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# Beliefs and Private Monitoring

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This paper develops new recursive, *set based* methods for studying repeated games with private monitoring. For any *finite-state* strategy profile, we find necessary and sufficient conditions for whether there exists a distribution over initial states such that the strategy, together with this distribution, form a correlated sequential equilibrium (CSE). Also, for any given correlation device for determining initial states (including degenerate cases where players' initial states are common knowledge), we provide necessary and sufficient conditions for the correlation device and strategy to be a CSE, or in the case of a degenerate correlation device, for the strategy to be a sequential equilibrium. We also consider several applications. In these, we show that the methods are computationally feasible, and how to construct and verify equilibria in a secret price-setting game.

Key words: Repeated games, Private monitoring

JEL Codes: C72, C73, D82

#### 1. INTRODUCTION

This paper develops new methods for studying repeated games with private monitoring. In particular, we develop tools that allow us to answer when a particular strategy is consistent with equilibrium. For an important subclass of strategies—those which can be represented as finite automata—we provide readily checkable and computable necessary and sufficient conditions for equilibrium.

The importance of these methods is as follows: while checking the equilibrium conditions in *public*-monitoring games and perfect public equilibria is relatively simple, for games with *private* monitoring, for almost all strategies, checking the equilibrium conditions has previously been considered difficult if not impossible. For instance, consider the following repeated game with private monitoring taken from Mailath and Morris (2002): two partners, privately, either cooperate or defect, and in each period each, privately, has either a good or a bad outcome. While each player can neither observe his partner's action nor his partner's outcome, outcomes are correlated: the vector of joint outcomes is a probabilistic function of the vector of joint actions. (A player cooperating makes it more likely that both players have a good outcome.)

At issue is that even for the simplest games, such as the one presented above, and even the simplest strategies, such as tit-for-tat, there are an infinite number of possible histories where incentives must be checked, and to check incentives one must calculate beliefs for all of them. (This difficulty is not confined to the example above. See, *e.g.* the work of Kandori, 2002 and Mailath and Samuelson, 2006, Chapter 12.) In this paper, for a very large class of strategies,

we resolve this issue by showing the necessity and sufficiency of checking incentives only for "extreme beliefs" (as opposed to checking incentives for all possible histories).

The focus of our analysis is strategies that can be represented by a finite automaton (*finite-state strategies*). A key point (first made by Mailath and Morris, 2002) is that if all players' strategies are finite automata, a particular player's private history is relevant only to the extent that it gives him information regarding the private states of his opponents. This lets us summarize a player's history as a *belief* over a finite state space, a much smaller object than the belief over the private histories of opponents (a point also made by Mailath and Morris, 2002). Moreover, unlike the set of possible private *histories*, the set of possible private *states* for one's opponents does not grow over time.

While many private histories may put a player in the same state of his automaton, they will, in general, induce different beliefs regarding the state of his opponents. Given this, there are two advantages to working with *sets* of beliefs representing all possible beliefs a player can have in a given private state. One is that it is necessary and sufficient to check incentives only for extreme points of those sets instead of looking at beliefs after all histories. The other advantage is that these sets can be readily calculated using recursive methods (operators from sets to sets) that we describe and demonstrate computationally.

Fixed points of our main set based operator represent the beliefs a player can have regarding his opponents' states "in the long run". We show that if incentives hold for extreme points of these sets, one can always use an initial correlation device to, in effect, start the game off as if it had been already running for a long time. This technique alleviates a fundamental difficulty associated with games with private monitoring: the continuation of (sequential) equilibrium play in a game with private monitoring is not a sequential equilibrium but rather a correlated equilibrium in which private histories function as the correlation device. But as Kandori (2002) notes, the correlation device becomes increasingly more complex over time. Using randomization or exogenous correlation in period 0 of the game to make it easier to satisfy incentives and hence support an equilibrium has been suggested by Sekiguchi (1997), Compte (2002), and Ely (2002). We present a robust way of applying this method to construct a family of correlated sequential equilibria.

Our main results are presented as follows. In Section 2, we present our model, a standard repeated game with private monitoring, with finiteness and full support (all signals seen with positive probability) as its only restrictive assumptions. We also present the subclass of strategies we study—*finite-state strategies*, or strategies that can be represented as finite automata.

In Section 3, we show a necessary and sufficient condition for a given correlation device (choosing initial states of players) and a profile of finite automata to form a *CSE* (Theorem 1). That condition involves checking incentive constraints on only the extreme points of a fixed point of our set operator (based on Bayes' rule) which we describe how to compute. Computation is feasible since we show (Lemma 2) that the extreme points of the belief sets of a given iteration are a function only of the extreme points of the belief sets of the previous iteration. Next, we show necessary and sufficient conditions for the *existence of a starting correlation device* such that if coupled with a given automaton they form a CSE (Theorem 2)—they involve checking incentives at extreme beliefs of a fixed point of a related operator. The result implies that the best hope for incentives to hold is to start the players *as if* the game has been played for a long time (without telling them what the outcomes were, but only in which state they should be now). We also show how to verify which starting conditions can support a CSE and which cannot. Since we can apply these results to arbitrary correlation devices, and in particular, to degenerate

ones, we can answer if a particular strategy profile is a *sequential* equilibrium—a correlated equilibrium with a degenerate correlation device.

In Section 4, we present two applications of our methods. We start with the partnership game described above and demonstrate that the methods are easy to apply computationally, and that allows us to gain new intuition regarding how private monitoring affects incentives. In the second application, we consider tacit collusion in a duopoly with competition in prices (with private prices and quantities) and show that one-period price wars are more robust to private monitoring than two-period price wars.

In Section 5, we conclude. Additional results are in an online appendix.

Our results complement the existing literature on the construction of belief-free equilibria (e.g. the work of Ely and Välimäki, 2002; Piccione, 2002; Ely, Hörner and Olszewski, 2005; and Kandori and Obara, 2006) in which players use mixed strategies and their best responses are independent of their beliefs about the private histories of their opponents. In contrast to belief-free equilibria, the equilibria we construct are belief dependent; players' best responses do depend on their beliefs. (For earlier work on constructing belief-dependent sequential equilibria, see Bhaskar and Obara, 2002 and Mailath and Morris, 2002. The first paper constructs a particular equilibrium for an almost-perfect monitoring prisoner's dilemma game. The second describes a class of finite-monitoring equilibria in almost public-monitoring games.)

In terms of the focus on strategies instead of pay-offs, our work is closest to Mailath and Morris (2002, 2006). They consider robustness of particular classes of strategies—those that are equilibria in a public-monitoring game—to a perturbation of the game from public to private, yet almost-public monitoring. They show that strict equilibria in strategies that look back only a finite number of periods (a subclass of the strategies we study) are robust to such perturbations. They also show when infinite history-dependent strategies (partly covered by our analysis) are not robust.

Finally, in a recent paper, Kandori (2010) studies equilibria he calls "Weakly Belief-Free" and shows that in some games, they can achieve higher pay-offs than any belief-free equilibrium. The definition of these equilibria can be translated to our language as follows: incentive constraints have to hold for initial beliefs and for all extreme beliefs obtained after *one iteration* of our operator on the set of *all* possible beliefs (in contrast, the belief-free equilibria check incentives for zero iterations, and our CSE check them after infinitely many iterations).

#### 2. THE MODEL

Consider the game,  $\Gamma^{\infty}$ , defined by the infinite repetition of a stage game,  $\Gamma$ , with N players,  $i=1,\ldots,N$ , each able to take actions  $a_i \in A_i$ . Assume that with probability P(y|a), a vector of private outcomes  $y=(y_1,\ldots,y_N)$  (each  $y_i \in Y_i$ ) is observed conditional on the vector of private actions  $a=(a_1,\ldots,a_N)$ , where for all (a,y), P(y|a)>0 (full support). Further assume that  $A=A_1\times\cdots\times A_N$  and  $Y=Y_1\times\cdots\times Y_N$  are both finite sets, and let  $H_i=A_i\times Y_i$ .

The current period pay-off to player i is denoted  $u_i: H_i \to R$ . That is, player i's pay-off is a function of his own current-period action and private outcome. If player i receives pay-off stream  $\{u_{i,t}\}_{t=0}^{\infty}$ , his lifetime discounted pay-off is  $(1-\beta)\sum_{t=0}^{\infty}\beta^t u_{i,t}$ , where  $\beta \in (0,1)$ . As usual, players care about the expected value of lifetime discounted pay-offs.

Let  $h_{i,t} = (a_{i,t}, y_{i,t})$  denote player *i*'s private action and outcome at date  $t \in \{0, 1, ...\}$ , and  $h_i^t = (h_{i,0}, ..., h_{i,t-1})$  denote player *i*'s private history up to, but not including, date *t*. A (behaviour) *strategy* for player *i*,  $\sigma_i = \{\sigma_{i,t}\}_{t=0}^{\infty}$ , is then, for each date *t*, a mapping from player *i*'s private history  $h_i^t$  to his probability of taking any given action  $a_i \in A_i$  in period *t*. Let  $\sigma$  denote the joint strategy  $\sigma = (\sigma_1, ..., \sigma_N)$  and  $\sigma_{-i}$  denote the joint strategy of all players other than

player i or  $\sigma_{-i} = (\sigma_1, \dots, \sigma_{i-1}, \sigma_{i+1}, \dots, \sigma_N)$ . (Throughout the paper, we use notation -i to refer to all players but player i.)

#### 2.1. Finite-state strategies

In this paper, we restrict attention to equilibria in *finite-state strategies*, or strategies that can be described as finite automata. (However, we allow deviation strategies to be unrestricted.) A finite-state strategy for player i is defined by four objects: (1) a finite private state space  $\Omega_i$  (with  $D_i$  elements  $\omega_i$ ), (2) a function  $p_i(a_i|\omega_i)$  giving the probability of each action  $a_i$  for each private state  $\omega_i \in \Omega_i$ , (3) a deterministic transition function  $\omega_i^+$ :  $\Omega_i \times H_i \to \Omega_i$  determining next period's private state as a function of this period's private state, player i's private action  $a_i$ , and his private outcome  $y_i$ , and (4) an initial state,  $\omega_{i,0}$ .<sup>2</sup> Given this set-up,  $\sigma_{i,0}(a_i) = p_i(a_i|\omega_{i,0})$ ,  $\sigma_{i,1}(a_{i,0}, y_{i,0})(a_i) = p_i(a_i|\omega_i^+(\omega_{i,0}, a_{i,0}, y_{i,0}))$ , and so on.<sup>3</sup> Note that each player's automaton  $\psi_i$  describes play both on and off the equilibrium path. We impose no requirement on the transition rule  $\omega_i^+$  that all states can be reached on the path of play.

Throughout the paper, we repeatedly make a distinction between a finite-state strategy's *automaton* (objects 1 through 3) and object 4, player *i*'s initial state,  $\omega_{i,0}$ . Let  $\psi_i = (\Omega_i, p_i, \omega_i^+)$  denote agent *i*'s automaton. The collection of automata over all players  $\psi = \{\psi_1, \dots, \psi_N\}$  is referred to as the *joint automaton*. Finally, let the number of joint states  $D = \prod_{i \leq N} D_i$ , and the number of joint states for players other than player i,  $D_{-i} = \prod_{j \neq i} D_j$ .

#### 2.2. Beliefs

Since our solution concept will be CSE, allow player i's initial beliefs over the initial state of his opponents,  $\omega_{-i,0}$ , to be possibly non-degenerate. In particular, let player i's beliefs about the initial state of his opponents,  $\mu_{i,0}$ , be a point in the  $(D_{-i}-1)$ -dimensional unit-simplex, denoted  $\Delta^{D_{-i}}$ . Taking as given  $\mu_{i,0}$ , the assumption of full support (P(y|a)>0 for all (a,y)) implies that the beliefs of player i regarding his opponents' private histories,  $h^t_{-i}$ , are always pinned down by Bayes' rule. But since the continuation strategies of players -i depend only on their current joint state,  $\omega_{-i,t}$ , to verify player i's incentive constraints after any given private history  $h^t_i$ , we need not directly consider player i's beliefs regarding  $\omega_{-i,0}$  and  $h^t_{-i}$ . Instead, we need focus only on player i's beliefs regarding his opponents' current state,  $\omega_{-i,t}$ . This is a much smaller object, and, importantly, its dimension does not grow over time.

For a particular initial belief,  $\mu_{i,0}$ , and private history,  $h_i^t$ , player i's belief over  $\omega_{-i,t}$  is, like  $\mu_{i,0}$ , simply a point in the  $(D_{-i}-1)$ -dimensional unit-simplex. Let  $\mu_{i,t}(\mu_{i,0},h_i^t)$  denote player i's belief at the beginning of period t about  $\omega_{-i,t}$  after private history  $h_i^t$  given initial beliefs  $\mu_{i,0}$ . Let  $\mu_{i,t}(\mu_{i,0},h_i^t)(\omega_{-i})$  denote the probability assigned to the particular state  $\omega_{-i}$ .

Beliefs  $\mu_{i,t}(\mu_{i,0}, h_i^t)$  can be defined recursively using Bayes' rule. Let  $B_i(m_i, h_i | \psi_{-i}) \in \Delta^{D_{-i}}$  denote the belief of player i over the state of his opponents at the beginning of period t if his beliefs over his opponents' state at period t-1 were  $m_i \in \Delta^{D_{-i}}$  and he subsequently

The restriction to deterministic transitions is for notational convenience only. All our methods and results apply to automata with non-deterministic transitions.

<sup>3.</sup> For a useful discussion of the validity of representing strategies as finite-state automata in the context of games with private monitoring, see Mailath and Morris (2002) and Mailath and Samuelson (2006).

<sup>4.</sup> Note that if the joint automaton of player i's opponents,  $\psi_{-i}$ , has  $D_{-i}$  states but only  $j < D_{-i}$  are used on path, the beliefs of player i are in the (j-1)-dimensional unit-simplex (rather than in the  $(D_{-i}-1)$ -dimensional unit-simplex). This implies off path states impose a lower computational burden than on path states.

observed  $h_i = (a_i, y_i)$ . This posterior belief can be written out explicitly (from Bayes' rule) as:

$$B_{i}(m_{i}, h_{i}|\psi_{-i})(\omega'_{-i}) = \frac{\sum_{\omega_{-i}} m_{i}(\omega_{-i}) H_{i}(\omega_{-i}, \omega'_{-i}, h_{i}|\psi_{-i})}{\sum_{\omega_{-i}} m_{i}(\omega_{-i}) F_{i}(\omega_{-i}, h_{i}|\psi_{-i})},$$

where

$$F_{i}(\omega_{-i}, h_{i}|\psi_{-i}) = \sum_{(a_{-i}, y_{-i})} p_{-i}(a_{-i}|\omega_{-i}) P(y_{i}, y_{-i}|a_{i}, a_{-i}),$$

$$H_{i}(\omega_{-i}, \omega'_{-i}, h_{i}|\psi_{-i}) = \sum_{h_{-i} \in G_{-i}(\omega_{-i}, \omega'_{-i}|\psi_{-i})} p_{-i}(a_{-i}|\omega_{-i}) P(y_{i}, y_{-i}|a_{i}, a_{-i}),$$

and

$$G_{-i}(\omega_{-i}, \omega'_{-i} | \psi_{-i}) = \{h_{-i} = (a_{-i}, y_{-i}) | \omega^+_{-i}(\omega_{-i}, a_{-i}, y_{-i}) = \omega'_{-i}\}$$

or  $G_{-i}$  is the set of  $(a_{-i}, y_{-i})$  pairs which cause players -i to transit from state  $\omega_{-i}$  to state  $\omega'_{-i}$ .

To define beliefs recursively, let  $B_i^s(m_i, h_i^s|\psi_{-i}) = B_i(B_i^{s-1}(m_i, h_i^{s-1}|\psi_{-i}), h_{i,s-1}|\psi_{-i})$ , where  $B_i^1(m_i, h_i|\psi_{-i}) = B_i(m_i, h_i|\psi_{-i})$ . Then,  $\mu_{i,t}(\mu_{i,0}, h_i^t) = B_i^t(\mu_{i,0}, h_i^t|\psi_{-i})$ . Note that  $B_i(m_i, h_i|\psi_{-i})$  does not depend on  $\sigma_i$  at all, and thus player i's beliefs are the same regardless of whether or not player i is playing a finite-state strategy.

#### 2.3. Equilibrium

Consider player i following an arbitrary strategy  $\sigma_i$ , while players -i follow a finite-state strategy  $\sigma_{-i}$  defined by  $(\omega_{-i,0}, \psi_{-i})$ . That is, players -i are restricted to finite-state strategies, but player i is not. Let  $V_{i,t}(h_i^t, \omega_{-i}|\sigma_i, \psi_{-i})$  denote the lifetime expected discounted pay-off to player i conditional on his private history  $h_i^t$ , and players -i being in state  $\omega_{-i}$ . Thus,

$$V_{i,t}(h_i^t, \omega_{-i} | \sigma_i, \psi_{-i}) = \sum_{a = (a_i, a_{-i})} (\sigma_{i,t}(h_i^t)(a_i) p_{-i}(a_{-i} | \omega_{-i})) \left( \sum_{y} P(y|a) [(1 - \beta) u_i(a_i, y_i) + \beta V_{i,t+1}((h_i^t, (a_i, y_i)), \omega_{-i}^+(\omega_{-i}, a_{-i}, y_{-i}) | \sigma_i, \psi_{-i})] \right).$$

For arbitrary beliefs  $m_i \in \Delta^{D_{-i}}$ , let

$$EV_{i,t}(h_i^t, m_i | \sigma_i, \psi_{-i}) = \sum_{\omega_{-i}} m_i(\omega_{-i}) V_{i,t}(h_i^t, \omega_{-i} | \sigma_i, \psi_{-i}).$$

Player *i*'s expected pay-off given correct beliefs  $\mu_{i,t}(\mu_{i,0}, h_i^t)$  is then  $EV_{i,t}(h_i^t, \mu_{i,t}(\mu_{i,0}, h_i^t)|\sigma_i, \psi_{-i})$ .

If  $\sigma_i$  is a finite-state strategy (defined by  $(\omega_{i,0}, \psi_i)$ ), let  $\omega_{i,t}(\omega_{i,0}, h_i^t)$  denote the private state for player i at date t implied by initial state  $\omega_{i,0}$ , transition rule  $\omega_i^+(\omega_i, a_i, y_i)$ , and history  $h_i^t = ((a_{i,0}, y_{i,0}), \dots, (a_{i,t-1}, y_{i,t-1}))$ . Then, for all  $(h_i^t, \hat{h}_i^t)$  such that  $\omega_{i,t}(\omega_{i,0}, h_i^t) = \omega_{i,t}(\omega_{i,0}, \hat{h}_i^t)$ ,  $V_{i,t}(h_i^t, \omega_{-i} | \sigma_i, \psi_{-i}) = V_{i,t}(\hat{h}_i^t, \omega_{-i} | \sigma_i, \psi_{-i})$ . Given this, we can write player i's lifetime payoff, conditional on  $\omega_{-i}$ , as a function of his current private state  $\omega_i$  as opposed to depending

directly on his private history,  $h_i^t$ . Thus, we define  $v_i(\omega_i, \omega_{-i}|\psi_i, \psi_{-i}) \equiv V_{i,t}(h_i^t, \omega_{-i}|\sigma_i, \psi_{-i})$  for any  $h_i^t$  such that  $\omega_i = \omega_{i,t}(\omega_{i,0}, h_i^t)$ . Then we denote player i's expected pay-off, now a function of his current state,  $\omega_i$ , and his beliefs over his opponents' state,  $\omega_{-i}$ , as

$$Ev_i(\omega_i, m_i | \psi_i, \psi_{-i}) = \sum_{\omega_{-i}} m_i(\omega_{-i}) v_i(\omega_i, \omega_{-i} | \psi_i, \psi_{-i}).$$

Definition 1. A probability distribution over initial states,  $x \in \Delta^D$ , and joint automaton,  $\psi$ , form a CSE of  $\Gamma^{\infty}$  if for all  $i, t, h_i^t, \omega_{i,0}$  such that  $\sum_{\omega_{-i,0}} x(\omega_{i,0}, \omega_{-i,0}) > 0$ , and arbitrary  $\hat{\sigma}_i$ ,

$$Ev_i(\omega_{i,t}(\omega_{i,0},h_i^t),\mu_{i,t}(\mu_{i,0}(x,\omega_{i,0}),h_i^t)|\psi_i,\psi_{-i}) \ge EV_{i,t}(h_i^t,\mu_{i,t}(\mu_{i,0}(x,\omega_{i,0}),h_i^t)|\hat{\sigma}_i,\psi_{-i}),$$

where 
$$\mu_{i,0}(x,\omega_{i,0})(\omega_{-i,0}) = x(\omega_{i,0},\omega_{-i,0})/\sum_{\overline{\omega}_{-i,0}} x(\omega_{i,0},\overline{\omega}_{-i,0}).$$

There are two difficulties in verifying whether a given  $(x, \psi)$  form a CSE. First, there are infinitely many deviation strategies. Second, to verify the IC constraints, we need to know the beliefs players have on and off path after each element of the infinite set of possible private histories. The first difficulty is shared by all repeated game models and, as usual, it is solved by using the one-shot deviation principle. The resolution of the second difficulty is the main focus of this paper.

**Lemma 1** (One-shot Deviation Principle). Suppose a correlation device x and joint automaton  $\psi$  satisfy for all i,  $h_i^t$ ,  $\hat{a}_i$ , and  $\omega_{i,0}$  such that  $\sum_{\omega_{-i,0}} x(\omega_{i,0},\omega_{-i,0}) > 0$ ,

$$Ev_{i}(\omega_{i,t}(\omega_{i,0}, h_{i}^{t}), \mu_{i,t}(\mu_{i,0}(x, \omega_{i,0}), h_{i}^{t})|\psi_{i}, \psi_{-i})$$

$$\geq \sum_{\omega_{-i}} \mu_{i,t}(\mu_{i,0}(x, \omega_{i,0}), h_{i}^{t})(\omega_{-i}) \left[ \sum_{a_{-i}} p_{-i}(a_{-i}|\omega_{-i}) \sum_{y} P(y|\hat{a}_{i}, a_{-i}) \right]$$

$$\times \left[ (1 - \beta)u_{i}(\hat{a}_{i}, y_{i}) + \beta v_{i}(\omega_{i}^{+}(\omega_{i,t}(\omega_{i,0}, h_{i}^{t}), \hat{a}_{i}, y_{i}), \omega_{-i}^{+}(\omega_{-i}, a_{-i}, y_{-i})|\psi_{i}, \psi_{-i}) \right].$$

Then,  $(x, \psi)$  form a CSE. That is, it is sufficient to check that player i does not wish to deviate once and then revert to playing according to his automaton  $\psi_i$ .

*Proof.* Mailath and Samuelson (2006, p. 397).

#### 3. VERIFYING EQUILIBRIA

We now turn to the main methodological contribution of the paper: set based methods delivering first, necessary and sufficient conditions for when a joint automaton  $\psi$ , when coupled with *particular* correlation device x, forms a CSE (Theorem 1), and second, necessary and sufficient conditions for whether there exists *any* correlation device x such that a joint automaton  $\psi$ , when coupled with x, forms a CSE. That is, our second main result (Theorem 2) regards whether the joint automaton  $\psi$  itself is consistent with equilibrium.<sup>5</sup>

<sup>5.</sup> Note that if a given joint automaton,  $\psi$ , is not consistent with equilibrium, there may still exist a different joint automaton,  $\hat{\psi}$ , with the same on-path play but different off-path play that is consistent with equilibrium. We provide partial answers to this problem in the online appendix.

Rather than considering separately the beliefs  $m_i \in \Delta^{D_{-i}}$  that a player will have after some private history, it is useful to consider *sets* of beliefs. In particular, let  $M_i(\omega_i) \subset \Delta^{D_{-i}}$  denote a closed convex set of beliefs, and  $M_i$  be a collection of  $D_i$  sets  $M_i(\omega_i)$ , one for each  $\omega_i$ . Let  $\mathcal{M}$  denote the space of such collections of sets  $M_i$ . To define the distance between two elements  $M_i$  and  $M_i' \in \mathcal{M}$ , first let the distance between two beliefs  $m_i$  and  $m_i' \in \Delta^{D_{-i}}$  be defined by the sup norm (or Chebyshev distance) denoted  $|m_i, m_i'| = \max_{\omega_{-i}} |m_i(\omega_{-i}) - m_i'(\omega_{-i})|$ . Next, for a belief  $m_i$  and a non-empty closed set  $A \subset \Delta^{D_{-i}}$ , let the distance between them (the Hausdorff distance) be defined as  $|m_i, A| = \min_{m_i' \in A} |m_i, m_i'|$ . For two non-empty closed sets  $(A, A') \subset \Delta^{D_{-i}}$ , the Hausdorff distance between them is defined as  $|A, A'| = \max\{\max_{m_i \in A} |m_i, A'|, \max_{m_i' \in A'} |m_i', A|\}$ . If A is non-empty, let  $|A, \emptyset| = 1$  and  $|\emptyset, A| = 1$ . Finally, let  $|\emptyset, \emptyset| = 0$ . (Note that for non-empty A and A',  $|A, A'| \leq 1$ .) Then the distance between two collections of belief sets  $M_i$ ,  $M_i' \in \mathcal{M}$  is defined as  $|M_i, M_i'| = \max_{\omega_i} |M_i(\omega_i), M_i'(\omega_i)|$ .

We begin by constructing two related operators from  $\mathcal{M}$  to  $\mathcal{M}$ , where fixed points of these operators will be a focus of our main results. Let the one-step operator  $T(M_i)$  be defined as<sup>6</sup>

$$T(M_i) = \{T(M_i)(\omega_i') | \omega_i' \in \Omega_i\},\,$$

where

$$T(M_i)(\omega_i') = \operatorname{co}(\{m_i' | \text{ there exists } \omega_i \in \Omega_i, m_i \in M_i(\omega_i) \text{ and } (a_i, y_i) \in G_i(\omega_i, \omega_i' | \psi_i)$$
  
such that  $m_i' = B_i(m_i, a_i, y_i | \psi_{-i})\}),$ 

where co() denotes the convex hull and  $G_i(\omega_i, \omega_i'|\psi_i)$  is the set of  $(a_i, y_i)$  such that  $\omega_i^+$   $(\omega_i, a_i, y_i) = \omega_i'$ . The T operator works as follows: suppose one takes as given the sets of "allowable" beliefs player i can have over the private state of the other players,  $\omega_{-i}$ , last period. For any given such allowable belief, Bayesian updating then implies what player i should believe about  $\omega_{-i}'$  this period for each realization of  $(a_i, y_i)$ , generating a collection of allowable belief sets. That is, if there exists a way to choose player i's state last period,  $\omega_i$ , the beliefs of player i over the private states of his opponents last period consistent with  $m_i \in M_i(\omega_i)$ , and a new realization of  $(a_i, y_i)$  such that Bayesian updating delivers beliefs  $m_i'$ , then  $m_i' \in T(M_i)(\omega_i^+(\omega_i, a_i, y_i))$ . In effect, the T operator gives, for a particular collection of belief sets  $M_i$ , the belief sets associated with all possible successor beliefs generated by new data and interpreted through  $\sigma_{-i}$  (as well as all convex combinations of such beliefs). Note that since  $B_i$  and  $G_i$  depend only on the joint automaton  $\psi$ , as opposed to starting conditions, x, the T operator retains this property as well.

Next, let the operator  $T^{U}(M_{i})$  (U for union) be

$$T^{U}(M_{i}) = \{T^{U}(M_{i})(\omega_{i}) | \omega_{i} \in \Omega_{i}\}, \text{ where } T^{U}(M_{i})(\omega_{i}) = \operatorname{co}(T(M_{i})(\omega_{i}) \cup M_{i}(\omega_{i})).$$

In words, the  $T^U$  operator calculates for every state  $\omega_i$ , the convex hull of the union of the prior beliefs player i could hold last period,  $M_i(\omega_i)$ , and all the posterior beliefs he can hold in that same state,  $T(M_i)(\omega_i)$ .

We note here that the T and  $T^U$  operators are relatively easy to operationalize. In particular, the following lemma implies that the extreme points of the collection of sets  $T(M_i)$  and  $T^U(M_i)$  can be calculated using only the extreme points of the collection of sets  $M_i$ .

**Lemma 2**. If  $M_i(\omega_i)$  is closed and convex for all  $\omega_i$ , then  $T(M_i)(\omega_i)$  and  $T^U(M_i)(\omega_i)$  are both closed and convex for all  $\omega_i$ . Next, if  $m_i$  is an extreme point of  $T^U(M_i)(\omega_i)$  but not  $T(M_i)(\omega_i)$ ,

<sup>6.</sup> The T operator depends on  $\psi_{-i}$  and varies across players (as does  $\mathcal{M}$ ), but to conserve notation, we write  $T(M_i)$  rather than  $T_i(M_i|\psi_{-i})$ .

then  $m_i$  is an extreme point of  $M_i(\omega_i)$ . Finally, if  $m_i$  is an extreme point of both  $T(M_i)(\omega_i)$  and  $T^U(M_i)(\omega_i)$ , then there exists  $\hat{m}_i$ ,  $\hat{\omega}_i$ ,  $h_i$  such that  $m_i = B_i(\hat{m}_i, h_i | \psi_{-i})$ ,  $h_i \in G_i(\hat{\omega}_i, \omega_i | \psi_i)$  and  $\hat{m}_i$  is an extreme point of  $M_i(\hat{\omega}_i)$ .

*Proof.* See Appendix.

# 3.1. Fixed points of T and $T^U$

Our results rely on properties of the fixed points of T and  $T^U$ . We write  $M_i^0 \subset M_i^1$  if  $M_i^0(\omega_i) \subset M_i^1(\omega_i)$  for all  $\omega_i$ . Furthermore, we write  $M_i$  is non-empty if there exists a private state  $\omega_i$  such that  $M_i(\omega_i)$  is non-empty.

Both T and  $T^U$  are monotonic operators (i.e. if  $M_i^0 \subset M_i^1$ , then  $T(M_i^0) \subset T(M_i^1)$  and  $T^U(M_i^0) \subset T^U(M_i^1)$ ). By construction,  $M_i \subset T^U(M_i)$  for all  $M_i \in \mathcal{M}$ . Since  $M_i \subset T^U(M_i)$ , and  $T^U(M_i) \subset T^U(T^U(M_i))$  (from monotonicity), the sequence  $\{M_i, T^U(M_i), T^U(T^U(M_i)), \ldots\}$  converges. That  $B_i$  is continuous implies  $T^U$  is continuous and thus this limit is a fixed point of  $T^U$ . Call this fixed point  $M_i^{*U}(M_i)$ . Next note that if  $M_i \subset T(M_i)$ , then  $T(M_i) = T^U(M_i)$ . This implies if  $M_i \subset T(M_i)$ , the sequence  $\{M_i, T(M_i), T(T(M_i)), \ldots\}$  also converges to  $M_i^{*U}(M_i)$ .

#### 3.2. When is a pair $(x, \psi)$ a CSE?

For an arbitrary correlation device, x, let the belief sets  $M_{i,0}(x,\omega_i) \in \Delta^{D_i}$  be defined such that

$$M_{i,0}(x,\omega_i) = \{\mu_{i,0}(x,\omega_i)\}\$$

for all  $\omega_i$  such that  $\sum_{\omega_{-i}} x(\omega_i, \omega_{-i}) > 0$ . Otherwise, let  $M_{i,0}(x, \omega_i) = \emptyset$ . That is, for all  $\omega_i$ , if  $\omega_i$  occurs with positive probability under distribution x,  $M_{i,0}(x, \omega_i)$  is the single point belief set consisting of what player i believes about  $\omega_{-i}$  when his initial state is  $\omega_i$ . Let  $M_{i,0}(x)$  be a collection of  $D_i$  sets  $M_{i,0}(x, \omega_i)$ , one for each  $\omega_i$ , and (with some abuse of notation)  $M_i^{*U}(x) \equiv M_i^{*U}(M_{i,0}(x))$ .

**Theorem 1.** A correlation device x and a joint automaton  $\psi$  form a CSE if and only if the incentive compatibility conditions

$$Ev_{i}(\omega_{i}, m_{i} | \psi_{i}, \psi_{-i}) \geq \sum_{\omega_{-i}} m_{i}(\omega_{-i}) \left[ \sum_{a_{-i}} p_{-i}(a_{-i} | \omega_{-i}) \sum_{y} P(y | \hat{a}_{i}, a_{-i}) [(1 - \beta)u_{i}(\hat{a}_{i}, y_{i}) + \beta v_{i}(\omega_{i}^{+}(\omega_{i}, \hat{a}_{i}, y_{i}), \omega_{-i}^{+}(\omega_{-i}, a_{-i}, y_{-i}) | \psi_{i}, \psi_{-i}) ] \right]$$

$$(1)$$

hold for all i,  $\hat{a}_i$ ,  $\omega_i$ , and  $m_i$  such that  $m_i$  is an extreme point of  $M_i^{*U}(x)$ .

*Proof.* If: since incentive compatibility conditions (1) are linear in beliefs, then if they hold for the extreme beliefs of  $M_i^{*U}(x)$ , they hold for all beliefs in these sets. By monotonicity,  $(T^U)^t(M_{i,0}(x)) \subset M_i^{*U}(x)$  for all  $t \geq 0$ , so incentives hold in the first period for all initial signals and in all subsequent periods for all possible continuation histories.

Only if: suppose that incentive compatibility conditions (1) are violated for some state  $\omega_i$  and extreme belief  $m_i \in M_i^{*U}(x)(\omega_i)$ . Since the incentive conditions (1) are continuous in beliefs and are weak inequalities, there exists an  $\varepsilon > 0$  such that for all beliefs  $m_i'$  such that  $|m_i', m_i| < \varepsilon$ , incentives are violated in state  $\omega_i$  with beliefs  $m_i'$ .

Now, by definition of  $T^U$ , for every t and  $\omega_i$ , every extreme point of  $(T^U)^t(M_{i,0}(x))(\omega_i)$  is either an extreme point of  $(T^U)^{t-1}(M_{i,0}(x))(\omega_i)$  or an extreme point of  $T((T^U)^{t-1}(M_{i,0}(x)))(\omega_i)$ . Therefore, we can find an initial state  $\omega_{i,0}$  and a private history  $h_i^t$  such that player i after  $h_i^t$  is in state  $\omega_i$  and his beliefs  $\mu_{i,t}(\mu_{i,0},h_i^t)$  satisfy  $|\mu_{i,t}(\mu_{i,0},h_i^t),m_i|<\varepsilon$  (using that  $(T^U)^n(M_{i,0}(x))\to M^{*U}(x)$ ). Thus  $(x,\psi)$  are not a CSE.  $\parallel$ 

3.3. When does there exist an x such that  $(x, \psi)$  is a CSE?

For a joint automaton  $\psi = (\Omega, p, \omega^+)$ , denote the Markov transition matrix on the joint state  $\omega \in \Omega$  by

$$\tau(\omega, \omega')(\psi) = \sum_{(a,y) \text{ s.t. } (a_i, y_i) \in G_i(\omega_i, \omega_i' | \psi_i) \text{ for all } i} P(y|a) \prod_i p_i(a_i | \omega_i).$$
 (2)

Since  $\tau(\psi)$  defines a finite-state Markov chain, it has at least one invariant distribution,  $\pi \in \Delta^D$ .

**Lemma 3**. Let  $\pi$  be an invariant distribution of the Markov process  $\tau(\psi)$ . Then for all i,  $M_{i,0}(\pi) \subset T(M_{i,0}(\pi))$ .

*Proof.* See Appendix.

The basic idea behind the proof of Lemma 3 is that beliefs drawn from an invariant distribution are an average, and thus a convex combination, of beliefs which condition on additional information. Since the T operator is the convex hull of all possible posteriors from given priors, and the average posterior belief is the prior belief, the convex hull of the set of possible posterior beliefs must contain the prior belief. Lemma 3 then implies that  $T(M_{i,0}(\pi)) = T^U(M_{i,0}(\pi))$  and that  $\{M_{i,0}(\pi), T(M_{i,0}(\pi)), T(T(M_{i,0}(\pi))), \ldots\}$  converges to  $M_i^{*U}(\pi)$ .

**Lemma 4**. For a given  $(x, \psi)$  let  $\pi = \lim_{t \to \infty} \frac{1}{t+1} \sum_{n=0}^{t} x \tau(\psi)^n$ . Then

- (a) The limit exists and is an invariant distribution of  $\tau(\psi)$ .
- (b)  $M_{i,0}(\pi) \subset M_i^{*U}(x)$ .

*Proof.* See Appendix.

Part (b) of Lemma 4 states that the *initial* beliefs player i can have if initial states are drawn from the invariant distribution of  $\tau(\psi)$  defined in part (a) (the sets  $M_{i,0}(\pi)$ ) are always contained in the set of *all* beliefs player i can have over all dates when starting with the arbitrary correlation device x,  $M_i^{*U}(x)$ . The intuition of Lemma 4 is similar to Lemma 3: the beliefs  $M_{i,0}(\pi)$  correspond to drawing initial states from a random time from the Markov chain  $\tau(\psi)$  and hence are a convex combination of beliefs that condition on both calendar time and the realized history, which in turn are contained in  $M_i^{*U}(x)$ .

**Theorem 2.** For a given joint automaton,  $\psi$ , there exists a correlation device x such that  $(x, \psi)$  form a CSE if and only if for some invariant distribution  $\pi$  of  $\tau(\psi)$ , incentives hold (i.e. condition (1) from Theorem 1) for all i,  $\omega_i$  and  $m_i$  which is an extreme point of  $M_i^{*U}(\pi)(\omega_i)$ .

<sup>7.</sup> The authors thank an anonymous referee for correctly suggesting that one of our sufficient conditions from a previous version of this paper—that incentives hold for all extreme points of  $M_i^{*U}(\pi)$  for an invariant distribution  $\pi$ —was most likely also necessary.

*Proof.* If: let  $x = \pi$ . From Lemma 3 (and the monotonicity of T), the time zero beliefs of each player i,  $M_{i,0}(\pi,\omega_{i,0}) \in M_i^{*U}(\pi)(\omega_{i,0})$  for each  $\omega_{i,0}$  drawn with positive probability. Moreover, the subsequent beliefs for each player i are elements of  $M_i^{*U}(\pi)(\omega_{i,t})$  for each date t and private history  $h_i^t$ , where  $\omega_{i,t}$  is player i's state at date t after private history  $h_i^t$ .

Suppose condition (1) holds, for all i,  $\hat{a}_i$ ,  $\omega_i$ , and extreme points of  $M_i^{*U}(\pi)(\omega_i)$ , where  $m_i$  and  $\hat{m}_i$  are two such points. Then since equation (1) is linear in these beliefs, for all  $\alpha \in [0, 1]$ , condition (1) holds for beliefs  $\alpha m_i + (1 - \alpha)\hat{m}_i$ , again for all i,  $\hat{a}_i$ , and  $\omega_i$ . Thus, incentives hold for all dates t and private histories  $h_i^t$  if initial states are drawn according to  $\pi$ .

Only if, suppose there exists a correlation device x such that  $(x, \psi)$  form a CSE, but for all invariant distributions  $\pi$  of  $\tau(\psi)$ ,  $(\pi, \psi)$  does not form a CSE. That  $(x, \psi)$  forms a CSE implies, by Theorem 1, that incentives hold for all i,  $\omega_i$  and  $m_i$  which is an extreme point of  $M_i^{*U}(x)$ . Let

$$\pi = \lim_{t \to \infty} \frac{1}{t} \sum_{n=0}^{t-1} x \tau^n.$$

By Lemma 4,  $\pi$  is an invariant distribution of  $\tau(\psi)$  and  $M_{i,0}(\pi) \subset M_i^{*U}(x)$ . Since  $T^U$  is a monotone operator:

$$(T^{U})^{n}(M_{i,0}(\pi)) \subset (T^{U})^{n}(M_{i}^{*U}(x)) = M_{i}^{*U}(x)$$

and so in the limit:

$$M_i^{*U}(\pi) \subset M_i^{*U}(x)$$
.

Applying Theorem 1, this implies that  $(\pi, \psi)$  is also a CSE, a contradiction.

#### 3.4. Strategies with unique invariant distributions

In the previous section, we showed that a joint automaton  $\psi$  is consistent with equilibrium if and only if it is a CSE to have initial private states drawn from an invariant distribution of  $\tau(\psi)$ . Verifying for a particular invariant distribution  $\pi$  of  $\tau(\psi)$  whether  $(\pi, \psi)$  form a CSE then involves calculating  $M_i^{*U}(\pi) \equiv \lim_{s \to \infty} T^s(M_{i,0}(\pi))$  and checking incentives at its extreme points. A second method involves calculating  $\overline{M_i} \equiv \lim_{s \to \infty} T^s(\overline{\Delta_i})$  (where  $\overline{\Delta_i}$  denotes the collection of  $D_i$ ,  $D_{-i} - 1$ -dimensional unit simplexes) and checking incentives at its extreme points. Since the set inclusion relationship,  $\subset$ , defines a complete lattice on the space of  $D_i$  closed subsets of  $\Delta^{D_{-i}}$ ,  $\overline{M_i}$  is the largest fixed point of T and all other fixed points of T are subsets of it (by Tarski's fixed point theorem). Thus, if incentives hold at the extreme points of  $\overline{M_i}$  (for all i), or incentives hold at the extreme points of any point in the sequence  $\{T^s(\overline{\Delta_i})\}_{s=0}^{\infty}$ ,  $(\pi, \psi)$  is a CSE for any invariant distribution  $\pi$  of  $\tau(\psi)$ . But this only establishes a sufficient condition for equilibrium. Here, we show that if  $\tau(\psi)$  is a regular matrix (i.e. there exists an s such that  $\tau(\psi)^s$  has all non-zero entries), then incentives holding at the extreme points of  $\overline{M_i}$  is necessary as well. (Note that if  $\tau(\psi)$  is a regular matrix, then all joint states are reached on path.)

**Lemma 5**. Suppose  $\tau(\psi)$  is a regular matrix. Then  $\overline{M}_i$  is the unique non-empty fixed point of T and for all non-empty  $M_i \in \mathcal{M}$ ,  $\lim_{n \to \infty} T^n(M_i) = \overline{M}_i$ .

*Proof.* See Appendix.

**Corollary 1** (of Theorem 2). If  $\tau(\psi)$  is a regular matrix, then there exists a correlation device x such that  $(x, \psi)$  form a CSE if and only if incentives hold (i.e. condition (1) from Theorem 1) for all i and  $m_i$  such that  $m_i$  is an extreme point of  $\overline{M}_i$ .

*Proof.* Lemma 5, Lemma 3 and that for all  $M_i$  such that  $M_i \subset T(M_i)$ ,  $T(M_i) = T^U(M_i)$  imply  $M_i^{*U}(\pi) = \overline{M}_i$ , where  $\pi$  is the unique invariant distribution of  $\tau(\psi)$ . Theorem 2 then implies the result.

#### 3.5. Which starting conditions work?

For a given joint automaton  $\psi$ , Theorem 2 gives us necessary and sufficient conditions for the existence of a correlation device x such that  $(x, \psi)$  form a CSE. Suppose we find a  $\psi$  that satisfies these conditions. A natural question is then, what x can be used to start the strategies without violating incentive constraints? From the proof of Theorem 2, we know that at least one of the invariant distributions of  $\tau(\psi)$  can be used.

One can use Theorem 1 to verify for any x, whether  $(x, \psi)$  is a CSE. That requires computing a fixed point of  $T^U$  for every such x. We now show that one can compute once a fixed point of a related operator and use it to evaluate any x.

In particular, define  $M_i^I(\omega_i)$  to be the set of beliefs such that incentives hold in the current period for all beliefs  $m_i \in M_i^I(\omega_i)$  if player i is in state  $\omega_i$  and plans to follow the strategy in the future. Clearly, a necessary condition for  $(x, \psi)$  to be a CSE is that  $M_{i,0}(x) \subset M_i^I$  since otherwise incentives would be violated in the first period. We need to ensure, however, that incentives are satisfied not only for a particular belief generated by the correlation device but also for all possible successors of that belief, and successors of those beliefs, and so on.

Define the operator  $T^{I}(M_{i})$  (I for incentives) as

$$T^{I}(M_{i}) = \{T^{I}(M_{i})(\omega_{i}) | \omega_{i} \in \Omega_{i}\}, \text{ where}$$

$$T^{I}(M_{i})(\omega_{i}) = \operatorname{co}(\{m_{i} | m_{i} \in M_{i}(\omega_{i}) \text{ and for all } (a_{i}, y_{i}),$$

$$B_{i}(m_{i}, a_{i}, y_{i} | \psi_{-i}) \in M_{i}(\omega^{+}(\omega_{i}, a_{i}, y_{i}))\}.$$
(3)

In words,  $T^I$  eliminates an element of  $M_i(\omega_i)$  if there exists a private history  $(a_i, y_i)$  and a successor belief which is not in  $M_i(\omega_i^+(\omega_i, a_i, y_i))$ . Clearly,  $T^I$  is monotone and  $T^I(M_i) \subset M_i$  for any  $M_i$ . Thus, the sequence  $\{(T^I)^n(M_i^I)\}_{n=0}^{\infty}$ 

Clearly,  $T^I$  is monotone and  $T^I(\dot{M}_i) \subset M_i$  for any  $M_i$ . Thus, the sequence  $\{(T^I)^n(M_i^I)\}_{n=0}^{\infty}$  (starting with the set of beliefs such that incentives hold in the first period), represents a sequence of (weakly) ever smaller collection of sets, guaranteeing that the limit, denoted  $M_i^{*I}$ , exists. Importantly,  $M_i^{*I}$  can be computed independently of x, allowing us to then evaluate all correlation devices to this benchmark:

**Corollary 2** (of Theorem 1). A correlation device x and a joint automaton  $\psi$  form a CSE if and only if for all i,  $M_{i,0}(x) \subset M_i^{*I}$ 

*Proof.* For any  $M_i$ , by the definition of  $T^I$ , we have

$$M_i \subset M_i^{*I} \iff M_i^{*U}(M_i) \subset M_i^{*I}$$

hence by Theorem 1,  $(x, \psi)$  form a CSE if and only if  $M_{i,0}(x) \subset M_i^{*I}$ .

Since the set of correlated equilibria is convex, if  $(x, \psi)$  and  $(x', \psi)$  are CSE, so is  $(x'', \psi)$  for any x'' which is a convex combination of x and x'. Finally, for belief-free equilibria (such as those in Ely and Välimäki, 2002), the conditions of the corollary hold automatically since  $M_i^{*I} = \overline{\Delta}_i$  or that incentives hold, by construction, for all beliefs.

#### 4. APPLICATIONS

In this section, we attempt to show that these methods are useful in analysing interesting economic applications.

# 4.1. A repeated partnership game (Mailath and Morris, 2002)

In this example, we use the repeated partnership game of Mailath and Morris (2002) to show that (a) one can use our methods to easily compute the relevant belief sets to verify incentive conditions, (b) analyse which starting conditions work, (c) do comparative statics regarding model parameters, and (d) investigate that histories are problematic when parameters are such that a strategy is not an equilibrium.

We also highlight two somewhat surprising results. First, we show that sometimes tit-for-tat coordination works if both players start in the bad state but not when both players start in the good state. Second, we compute an example where knowing too well the state of one's opponent can be bad for incentives. If a player has less knowledge about the state of his opponent (because of stochastic starting conditions or less predictable consumers or less correlated private signals), it can make it easier to satisfy incentives.<sup>8</sup>

**4.1.1. The partnership game.** Consider the two player partnership game in which each player  $i \in \{1,2\}$  can take action  $a_i \in \{C,D\}$  (cooperate or defect) and each can realize a private outcome  $y_i \in \{G,B\}$  (good or bad). The P(y|a) function is such that if m players cooperate, then with probability  $p_m(1-\epsilon)^2+(1-p_m)\epsilon^2$ , both players realize the good private outcome. With probability  $(1-\epsilon)\epsilon$ , player 1 realizes the good outcome, while player 2 realizes the bad. (Likewise, with this same probability, player 2 realizes the good outcome and player 1 the bad.) Finally, with probability  $p_m\epsilon^2+(1-p_m)(1-\epsilon)^2$ , both players realize the bad outcome. Essentially, this game is akin to one in which  $p_m$  determines the probability of an unobservable underlying outcome and  $\epsilon$  is the probability that player i's outcome differs from this underlying outcome. Thus, when  $\epsilon=0$ , outcomes are public, and when  $\epsilon$  approaches 0, outcomes are almost public. Pay-offs are determined by specifying  $\beta$  and for each player i, the vector  $\{u_i(C,G),u_i(C,B),u_i(D,G),u_i(D,B)\}$ .

**4.1.2. Tit-for-tat.** Next, consider perhaps the simplest non-trivial pure strategy: tit-for-tat. That is, let each player i play C if his private outcome was good in the previous period and D otherwise. This is a two-state strategy with  $\Omega_i = \{R, P\}$  for "reward" and "punish". For  $i \in \{1, 2\}$ ,  $p_i(C|R) = 1$ ,  $p_i(D|P) = 1$ ,  $\omega_i^+(\omega_i, a_i, G) = R$ ,  $\omega_i^+(\omega_i, a_i, B) = P$  for  $\omega_i \in \{R, P\}$ , and  $a_i \in \{C, D\}$ . Since every joint state can be reached from every other joint state with positive probability,  $\tau(\psi)$  is a regular matrix and Corollary 1 of Theorem 2 applies and thus tit-for-tat is compatible with equilibrium if and only if incentives hold for the extreme points of the unique non-empty fixed point of T,  $\overline{M}_i$ . Since the number of states of i's opponent  $D_{-i} = 2$ , the set  $\overline{M}_i(\omega_i)$  is simply a closed interval specifying the range of probabilities that player -i is in state R, given that player i is in state  $\omega_i \in \{R, P\}$ . Operator T maps a collection of two intervals (one for each  $\omega_i$ ) to a collection of two intervals.

For  $\beta = 0.9$ ,  $p_0 = 0.3$ ,  $p_1 = 0.55$ , and  $p_2 = 0.9$  and a pay-off of 1 for receiving a good outcome and a pay-off of -0.4 for cooperating, we can easily verify that the static game is a prisoner's dilemma and that tit-for-tat is an equilibrium of the public outcome ( $\epsilon = 0$ ) game, starting from either both players in state R or both players in state P. For  $\epsilon > 0$ , beliefs matter

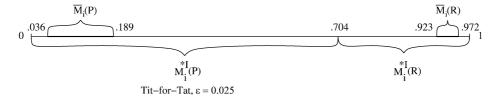


FIGURE 1
Belief Sets for Tit-for-Tat

and to check equilibrium conditions, one must construct the intervals  $\overline{M}_i(\omega_i)$ . The procedure of iterating the T mapping is relatively easily implemented on a computer. For  $\epsilon = 0.025$ , the procedure converges (in less than a second) to these intervals:  $\overline{M}_i(R) = [0.923, 0.972]$ , and  $\overline{M}_i(P) = [0.036, 0.189]$  (see Figure 1).

Again, tit-for-tat is compatible with equilibrium if and only if each player indeed wishes to play C when he believes his opponent is in state R with either probability 0.923 or 0.972 and indeed wishes to play D when he believes his opponent is in state R with either probability 0.036 or 0.189 (assuming a reversion to path play after a deviation). This is a matter of simply checking equation (1) for each of these four beliefs, and it holds in this case, thus there exist starting conditions such that tit-for-tat is an equilibrium.

In particular, Theorem 2 delivers one such starting condition. If both players follow the equilibrium, the transition matrix  $\tau(\psi)$  between joint state  $\omega \in \Omega = \{RR, RP, PR, PP\}$  and  $\omega' \in \Omega$  implies a unique invariant distribution  $\pi = (0.659, 0.038, 0.038, 0.264)$ . If one chooses the correlation device  $x = \pi$ , then if player  $i \in \{1, 2\}$  has R as his initial recommended state, he believes his opponent's initial recommended state is R with probability 0.945 = 0.659/(0.659 + 0.038). Likewise, if his initial recommended state is R, he believes his opponent's initial recommended state is R with probability 0.127 = 0.038/(0.038 + 0.264). Note that Lemma 4 implies the belief of player i after recommendation R,  $\mu_{i,0}(R) = 0.945 \in \overline{M}_i(R)$  and likewise,  $\mu_{i,0}(P) = 0.127 \in \overline{M}_i(P)$ . Thus, the correlation device  $x = \pi$  and tit-for-tat form a CSE.

Are there any other starting conditions for which tit-for-tat is an equilibrium? Using the  $T^I$  operator, one can also readily calculate the sets  $M_i^{*I}$  for players  $i \in \{1,2\}$ . In this example,  $M_i^{*I}(R) = [0.704, 1]$  and  $M_i^{*I}(P) = [0, 0.704]$ . Corollary 2 then implies any correlation device x that delivers conditional beliefs  $\mu_{i,0}(R) \in [0.704, 1]$  and  $\mu_{i,0}(P) \in [0, 0.704]$ , together with tit-for-tat, forms a CSE. Thus, starting each player off in state  $\omega_i = R$  with certainty (or x puts all mass on  $\omega = RR$ ) and following tit-for-tat is a *sequential* equilibrium since  $M_{i,0}(x,R) = \{1\} \subset M_i^{*I}(R)$  and  $M_{i,0}(x,P) = \emptyset \subset M_i^{*I}(P)$ . Likewise, starting each player off in state P (x puts all weight on x000 and x101 salso a sequential equilibrium since x101 salso a sequential equilibrium since x102 cm x103 salso a sequential equilibrium since x104 salt x105 salt x106 salt x107 salt x108 salt x109 cm x109 cm x110 salt x109 salt x110 salt x111 salt x111

If  $\epsilon$  is increased to  $\epsilon = 0.04$ , then the intervals  $\overline{M}_i(\omega_i)$  shift towards the middle and widen and tit-for-tat ceases to be equilibrium for any starting conditions. From Mailath and Morris (2002), we know that in this example, for sufficiently small  $\epsilon$ , tit-for-tat is an equilibrium, and obviously for sufficiently high  $\epsilon$ , it is not. Our analysis of this example allows us to go further: to establish exactly for which  $\epsilon$ 's the profile is an equilibrium. That is, our methods allow us

The Matlab code for checking arbitrary finite-state strategies for arbitrary games can be found on the authors' Web sites.

to consider whether any proposed strategy is an equilibrium strategy regardless of whether the outcomes are nearly public.

Next, rather than increasing  $\epsilon$  from  $\epsilon = 0.025$  to  $\epsilon = 0.04$ , instead consider keeping  $\epsilon = 0.025$  and decreasing the cost to cooperating from 0.4 to 0.357. Since the belief sets  $\overline{M}_i$  do not depend on pay-offs, they are still represented by Figure 1. Further, for these new pay-offs, incentives continue to hold at the extreme points of  $\overline{M}_i(R)$  and  $\overline{M}_i(P)$ , ensuring that letting the correlation device x on initial recommended states be the invariant distribution  $\pi = (0.659, 0.038, 0.038, 0.264)$  remains a correlated equilibrium. However, given this change in pay-offs, letting x be such that both players start off in state R with certainty is now no longer a sequential equilibrium. In fact for these pay-offs, the only sequential equilibrium associated with tit-for-tat is for both players to start off in state P with certainty, which delivers the worst pay-off over all ways of starting up a tit-for-tat equilibrium.

How can starting off with too much certainty be a problem? The difficulty with starting each player off in the reward state with certainty is that while each player is willing to cooperate in the first period, each is unwilling to defect in the second period, as tit-for-tat calls for, if he sees a bad outcome in the first period. The problem is that the certainty that one's opponent was in state R in the first period makes the player in the second period (after a bad outcome in the first period) insufficiently confident that his opponent is also in state P. In particular, his belief in period 2 that his opponent is in state R,  $B_i(m_{i,0} = 1, h_i = (C, B)|\psi_{-i}) = 0 \cdot 203$ , which is outside of  $\overline{M}_i(P) = [0.036, 0.189]$ . On the other hand, if the correlation device x = (0.8, 0.03, 0.03, 0.14) on the initial states  $\Omega = \{RR, RP, PR, PP\}$ , then if player i receives recommended state  $\omega_{i,0} = R$ , he believes his opponent is in state R with probability  $m_{i,0} = 0.8/0.83 = 0.964$ . Then,  $B_i(m_{i,0} = 0.964, h_i = (C, B)|\psi_{-i}) = 0.185$ , which is sufficiently low such that tit-for-tat is again a correlated equilibrium. (In fact, one can use our methods to find the correlation device x that delivers the best symmetric equilibrium pay-off associated with any given strategy. In this case, this is approximately x = (0.8, 0.03, 0.03, 0.04).)

Finally, in an online appendix, we demonstrate our methods are not confined to two-state strategies by considering for this game a strategy that we label "tit for tat-tat" (cooperate only if one has observed a good outcome in the last two periods). This is a three-state strategy that nevertheless is computed in seconds.

#### 4.2. Secret price cuts

In this section, we study a secret price cutting game with a rich action and signal space. First, we show that a natural strategy from the public-monitoring game, namely Taking Turns, is not going to work with private monitoring. Second, we show that one-period price wars can support collusion, but they may require random correlated starting conditions. Finally, we show an example with two-period price wars that support collusion, while one-period ones are not enough. In that example, if customer behaviour is more predictable, it is more difficult to sustain collusion in the private-monitoring case. It also suggests that strategies with two-period punishments are much more fragile to private monitoring than one-period punishments.

**4.2.1.** A Bertand pricing game. Consider a repeated Bertrand duopoly game. At each date, each of two players (firms) privately chooses a price  $a_i \in \{0, 0.01, 0.02, \dots, 4\}$ . A player's private outcome is his number of customers  $y_i \in Y_i = \{0, 1, 2, 3, 4, 5\}$ . With probability  $(1 - \epsilon)$ , the total number of customers,  $y_1 + y_2 = 5$ , and with probability  $\epsilon/10$ , the total number of customers is any particular element of  $\{0, 1, 2, 3, 4, 6, 7, 8, 9, 10\}$ . If both players choose the same price, each customer flips a fair coin to determine from which firm he buys. If the firms choose different prices, each customer chooses the lower price firm with probability  $1 - \delta$ . (If the total

number of customers is more than five, and these coin flips imply one player selling to more than five customers, that player is assumed to have exactly five customers, with the other player selling to the other customers.) Production is assumed to have a constant marginal cost  $c \ge 0$  so  $u_i(a_i, y_i) = (a_i - c) * y_i$ . If  $\delta = 0$ , and as the grid on prices gets infinitely fine, the unique stage game Nash equilibrium is for both firms to choose price  $a_i = c$ . If  $\epsilon$  and  $\delta$  are each strictly positive, all joint outcomes  $(y_1, y_2)$  occur with positive probability for all  $(a_1, a_2)$  and this game fits in our framework.

**4.2.2. Taking turns.** Consider the following three-state strategy: in state Me, player i chooses  $a_i = 3.99$ , while in state You, player i chooses  $a_i = 4$ . In state P (Punishment), player i chooses  $a_i = 0$ . If in state Me, player i receives 3 or more customers, he transits to state You, otherwise he transits to state P. If in state You, player i receives 2 or fewer customers, he transits to state Me, otherwise he transits to state P. Finally, if in state P, player i receives 0, 1, 4, or 5 customers, he stays in state P, if he receives 2 customers, he transits to state Me and if he receives 3 customers, he transits to state You.

If  $\beta = 0.95$ ,  $\delta = 0.05$ , and c = 1, for the game with public monitoring ( $\epsilon = 0$ ), this strategy is a perfect public equilibrium when one player starts in state Me and the other in state You. As long as the lower price firm gets a majority of the customers (a high probability event), both players choose a high price (with one slightly undercutting the other) and take turns regarding which one receives most of the customers. In the unlikely event that a firm receives a majority of the customers out of turn, a price war ensues. In a price war, each firm has the incentive to charge  $a_i = 0$  since this maximizes the probability that customers will be split as evenly as possible, causing the price war to end.

First, note that the conditions for Lemma 5 hold in this example, thus checking incentives at the extreme points of the largest fixed point of T,  $\overline{M}_i(\omega_i)$  is necessary and sufficient for the existence of starting conditions such that Taking Turns is a correlated equilibrium. But here, when  $\overline{M}_i(Me)$  and  $\overline{M}_i(P)$  are calculated, their intersection is non-empty. Thus, for the incentive conditions to be satisfied, each player must be indifferent between following the continuation strategy associated with state Me and the continuation strategy associated with state P for all points in this non-empty intersection, which is not the case here. One reason the non-empty intersection of  $\overline{M}_i(Me)$  and  $\overline{M}_i(P)$  occurs in this game is that if player 1 is in state Me and receives  $y_1 = 2$  customers, he transits to state P, while if he is in state P and receives  $p_1 = 2$  customers, he transits to state  $p_1 = 2$  outcomes, he will be in state  $p_2 = 2$  outcomes, he will be in state  $p_3 = 2$  outcomes, he will be in state  $p_4 = 2$  outcomes, he will be in

Such state-dependent transitions appear (at least to us) to be essential to any turn-taking equilibrium with public monitoring. That is, which outcomes require a transition to a given state would typically rely on whose turn it was to win the majority of customers last period (or whether the players are currently in the punishment state if such a state is also used). But, certainly for this example and we suspect more generally, these state-dependent transitions make the strategy not an equilibrium with private monitoring.

**4.2.3. High equal prices with price wars.** Now consider a different strategy. In state R (Reward), each firm chooses  $a_i = 4$  and in state P (Punish), each firm chooses  $a_i = 0$ . From any state, if  $y_i \in \{0, 5\}$  (a firm sells to either zero or five customers), it transits to state P in the next

period regardless of its price  $a_i$ . If  $y_i \in \{1, 2, 3, 4\}$ , from any state, it transits to state R tomorrow. In words, each firm sets a price of four unless last period it had an extreme number of customers. If  $\epsilon = 0$  or the total number of customers is certain to be five, this is a game of public monitoring, and this strategy is a public equilibrium as long as  $\delta$ , the probability that a customer chooses the high-price firm, is not too high (or for  $\beta$  near 1,  $\delta \le 0.06$ ).

If  $\epsilon \leq 0.04$  (with  $\beta = 0.95$ ,  $\delta = 0.05$ , and c = 1), unlike taking turns, there exists a correlation device such that this strategy is also an equilibrium of the private-monitoring game (specifically, drawing initial states from the unique invariant distribution, where joint state  $\omega \in \{RR, RP, PR, PP\}$  is drawn with probability (0.90, 0.01, 0.01, 0.08)). Interestingly, however, for these parameters, there exists no *deterministic* correlation device such that this is an equilibrium. Starting one player in state R and the other in state P is obviously not an equilibrium. However, for less obvious reasons, starting both in state R or both in state R is also not an equilibrium. For  $\epsilon = 0.04$ ,  $\overline{M}_i(R) = [0.263, 0.994]$  and  $\overline{M}_i(P) = [0.016, 0.124]$ , relatively wide but non-overlapping belief sets, and incentives hold on their extreme points. However, if both players start off in state R with certainty, while  $M_i^{*U}(P) = \overline{M}_i(P)$ ,  $M_i^{*U}(R) = [0.104, 1.000] \neq \overline{M}_i(R)$ . The interval  $M_i^{*U}(R)$  has not only a higher upper bound than  $\overline{M}_i(R)$ , but also a smaller lower bound. At this reduced lower bound, incentives do not hold.

Which histories create the problem? Specifically, the lower bound of  $M_i^{*U}(R)$  is generated by assuming player i believes his opponent is in state R with probability 1, sets  $a_i = 0$  and receives one customer (i.e.  $B_i(m_i = 1, h_i = (0, 1)|\psi_{-i}) = 0.104$ ). Bayesian updating essentially depends on reconciling the player's observations with its possible explanations and the most likely explanation for player i receiving only one customer when he undercut his opponent is that the total number of customers was actually only one and this customer chose the lower price, putting player -i in state P (which happens with probability 1 - 0.104). On the other hand, if player i is only 99.4% certain that his opponent is in state R (the upper bound of  $\overline{M}_i(R)$ ), then if he sets  $a_i = 0$  and receives one customer, he now believes his opponent is in state R with probability  $0.265 \in \overline{M}_i(R)$  and incentives hold. This change in updating occurs since the small amount of doubt leaves another explanation for player i receiving only one customer—his opponent was actually in state P and thus both set a price of zero, and thus it is more likely his opponent received a positive number of customers. A similar explanation rules out both players starting out in state P with certainty.

**4.2.4. Two-period price wars.** For this game, if the marginal cost of production c=0, one can show analytically that the two-state strategy considered in the previous section is not an equilibrium of the  $\epsilon=0$  public game. A price war of possibly only one period of zero profits (as opposed to negative profits if c>0) is an insufficient punishment to hinder slightly undercutting one's opponent. In this section, we show that a minimum two-period punishment can be an equilibrium, but that the co-ordination necessary for two-period punishments implies that the number of customers must be very close to public information.

Consider the following three-state strategy: In state R, each firm chooses  $a_i = 4$  and in states P1 and P2, each firm chooses  $a_i = 0$ . From any state, if  $y_i \in \{0, 5\}$  (a firm sells to either 0 or 5 customers), it transits to state P1 in the next period regardless of its price  $a_i$ . On the other hand, if  $y_i \in \{1, 2, 3, 4\}$ , it transits to state R tomorrow if today's state was R or P2 and transits to state P2 tomorrow if today's state was P1. In words, each firm sets a price of zero unless in each of the last two periods, it had an interior number of customers. If  $\epsilon = 0$ , or the total number of customers is certain to be 5, this is a game of public monitoring, and this strategy is a public equilibrium as long as  $\delta$ , the probability that a customer chooses the high-price firm, is not too high (or for  $\beta$  near 1,  $\delta \le 0.16$ ).

From Mailath and Morris (2002), we then know for any given  $\beta$  and  $\delta$ , there exists an  $\overline{\epsilon} > 0$  such that for all  $0 < \epsilon \le \overline{\epsilon}$ , this strategy is also an equilibrium of the private-monitoring game with an uncertain number of customers. However, assuming  $\beta = 0.95$ , if  $\delta = 0.1$  (or the customer chooses the lower price with probability 0.9), our computation method shows that for the above strategy to be an equilibrium, one needs  $\epsilon < 4 \times 10^{-7}$ , or there must be less than four chances in 10 million that the number of customers differs from five. For smaller  $\delta$  (or for higher probabilities that consumers choose the lower price),  $\epsilon$  must be even *lower*. If  $\delta = 0.05$  (or the customer chooses the lower price with probability 0.95), equilibrium requires  $\epsilon < 4 \times 10^{-9}$ , or there must be less than four chances in a billion that the number of customers differs from five.

The reason  $\epsilon$  must be so small (and small relative to  $\delta$ ) again comes from a player's off path Bayesian updating. For instance, suppose  $\epsilon$  and  $\epsilon/\delta$  are both positive but infinitesimal. Then, regardless of a player's action and regardless of his beliefs regarding his opponent's state (and thus his action) if he receives 0 or 5 customers, he concludes his opponent also received 0 or 5 customers, and if he receives one through 4 customers, he concludes his opponent did as well. This guarantees that regardless of starting states and actions taken, within two periods, each player is convinced the other player is in the same state he is. (More formally, in the limit as  $\epsilon \to 0$  for a given  $\delta > 0$ ,  $\overline{M}_i(R) = \{(1,0,0)\}$ ,  $\overline{M}_i(P1) = \{(0,1,0)\}$ , and  $\overline{M}_i(P2) = \{(0,0,1)\}$ .) On the other hand, if  $\epsilon$  and  $\delta/\epsilon$  are both positive and not infinitesimal, very different Bayesian updating occurs.

Suppose  $\delta = 0.1$  and  $\epsilon = 10^{-8}$  (which is too high for this strategy to be an equilibrium). What goes wrong? Again, one feature of our computation method is that it points out at exactly which state,  $\omega_i$ , and which extreme belief in  $\overline{M}_i(\omega_i)$  incentives fail to hold. For these parameters, incentives fail to hold for an extreme point in  $\overline{M}_i(P2)$  when player i believes his opponent is in state R with (approximately) 50% probability and state P2 with (approximately) 50% probability. Here, with this level of doubt, player i is unwilling to play  $a_i = 0$ , preferring a higher price.

Further, as in the previous example, our methods allow one to trace how an extreme belief can be supported. This particular extreme belief (player i is in state P2 but believes his opponent is 50/50 in R or P2) is generated as follows: suppose player i is in state R, believes his opponent is also in state R (with certainty), deviates and plays  $a_i = 0$ , and receives 0 customers, putting him in state P1 tomorrow. One possibility is that the number of customers was 5, but each of them chose the higher price firm. This happens with probability  $\delta^5 * (1 - \epsilon)$  which is about  $3 \cdot 1 \times 10^{-7}$ , or one in  $3 \cdot 1$  million. In this scenario, player i's opponent had 5 customers and is in state P1 tomorrow. A second possibility is that the number of customers was 1 and this single customer chose the higher price firm. This happens with probability  $\delta * (\epsilon/10)$ , which is  $1 \cdot 25 \times 10^{-6}$ , or one in eight hundred thousand. In this second scenario, player 1's opponent had one customer and is in state R tomorrow. The ratio of these events is  $0 \cdot 00016$  (or one in 625), which closely matches the actual posterior of player i given this scenario. And given he is in state P1 and believes his opponent is in state P1 with probability  $0 \cdot 99984$  and state R with probability  $0 \cdot 00016$ , he wishes to follow the strategy and play  $a_i = 0$ .

But from this state and belief, suppose player i then chooses an intermediate price  $a_i \in \{0.01, \ldots, 3.99\}$  and receives three customers, putting player i in state P2 the following period. How does he account for this event? One possibility is that his opponent was in state P1 (and thus played  $a_{-i} = 0$ ) and four out of five customers chose the higher price firm, putting player -i in state P2 tomorrow. This happens with probability  $0.999984*5*\delta^4*(1-\delta)*(1-\epsilon)$ , which is about 0.00003. Another possibility is that his opponent was in state R (and thus played  $a_{-i} = 4$ ) and only one out of five customers chose the higher price firm, putting player -i in state R tomorrow. This happens with probability  $0.00016*5*(1-\delta)^4*\delta*(1-\epsilon)$ , which is also

about 0.00003. Since the ratio of these two events is near 1, from state P2, player i now believes player -i is in state R with (about) 50% probability and state P2 with 50% probability.

#### 5. CONCLUDING REMARKS

Beyond using our methods directly to compute equilibria, one can extend and apply these methods in several ways.

First, as shown in a recent paper by Kandori and Obara (2010), one can use set-based methods similar to ours to study strategies that can be represented by finite automata on the equilibrium path but can be much more complicated off the equilibrium path. For example, they allow the strategy off the equilibrium path to be a function of beliefs over other players' states, which implies an infinite number of the automaton states (since players believe that others are always on the equilibrium path, the beliefs are still manageable).

Second, one can prove that if incentives hold strictly (uniformly bounded) for all extreme beliefs of the fixed point operator  $T^U$ , then this CSE is robust to small perturbations of the stage game pay-offs or the discount factor. The reasoning is as follows: first, the  $T^U$  operator and the initial belief sets  $M_{i,0}(x)$  are independent of the pay-offs. Hence, the fixed point is independent. Second, the incentive constraints are continuous in the stage-game pay-offs and the discount factor. Hence, if for the given game the incentives hold strictly for all extreme beliefs of the fixed point of the  $T^U$  operator, they also hold weakly for small perturbations of the pay-offs or the discount factor. Then, Theorem 1 implies that for the perturbed game, the same  $(x, \psi)$  are a CSE. Similar arguments can be used for perturbations of the monitoring technology (the P(y|a) function) to study robustness to changes in monitoring.

#### APPENDIX A

Proof of Lemma 2

*Proof.* First, recall that  $T(M_i)(\omega_i)$  is convex from the definition of T. Next, from its definition, we can express  $T(M_i)(\omega_i')$  as

$$T(M_i)(\omega_i') = \operatorname{co}(\cup_{\omega_i, h_i \in G_i(\omega_i, \omega_i' | \psi_i)} T(M_i)(\omega_i, h_i)(\omega_i')),$$

where  $T(M_i)(\omega_i,h_i)(\omega_i')=\{m_i'|$  there exists  $m_i\in M_i(\omega_i)$  such that  $m_i'=B_i(m_i,h_i|\psi_{-i})\}$ . Next, note that  $B_i(m_i,h_i|\psi_{-i})(\omega_i')$  is continuous in  $m_i$  on the whole domain  $m_i\in\Delta^{D_{-i}}$  and  $M_i(\omega_i)$  is closed (and bounded). Since  $T(M_i)(\omega_i,h_i)(\omega_i')$  is an image of a closed and bounded set under a continuous mapping, it is closed (and bounded) as well. As a finite union of closed sets,  $T(M_i)(\omega_i')$  is closed as well. The same reasoning applies to the  $T^U$  operator. The observation that if  $m_i$  is an extreme point of  $T^U(M_i)(\omega_i)$  but not  $T(M_i)(\omega_i)$ , then  $m_i$  is an extreme point of  $T^U$ .

For the last part of the lemma, we use an important property of the non-linear function  $B_i(m_i, h_i | \psi_{-i})(\omega_{-i})$ . For all  $\omega'_{-i}, m_i^2, m_i^2$ ,  $h_i$  and  $\alpha \in (0, 1)$ ,

$$B_i(\alpha m_i^1 + (1-\alpha) m_i^2, h_i | \psi_{-i})(\omega_{-i}') = \alpha' B_i(m_i^1, h_i | \psi_{-i})(\omega_{-i}') + (1-\alpha') B_i(m_i^2, h_i | \psi_{-i})(\omega_{-i}')$$

for some  $\alpha' \in (0,1)$ . That is, the posterior of a convex combination of beliefs  $m_i^1$  and  $m_i^2$  is a convex combination of their posteriors, albeit with different weights. To see this, algebraic manipulation delivers

$$\begin{split} B_{i}(\alpha m_{i}^{1} + (1-\alpha)m_{i}^{2}, h_{i}|\psi_{-i})(\omega'_{-i}) \\ &= \frac{\alpha \sum_{\omega_{-i}} m_{i}^{1}(\omega_{-i})F_{i}(\omega_{-i}, h_{i}|\psi_{-i})}{\sum_{\omega_{-i}} (\alpha m_{i}^{1}(\omega_{-i}) + (1-\alpha)m_{i}^{2}(\omega_{-i}))F_{i}(\omega_{-i}, h_{i}|\psi_{-i})} B_{i}(m_{i}^{1}, h_{i}|\psi_{-i})(\omega'_{-i}) \\ &+ \frac{(1-\alpha)\sum_{\omega_{-i}} m_{i}^{2}(\omega_{-i})F_{i}(\omega_{-i}, h_{i}|\psi_{-i})}{\sum_{\omega_{-i}} (\alpha m_{i}^{1}(\omega_{-i}) + (1-\alpha)m_{i}^{2}(\omega_{-i}))F_{i}(\omega_{-i}, h_{i}|\psi_{-i})} B_{i}(m_{i}^{2}, h_{i}|\psi_{-i})(\omega'_{-i}). \end{split}$$

Note

$$\frac{\alpha \sum_{\omega_{-i}} m_i^1(\omega_{-i}) F_i(\omega_{-i}, h_i | \psi_{-i})}{\sum_{\omega_{-i}} (\alpha m_i^1(\omega_{-i}) + (1 - \alpha) m_i^2(\omega_{-i})) F_i(\omega_{-i}, h_i | \psi_{-i})} + \frac{(1 - \alpha) \sum_{\omega_{-i}} m_i^2(\omega_{-i}) F_i(\omega_{-i}, h_i | \psi_{-i})}{\sum_{\omega_{-i}} (\alpha m_i^1(\omega_{-i}) + (1 - \alpha) m_i^2(\omega_{-i})) F_i(\omega_{-i}, h_i | \psi_{-i})} = 1.$$

Further, examination of the first quotient has the numerator strictly positive and strictly less than the denominator. So indeed

$$\alpha'(\alpha, m_i^1, m_i^2) = \frac{\alpha \sum_{\omega_{-i}} m_i^1(\omega_{-i}) F_i(\omega_{-i}, h_i | \psi_{-i})}{\sum_{\omega_{-i}} (\alpha m_i^1(\omega_{-i}) + (1 - \alpha) m_i^2(\omega_{-i})) F_i(\omega_{-i}, h_i | \psi_{-i})} \in (0, 1).$$

Now take any  $m_i$  which is an extreme point of  $T(M_i)(\omega_i)$  and suppose that for all collections  $(m_i', \omega_i', h_i')$  such that  $m_i = B_i(m_i', h_i'|\psi_{-i}), m_i' \in M_i(\omega_i')$  and  $h_i' \in G_i(\omega_i', \omega_i)$ , the belief  $m_i'$  is not an extreme point of  $M_i(\omega_i')$ . That implies that there exist two priors  $(m_i^0, m_i^1)$  that are extreme points of  $M_i(\omega_i)$  such that  $m_i'$  is a strict convex combination of them. There are three possibilities: (1)  $B_i(m_i', h_i'|\psi_{-i}) = B_i(m_i^0, h_i'|\psi_{-i})$  or (2)  $B_i(m_i', h_i'|\psi_{-i}) = B_i(m_i^1, h_i'|\psi_{-i})$  or (3)  $B_i(m_i', h_i'|\psi_{-i})$  is a strict convex combination of  $B_i(m_i^0, h_i'|\psi_{-i})$  and  $B_i(m_i^1, h_i'|\psi_{-i})$ . In the first two cases, we have then found the priors that lead to the posterior  $m_i$ , a contradiction. In the third case,  $m_i$  is not an extreme point of  $T(M_i)(\omega_i)$ , again a contradiction.

Proof of Lemma 3

*Proof.* For  $\omega_i$  such that  $\sum_{\overline{\omega}_{-i}} \pi(\omega_i, \overline{\omega}_{-i}) > 0$ , let  $m_i^0(\omega_i)(\omega_{-i}) = \frac{\pi(\omega_i, \omega_{-i})}{\sum_{\overline{\omega}_{-i}} \pi(\omega_i, \overline{\omega}_{-i})}$ . That is,  $m_i^0(\omega_i)$  is the single point in the set  $M_{i,0}(\pi,\omega_i)$ . Since  $\pi$  is an invariant distribution, for all  $\omega = (\omega_i, \omega_{-i})$ 

$$\begin{split} m_i^0(\omega_i)(\omega_{-i}) &= \frac{\sum_{\omega^0} \pi(\omega^0) \sum_{h_i \in G_i(\omega_i^0,\omega_i|\psi_i)} \sum_{h_{-i} \in G_i(\omega_{-i}^0,\omega_{-i}|\psi_{-i})} p_i(a_i|\omega_i^0) p_{-i}(a_{-i}|\omega_{-i}^0) P(y|a)}{\sum_{\omega^0} \pi(\omega^0) \sum_{h_i \in G_i(\omega_i^0,\omega_i|\psi_i)} \sum_{h_{-i}} p_i(a_i|\omega_i^0) p_{-i}(a_{-i}|\omega_{-i}^0) P(y|a)} \\ &= \frac{\sum_{\omega_i^0} \sum_{h_i \in G_i(\omega_i^0,\omega_i|\psi_i)} p_i(a_i|\omega_i^0) \sum_{\omega_{-i}^0} \pi(\omega^0) H_i(\omega_{-i}^0,\omega_{-i},h_i|\psi_{-i})}{\sum_{\omega_i^0} \sum_{h_i \in G_i(\omega_i^0,\omega_i|\psi_i)} p_i(a_i|\omega_i^0) \sum_{\omega_{-i}^0} \pi(\omega^0) F_i(\omega_{-i}^0,h_i|\psi_{-i})}. \end{split}$$

Next, note that

$$B_i(m_i^0(\omega_i^0),h_i|\psi_{-i})(\omega_{-i}) = \frac{\sum_{\omega_{-i}^0} \pi(\omega_i^0,\omega_{-i}^0) H_i(\omega_{-i}^0,\omega_{-i},h_i|\psi_{-i})}{\sum_{\omega^0} \pi(\omega_i^0,\omega_{-i}^0) F_i(\omega_{-i}^0,h_i|\psi_{-i})}.$$

We wish to show for all  $\omega_i$ ,  $m_i^0(\omega_i)$  is a convex combination of  $B_i(m_i^0, h_i|\psi_{-i})$  over all  $(\omega_i^0, h_i)$  such that  $h_i \in G_i(\omega_i^0, \omega_i|\psi_i)$ . For all  $(\omega_i^0, h_i)$  such that  $h_i \in G_i(\omega_i^0, \omega_i|\psi_i)$ , let

$$\alpha(\omega_i^0,h_i|\omega_i) = \frac{p_i(a_i|\omega_i^0)\sum_{\omega_{-i}^0}\pi(\omega_i^0,\omega_{-i}^0)F_i(\omega_{-i}^0,h_i|\psi_{-i})}{\sum_{\overline{\omega}_i^0}\sum_{\overline{h}_i\in G_i(\overline{\omega}_i^0,\omega_i|\psi_i)}p_i(\overline{a}_i|\overline{\omega}_i^0)\sum_{\omega_{-i}^0}\pi(\overline{\omega}_i^0,\omega_{-i}^0)F_i(\omega_{-i}^0,\overline{h}_i|\psi_{-i})}.$$

Since the denominator of  $\alpha(\omega_i^0,h_i|\omega_i)$  is the sum of the numerators over all  $(\omega_i^0,h_i)$  such that  $h_i \in G_i(\omega_i^0,\omega_i|\psi_i)$ , it is clear that  $\sum_{\omega_i} \sum_{h_i \in G_i(\omega_i^0,\omega_i|\psi_i)} \alpha(\omega_i^0,h_i|\omega_i) = 1$ .

Next, for a given  $\omega_i$  and  $\omega_{-i}$ , consider

$$\begin{split} \sum_{\omega_{i}^{0}} \sum_{h_{i} \in G_{i}(\omega_{i}^{0}, \omega_{i} | \psi_{i})} & \alpha(\omega_{i}^{0}, h_{i} | \omega_{i}) B_{i}(m_{i}^{0}(\omega_{i}), h_{i} | \psi_{-i})(\omega_{-i}) \\ &= \sum_{\omega_{i}^{0}} \sum_{h_{i} \in G_{i}(\omega_{i}^{0}, \omega_{i} | \psi_{i})} \frac{p_{i}(a_{i} | \omega_{i}^{0}) \sum_{\omega_{-i}^{0}} \pi(\omega_{i}^{0}, \omega_{-i}^{0}) F_{i}(\omega_{-i}^{0}, h_{i} | \psi_{-i}) B_{i}(m_{i}^{0}(\omega_{i}), h_{i} | \psi_{-i})(\omega_{-i})}{\sum_{\overline{\omega_{i}^{0}}} \sum_{\overline{h}_{i} \in G_{i}(\overline{\omega_{i}^{0}}, \omega_{i} | \psi_{i})} p_{i}(\overline{a}_{i} | \overline{\omega_{i}^{0}}) \sum_{\omega_{-i}^{0}} \pi(\overline{\omega_{i}^{0}}, \omega_{-i}^{0}) F_{i}(\omega_{-i}^{0}, \overline{h}_{i} | \psi_{-i})} \\ &= \frac{\sum_{\omega_{i}^{0}} \sum_{h_{i} \in G_{i}(\omega_{i}^{0}, \omega_{i} | \psi_{i})} p_{i}(a_{i} | \omega_{i}^{0}) \sum_{\omega_{-i}^{0}} \pi(\omega^{0}) H_{i}(\omega_{-i}^{0}, \omega_{-i}, h_{i} | \psi_{-i})}{\sum_{\omega_{i}^{0}} \sum_{h_{i} \in G_{i}(\omega_{i}^{0}, \omega_{i} | \psi_{i})} p_{i}(a_{i} | \omega_{i}^{0}) \sum_{\omega_{-i}^{0}} \pi(\omega^{0}) F_{i}(\omega_{-i}^{0}, h_{i} | \psi_{-i})} \\ &= m_{i}^{0}(\omega_{i})(\omega_{-i}). \end{split}$$

Proof of Lemma 4

*Proof.* First, that the limit exists and  $\pi$  is a stationary distribution of  $\tau$  is a standard result on Markov chains (see, e.g. Theorem 11.1 in Stokey and Lucas).

Next, define

$$\pi_t = \frac{1}{t+1} \sum_{n=0}^t x \tau^n.$$

Note that  $\pi_t$  is a probability distribution over joint states for any t (it is the distribution over joint states given

starting correlation device x and the transition matrix  $\tau$ , averaged over periods  $\{0, \dots, t\}$ ). We prove by induction that for all t,  $M_{i,0}(\pi_t) \subset (T^U)^t(M_{i,0}(x))$  and  $M_{i,0}(x\tau^t) \subset (T^U)^t(M_{i,0}(x))$  (where  $(T^U)^0$ (M) = M).

For t = 0, all these collections of sets are equal, so the claim is true. Now, suppose the claim is true for t - 1.

Let  $m_i^t(\omega_i)(\omega_{-i}) = \frac{\pi_t(\omega_i, \omega_{-i})}{\sum_{\overline{\omega}_{-i}} \pi_t(\omega_i, \overline{\omega}_{-i})}$  be the belief player i assigns to players -i being in state  $\omega_{-i}$  conditional on

observing that the correlation device  $\pi_t$  puts him in state  $\omega_i$ . Also let  $\hat{m}_i^t(\omega_i)(\omega_{-i}) = \frac{(x\tau^t)(\omega_i,\omega_{-i})}{\sum_{\overline{\omega}_i}(x\tau^t)(\omega_i,\overline{\omega}_{-i})}$  (analogous belief for correlation device  $x\tau^t$ ). Note that

$$\pi_t = \frac{\sum_{n=0}^t x \tau^n}{t+1} = \frac{t \pi_{t-1} + x \tau^t}{t+1},$$

that is,  $\pi_t$  is a weighted average of distributions  $\pi_{t-1}$  and  $x\tau^t$ .

By the same calculation as in Lemma 3,  $\hat{m}_i^t(\omega_i)(\omega_{-i})$  is a convex combination of posterior beliefs  $B_i(\hat{m}_i^{t-1}, h_i|\psi_{-i})$ over all  $(\omega_i^{t-1}, h_i)$  such that  $h_i \in G_i(\omega_i^{t-1}, \omega_i | \psi_{-i})$ . The intuition is that  $\hat{m}_i^t(\omega_i)(\omega_{-i})$  can be thought of as beliefs player i has after learning that at time t he is in state  $\omega_i$  but not knowing his history of the game so far. If he knew that his belief last period was  $\hat{m}_{i}^{t-1}$  he could then compute his posterior using that prior and averaging over all one-period

histories that according to the equilibrium path could have brought him to the current state  $\omega_i$ . Since by the inductive hypothesis all priors  $\hat{m}_i^{t-1}(\omega_i) \in (T^U)^{t-1}(M_{i,0}(x))(\omega_i)$ , all such posteriors  $\hat{m}_i^t(\omega_i) \in (T^U)^{t-1}(M_{i,0}(x))(\omega_i)$ , all such posteriors  $\hat{m}_i^t(\omega_i)$  $T((T^U)^{t-1}(M_{i,0}(x)))(\omega_i) \subset (T^U)^t(M_{i,0}(x))(\omega_i).$ 

Finally, since the correlation device  $\pi_t$  draws joint states either according to  $\pi_{t-1}$  (with probability  $\frac{t}{t+1}$ ) or  $x\tau^t$ (with probability  $\frac{1}{t+1}$ ), the posterior satisfies

$$\begin{split} m_i^t(\omega_i)(\omega_{-i}) &= \frac{\pi_t(\omega_i, \omega_{-i})}{\sum_{\overline{\omega}_{-i}} \pi_t(\omega_i, \overline{\omega}_{-i})} \\ &= \frac{\frac{t}{t+1} \pi_{t-1}(\omega_i, \omega_{-i}) + \frac{1}{t+1} (x \tau^t)(\omega_i, \omega_{-i})}{\sum_{\overline{\omega}_{-i}} \pi_t(\omega_i, \overline{\omega}_{-i})} \\ &= \frac{t}{t+1} \frac{\sum_{\overline{\omega}_{-i}} \pi_{t-1}(\omega_i, \overline{\omega}_{-i})}{\sum_{\overline{\omega}_{-i}} \pi_t(\omega_i, \overline{\omega}_{-i})} m_i^{t-1}(\omega_i)(\omega_{-i}) \\ &+ \frac{1}{t+1} \frac{\sum_{\overline{\omega}_{-i}} (x \tau^t)(\omega_i, \overline{\omega}_{-i})}{\sum_{\overline{\omega}_{-i}} \pi_t(\omega_i, \overline{\omega}_{-i})} \hat{m}_i^t(\omega_i)(\omega_{-i}). \end{split}$$

Since the coefficients on the two beliefs are positive and add up to one,  $m_i^t(\omega_i)(\omega_{-i})$  is a convex combination of the beliefs  $m_i^{t-1}(\omega_i)(\omega_{-i})$  and  $\hat{m}_i^t(\omega_i)(\omega_{-i})$ . Since we have shown that  $\hat{m}_i^t(\omega_i) \in (T^U)^t(M_{i,0}(x))(\omega_i)$  and by the inductive hypothesis,

$$m_i^{t-1}(\omega_i) \in (T^U)^{t-1}(M_{i,0}(x))(\omega_i) \subset (T^U)^t(M_{i,0}(x))(\omega_i),$$

we conclude that  $m_i^t(\omega_i) \subset (T^U)^t(M_{i,0}(x))(\omega_i)$ , which finishes the proof of induction.

As  $M_{i,0}(\pi_t) \subset (T^U)^t(M_{i,0}(x))$  for all t, it also holds in the limit, so indeed  $M_{i,0}(\pi) \subset M_i^{*U}(M_{i,0}(x))$ .

Proof of Lemma 5

*Proof.* That  $\tau(\psi)$  is a regular matrix implies that there exists an L such that for any joint states  $\omega$  and  $\omega'$ , the players on equilibrium path move with a positive probability from state  $\omega$  to  $\omega'$  in exactly L periods. That implies that for any non-empty  $M_i$  (i.e. that there exists at least one  $\omega_i$  such that  $M_i(\omega_i)$  is non-empty), the set  $T^n(M_i)(\omega_i)$  is non-empty for all  $\omega_i \in \Omega_i$  for any  $n \geq L$ .

Next, let  $\mathcal{H}(h_i)$  denote the  $D_{-i} \times D_{-i}$  matrix  $H_i(\omega_{-i}, \omega'_{-i}, h_i | \psi_{-i})$  where rows correspond to  $\omega_{-i}$  and the columns to  $\omega'_{-i}$ . We note that the matrix  $\mathcal{H}(h_i)$  has all entries between 0 and 1 and that the rows add up to at most 1, so that if some element is positive, all other elements are strictly bounded away from 1.

Since  $\tau(\psi)$  is a regular matrix and we have assumed that the set of signals players -i observe with positive probability does not depend on player i actions (full support) for all  $h_{i,1}, \ldots, h_{i,L}$  all elements of the matrix  $\mathcal{H}(h_{i,L}) * \cdots * \mathcal{H}(h_{i,1})$  contain no zeros (since player i assigns positive probability to the other players moving from any state to any state in L periods on the equilibrium path). Let  $\varepsilon > 0$  be the lower bound on the elements of that matrix (it exists since L and the set of  $h_i$  are finite).

The rest of the proof has two steps. Let beliefs  $m_i^{E0}$  and  $m_i^{E1}$  be such that  $m_i^{E0}(\omega_{-i}^0) = 1$  and  $m_i^{E1}(\omega_{-i}^1) = 1$ . That is,  $m_i^{E0}$  puts all probability on state  $\omega_{-i}^0$  and  $m_i^{E1}$  puts all weight on state  $\omega_{-i}^1$ . First, we show that for all  $\{h_{i,n}\}_{n=0}^{\infty}$ ,  $\lim_{n\to\infty} |B_i^n(m_i^{E0}, h_i^n|\psi_{-i}), B_i^n(m_i^{E1}, h_i^n|\psi_{-i})| = 0$ . Next, we show that this implies  $\lim_{n\to\infty} T^n(M_i) = \overline{M}_i$  for all non-empty  $M_i \in \mathcal{M}$ .

Step 1:

Recall from Lemma 1 that

$$B_i(m_i,h_i|\psi_{-i})(\omega_{-i}') = \frac{\sum_{\omega_{-i}} m_i(\omega_{-i}) H_i(\omega_{-i},\omega_{-i}',h_i|\psi_{-i})}{\sum_{\omega_{-i}} m_i(\omega_{-i}) F_i(\omega_{-i},h_i|\psi_{-i})}.$$

Let  $B_i(m_i, h_i | \psi_{-i})$  denote the vector  $B_i(m_i, h_i | \psi_i)(\omega'_{-i})$  and  $F_i(h_i | \psi_{-i})$  denote the vector  $F_i(\omega_{-i}, h_i | \psi_{-i})$ . We can then re-write Bayes' rule in the matrix form as

$$B_{i}(m_{i}, h_{i}|\psi_{-i}) = \underbrace{\frac{1}{m_{i} \cdot F_{i}(h_{i}|\psi_{-i})}}_{\text{scalar}} m_{i} \mathcal{H}(h_{i}), \tag{A.1}$$

where  $m_i$  is a row vector with elements  $m_i(\omega_{-i})$ .

If player i starts with prior  $m_i^0$  and observes  $(h_{i,L},\ldots,h_{i,1})$  (with  $h_{i,1}$  being the most recent observation), then his posterior beliefs after L periods are

$$\begin{split} B_i^L(m_i^0, h_{i,L}, \dots, h_{i,1} | \psi_{-i}) \\ &= \frac{1}{B_i^{L-1}(m_i^0, h_{i,L}, \dots, h_{i,2} | \psi_{-i}) \cdot F_i(h_{i,1} | \psi_{-i})} B_i^{L-1}(m_i^0, h_{i,L}, \dots, h_{i,2} | \psi_{-i}) \mathcal{H}(h_{i,1}) \\ &= \frac{1}{(m_i^0 \mathcal{H}(h_{i,L}) \dots \mathcal{H}(h_{i,2})) \cdot F_i(h_{i,1} | \psi_{-i})} m_i^0 \mathcal{H}(h_{i,L}) \dots \mathcal{H}(h_{i,1}). \end{split}$$

This implies that for  $j \in \{0,1\}$ ,  $B_i^L(m_i^{E_j},h_{i,L},\ldots,h_{i,1}|\psi_{-i})$  is equal to the  $\omega_{-i}^j$  row of matrix

$$\frac{1}{(m_i^{E_j}\mathcal{H}(h_{i,L}),\ldots,\mathcal{H}(h_{i,2}))\cdot F_i(h_{i,1}|\psi_{-i})}\mathcal{H}(h_{i,L}),\ldots,\mathcal{H}(h_{i,1}).$$

For a matrix Q, let  $R_l^Q = \sum_k q_{lk}$  be the sum of the elements of row l of this matrix. Denote by R(Q) a matrix obtained by dividing each element of matrix Q by the corresponding  $R_l^Q$ , that is, if B = R(Q), then  $b_{lk} = \frac{q_{lk}}{R_l^Q}$ . By definition, the rows of R(Q) add up to 1. Hence,  $R(\mathcal{H}(h_{i,L}), \dots, \mathcal{H}(h_{i,1}))$  is a probability matrix and the posterior belief  $B_l^L(m_i^{E0}, h_{i,L}, \dots, h_{i,1} | \psi_{-i})$  is equal to the  $\omega_{-i}^0$  row of  $R(\mathcal{H}(h_{i,L}), \dots, \mathcal{H}(h_{i,1}))$ .

Let  $d_k(Q)$  be the difference between the largest and smallest elements of Q's column k:  $d_k(Q) = \max_{l,j} (q_{lk} - q_{jk})$ ). Let d(Q) be the vector of these differences. Then  $\max_{\omega'_{-i}} d(R(\mathcal{H}(h_{i,L}), \ldots, \mathcal{H}(h_{i,1})))(\omega'_{-i})$  is the maximum distance of the posterior beliefs  $B_i^L(m_i^{E0}, h_{i,L}, \ldots, h_{i,1}|\psi_{-i})$  and  $B_i^L(m_i^{E1}, h_{i,L}, \ldots, h_{i,1}|\psi_{-i})$  over all extreme priors,  $m_i^{E0}$  and  $m_i^{E1}$ . To continue, we invoke the following technical lemma (proven below):

#### **Technical Lemma:**

Suppose that  $\{Q_n\}_{n=1}^{\infty}$  is a sequence of square matrices with all elements  $q_{nij} \in (\varepsilon, 1-\varepsilon)$  for some  $\varepsilon > 0$ . Then there exists a  $\delta \in (0,1)$  such that for every n

$$d(R(Q_n, ..., Q_1)) \le \delta d(R(Q_{n-1}, ..., Q_1)) \le \delta^{n-1} d(R(Q_1)),$$

i.e. the distance between the normalized rows of  $Q_n, \ldots, Q_1$  contracts by a factor of at least  $\delta$  as we left-multiply it by another matrix from the sequence.

Now, since there exists  $L \ge 1$  and  $\varepsilon > 0$  such that for all  $(h_{i,L}, \dots, h_{i,1})$ , all elements of  $\mathcal{H}(h_{i,L}), \dots, \mathcal{H}(h_{i,1})$  are bounded between  $(\varepsilon, 1 - \varepsilon)$ , this technical lemma implies that there exists a  $\delta \in (0,1)$  such that for any integer n:

$$d(R(\mathcal{H}(h_{i,nL}),\ldots,\mathcal{H}(h_{i,1}))) \leq \delta d(R(\mathcal{H}(h_{i,(n-1)L}),\ldots,\mathcal{H}(h_{i,1}))) \leq \delta^{n-1}\mathbf{1},$$

where 1 is a vector of ones (of length  $D_{-i}$ ). Therefore, for any  $\varepsilon'$ , we can find n large enough so that for any history of length nL and any two extreme priors,  $m_i^{E0}$  and  $m_i^{E1}$ , the distance between the posteriors will be less than  $\varepsilon'$ . So, for every history  $h_i^n$ , as  $n \to \infty$ , the posteriors converge to the same belief for all extreme priors.

Step 2:

As we have shown in the proof of Lemma 3, beliefs  $B_i(m_i^0, h_i|\psi_{-i})$  are a convex combination of beliefs  $B_i(m_i^E, h_i|\psi_{-i})$  of all extreme priors  $m_i^E$ . Applying this reasoning iteratively (that if prior belief  $m_i$  is a convex combination of priors  $m_i'$  and  $m_i''$ , then after applying  $B_i$ , the posterior of  $m_i$  is a convex combination of the posteriors of  $m_i'$  and  $m_i''$ ), we get that for any history sequence, the posteriors after all possible beliefs are convex combinations of posteriors  $B_i^L(m_i^E, h_{i,L}, \ldots, h_{i,1}|\psi_{-i})$ . Since for any sequence  $\{h_i^L\}_{L=1}^{\infty}$ , for all  $m_i^E$ , the posteriors  $B_i^L(m_i^E, h_{i,L}, \ldots, h_{i,1}|\psi_{-i})$  converge, the same is true for posteriors after arbitrary priors. In other words, after long enough histories, the posteriors depend (almost) only on the history and not on the prior.

As we described in the text, by the Tarski's fixed point theorem, T has at least one fixed point,  $\overline{M}_i$ . Now, suppose that there exists a collection of sets  $M_i^0$  such that  $\lim_{n\to\infty} T^n(M_i^0) \neq \overline{M}_i$  (either because the sequence  $\{T^n(M_i^0)\}_{n=0}^{\infty}$  converges to something else or does not converge at all).

By monotonicity of T, for all n,  $T^n(M_i^0) \subset T^n(\overline{\Delta}_i)$ . Since  $T^n(\overline{\Delta}_i)$  converges to  $\overline{M}_i$ , for any  $\varepsilon > 0$ , we can find n large enough so that for all  $\omega_i \in \Omega_i$  and all  $m_i \in T^n(M_i^0)(\omega_i)$ ,  $|m_i, \overline{M}_i(\omega_i)| < \varepsilon$ . That is, the sets  $T^n(M_i^0)$  cannot "stick out" of  $\overline{M}_i$  in the limit.

So the only remaining possibility for  $\lim_{n\to\infty} T^n(M_i^0) \neq \overline{M_i}$  is that there exists  $\varepsilon>0$  such that for all n', we can have that  $n\geq n'$  and a state  $\omega_i^n$  such that  $\max_{m_i\in\overline{M_i}(\omega_i^n)}|T^n(M_i^0)(\omega_i^n),m_i|>\varepsilon$  (in words, that the set  $\overline{M_i}(\omega_i^n)$  strictly "sticks out" of the set  $T^n(M_i^0)(\omega_i^n)$  even for arbitrarily large n). If so, then we can find an extreme belief  $m_i^n\in\overline{M_i}(\omega_i^n)$  that satisfies  $|m_i^n,T^n(M_i^0)(\omega_i^n)|>0$ . Fix n' such that the distance between  $B_i^n(m_i^{E0},h_i^n|\psi_{-i})$  and  $B_i^n(m_i^{E1},h_i^n|\psi_{-i})$  is uniformly bounded by  $\varepsilon/2$  for all histories  $h_i^n$  (for all n>n') and all extreme points  $m_i^{E0},m_i^{E1}$ . Since  $\lim_{n\to\infty}T^n(\overline{\Delta_i})=\overline{M_i}$ , we can find a history  $h_i^n$  and a prior  $m_i^{E0}$  such that  $|B_i^n(m_i^{E0},h_i^n|\psi_{-i}),m_i^n|\leq \varepsilon/2$  and a starting state  $\omega_i^0$  such that after that history, player i is in the state  $\omega_i^n$ . Now, take any prior  $m_i^0\in M_i^0(\omega_i^0)$ . It is a convex combination of the priors  $m_i^E$ . Moreover, after the history  $h_i^n$ , the posterior  $B_i^n(m_i^0,h_i^n|\psi_{-i})\in T^n(M_i^0)(\omega_i^n)$  and it is a convex combination of the posteriors  $B_i^n(m_i^E,h_i^n|\psi_{-i})$ . (The last claim follows from inspection of (A1)—see also Lemma 2.) Therefore,

$$|B_i^n(m_i^0,h_i^n|\psi_{-i}),B_i^n(m_i^{E0},h_i^n|\psi_{-i})| \leq \max_{m_i^{E1},m_i^{E2}}|B_i^n(m_i^{E1},h_i^n|\psi_{-i}),B_i^n(m_i^{E2},h_i^n|\psi_{-i})| \leq \varepsilon/2.$$

Using the triangle inequality,  $|B_i^n(m_i^0, h_i^n|\psi_{-i}), m_i^n| \le \varepsilon$  but that contradicts that  $|m_i^n, T^n(M_i^0)(\omega_i^n)| > \varepsilon$ .

Proof of Technical Lemma

*Proof.* Consider a general multiplication:  $Q = Q_n, \dots, Q_1$ . Let  $C = Q_n, F = Q_{n-1}, B = Q_{n-2}, \dots, Q_1$ . Also, let G = FB, so that Q = CG = CFB. By assumption all the elements of C and F are bounded from below by  $\varepsilon > 0$ , but we do not know that about B or G.

For arbitrary matrix A, let  $R_k^A$  be the sum of elements in row k of that matrix. Then

$$R_i^Q = \sum_j q_{ij} = \sum_j \left(\sum_k c_{ik} g_{kj}\right) = \sum_k c_{ik} \sum_j g_{kj} = \sum_k c_{ik} R_k^G.$$

Moreover,

$$\frac{q_{ij}}{R_i^Q} = \sum_k \Gamma_k^i \frac{g_{kj}}{R_k^G},$$

where

$$\Gamma_k^i = \frac{c_{ik} R_k^G}{\sum_l c_{il} R_l^G}.$$

In words, the elements of  $R(Q_nG)$  are a weighted average of elements of R(G) (note that  $\sum_k \Gamma_k^i = 1$ ). We now bound the weights  $\Gamma_k^i$  uniformly away from zero for all G. To this end, bound

$$\Gamma_k^i = \frac{c_{ik} R_k^G}{\sum_l c_{il} R_l^G} > c_{ik} \frac{R_k^G}{\sum_l R_l^G}.$$

Next,

$$\begin{split} \frac{R_i^G}{\sum_{l} R_l^G} &= \frac{\sum_{k} f_{ik} R_k^B}{\sum_{l} \sum_{k} f_{lk} R_k^B} = \frac{\sum_{k} f_{ik} R_k^B}{\sum_{k} \sum_{l} f_{lk} R_k^B} = \frac{\sum_{k} f_{ik} R_k^B}{\sum_{k} R_k^B L_k^F} \\ &= \sum_{k} \frac{f_{ik}}{L_k^F} \frac{L_k^F R_k^B}{\sum_{k} R_k^B L_k^F} = \sum_{k} \frac{f_{ik}}{L_k^F} \gamma_k, \end{split}$$

where  $L_k^F$  is the sum of elements of column k of matrix F and

$$\gamma_{k} = \frac{L_{k}^{F} R_{k}^{B}}{\sum_{k} R_{k}^{B} L_{k}^{F}} \in [0, 1].$$

Note that for any matrices F and B,  $\sum_{k} \gamma_{k} = 1$ .

Therefore, we can find a bound  $\varepsilon_L \in (0, \frac{1}{2})$  that depends only on F and C:

$$\Gamma_k^i \ge c_{ik} \frac{R_k^G}{\sum_l R_l^G} \ge \varepsilon \min_k \frac{f_{ik}}{L_k^F} > \varepsilon_L,$$

where  $\varepsilon_L$  can be chosen independently of i and k.

To finish the proof, we show how to choose  $\delta$ . Consider any column k. Any element of column k of matrix  $R(Q_n,\ldots,Q_1)$  is a weighted average of elements in the same column of  $R(Q_{n-1},\ldots,Q_1)$ , with the weights bounded uniformly away from zero by  $\varepsilon_L$ . Suppose that the largest and smallest elements of column k of  $R(Q_{n-1},\ldots,Q_1)$  are equal to  $q_h$  and  $q_l$ , respectively. Then

$$d_k(R(Q_n,\ldots,Q_1)) \leq (1-\varepsilon_L)q_h + \varepsilon_Lq_l - (\varepsilon_Lq_h + (1-\varepsilon_L)q_l) = (1-2\varepsilon_L)d_k(R(Q_{n-1},\ldots,Q_1)).$$

So we can pick  $\delta = (1 - 2\varepsilon_L)$ .

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#### Supplementary Data

Supplementary data are available at Review of Economic Studies online.

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