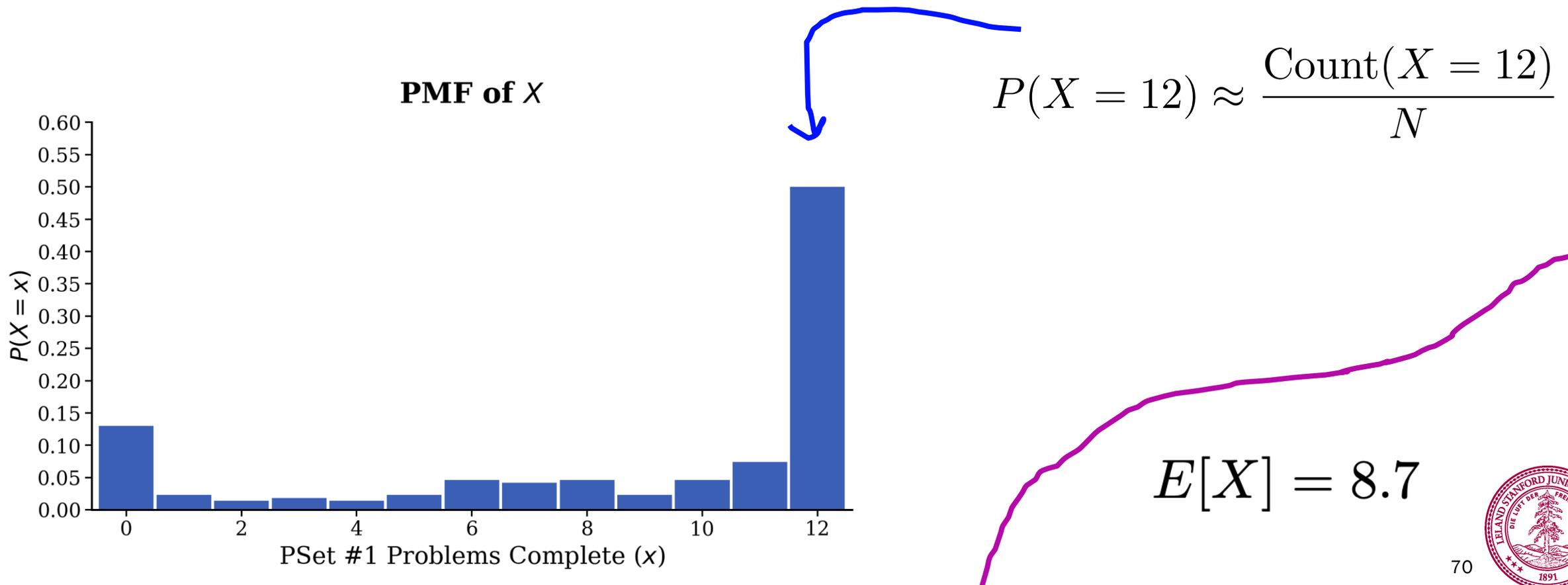
Expectation is a single number  
summary...

Expectation leaves much to be  
desired...

# Expectation vs PMF

Let  $X$  be the number of problems that a randomly selected student has completed, as of 11a today.

$X$  takes on values, with uncertainty.  $X$  is a random variable.



Why People Care?

# Properties of Expectation (proof later)

---

## Linearity:

$$E[aX + b] = aE[X] + b$$

- Consider  $X = 6$ -sided die roll, Winnings =  $2X - 1$ .
- $E[X] = 3.5$        $E[2X-1] = 6$

**Expectation of a sum** is the sum of expectations

$$E[X + Y] = E[X] + E[Y]$$

## Unconscious statistician:

$$E[g(X)] = \sum_{x \in X} g(x) P(X = x)$$



# Law of the Unconscious Statistician (LOTUS)

$$E[g(X)] = \sum_x g(x)P(X = x)$$

This lets you get the expectation of **any** function of a random variable.

**Examples:**

$$E[X^2] = \sum_x x^2 \cdot P(X = x)$$

$$E[\sin(X)] = \sum_x \sin(x) \cdot P(X = x)$$

$$E[\sqrt{X}] = \sum_x \sqrt{x} \cdot P(X = x)$$

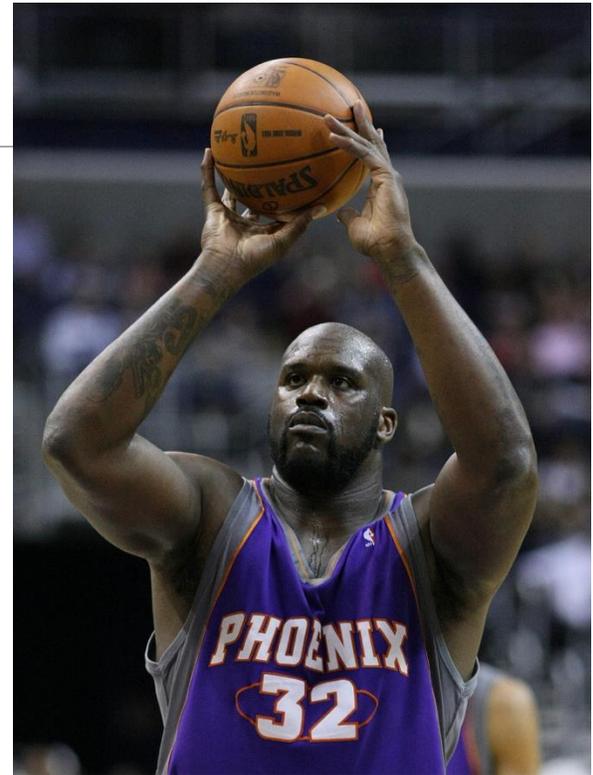
# Expectation of Classic Random Variables

# Expected Value of Free Throws

In basketball, players sometimes get a chance to shoot a free throw. If they make it, the team gets 1 point; otherwise they get no points.

Some players are not very good at free throws, such as Shaq. While in the NBA, Shaq made only 53% of his free throws.

Let  $X$  be the points gained from Shaq attempting a free throw. What is  $E[X]$ ?

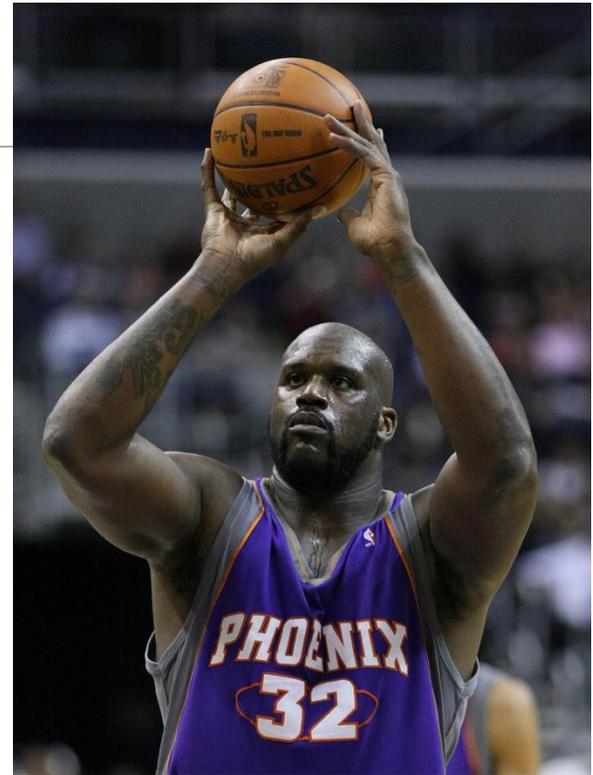


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$$X \sim \text{Bern}(p = 0.53)$$

$$\begin{aligned} E[X] &= 0 \cdot P(X = 0) + 1 \cdot P(X = 1) \\ &= 0 \cdot 0.47 + 1 \cdot 0.53 = \underline{0.53} \end{aligned}$$

For Bernoulli random variables,  $E[X] = p$  (always)



# With Classic RVs, You Get Expectations For Free Too!

Course Reader for CS109

Search book...

Notation Reference  
Core Probability Reference  
Random Variable Reference  
Python Reference  
Calculators

Part 1: Core Probability

- Counting
- Combinatorics
- Definition of Probability
- Equally Likely Outcomes
- Probability of or
- Conditional Probability
- Independence
- Probability of and
- Law of Total Probability
- Bayes' Theorem
- Log Probabilities
- Many Coin Flips
- Applications
  - Enigma Machine
  - Serendipity
  - Random Shuffles
  - Random Graphs
  - Bacteria Evolution

Part 2: Random Variables

- Random Variables
- Probability Mass Functions

## Random Variable Reference

### Discrete Random Variables

#### Bernoulli Random Variable

**Notation:**  $X \sim \text{Bern}(p)$

**Description:** A boolean variable that is 1 with probability  $p$

**Parameters:**  $p$ , the probability that  $X = 1$ .

**Support:**  $x$  is either 0 or 1

**PMF equation:** 
$$P(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \end{cases}$$

**PMF (smooth):**  $P(X = x) = p^x(1 - p)^{1-x}$

**Expectation:**  $E[X] = p$

**Variance:**  $\text{Var}(X) = p(1 - p)$

**PMF graph:**

Parameter  $p$ :

Value of X	Probability
0	0.2
1	0.8



# We Can Now Calculate Expectation of Binomial

---

$$X \sim \text{Bin}(n, p)$$



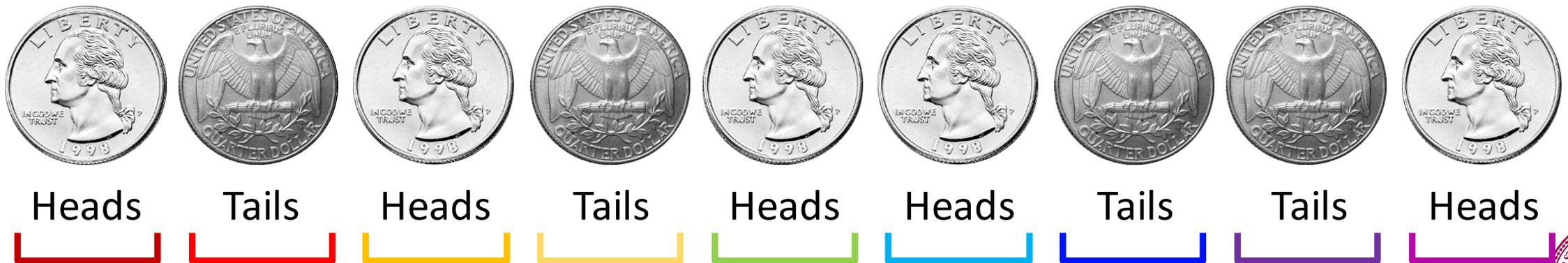
# We Can Now Calculate Expectation of Binomial

$$X \sim \text{Bin}(n, p)$$

Let  $Y_i$  be 1 if trial  $i$  was a success, otherwise 0, with  $i$  from 1 to  $n$ .  $Y_i \sim \text{Bern}(p)$ .

The Binomial

...is a sum of Bernoulli random variables



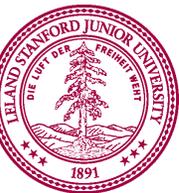
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$$\mathbf{E}[X] = \mathbf{E} \left[ \sum_{i=1}^n Y_i \right] \quad \text{Since } X = \sum_{i=1}^n Y_i$$



# We Can Now Calculate Expectation of Binomial

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Expectation of sum

**Expectation of a sum** is the sum of expectations:  $E[X + Y] = E[X] + E[Y]$



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$$\text{Since } X = \sum_{i=1}^n Y_i$$

$$= \sum_{i=1}^n \mathbf{E}[Y_i]$$

Expectation of sum

$$= \sum_{i=1}^n p$$

Expectation of Bernoulli

$$= n \cdot p$$

Sum  $n$  times

True for every binomial  
ever



# You Get So Much For Free!

## Binomial Random Variable

**Notation:**  $X \sim \text{Bin}(n, p)$

**Description:** Number of "successes" in  $n$  identical, independent experiments each with probability of success  $p$ .

**Parameters:**  $n \in \{0, 1, \dots\}$ , the number of experiments.  
 $p \in [0, 1]$ , the probability that a single experiment gives a "success".

**Support:**  $x \in \{0, 1, \dots, n\}$

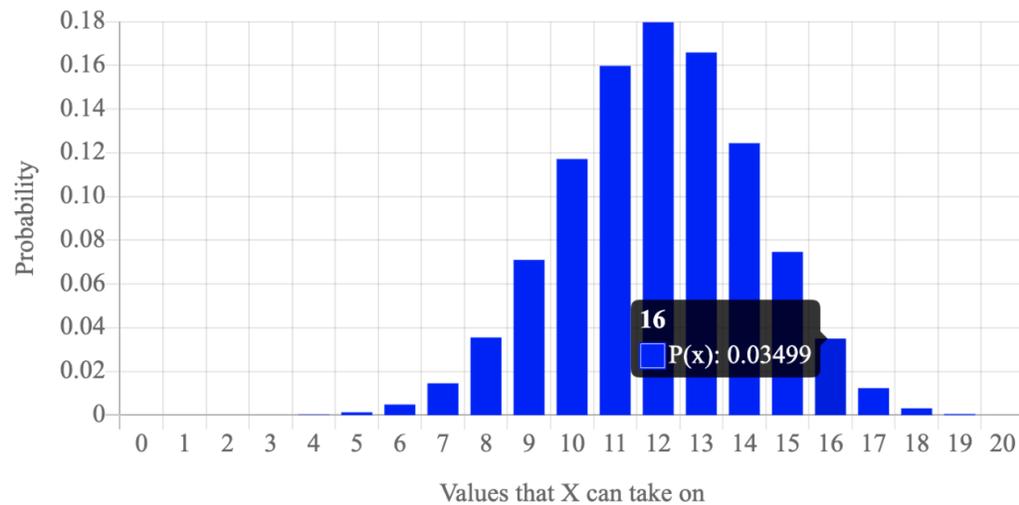
**PMF equation:**  $\Pr(X = x) = \binom{n}{x} p^x (1 - p)^{n-x}$

**Expectation:**  $E[X] = n \cdot p$

**Variance:**  $\text{Var}(X) = n \cdot p \cdot (1 - p)$

**PMF graph:**

Parameter  $n$ :  Parameter  $p$ :

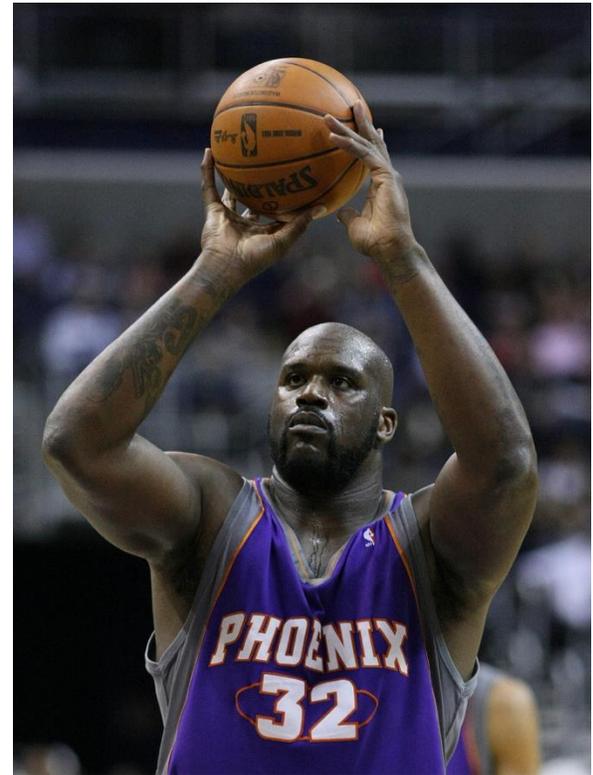


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Let  $Y$  be the points gained from Shaq attempting **500** free throws. What is  $E[Y]$ ?



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$$E[Y] = n \cdot p = 500 \cdot 0.53 = 265$$



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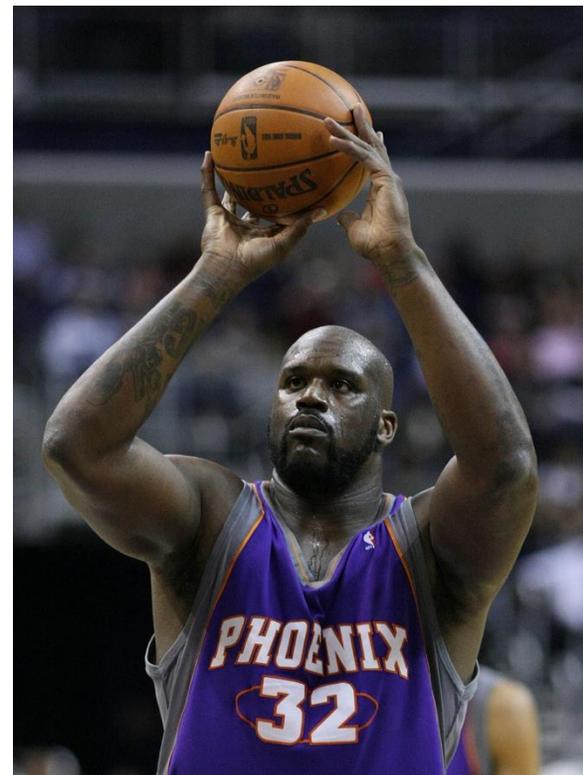
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$$Y \sim \text{Bin}(n = 500, p = 0.53)$$

$$E[Y] = n \cdot p = 500 \cdot 0.53 = 265$$

Challenge: If Shaq was 10% better at shooting free throws, how many *more* free throws would you expect him to make, out of 500?



# Expected Value of The Geometric

$$\text{If } X \sim \text{Geo}(p), \text{ then } E[X] = \frac{1}{p}$$

This definition has intuition built in:

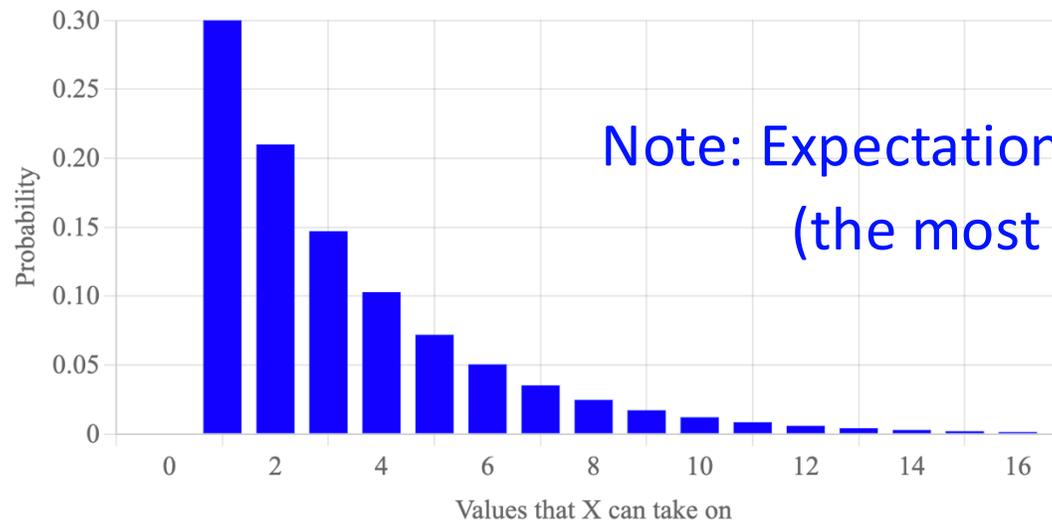
- If Shaq makes about half his free throws, then on average, it will take him two shots to make one free throw.  $E[X] = (1/2)^{-1} = 2$ .

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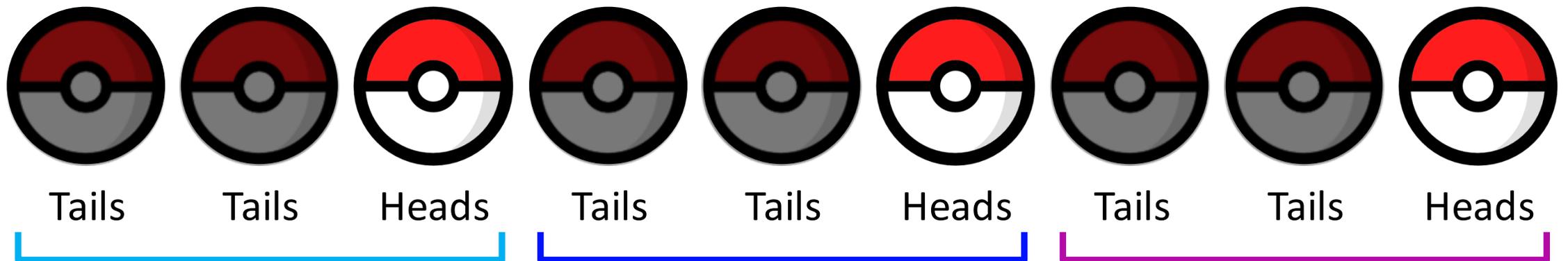
Note: Expectation is often **not** the mode  
(the most likely outcome)

# Expected Value of The Negative Binomial

We can derive using the **sum of expectations** property, similar to binomials.

The Negative Binomial

...is a sum of Geometric random variables



# Expected Value of The Negative Binomial

We can derive using the **sum of expectations** property, similar to binomials.

Let  $X_i \sim \text{Geo}(p)$ , for each  $i$  from 1 to  $r$ .

$$E[X_i] = \frac{1}{p}$$

Let  $Y \sim \text{NegBin}(r, p)$ .

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$$\text{Let } X_i \sim \text{Geo}(p), \text{ for each } i \text{ from } 1 \text{ to } r. \quad E[Y] = E \left[ \sum_{i=1}^r X_i \right]$$
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$$= \sum_{i=1}^r E[X_i]$$

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$$\begin{aligned} E[Y] &= E \left[ \sum_{i=1}^r X_i \right] \\ &= \sum_{i=1}^r E[X_i] \\ &= \sum_{i=1}^r \frac{1}{p} = \frac{r}{p} \end{aligned}$$

# Expectations of Classic Random Variables

$$X \sim \text{Geo}(p)$$

$$E[X] = \frac{1}{p}$$

$$X \sim \text{Bern}(p)$$

$$E[X] = p$$

$$Y \sim \text{NegBin}(r, p)$$

$$E[Y] = \frac{r}{p}$$

$$Y \sim \text{Bin}(n, p)$$

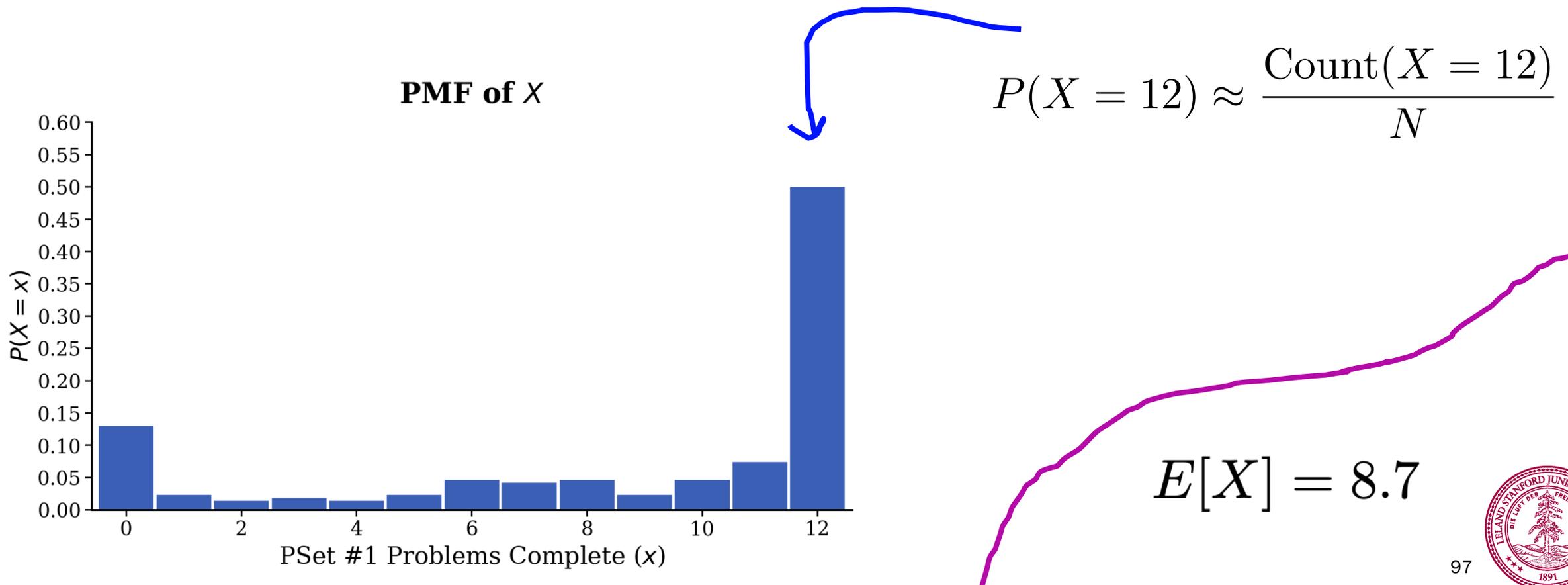
$$E[Y] = n \cdot p$$

Expectation is easy to work with, but  
still leaves much to be desired

# Expectation vs PMF

Let  $X$  be the number of problems that a randomly selected student has completed, as of 11a today.

$X$  takes on values, with uncertainty.  $X$  is a random variable.



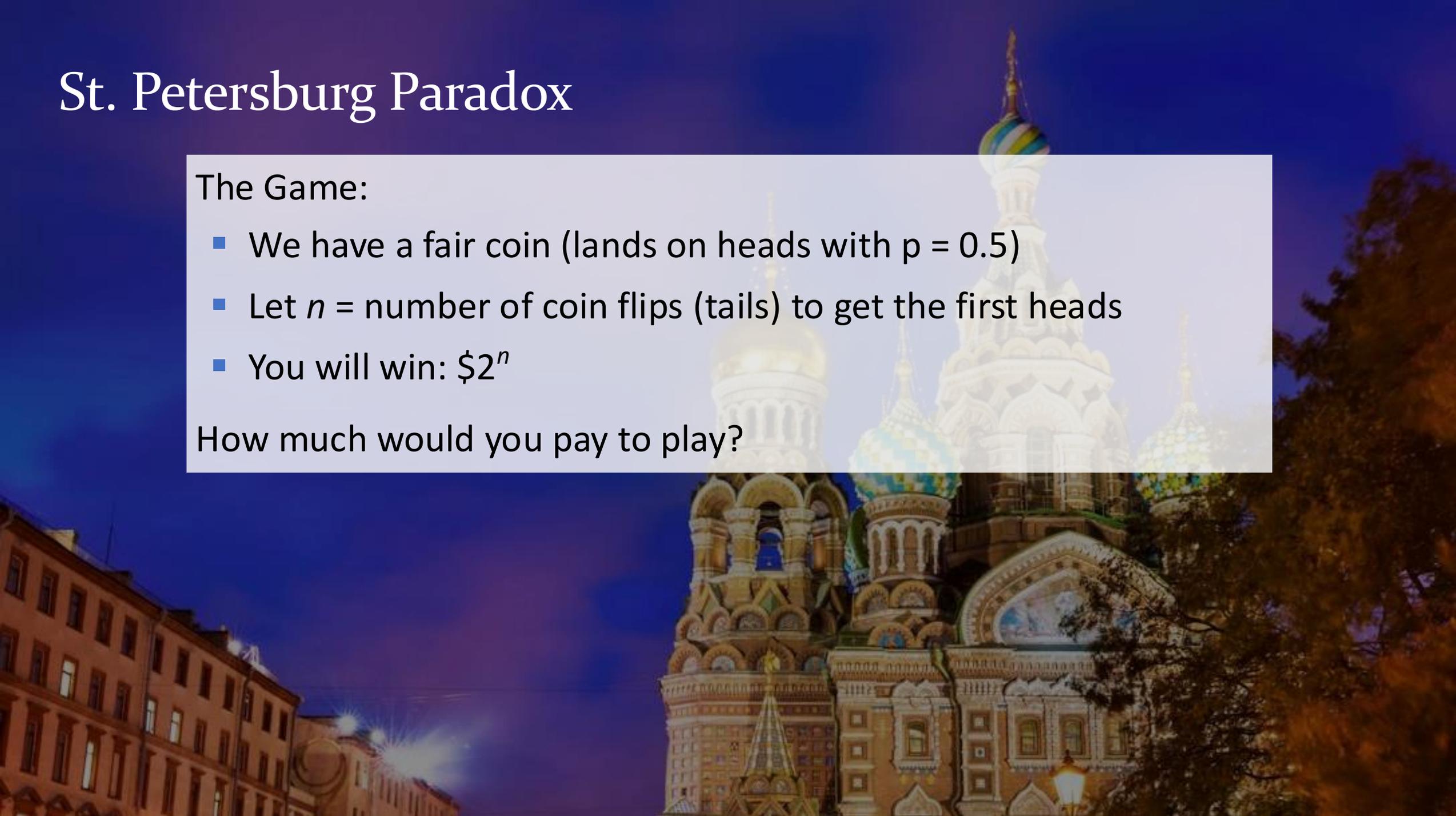
If extra time...

# St. Petersburg Paradox

The Game:

- We have a fair coin (lands on heads with  $p = 0.5$ )
- Let  $n$  = number of coin flips (tails) to get the first heads
- You will win:  $\$2^n$

How much would you pay to play?



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How much would you pay to play?

Let  $X$  be your winnings.

$$E[X] = \left(\frac{1}{2}\right)^1 2^1 + \left(\frac{1}{2}\right)^2 2^2 + \left(\frac{1}{2}\right)^3 2^3 + \dots = \sum_{i=0}^{\infty} 1 = \infty$$

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What if you could play this game for only \$1000...but just once?