

Stat 209
Week 9

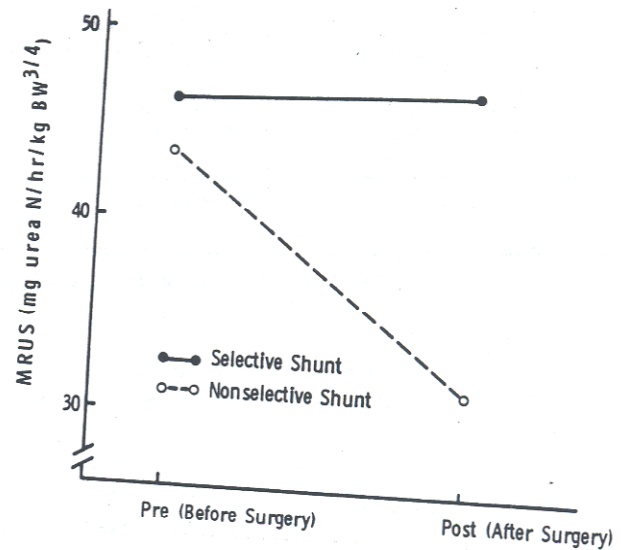
Pre-post, 2 groups Repeated Measures ANOVA

Brogan -
Kutner
example

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1. Pre and Post Maximal Rate of Urea Synthesis Level (mg urea N/hr/kg BW^{3/4}) and Sample Cell Means, by Group

Group	Subject	Pre	Post
Selective Shunt (new operation)	1	51	48
	2	35	55
	3	66	60
	4	40	35
	5	39	36
	6	46	43
	7	52	46
	8	42	54
Mean		$\hat{\mu}_{11} = 46.375$	$\hat{\mu}_{12} = 47.125$
Nonselective Shunt (standard operation)	9	34	16
	10	40	36
	11	34	16
	12	36	18
	13	38	32
	14	32	14
	15	44	20
	16	50	43
	17	60	45
	18	63	67
	19	50	36
	20	42	34
	21	43	32
Mean		$\hat{\mu}_{21} = 43.538$	$\hat{\mu}_{22} = 31.462$



analysis on back

Bock, DR MSMBR text

EXAMPLE 7.1-1 (Mixed-model analysis of vocabulary growth) Data for this example are drawn from test results on file in the Records Office of the Laboratory School of the University of Chicago. They consist of scores, obtained from a cohort of pupils at the eighth through eleventh grade level, on alternative forms of the vocabulary section of the Cooperative Reading Tests [Davis, 1950]. Since these data cover an age range in which physical growth is beginning to decelerate, it is of interest to inquire whether a similar deceleration can be observed in the acquisition of new vocabulary.

Growth Curves (Group) T=4

Table 7.2-5 MIXED-MODEL ANALYSIS OF VARIANCE OF SEX EFFECTS IN THE VOCABULARY-SCALED SCORES

Source	df	ss	F	p
Constant	1	ssm = 1,644.90		
Sex	1	ssb = .85	.06	> .5
Occasions	3	ssc = 194.18		
Sex x occasions	3	ssbc = 2.79	1.12	> .1
Subjects within groups	62	ssa = 873.00		
Occasions x subjects within groups	186	sse = 152.17		
Total	256	sst = 2,867.90		

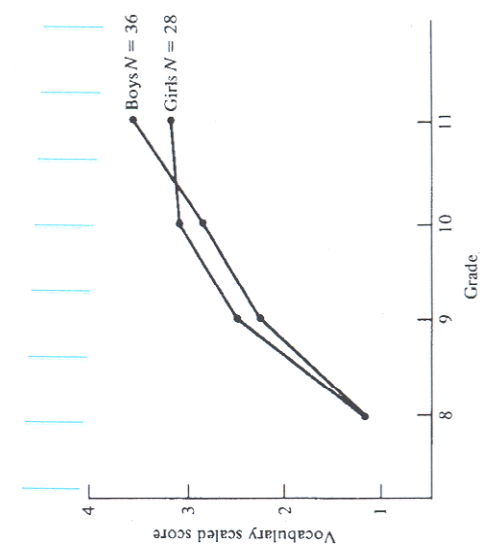


FIGURE 7.2-1 Average vocabulary scores of boys and girls in a cohort from the University of Chicago Laboratory School (longitudinal data).

Repeated Measures Brogan-Kutner ix p. 2

D. Rogosa

model

$$X_{ijk} = \mu + \alpha_i + \Pi_{k(i)} + \beta_j + \alpha\beta_{ij} + \beta\Pi_{jk(i)} + \epsilon_{m(ijk)} \quad (3.1)$$

$j = 1, 2$ (pretest = 1, posttest = 2),

$i = 1, 2$ (group 1 = 1, group 2 = 2),

$k = 1, 2, \dots, n_i, m = 1,$

where X_{ijk} is the observed value of subject k within group i at time j ,

μ is the overall mean,

α_i is the effect of group i ,

$\Pi_{k(i)}$ is the effect of subject k nested within group i ,

β_j is the effect of the repeated-measures variable j (i.e., pretest and posttest),

$\alpha\beta_{ij}$ is the interaction of group i with level j of the repeated measures factor,

$\beta\Pi_{jk(i)}$ is the interaction of subject k within group i with level j of the repeated-measures factor.

2. Repeated-Measures Analysis of Variance for Maximal Rate of Urea Synthesis Level

Source of Variation	df	Sum of Squares	Mean Squares	F Ratio
Between Subjects	20 (n - 1)			
Groups	1	847.48	847.48 (MS _G)	3.63 (MS _G /MS _E)
Subjects Within Groups	19 (n - 2)	4440.00	233.68 (MS _E)	
Within Groups	21 (n)			
Pre/Post	1	317.69	317.69 (MS _P)	8.86 (MS _P /MS _{PE})
Groups x Pre/Post	1	407.41	407.4 (MS _{GP})	11.36 (MS _{GP} /MS _{PE})
(Pre/Post) x Subjects	19 (n - 2)	681.21	35.85 (MS _{PE})	
Within Groups				

Did the groups change differentially?

SAS or minitab does it

(R has problem w/ imbalance anova TBD /error

```
proc glm data=brogk;
class grp;
model m1--m2 = grp /nouni;
repeated Time 2 (1 2) / summary printe;
run;
```

OUTPUT (selected)

The SAS System 16:13 Tuesday, May 16, 2000 35

The GLM Procedure
Repeated Measures Analysis of Variance
Tests of Hypotheses for Between Subjects Effects

Source	DF	Type III SS	Mean Square	F Value	Pr > F
grp	1	847.476190	847.476190	3.63	0.0721
Error	19	4440.000000	233.684211		

The GLM Procedure
Repeated Measures Analysis of Variance
Univariate Tests of Hypotheses for Within Subject Effects

Source	DF	Type III SS	Mean Square	F Value	Pr > F
Time	1	317.6932234	317.6932234	8.86	0.0078
Time*grp	1	407.4075092	407.4075092	11.36	0.0032
Error(Time)	19	681.2115385	35.8532389		

Repeated Measures Anova - more BK

Week 9
Stat 209

Brogan-Kutner Data see <http://www-stat.stanford.edu/~rag/ed351longit/brogkut.dat>

```
# Cell means
> tapply(urea, list(method, prepost), mean)
      1      2
1 46.37500 47.12500
2 43.53846 31.46154
```

cf main BK
handout

data observations
as rows (42)

```
# Recreate repeated measures anova (nesting)
# within-groups anova to obtain the 2 error terms
```

Do repeated measures anova
by crossed designs on subsets.

```
#within group 1 subjXtime
> bkrepaovW1 = aov(urea[method == "1"] ~ as.factor(prepost[method == "1"])*as.factor(subj[method == "1"]))
> summary(bkrepaovW1)
```

	Df	Sum Sq	Mean Sq	Sq
as.factor(prepost[method == "1"])	1	2.25	2.25	
as.factor(subj[method == "1"])	7	915.00	130.71	
piece of subjects within groups				Between subjects error term
as.factor(prepost[method == "1"]):as.factor(subj[method == "1"])	7	331.75	47.39	
piece of subjectsxrepeated measure				Within subjects error term

```
#within group 2 subjXtime
> bkrepaovW2 = aov(urea[method == "2"] ~ as.factor(prepost[method == "2"])*as.factor(subj[method == "2"]))
> summary(bkrepaovW2)
```

	Df	Sum Sq	Mean Sq	Sq
as.factor(prepost[method == "2"])	1	948.0	948.0	
as.factor(subj[method == "2"])	12	3525.0	293.7	
piece of subjects within groups				Between subjects error term
as.factor(prepost[method == "2"]):as.factor(subj[method == "2"])	12	349.5	29.1	
piece of subjectsxrepeated measure				Within subjects error term

```
#
# 915 + 3525 = 4440 (and 7 + 12 = 19df) Between subjects SS error term
# 331.7 + 349.5 = 681.2 (and 7 + 12 = 19df) Within subjects SS error term
```

```
# ignore within-subjects, get
> bkrepaovBase = aov(urea ~ as.factor(prepost)*as.factor(method))
> summary(bkrepaovBase)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
as.factor(prepost)	1	542.9	542.9	4.0282	0.05190	#repeated measure (Within subj part)
as.factor(method)	1	847.5	847.5	6.2884	0.01654	* #Group (Between subjects part)
as.factor(prepost):as.factor(method)	1	407.4	407.4	3.0230	0.09019	#GroupxRepeated measure Interaction
Residuals	38	5121.2	134.8			(Within subjects part)

```
# Brogan-Kutner Section 5 Equivalences
# Groups, pooling over occasion
```

```
> sumtime = pre + post
> t.test(sumtime ~ as.factor(method), var.equal = TRUE)
Two Sample t-test data: sumtime by as.factor(method)
t = 1.9044, df = 19, p-value = 0.07212
95 percent confidence interval: -1.832786 38.832786
mean in group 1 mean in group 2
 93.5      75.0
> 1.904^2 [1] 3.625216 # matches F-stat for Groups (bet subj)
```

subj as rows format

```
> bksubj
pre post method
1 51 48 1
2 35 55 1
3 66 60 1
4 40 35 1
5 39 36 1
6 46 43 1
7 52 46 1
8 42 54 1
9 34 16 2
10 40 36 2
11 34 16 2
12 36 18 2
13 38 32 2
14 32 14 2
15 44 20 2
16 50 43 2
17 60 45 2
18 63 67 2
19 50 36 2
20 42 34 2
21 43 32 2
```

```
> imp = post - pre
> t.test(imp ~ as.factor(method), var.equal = TRUE)
Two Sample t-test data: imp by as.factor(method)
t = 3.3709, df = 19, p-value = 0.003209
95 percent confidence interval: 4.862645 20.791201
mean in group 1 mean in group 2
 0.75000      -12.07692
> 3.3709^2 [1] 11.36297 # matches F-stat for Groups X prepost
BK p.232
```

main event
differential
change
by
t-test

```
> bkrepaov1 = aov(urea ~ as.factor(prepost)*as.factor(method) + Error(as.factor(subj)))
> summary(bkrepaov1)
Error: as.factor(subj)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(method)	1	847.5	847.5	3.6266	0.07212
Residuals	19	4440.0	233.7		

Error: Within

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(prepost)	1	542.88	542.88	15.142	0.0009823 ***
as.factor(prepost):as.factor(method)	1	407.41	407.41	11.363	0.0032085 **
Residuals	19	681.21	35.85		

subj x prepost x method

R does the
repeated meas
design
See Baron + Li

Sequential SS issue
w/ prepost SS

```
# Brogan-Kutner Data see http://www-stat.stanford.edu/~rag/ed351longit/brogkut.dat
```

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# Cell means
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piece of subjects within groups				Between subjects error term
as.factor(prepost[method == "1"]):as.factor(subj[method == "1"])	7	331.75	47.39	
piece of subjectsxrepeated measure within group interaction				Within subjects error term

```
#within group 2 subjXtime
> bkrepaovW2 = aov(urea[method == "2"] ~ as.factor(prepost[method == "2"])*as.factor(subj[method == "2"]))
> summary(bkrepaovW2)
```

	Df	Sum Sq	Mean Sq	
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as.factor(subj[method == "2"])	12	3525.0	293.7	
piece of subjects within groups				Between subjects error term
as.factor(prepost[method == "2"]):as.factor(subj[method == "2"])	12	349.5	29.1	
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```
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```

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as.factor(method)	1	847.5	847.5	6.2884	0.01654	* #Group (Between subjects part)
as.factor(prepost):as.factor(method)	1	407.4	407.4	3.0230	0.09019	. #GroupxRepeated measure Interaction
Residuals	38	5121.2	134.8			(Within subjects part)

```
# Brogan-Kutner Section 5 Equivalences
```

```
# Groups, pooling over occasion
```

```
> sumtime = pre + post
> t.test(sumtime ~ as.factor(method), var.equal = TRUE)
Two Sample t-test data: sumtime by as.factor(method)
t = 1.9044, df = 19, p-value = 0.07212
95 percent confidence interval: -1.832786 38.832786
mean in group 1 mean in group 2
      93.5      75.0
```

```
> 1.904^2 [1] 3.625216 # matches F-stat for Groups (bet subj)
```

```
> imp = post - pre
> t.test(imp ~ as.factor(method), var.equal = TRUE)
Two Sample t-test data: imp by as.factor(method)
t = 3.3709, df = 19, p-value = 0.003209
95 percent confidence interval: 4.862645 20.791201
mean in group 1 mean in group 2
      0.75000      -12.07692
```

```
> 3.3709^2 [1] 11.36297 # matches F-stat for Groups X prepost
```

```
> t.test(imp)
One Sample t-test data: imp
t = -3.1581, df = 20, p-value = 0.004947
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval: -11.939835 -2.441117
mean of x -7.190476
```

```
> 3.1581^2 [1] 9.973596 # equiv to prepost, no differential change
BK p.232
```

```
> bksubj
pre post method
1 51 48 1
2 35 55 1
3 66 60 1
4 40 35 1
5 39 36 1
6 46 43 1
7 52 46 1
8 42 54 1
9 34 16 2
10 40 36 2
11 34 16 2
12 36 18 2
13 38 32 2
14 32 14 2
15 44 20 2
16 50 43 2
17 60 45 2
18 63 67 2
19 50 36 2
20 42 34 2
21 43 32 2
```

```
> bkrepaov1 = aov(urea ~ as.factor(prepost)*as.factor(method)+ Error(as.factor(subj)))
```

```
> summary(bkrepaov1)
```

```
Error: as.factor(subj)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(method)	1	847.5	847.5	3.6266	0.07212
Residuals	19	4440.0	233.7		

```
Error: Within
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
as.factor(prepost)	1	542.88	542.88	15.142	0.0009823	***Type III SS(prepost) = 317
as.factor(prepost):as.factor(method)	1	407.41	407.41	11.363	0.0032085	**
Residuals	19	681.21	35.85			

BK lmer

Update of BK repeated measures analysis

```
> library(lme4)
> #note brogkutlong restarts subject numbering at 1 for each method; brogkutlong2 numbers sequentially
> bk = read.table(file="http://www-stat.stanford.edu/~rag/stat222/brogkutlong2.dat", header = T)
> attach(bk)
> bklist = lmList(outcome ~ time|subject, data = bk) # getting difference scores the hard way
> bklist
```

data in stat222 web

```
Call: lmList(formula = outcome ~ time | subject, data = bk)
Coefficients:
```

	(Intercept)	time
1	54	-3
2	15	20
...		
21	54	-11

truncated

```
# if you want the "intercept" to be level at time=1 (pretest) the
> t1 = time - 1
> bklist1 = lmList(outcome ~ t1|subject, data = bk)
```

better version

```
> library(lattice) # make a plot for individual subjects
> xyplot(outcome ~ time|subject, groups = method, type = c("p","r"), data = bk)
```

fun plots

```
# the repeated measures anova, shown in previous analysis
> bkrepav1 = aov(outcome ~ as.factor(time)*as.factor(method) + Error(as.factor(subject)))
> summary(bkrepav1)
```

Error: as.factor(subject)

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(method)	1	847	847.5	3.627	0.0721
Residuals	19	4440	233.7		

unequal group sizes makes non-orthog design

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Error: Within
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
as.factor(time)	1	542.9	542.9	15.14	0.000982 ***
as.factor(time):as.factor(method)	1	407.4	407.4	11.36	0.003209 **
Residuals	19	681.2	35.9		

bigger than Type II matches SAS, publication

```
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

as noted R does Type I SS, Type III SS for time is 317 (SAS etc); interaction is prime concern, that (407) matches SAS PROC GLM

```
#so let's try an lmer model: level 1 outcome ~ time; level 2 slope (diff score) depends on method
> bklmera = lmer(outcome ~ I(time - 1) + I(time-1):as.factor(method) + (time|subject), data = bk)
> summary(bklmera)
```

Linear mixed model fit by REML
Formula: outcome ~ I(time - 1) + I(time - 1):as.factor(method) + (time | subject)

Data: bk

	AIC	BIC	logLik	deviance	REMLdev
	305.7	317.9	-145.9	301.1	291.7

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
subject	(Intercept)	35.000	5.9161	
	time	21.455	4.6320	0.220
	Residual	25.125	5.0124	

Number of obs: 42, groups: subject, 21
Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	44.619	2.112	21.130
I(time - 1)	-5.672	1.902	-2.981
I(time - 1):as.factor(method)1	6.378	1.902	3.354

Correlation of Fixed Effects:
(Intr) I(t-1)

```
I(time - 1) 0.028
I(-1):s.(1) 0.000 0.238
# so interaction matches F-statistic from repeated measures anova
> 3.354^2
[1] 11.24932
```

```
# AND lmer gets the occasions (time) term "correct" in the test statistic
> 2.981^2
[1] 8.886361
```

Type III SAS

```
# this matches SAS F-statistic in publication (and SAS) repeated measures output of 8.86 for pre/post (time)
# whereas the aov above has F-statistic 15.1
```

SS not comparable with anova because here were are modeling level 1 params, not outcome
So before looking at other small details, let us declare an lmer victory over non-orthogonal designs

extended version posted bklmer

lmer rules

Model

Level 1 within subject
 $t1 = time - 1$ better this way

$$Y = \alpha_0 + \alpha_1 t1 + \epsilon \quad \alpha_0 = \text{pre}$$

Level 2 $\alpha_0 = \gamma_{00} + u_0$ $\alpha_1 = \gamma_{10} + \gamma_{11} \text{method} + u_1$

Combined

$$Y = \gamma_{00} + \gamma_{10} t1 + \gamma_{11} t1 \cdot \text{method} + (u_0 + u_1 t1 + \epsilon)$$