# EE378B Homework 7 Solution

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## 1 Part (a)

Direct application of Wielandt-Hoffman inequality gives

$$\sum_{i=1}^{n} \left| \lambda_i \left( \boldsymbol{W}_n / \sqrt{n} \right) - \lambda_i \left( \boldsymbol{W}_n^K / \sqrt{n} \right) \right|^2 \le \frac{1}{n} \left\| \boldsymbol{W}_n - \boldsymbol{W}_n^K \right\|_F^2 = \frac{1}{n} \sum_{i \ne j} \xi_{ij}, \tag{1}$$

where  $\xi_{ij}$  are i.i.d. random variables having the same distribution as  $W_{12}^2 \mathbf{1}_{|W_{12}|>K}$ . Therefore by SLLN with probability 1,

$$\limsup_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} \left| \lambda_i \left( \mathbf{W}_n / \sqrt{n} \right) - \lambda_i \left( \mathbf{W}_n^K / \sqrt{n} \right) \right|^2 \leq \limsup_{n \to \infty} \frac{1}{n^2} \sum_{i \neq j} \xi_{ij}$$
$$= \lim_{n \to \infty} \frac{n-1}{n} \cdot \frac{1}{n(n-1)} \sum_{i \neq j} \xi_{ij}$$
$$= \mathbb{E} \left[ W_{12}^2 \mathbf{1}_{|W_{12}| > K} \right]. \tag{2}$$

### 2 Part (b)

For simplicity we write

$$\frac{1}{n}\sum_{i=1}^{n}\left|\lambda_{i}\left(\boldsymbol{W}_{n}/\sqrt{n}\right)-\lambda_{i}\left(\boldsymbol{W}_{n}^{K}/\sqrt{n}\right)\right|^{2}=\Delta_{n,K}.$$
(3)

As a consequence of part (a) we know  $\limsup_{n\to\infty} \Delta_{n,K} \leq \Delta_K$ . By Markov's inequality

$$\frac{1}{n} \# \left\{ i : \left| \lambda_i \left( \boldsymbol{W}_n / \sqrt{n} \right) - \lambda_i \left( \boldsymbol{W}_n^K / \sqrt{n} \right) \right| \ge \epsilon \right\} \le \frac{\Delta_{n,K}}{\epsilon^2}, \tag{4}$$

which allows us to derive that

$$F_n^K(x-\epsilon) - F_n(x) = \frac{1}{n} \sum_{i=1}^n \left( \mathbbm{1}_{\lambda_i(\mathbf{W}_n^K/\sqrt{n}) \le x-\epsilon} - \mathbbm{1}_{\lambda_i(\mathbf{W}_n/\sqrt{n}) \le x} \right)$$
  
$$\leq \frac{1}{n} \sum_{i=1}^n \mathbbm{1}_{\lambda_i(\mathbf{W}_n^K/\sqrt{n}) \le x-\epsilon, \lambda_i(\mathbf{W}_n/\sqrt{n}) > x}$$
  
$$\leq \frac{1}{n} \# \left\{ i : \left| \lambda_i \left( \mathbf{W}_n/\sqrt{n} \right) - \lambda_i \left( \mathbf{W}_n^K/\sqrt{n} \right) \right| \right\}$$
  
$$\leq \frac{\Delta_{n,K}}{\epsilon^2}.$$
(5)

Similarly we bound  $F_n(x) - F_n^K(x + \epsilon) \leq \frac{\Delta_{n,K}}{\epsilon^2}$ . For the above two inequalities, take lim sup on both sides and rearrange terms, which gives the desired result

$$F^{K}(x-\epsilon) - \frac{\Delta_{K}}{\epsilon^{2}} \le \liminf_{n \to \infty} F_{n}(x) \le \limsup_{n \to \infty} F_{n}(x) \le F^{K}(x+\epsilon) + \frac{\Delta_{K}}{\epsilon^{2}}.$$
(6)

Finally, if  $F^K(x) \to F(x)$  for every  $x \in \mathbb{R}$  as  $K \to \infty$ , using the fact that  $\Delta_K \to 0$  since  $W_{12}$  has finite second moment, we obtain by taking  $K \to \infty$  that

$$F(x-\epsilon) \le \liminf_{n \to \infty} F_n(x) \le \limsup_{n \to \infty} F_n(x) \le F(x+\epsilon).$$
(7)

Whenever F is continuous at x, by taking  $\epsilon \to 0$  it is concluded that  $\lim_{n\to\infty} F_n(x) = F(x)$ .

#### Part (c) 3

In fact, it suffices to show that for all K with  $\beta_K = \operatorname{Var}\left(W_{12}\mathbb{1}_{|W_{12}| \leq K}\right)$ ,

$$F_n^K(x) \to F^K(x) := \frac{1}{2\beta_K \pi} \int_{-\infty}^x \sqrt{4\beta_K - t^2} \mathbb{1}_{|t| \le 2\sqrt{\beta_K}} \mathrm{d}t, \tag{8}$$

i.e., the semi-circular law with variance  $\beta_K$ . Then all conditions hold since  $\beta_K \to 1$  as  $K \to \infty$  and  $F^K(x) \to F(x)$  pointwise with F(x) being the semi-circular law CDF with variance 1. Let  $\tilde{\boldsymbol{W}}_n^K = \boldsymbol{W}_n^K + \alpha_K \boldsymbol{1} \boldsymbol{1}^\top$  where  $\alpha_K = -\mathbb{E}[W_{12}\mathbb{1}_{|W_{12}| \leq K}]$ . Then  $\mathbb{E}\tilde{\boldsymbol{W}}_n^K = \boldsymbol{0}$  and we can apply results that have been shown in class that

$$\tilde{F}_n^K(x) \to F^K(x),\tag{9}$$

where  $\tilde{F}_n^K(x)$  is the empirical CDF of eigenvalues of  $\tilde{W}_n^K/\sqrt{n}$ . The proof will be concluded once we can establish

$$\left|\tilde{F}_{n}^{K}(x) - F_{n}^{K}(x)\right| \to 0 \tag{10}$$

for any fixed K and  $x \in \mathbb{R}$ . By the generalized Weyl's inequality, we're able to get

$$\lambda_i(\tilde{\boldsymbol{W}}_n^K/\sqrt{n}) \ge \lambda_{i+1}(\boldsymbol{W}_n^K/\sqrt{n}) + \lambda_{n-1}(\alpha_K \mathbf{1} \mathbf{1}^\top/\sqrt{n}), \qquad i = 1, 2, \cdots, n-1;$$
(11)

$$\lambda_i(\tilde{\boldsymbol{W}}_n^K/\sqrt{n}) \le \lambda_{i-1}(\boldsymbol{W}_n^K/\sqrt{n}) + \lambda_2(\alpha_K \mathbf{1}\mathbf{1}^\top/\sqrt{n}), \qquad i = 2, 3, \cdots, n.$$
(12)

While  $\alpha_K \mathbf{1} \mathbf{1}^\top / \sqrt{n}$  is only a rank-1 matrix, it follows when  $n \geq 3$  that

$$\lambda_2(\alpha_K \mathbf{1}\mathbf{1}^\top/\sqrt{n}) = \lambda_{n-1}(\alpha_K \mathbf{1}\mathbf{1}^\top/\sqrt{n}) = 0,$$
(13)

which gives us

$$\lambda_i(\tilde{\boldsymbol{W}}_n^K/\sqrt{n}) \ge \lambda_{i+1}(\boldsymbol{W}_n^K/\sqrt{n}), \qquad i = 1, 2, \cdots, n-1;$$
(14)

$$\lambda_i(\tilde{\boldsymbol{W}}_n^K/\sqrt{n}) \le \lambda_{i-1}(\boldsymbol{W}_n^K/\sqrt{n}), \qquad i = 2, 3, \cdots, n.$$
(15)

By the first bound, one can deduce that

$$\tilde{F}_{n}^{K}(x) - F_{n}^{K}(x) = \frac{1}{n} \sum_{i=1}^{n} \left( \mathbb{1}_{\lambda_{i}(\tilde{\mathbf{W}}_{n}^{K}/\sqrt{n}) \leq x} - \mathbb{1}_{\lambda_{i}(\mathbf{W}_{n}^{K}/\sqrt{n}) \leq x} \right) \\ \leq \frac{1}{n} \sum_{i=1}^{n-1} \left( \mathbb{1}_{\lambda_{i}(\tilde{\mathbf{W}}_{n}^{K}/\sqrt{n}) \leq x} - \mathbb{1}_{\lambda_{i+1}(\mathbf{W}_{n}^{K}/\sqrt{n}) \leq x} \right) + \frac{1}{n} \\ \leq \frac{1}{n}.$$
(16)

The other side is similar, and therefore  $\left|\tilde{F}_n^K(x) - F_n^K(x)\right| \le 1/n \to 0$ . The proof is done.

#### 4 Part (d)

We can generate uniformly distributed (in the sense of Haar's measure) orthogonal matrices  $U \in \mathbb{R}^{n \times k}$  by first generate an n by k matrix with i.i.d. standard Gaussian entries and perform Gram-Schmidt orthogonalization. Since the distribution of Gaussian matrix is invariant under orthogonal transformation, the distribution of resulting orthogonal matrix U is also invariant, and thus is uniformly distributed under Haar's measure.

(i) The plots are shown in Figure 1.



Figure 1: Histograms of eigenvalues of Y when n = 500 (left top), n = 1000 (right top), n = 2000 (left down), n = 4000 (right down).

(ii) The result is different from Winger's semicircle's law whereas the histogram should look like a semicircle.

(iii) We can't apply since the entries  $Y_{ij}$  are correlated.

Remark from class: What happens for a larger choice of constant? Does it recover the semicircle law?