# EE376A - Information Theory Final, Thursday March 22nd

# Instructions:

- You have three hours, 12:15PM 3:15PM
- The exam has 5 questions, totaling 100 points.
- Please start answering each question on a new page of the answer booklet.
- You are allowed to carry the textbook, your own notes and other course related material with you. Electronic reading devices [including kindles, laptops, ipads, etc.] are allowed, provided they are used solely for reading pdf files already stored on them and not for any other form of communication or information retrieval.
- Calculators are allowed for numerical computations.
- You are required to provide a sufficiently detailed explanation of how you arrived at your answers.
- You can use previous parts of a problem even if you did not solve them.
- As throughout the course, entropy (H) and Mutual Information (I) are specified in bits.
- log is taken in base 2.
- Good Luck!

#### 1. Universal Compression (20 points)

In this problem, we describe a lossless compression scheme that asymptotically (for large n) achieves entropy for any iid source. Let  $x^n$  be a particular sequence, where each symbol is in alphabet  $\mathcal{X} = \{1, 2, 3, \ldots, |\mathcal{X}|\}$ . Let  $P_{x^n}$  be the empirical distribution of the sequence  $x^n$ . Consider the compressor C for the sequence  $x^n$ :

- In the first step, the compressor encodes the empirical distribution  $P_{x^n}$  of the sequence, using a fixed-length code.
- In the second step, the compressor outputs the index of the sequence in the type class  $\mathcal{T}(P_{x^n})$ , using  $\lceil \log_2 |\mathcal{T}(P_{x^n})| \rceil$  bits.
- (a) Describe the operations of the decoder D, when a sequence  $x^n$  is compressed using the compressor C.
- (b) Let  $L(x^n)$  be number of bits required to encode a sequence  $x^n$  using the compressor. Show that:

$$L(x^n) \le |\mathcal{X}| \log_2(n+1) + nH(P_{x^n}) + 2$$

(c) Let the sequence  $X^n$  be generated i.i.d according to the distribution q(x). We define the rate of the compressor to be R:

$$R = \frac{\mathbb{E}[L(X^n)]}{n}$$

Show that for any distribution q(x), the rate R converges to H(q) as  $n \to \infty$ .

(d) Let  $f : \mathcal{X} \to \mathbb{R}$  be an arbitrary function, and let  $\bar{f}(x^n) = \frac{1}{n} \sum_{i=1}^n f(x_i)$ . Show that it is possible to compute  $\bar{f}(x^n)$  from the compressed sequence without decoding it completely. How many bits of the compressed sequence need to be read for computing  $\bar{f}(x^n)$ ?

#### Solution:

- (a) The decoder decodes the empirical distribution  $P_{x^n}$  from the first fixed-length code, and then using the index in the second part to find the sequence in  $\mathcal{T}(P_{x^n})$ .
- (b) The number of types is at most  $(n+1)^{|\mathcal{X}|}$ , thus the fixed-length code is of length at most  $\lceil \log_2(n+1)^{|\mathcal{X}|} \rceil \leq |\mathcal{X}| \log_2(n+1)+1$ . We also know from class that  $|\mathcal{T}(P_{x^n})| \leq 2^{nH(P_{x^n})}$ , and thus the code in the second step has length at most  $\lceil \log_2 |\mathcal{T}(P_{x^n})| \rceil \leq nH(P_{x^n}) + 1$ . Summing up gives the desired answer.
- (c) Note that H(P) is concave in P, we have

$$R = \frac{\mathbb{E}[L(X^n)]}{n} \le \frac{|\mathcal{X}| \log_2(n+1) + 2}{n} + \mathbb{E}H(P_{X^n})$$
$$\le \frac{|\mathcal{X}| \log_2(n+1) + 2}{n} + H(\mathbb{E}P_{X^n})$$
$$= \frac{|\mathcal{X}| \log_2(n+1) + 2}{n} + H(q) \xrightarrow{n \to \infty} H(q)$$

On the other hand,  $R \geq H(q)$  for any lossless code with source distribution q(x), so the rate converges to H(q).

(d) We only need to know the type of  $P_{x^n}$  to compute  $\bar{f}(x^n)$ . Hence, only  $|\mathcal{X}|\log_2(n+1)|$ 1) + 1 bits at the beginning of the compressed sequence need to be read.

#### 2. Rate-Distortion function for pairs of random variables (20 points)

Let X, Y be independent sources, with rate distortion functions  $R_X(D)$  and  $R_Y(D)$ , corresponding to distortion functions  $d_X : \mathcal{X} \times \hat{\mathcal{X}} \to \mathbb{R}^+$  and  $d_Y : \mathcal{Y} \times \hat{\mathcal{Y}} \to \mathbb{R}^+$  respectively.

We want to perform lossy compression on the product source (X, Y), where the distortion measure  $d_{X,Y}$  is given by:

$$d_{X,Y}((x,y),(x',y')) = d_X(x,x') + d_Y(y,y')$$

Let R(D) be the rate distortion function corresponding to the product source (X, Y) and the distortion  $d_{X,Y}$ .

(a) Show that if X, Y are independent, then for any  $\hat{X}, \hat{Y}$ :

$$I(X,Y;\hat{X},\hat{Y}) \ge I(X;\hat{X}) + I(Y;\hat{Y})$$

(b) Show the following lower bound on R(D):

$$R(D) \ge \min_{D_1+D_2 \le D} [R_X(D_1) + R_Y(D_2)]$$

(c) Show that the lower bound on R(D) is achievable, i.e.,

$$R(D) \le \min_{D_1+D_2 \le D} \left[ R_X(D_1) + R_Y(D_2) \right]$$

(d) Let X, Y be independent binary random variables, distributed as  $X \sim Ber(0.5)$  and  $Y \sim Ber(0.3)$ . Find the value of R(D) for the product source (X, Y), for D = 0.4where  $d_X$  and  $d_Y$  are Hamming distortions.

(you can leave the final answers in terms of binary entropy function)

(e) Let X, Y be independent Gaussian random variables distributed as  $X \sim \mathcal{N}(0,1)$ and  $Y \sim \mathcal{N}(0,4)$ . Find the value of R(D) for the product source (X,Y), for D=4and mean square distortion:

$$d_{X,Y}((x,y),(x',y')) = (x-x')^2 + (y-y')^2$$

How many bits/symbol are used to describe X?

#### Solution:

(a) The following chain of inequalities holds:

$$\begin{split} I(X,Y;\hat{X},\hat{Y}) &= H(X,Y) - H(X,Y|\hat{X},\hat{Y}) \\ &= H(X) + H(Y) - H(X|\hat{X},\hat{Y}) - H(Y|X,\hat{X},\hat{Y}) \\ &\geq H(X) + H(Y) - H(X|\hat{X}) - H(Y|\hat{Y}) \\ &= I(X;\hat{X}) + I(Y;\hat{Y}). \end{split}$$

(b) Due to the additive structure of  $d_{X,Y}$ , we have

$$\begin{split} R(D) &= R^{(I)}(D) = \min_{p(\hat{x}, \hat{y}|x, y) : \mathbb{E}d_{X,Y}((x, y), (\hat{x}, \hat{y})) \le D} I(X, Y; \hat{X}, \hat{Y}) \\ &\geq \min_{p(\hat{x}, \hat{y}|x, y) : \mathbb{E}d_{X,Y}((x, y), (\hat{x}, \hat{y})) \le D} I(X; \hat{X}) + I(Y; \hat{Y}) \\ &\geq \min_{D_1 + D_2 \le D} \left( \min_{p(\hat{x}, \hat{y}|x, y) : \mathbb{E}d_X(x, \hat{x}) \le D_1} I(X; \hat{X}) + \min_{p(\hat{x}, \hat{y}|x, y) : \mathbb{E}d_Y(y, \hat{y}) \le D_2} I(Y; \hat{Y}) \right) \\ &= \min_{D_1 + D_2 \le D} \left( \min_{p(\hat{x}|x) : \mathbb{E}d_X(x, \hat{x}) \le D_1} I(X; \hat{X}) + \min_{p(\hat{y}|y) : \mathbb{E}d_Y(y, \hat{y}) \le D_2} I(Y; \hat{Y}) \right) \\ &= \min_{D_1 + D_2 \le D} R_X^{(I)}(D_1) + R_Y^{(I)}(D_2) \\ &= \min_{D_1 + D_2 \le D} R_X(D_1) + R_Y(D_2). \end{split}$$

(c) For any  $D_1, D_2 \ge 0$  with  $D_1 + D_2 \le D$ , let  $p^*(\hat{x}|x), p^*(\hat{y}|y)$  be the minimum achieving distributions of  $R_X^{(I)}(D_1), R_Y^{(I)}(D_2)$ , respectively. Now consider  $p(\hat{x}, \hat{y}|x, y) = p^*(\hat{x}|x)p^*(\hat{y}|y)$ , then  $\mathbb{E}d_{X,Y}((X,Y), (\hat{X}, \hat{Y})) = \mathbb{E}d_X(X, \hat{X}) + \mathbb{E}d_Y(Y, \hat{Y}) \le D_1 + D_2 \le D$ . Moreover,  $(X, \hat{X})$  is independent of  $(Y, \hat{Y})$ , and thus

$$R(D) = R^{(I)}(D) \le I(X, Y; \hat{X}, \hat{Y}) = I(X; \hat{X}) + I(Y; \hat{Y})$$
  
$$\le R_X^{(I)}(D_1) + R_Y^{(I)}(D_2) = R_X(D_1) + R_Y(D_2).$$

This inequality holds for any  $D_1 + D_2 \leq D$ , and the result follows. (d) By (b) and (c), we have

$$R(0.4) = \min_{D_1+D_2 \le 0.4} R_X(D_1) + R_Y(D_2)$$
  
=  $\min_{D_1+D_2 \le 0.4} H(0.5) - H(\min\{D_1, 0.5\}) + H(0.3) - H(\min\{D_2, 0.3\})$   
$$\geq \min_{D_1+D_2 \le 0.4} H(0.5) + H(0.3) - 2H(\frac{\min\{D_1, 0.5\} + \min\{D_2, 0.3\}}{2})$$
  
$$\geq \min_{D_1+D_2 \le 0.4} H(0.5) + H(0.3) - 2H(\frac{D_1 + D_2}{2})$$
  
$$\geq 1 + H(0.3) - 2H(0.2)$$

where we have used the fact that H(p) is increasing on  $p \in [0, \frac{1}{2}]$  and concave. The minimum is attained at  $D_1 = D_2 = 0.2$ .

(e) By (b) and (c), we have

$$R(4) = \min_{D_1 + D_2 \le 4} R_X(D_1) + R_Y(D_2) = \min_{D_1 + D_2 \le 4} \frac{1}{2} \log \frac{1}{\min\{D_1, 1\}} + \frac{1}{2} \log \frac{4}{\min\{D_2, 4\}}.$$

If  $D_1 \leq 1$ , by the convexity of  $x \mapsto -\log x$  we know that the minimum is achieved at  $D_1 = 1, D_2 = 3$ . If  $D_1 > 1$ , we have  $D_2 < 3$  and  $\log \frac{4}{D_2} > \log \frac{4}{3}$ . Hence,  $(D_1^*, D_2^*) = (1, 3)$ , and  $R(4) = \frac{1}{2} \log \frac{4}{3}$ . Note that  $R_X(D_1^*) = 0$  in this case, no bit is used to describe  $X_1$ .

## **3.** Compression with some help (25 points)

Consider the lossless source coding problem in Figure 1. The pair  $(X^n, Y^n)$  is generated by i.i.d. drawings of the finite alphabet pair (X, Y), that is  $p(x^n, y^n) = \prod_{i=1}^n p_{XY}(x_i, y_i)$ . We wish to transmit the source sequence  $X^n$  near-losslessly when  $Y^n$  is available at both the encoder and the decoder. Formally, a  $(2^{nR}, n)$  code is defined by an encoder  $m(x^n, y^n) \in \{1, 2, \ldots, 2^{nR}\}$  and a decoder  $\hat{X}^n(m, y^n)$ , and the probability of decoding error is defined as  $P_e = P\{\hat{X}^n \neq X^n\}$ , where  $\hat{X}^n = \hat{X}^n(m(X^n, Y^n), Y^n)$ . A rate R is achievable if there exists a sequence of codes with  $P_e \to 0$  as  $n \to \infty$ .

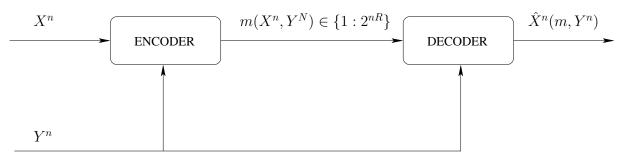


Figure 1: Conditional Lossless Source Coding

- (a) Prove that any rate R > H(X|Y) is achievable. [*Hint*: If  $y^n \in T^{(n)}_{\delta'}(Y)$  and  $x^n \in T^{(n)}_{\delta}(X|y^n)$  for appropriate  $\delta' < \delta$ , transmit the index of  $x^n$  in  $T^{(n)}_{\delta}(X|y^n)$ .]
- (b) Prove that any rate R < H(X|Y) is not achievable via the following steps:
  - i. For  $M = m(X^n, Y^n)$  argue why

$$I(X^n; M|Y^n) \le nR.$$

ii. Use the previous step and a relation that you know between conditional entropy and probability of error to deduce that if R < H(X|Y) then one cannot get  $P_e \to 0$  as  $n \to \infty$ .

Now we consider a simple instance of this problem and develop concrete schemes for achieving the optimal rate. Let X be a random variable uniformly distributed on  $\{0, 1\}^3$ , i.e., X is a sequence of 3 independent unbiased bits. Let  $Y = X \oplus Z$ , where Z is independent of X and is uniformly distributed on  $\{(0,0,0), (0,0,1), (0,1,0), (1,0,0)\}$  (set of binary triplets with at most one 1).

- (c) Give a scheme to losslessly compress X into 2 bits when Y is known at both the encoder and the decoder. Specifically, you should describe the encoder  $m(x, y) \in \{1, 2, 3, 4\}$  and a decoder  $\hat{X}(m, y)$  which satisfy  $\hat{X}(m(X, Y), Y) = X$ . Is this optimal?
- (d) Now, if only the decoder has access to Y, show that random variable X can still be losslessly compressed using 2 bits.

[*Hint*: Partition  $\mathcal{X}$  into 4 suitable subsets, and transmit the index of the subset.]

(e) In part (d), can we do better (with less) than 2 bits?

# Solution:

- (a) Fix any  $\delta > \delta' > 0$ . By strong AEP, with probability tending to 1, we have  $y^n \in T^{(n)}_{\delta'}(Y)$  and  $x^n \in T^{(n)}_{\delta}(X|y^n)$ . We consider the encoding/decoding scheme as follows:
  - Encoding: the compressor sends the index of the sequence  $x^n$  in  $T^{(n)}_{\delta}(X|y^n)$  if conditional typicality holds; otherwise, just send 1;
  - Decoding: find the sequence  $x^n$  in  $T^{(n)}_{\delta}(X|y^n)$  with the received index.

Note that this scheme has error probability tending to zero. Moreover,  $|T_{\delta}^{(n)}(X|y^n)| \leq 2^{n(1+\delta)H(X|Y)}$ , therefore the rate is at most  $R \leq (1+\delta)H(X|Y)$ . Since  $\delta > 0$  is arbitrary, any rate R > H(X|Y) is achievable.

(b) i. Note that  $H(M) \leq nR$  since  $M \in \{1, 2, \dots, 2^{nR}\}$ , we have

$$I(X^{n}; M|Y^{n}) = H(M|Y^{n}) - H(M|X^{n}, Y^{n}) = H(M|Y^{n}) \le H(M) \le nR.$$

ii. Let  $p_e = \mathbb{P}(\hat{X}^n \neq X^n)$ , Fano's inequality gives

$$I(X^{n}; M|Y^{n}) = H(X^{n}|Y^{n}) - H(X^{n}|M, Y^{n})$$
  

$$\geq H(X^{n}|Y^{n}) - H(X^{n}|\hat{X}^{n})$$
  

$$\geq nH(X|Y) - H(p_{e}) - np_{e} \log |\mathcal{X}|.$$

Combining with the previous question, we see that

$$R \ge H(X|Y) - \frac{H(p_e)}{n} - p_e \log |\mathcal{X}|$$

i.e., any R < H(X|Y) is impossible given  $p_e \to 0$ .

(c) Since the alphabet of Z has size  $|\mathcal{Z}| = 4$ , there exists a bijection f between  $\mathcal{Z}$  and  $\{1, 2, 3, 4\}$ . Define encoder  $m(x, y) = f(x \oplus y)$  and decoder  $\hat{X}(m, y) = f^{-1}(m) \oplus y$ . This definition is feasible since  $X \oplus Y = Z \in \mathcal{Z}$ . Clearly  $\hat{X}(m(x, y), y) = f^{-1}(f(x \oplus y)) \oplus y = x$ , and the rate is  $\log |\mathcal{Z}| = 2$ . This is not improvable, for

$$H(X|Y) = H(X) + H(Y|X) - H(Y) = H(X) + H(Z) - H(X \oplus Z) = 2.$$

- (d) Split  $\{0, 1\}^3$  into four groups:  $G_1 = \{(0, 0, 0), (1, 1, 1)\}, G_2 = \{(1, 0, 0), (0, 1, 1)\}, G_3 = \{(0, 1, 0), (1, 0, 1)\}, G_4 = \{(0, 0, 1), (1, 1, 0)\}$ . Upon receiving X, the encoder encodes the index of the group which X lies in. The decoder determines  $\hat{X}$  to be the closest symbol to the side information Y (in Hamming distance) in the given group. Clearly the rate is 2, and this is lossless because the symbols in each group have minimum distance 3 and can thus correct 1-bit error caused by Z.
- (e) No, because 2 bits are optimal even in the setting of (c), where the encoder also has the extra side information Y.

## 4. Channel Capacity (15 points)

Find the capacities of the following channels with the given channel transition matrices p(y|x). Also, give the capacity-achieving input distribution p(x). Justify your answers. (you can leave the final answers in terms of the binary entropy function)

(a)  $\mathcal{X} = \mathcal{Y} = \{0, 1, 2\}$   $p(y|x) = \begin{bmatrix} 0 & 1/3 & 2/3 \\ 2/3 & 0 & 1/3 \\ 1/3 & 2/3 & 0 \end{bmatrix}$ (b)  $\mathcal{X} = \mathcal{Y} = \{0, 1, 2\}$   $p(y|x) = \begin{bmatrix} 0 & 1/3 & 2/3 \\ 2/3 & 0 & 1/3 \\ 0 & 2/3 & 1/3 \end{bmatrix}$ (c)  $\mathcal{X} = \{0, 1\}, \mathcal{Y} = \{0, 1, 2\}$  $p(y|x) = \begin{bmatrix} 0 & 2/3 & 1/3 \\ 1/3 & 2/3 & 0 \end{bmatrix}$ 

## Solution:

(a) For any input distribution p(x), we have

$$I(X;Y) = H(Y) - H(Y|X) = H(Y) - H(\frac{1}{3}) \le \log 3 - H(\frac{1}{3})$$

with equality iff Y is uniformly distributed on  $\mathcal{Y}$ . Therefore, the capacity-achieving input distribution is  $p(x) = (\frac{1}{3}, \frac{1}{3}, \frac{1}{3})$ .

- (b) We can show that  $I(X;Y) \leq \log 3 H(\frac{1}{3})$  as in (a), with equality iff Y is uniformly distributed on  $\mathcal{Y}$ . This gives the capacity-achieving distribution  $p(x) = (0, \frac{1}{2}, \frac{1}{2})$ .
- (c) For input distribution (p, 1-p), we have  $Y \sim (\frac{1-p}{3}, \frac{2}{3}, \frac{p}{3})$ , and

$$\begin{split} I(X;Y) &= H(Y) - H(Y|X) = -\frac{1-p}{3}\log\frac{1-p}{3} - \frac{p}{3}\log\frac{p}{3} - \frac{2}{3}\log\frac{2}{3} - H(\frac{1}{3})\\ &\leq 2 \cdot \frac{\log 6}{6} - \frac{2}{3}\log\frac{2}{3} - H(\frac{1}{3}) = \frac{1}{3} \end{split}$$

where the inequality follows from the concavity of  $x \mapsto -x \log x$ . As a result, the capacity-achieving input distribution is  $p(x) = (\frac{1}{2}, \frac{1}{2})$ .

The capacity can also be computed by observing that the channel is a special case of BEC channel (erasure probability 2/3).

# 5. Information Theory and Statistics (20 points)

This problem illustrates an application of information-theoretic tools in statistics. Suppose we observe a sample  $X \sim \mathcal{N}(\theta, I_d)$ , where  $\theta \in \mathbb{R}^d$  is an unknown mean vector, and  $I_d$  denotes the  $d \times d$  identity matrix. An estimator  $\hat{\theta} = \hat{\theta}(X)$  is a function of X, and we want to find an estimator  $\hat{\theta}$  which is close to the true  $\theta$ . We consider the mean squared error  $l(\theta) = \mathbb{E}_{\theta} \| \hat{\theta}(X) - \theta \|_2^2$ , where the expectation is taken with respect to  $X \sim \mathcal{N}(\theta, I_d)$ .

(a) A natural estimator is  $\hat{\theta}(X) = X$ . What is  $l(\theta)$  in this case? What is the worst-case  $l(\theta)$  when  $\theta$  can be any value in  $\mathbb{R}^d$ ?

In the following, we show that this natural estimator is in fact a *minimax* estimator for estimating  $\theta$  under mean squared error. By minimax we mean that it achieves the minimum worst-case error possible for any estimator. For this we'll use ideas from channel capacity and rate-distortion. First, we state some results for multivariate Gaussian distributions. These can be derived using similar techniques as those used for univariate Gaussian.

• Capacity of multivariate AWGN channel: Consider a channel from  $\theta$  to X defined as  $X = \theta + Z$  where  $Z \sim \mathcal{N}(0, I_d)$  with power constraint  $\mathbb{E} \|\theta\|_2^2 \leq d\sigma^2$ . For this channel,

$$C = \frac{d}{2}\log(1+\sigma^2) \tag{1}$$

• Rate-distortion function for multivariate Gaussian source: Consider a source  $\theta \sim \mathcal{N}(0, \sigma^2 I_d)$  and distortion metric  $d(\theta, \hat{\theta}) = \mathbb{E} \|\theta - \hat{\theta}\|_2^2$ . For this setting,

$$R(D) = \frac{d}{2}\log\frac{d\sigma^2}{D}$$
(2)

(b) Assume that there exists an estimator  $\hat{\theta}$  with  $l(\theta) \leq D$  for any  $\theta \in \mathbb{R}^d$ . Argue why that implies that we must have  $R(D) \leq C$ , where C and R(D) are as defined in equations (1) and (2), respectively.

[*Hint:* Frame this as a joint source-channel coding problem with appropriate source and channel.]

- (c) Conclude from (b) that  $D \geq \frac{d\sigma^2}{1+\sigma^2}$ . Since that argument holds for any value of  $\sigma^2$ , further conclude that  $D \geq d$ .
- (d) Argue how your results in (b) and (c) imply that the estimator in (a) is a minimax estimator. Specifically, argue why no other estimator can achieve worst-case risk lower than that achieved by  $\hat{\theta}(X) = X$ .

#### Solution:

- (a) We have  $X_i \sim \mathcal{N}(\theta, 1)$  for each  $i = 1, 2, \cdots, d$ . Hence,  $l(\theta) = \sum_{i=1}^d \mathbb{E}_{\theta} (X_i \theta)^2 = d$ . Since  $l(\theta) = d$  for any  $\theta$ , so is the worst-case risk.
- (b) Consider the joint source-channel coding problem with source  $\theta \sim \mathcal{N}(0, \sigma^2 I_d)$  and channel  $x|\theta \sim \mathcal{N}(\theta, I_d)$ . The overall rate is 1, so  $R(D) \leq C$  follows from the joint source-channel coding theorem. Alternatively, we can also write

$$R(D) = \min_{p(\hat{\theta}|\theta): \mathbb{E}\|\hat{\theta}-\theta\|_2^2 \le D} I(\theta; \hat{\theta}) \le I(\theta; \hat{\theta}) \le I(\theta; X) \le \max_{p(\theta): \mathbb{E}\|\theta\|_2^2 \le d\sigma^2} I(\theta; X) = C$$

for  $\theta - X - \hat{\theta}$  forms a Markov chain.

(c) By (b) we have  $\frac{d}{2}\log\frac{d\sigma^2}{D} \leq \frac{d}{2}\log(1+\sigma^2)$ , which gives  $D \geq \frac{d\sigma^2}{1+\sigma^2}$ . This inequality holds for any  $\sigma^2$ , we choose  $\sigma^2 \to \infty$  to conclude that  $D \geq d$ .

(d) Part (c) shows that the worst-case risk for any estimator must be no smaller than D. Since the natural estimator  $\hat{\theta}(X) = X$  achieves the worst-case risk D, we conclude that this estimator is minimax.