

A low-angle photograph of the Statue of Liberty against a clear blue sky. The statue's right arm is raised, holding the torch with the flame. The crown with its seven spikes is visible. The statue's face is in profile, looking upwards and to the left.

# Independence

CS 109  
Lecture 5  
April 6th, 2016

# Today's Topics

**Last time:**

Conditional Probability

Bayes Theorem

**Today:**

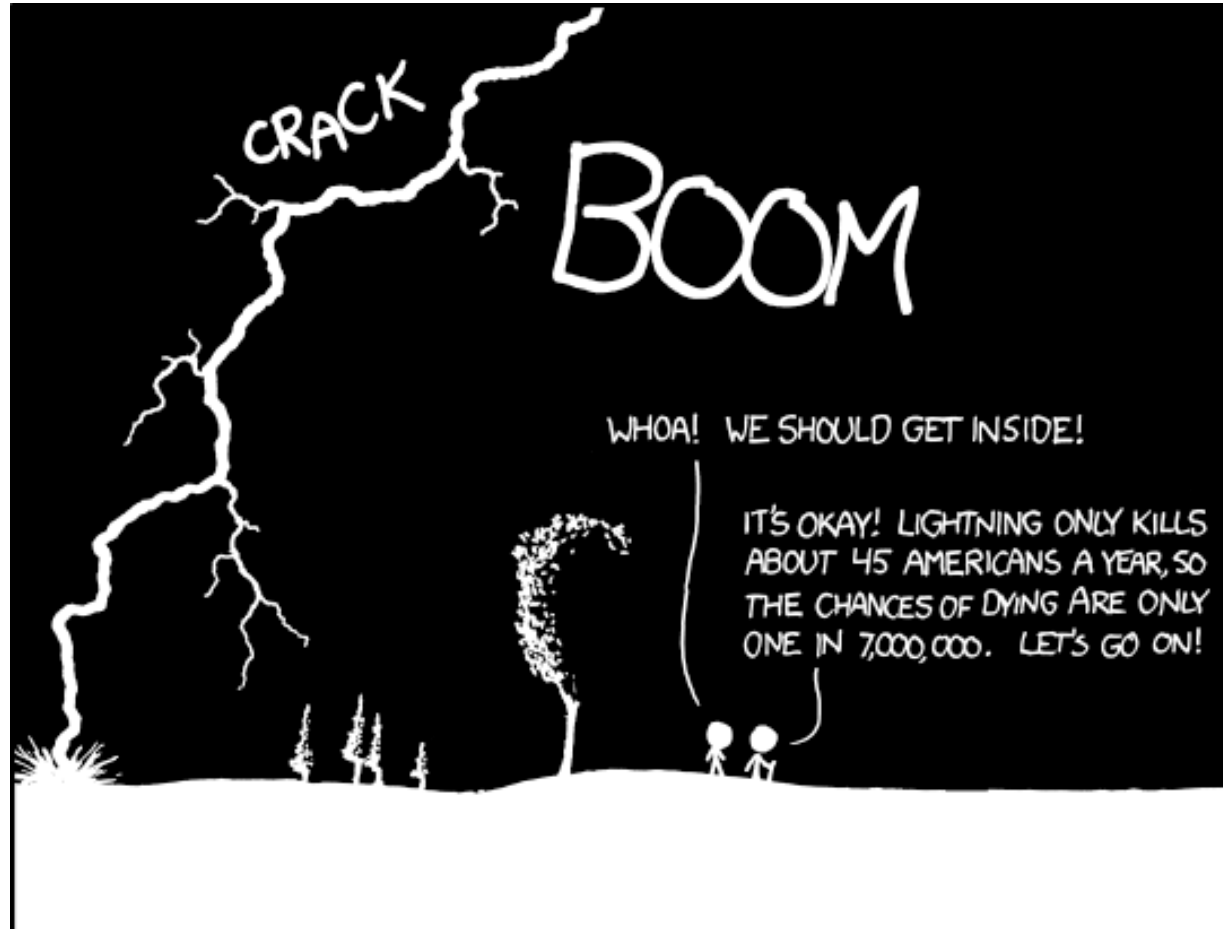
Independence

Conditional Independence

**Next time:**

Random Variables

# The Tragedy of Conditional Prob



THE ANNUAL DEATH RATE AMONG PEOPLE  
WHO KNOW THAT STATISTIC IS ONE IN SIX.

Thanks xkcd! <http://xkcd.com/795/>

# A Few Useful Formulas

- For any events A and B:

$$P(A \cap B) = P(B \cap A) \quad (\text{Commutativity})$$

$$\begin{aligned} P(A \cap B) &= P(A \mid B) P(B) && (\text{Chain rule}) \\ &= P(B \mid A) P(A) \end{aligned}$$

$$P(A \cap B^c) = P(A) - P(A \cap B) \quad (\text{Intersection})$$

$$P(A \cap B) \geq P(A) + P(B) - 1 \quad (\text{Bonferroni})$$

# Generality of Conditional Probability

- For any events A, B, and E, you can condition consistently on E, and these formulas still hold:

$$P(A \cap B \mid E) = P(B \cap A \mid E)$$

$$P(A \cap B \mid E) = P(A \mid B \cap E) P(B \mid E)$$

$$P(A \mid B \cap E) = \frac{P(B \mid A \cap E) P(A \mid E)}{P(B \mid E)} \quad (\text{Bayes' Thm.})$$

# BAE's Theorem?

$$P(A | B \ E) = \frac{P(B | A \ E) P(A | E)}{P(B | E)}$$



# Generality of Conditional Probability

- For any events A, B, and E, you can condition consistently on E, and these formulas still hold:

$$P(A \cap B \mid E) = P(B \cap A \mid E)$$

$$P(A \cap B \mid E) = P(A \mid B \cap E) P(B \mid E)$$

$$P(A \mid B \cap E) = \frac{P(B \mid A \cap E) P(A \mid E)}{P(B \mid E)} \quad (\text{Bayes' Thm.})$$

- Can think of E as “everything you already know”
- Formally,  $P(\bullet \mid E)$  satisfies 3 axioms of probability

# Our Still Misunderstood Friend

- Roll two 6-sided dice, yielding values  $D_1$  and  $D_2$ 
  - Let E be event:  $D_1 = 1$
  - Let F be event:  $D_2 = 1$
- What is  $P(E)$ ,  $P(F)$ , and  $P(EF)$ ?
  - $P(E) = 1/6$ ,  $P(F) = 1/6$ ,  $P(EF) = 1/36$
  - $P(EF) = P(E) P(F) \rightarrow$  E and F independent
- Let G be event:  $D_1 + D_2 = 5$   $\{(1, 4), (2, 3), (3, 2), (4, 1)\}$
- What is  $P(E)$ ,  $P(G)$ , and  $P(EG)$ ?
  - $P(E) = 1/6$ ,  $P(G) = 4/36 = 1/9$ ,  $P(EG) = 1/36$
  - $P(EG) \neq P(E) P(G) \rightarrow$  E and G dependent



# Independence

- Two events E and F are called **independent** if:

$$P(EF) = P(E) P(F)$$

Or, equivalently:  $P(E | F) = P(E)$

- Otherwise, they are called **dependent** events

- Three events E, F, and G independent if:

$$P(EFG) = P(E) P(F) P(G), \text{ and}$$

$$P(EF) = P(E) P(F), \text{ and}$$

$$P(EG) = P(E) P(G), \text{ and}$$

$$P(FG) = P(F) P(G)$$

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^C$  are independent

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent

- Proof:

$$P(E \cap F^c)$$

We want to show  
that this is equal  
to  $P(E)P(F^c)$

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent

- Proof:

$$P(E \cap F^c) = P(E) - P(EF)$$

$$\text{Since } P(E) = P(EF^c) + P(EF)$$

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^C$  are independent
- Proof:

$$\begin{aligned}P(E \cap F^C) &= P(E) - P(EF) \\ &= P(E) - P(E)P(F)\end{aligned}$$

Since we are told  
 $E$  and  $F$  are  
independent

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent
- Proof:

$$\begin{aligned}P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E)P(F) \\ &= P(E) [1 - P(F)]\end{aligned}$$

Factoring!

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent
- Proof:

$$\begin{aligned}P(E \cap F^c) &= P(E) - P(E \cap F) \\&= P(E) - P(E) P(F) \\&= P(E) [1 - P(F)] \\&= P(E) P(F^c)\end{aligned}$$

Yep, that's the complement

# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent
- Proof:

$$\begin{aligned}P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E)P(F) \\ &= P(E) [1 - P(F)] \\ &= P(E) P(F^c)\end{aligned}$$

So,  $E$  and  $F^c$  independent, implying that:

$$P(E | F^c) = P(E) = P(E | F)$$



# Independence

- Given independent events  $E$  and  $F$ , prove that  $E$  and  $F^c$  are independent

- Proof:

$$\begin{aligned}P(E \cap F^c) &= P(E) - P(EF) \\ &= P(E) - P(E)P(F) \\ &= P(E) [1 - P(F)] \\ &= P(E) P(F^c)\end{aligned}$$

So,  $E$  and  $F^c$  independent, implying that:

$$P(E | F^c) = P(E) = P(E | F)$$

- Intuitively, if  $E$  and  $F$  are independent, knowing whether  $F$  holds gives us no information about  $E$

# Generalized Independence

- General definition of Independence:

Events  $E_1, E_2, \dots, E_n$  are independent if for every subset with  $r$  elements (where  $r \leq n$ ) it holds that:

$$P(E_1, E_2, E_3, \dots, E_r) = P(E_1)P(E_2)P(E_3) \dots P(E_r)$$

- Example: outcomes of  $n$  separate flips of a coin are all independent of one another
  - Each flip in this case is called a “trial” of the experiment

# Two Dice

- Roll two 6-sided dice, yielding values  $D_1$  and  $D_2$ 
  - Let E be event:  $D_1 = 1$
  - Let F be event:  $D_2 = 6$
  - Are E and F independent? **Yes!**
- Let G be event:  $D_1 + D_2 = 7$ 
  - Are E and G independent? **Yes!**
  - $P(E) = 1/6$ ,  $P(G) = 1/6$ ,  $P(E \cap G) = 1/36$  [roll (1, 6)]
  - Are F and G independent? **Yes!**
  - $P(F) = 1/6$ ,  $P(G) = 1/6$ ,  $P(F \cap G) = 1/36$  [roll (1, 6)]
  - Are E, F and G independent? **No!**
  - $P(EFG) = 1/36 \neq 1/216 = (1/6)(1/6)(1/6)$

# Generating Random Bits

- A computer produces a series of random bits, with probability  $p$  of producing a 1.
  - Each bit generated is an independent trial
  - $E$  = first  $n$  bits are 1's, followed by a single 0
  - What is  $P(E)$ ?
- Solution
  - $P(\text{first } n \text{ 1's}) = P(1^{\text{st}} \text{ bit}=1) P(2^{\text{nd}} \text{ bit}=1) \dots P(n^{\text{th}} \text{ bit}=1)$   
 $= p^n$
  - $P(n+1 \text{ bit}=0) = (1 - p)$
  - $P(E) = P(\text{first } n \text{ 1's}) P(n+1 \text{ bit}=0) = p^n (1 - p)$

# Coin Flips

- Say a coin comes up heads with probability  $p$ 
  - Each coin flip is an independent trial
- $P(n \text{ heads on } n \text{ coin flips}) = p^n$
- $P(n \text{ tails on } n \text{ coin flips}) = (1 - p)^n$
- $P(\text{first } k \text{ heads, then } n - k \text{ tails}) = p^k (1 - p)^{n-k}$
- $P(\text{exactly } k \text{ heads on } n \text{ coin flips}) = ?$

# Explain...

$$P(\text{exactly } k \text{ heads on } n \text{ coin flips})? \quad \binom{n}{k} p^k (1-p)^{n-k}$$

---

Think of the flips as ordered:

Ordering 1: T, H, H, T, T, T....

Ordering 2: H, T, H, T, T, T....

And so on...

The coin flips are independent!

$$P(F_i) = p^k (1-p)^{n-k}$$

Let's make each ordering with  $k$  heads an event...  $F_i$

---

$P(\text{exactly } k \text{ heads on } n \text{ coin flips}) = P(\text{any one of the events})$


$P(\text{exactly } k \text{ heads on } n \text{ coin flips}) = P(F_1 \text{ or } F_2 \text{ or } F_3 \dots)$

Those events are mutually exclusive!

# Moment of Crystallization

Add vs Multiply?



A promotional image for the movie 'Batman v Superman: Dawn of Justice'. It shows Superman on the left, seen from the back, wearing his blue suit and red cape. He is looking towards Batman on the right, who is in his black tactical suit and cowl, with his right arm raised in a fist. The background is a dark, blue, textured surface with some debris or particles floating around. A bright light source is visible on the right side, creating a lens flare effect.

# Batman vs Superman



**COMING SOON**

**#BATMAN v SUPERMAN**



WARRNER BROS. PICTURES  
LIONS GATE FILMS

TM & © DC COMICS

SEE IT IN 3D

Warner Bros. Pictures  
Lions Gate Films

WARRNER BROS. PICTURES  
LIONS GATE FILMS

# Add vs Multiply



+

VS

x

# Add vs Multiply

## Probabilities

$P(AB)$	Generally:	$P(A)P(B A)$	multiply 
	Independent:	$P(A)P(B)$	
$P(A \cup B)$	Generally:	$P(A) + P(B) - P(AB)$	add 
	Mutually Exclusive:	$P(A) + P(B)$	

# Add vs Multiply

## Counting

A then B	Generally:	<i>Not handled</i>	<i>multiply</i>
	Independent:	$ A  B $	
$A \cup B$	Generally:	$ A  +  B  -  AB $	
	Mutually Exclusive:	$ A  +  B $	<i>add</i>

Next up...

# And vs Condition

$P(AB)$  vs  $P(A|B)$

$$P(AB) = P(A|B)P(B)$$

# Hash Tables

- $m$  strings are hashed (unequally) into a hash table with  $n$  buckets
  - Each string hashed is an independent trial, with probability  $p_i$  of getting hashed to bucket  $i$
  - $E$  = at least one string hashed to first bucket
  - What is  $P(E)$ ?
- Solution

To the chalk board!



# Yet More Hash Tables

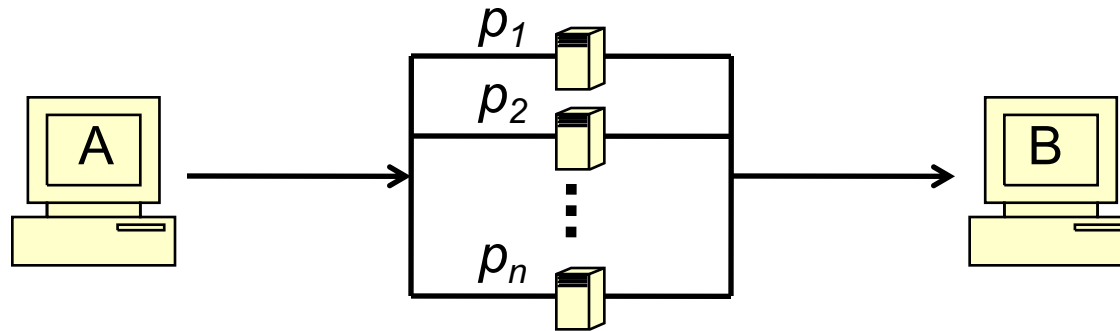
- $m$  strings are hashed (unequally) into a hash table with  $n$  buckets
  - Each string hashed is an independent trial, with probability  $p_i$  of getting hashed to bucket  $i$
  - $E =$  At least 1 of buckets 1 to  $k$  has  $\geq 1$  string hashed to it
- Solution
  - $F_i =$  at least one string hashed into  $i$ -th bucket
  - $P(E) = P(F_1 \cup F_2 \cup \dots \cup F_k) = 1 - P((F_1 \cup F_2 \cup \dots \cup F_k)^c)$   
 $= 1 - P(F_1^c F_2^c \dots F_k^c)$  (DeMorgan's Law)
  - $P(F_1^c F_2^c \dots F_k^c) = P(\text{no strings hashed to buckets 1 to } k)$   
 $= (1 - p_1 - p_2 - \dots - p_k)^m$
  - $P(E) = 1 - (1 - p_1 - p_2 - \dots - p_k)^m$

# No, Really, More Hash Tables

- $m$  strings are hashed (unequally) into a hash table with  $n$  buckets
  - Each string hashed is an independent trial, with probability  $p_i$  of getting hashed to bucket  $i$
  - $E =$  Each of buckets 1 to  $k$  has  $\geq 1$  string hashed to it
- Solution
  - $F_i =$  at least one string hashed into  $i$ -th bucket
  - $P(E) = P(F_1 F_2 \dots F_k) = 1 - P((F_1 F_2 \dots F_k)^c)$   
 $= 1 - P(F_1^c \cup F_2^c \cup \dots \cup F_k^c)$  (DeMorgan's Law)  
 $= 1 - P\left(\bigcup_{i=1}^k F_i^c\right) = 1 - \sum_{r=1}^k (-1)^{(r+1)} \sum_{i_1 < \dots < i_r} P(F_{i_1}^c F_{i_2}^c \dots F_{i_r}^c)$   
where  $P(F_{i_1}^c F_{i_2}^c \dots F_{i_r}^c) = (1 - p_{i_1} - p_{i_2} - \dots - p_{i_r})^m$

# Sending a Message Through Network

- Consider the following parallel network:



- $n$  independent routers, each with probability  $p_i$  of functioning (where  $1 \leq i \leq n$ )
  - $E$  = functional path from A to B exists. What is  $P(E)$ ?

- Solution:

- $$P(E) = 1 - P(\text{all routers fail})$$
$$= 1 - (1 - p_1)(1 - p_2)\dots(1 - p_n)$$
$$= 1 - \prod_{i=1}^n (1 - p_i)$$

Phew...

2 min pedagogical pause



# Digging Deeper on Independence

- Recall, two events  $E$  and  $F$  are called independent if

$$P(EF) = P(E) P(F)$$

- If  $E$  and  $F$  are independent, does that tell us whether the following is true or not:

$$P(EF \mid G) = P(E \mid G) P(F \mid G),$$

where  $G$  is an arbitrary event?

- In general, No!

# Not So Independent Dice

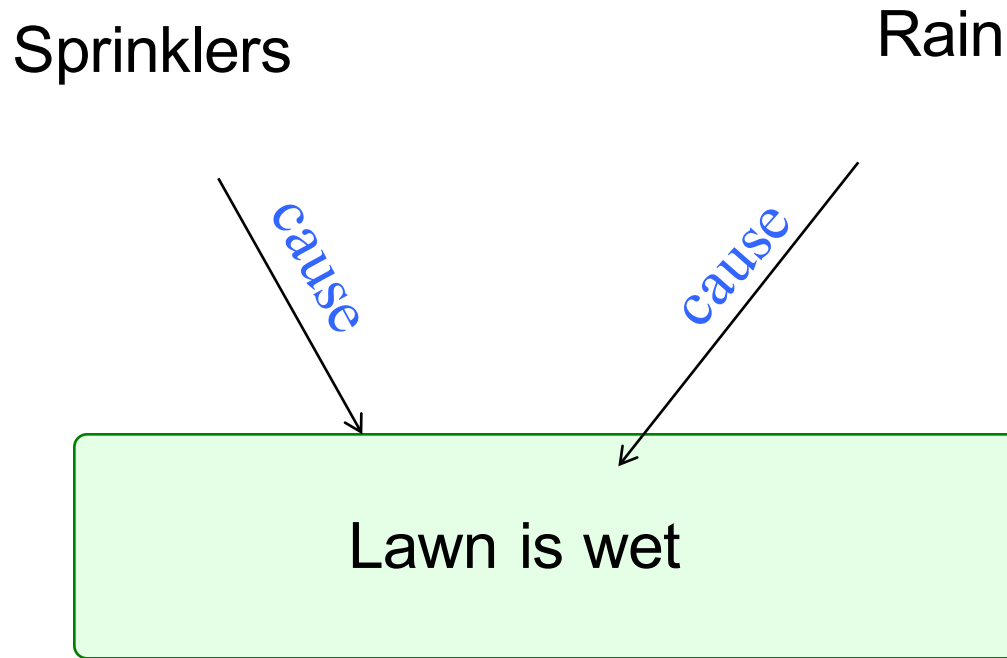
- Roll two 6-sided dice, yielding values  $D_1$  and  $D_2$ 
  - Let E be event:  $D_1 = 1$
  - Let F be event:  $D_2 = 6$
  - Let G be event:  $D_1 + D_2 = 7$
- E and F are independent
  - $P(E) = 1/6$ ,  $P(F) = 1/6$ ,  $P(EF) = 1/36$
- Now condition both E and F on G:
  - $P(E|G) = 1/6$ ,  $P(F|G) = 1/6$ ,  $P(EF|G) = 1/6$
  - $P(EF|G) \neq P(E|G) P(F|G) \rightarrow E|G$  and  $F|G$  dependent
- Independent events can become dependent by conditioning on additional information



# Explaining Away

- Say you have a lawn
  - It gets watered by rain or sprinklers
  - $P(\text{rain})$  and  $P(\text{sprinklers were on})$  are independent
  - Now, you come outside and see the grass is wet
    - You know that the sprinklers were on
    - Does that lower probability that rain was cause of wet grass?
  - This phenomena is called “explaining away”
    - One cause of an observation makes other causes less likely

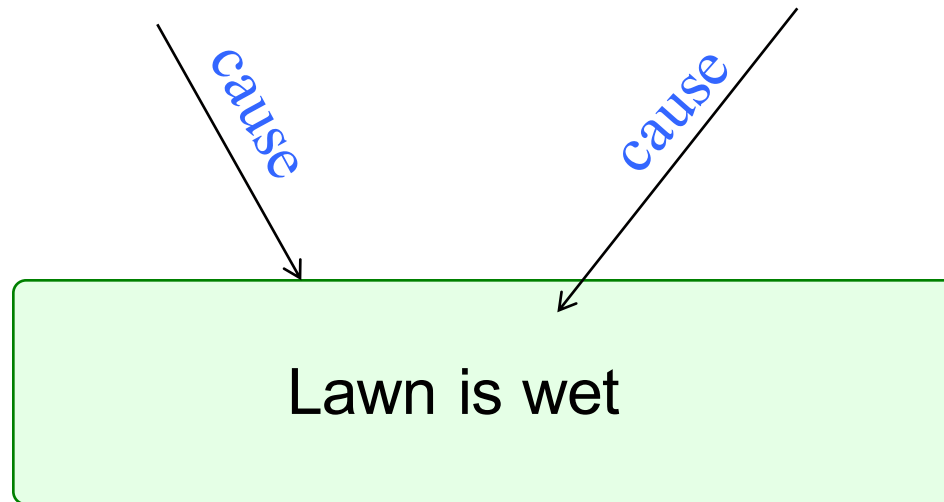
# Explaining Away



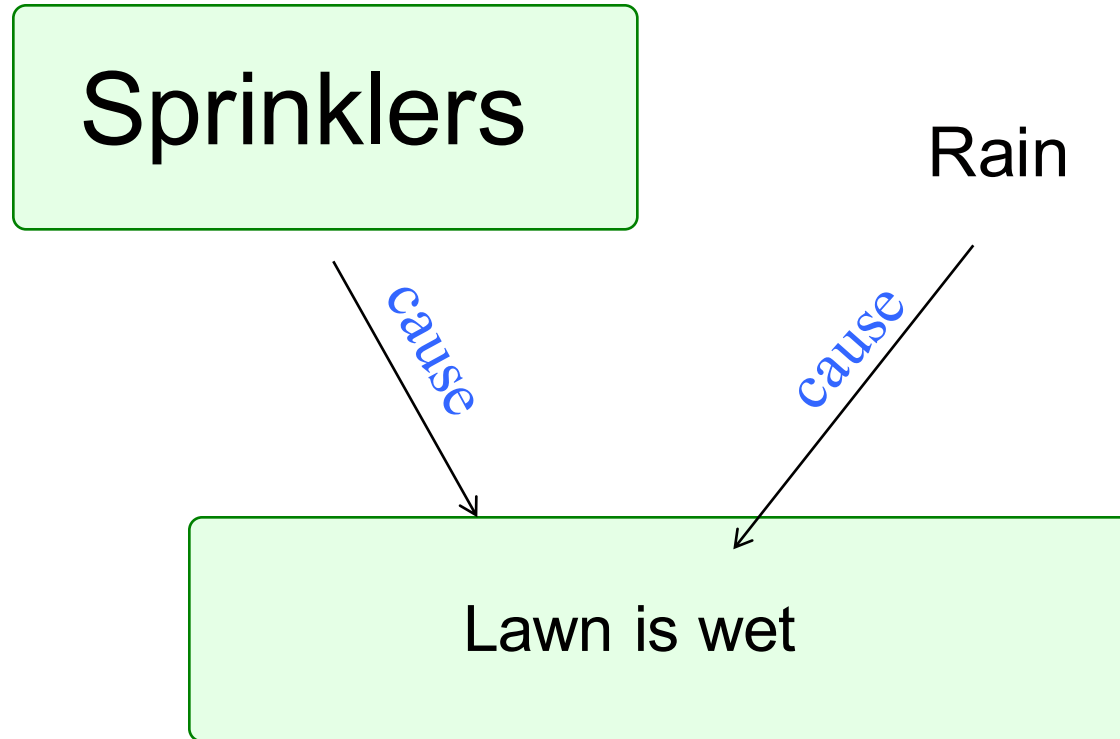
# Explaining Away

Sprinklers

Rain



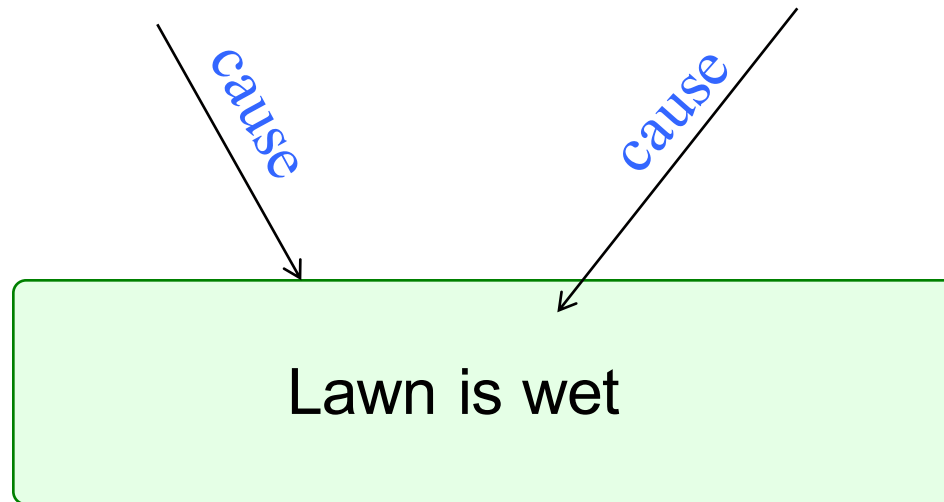
# Explaining Away



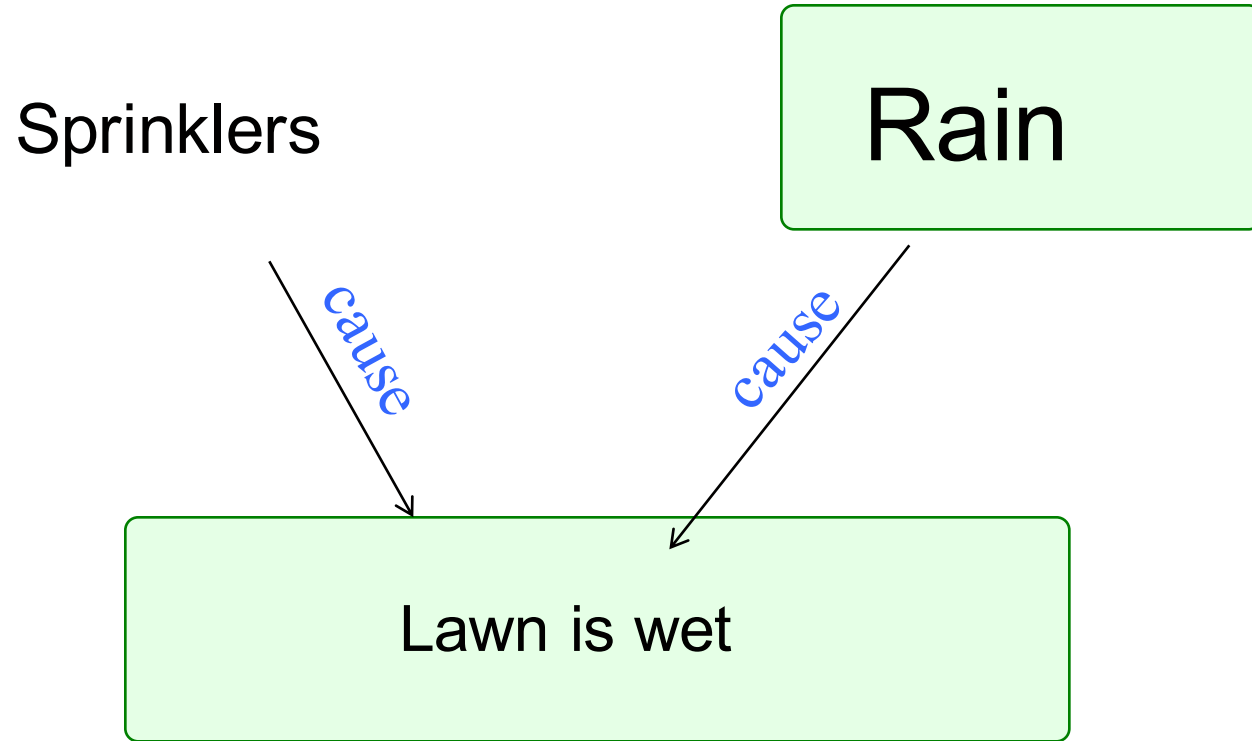
# Explaining Away

Sprinklers

Rain



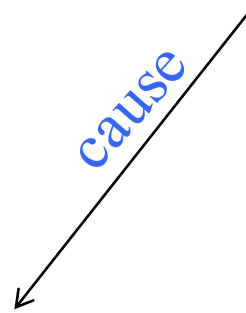
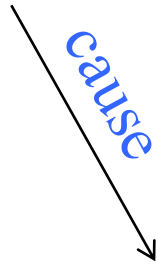
# Explaining Away



# Explaining Away

Sprinklers

Rain



Lawn is wet

# Conditioning Can Make Independence

- Consider a randomly chosen day of the week
  - Let A be event: It is not Monday
  - Let B be event: It is Saturday
  - Let C be event: It is the weekend
- A and B are dependent
  - $P(A) = 6/7$ ,  $P(B) = 1/7$ ,  $P(AB) = 1/7 \neq (6/7)(1/7)$
- Now condition both A and B on C:
  - $P(A|C) = 1$ ,  $P(B|C) = 1/2$ ,  $P(AB|C) = 1/2$
  - $P(AB|C) = P(A|C) P(B|C) \rightarrow A|C$  and  $B|C$  *independent*
- Dependent events can become independent by conditioning on additional information



# Conditional Independence

- Two events  $E$  and  $F$  are called conditionally independent given  $G$ , if

$$P(E F | G) = P(E | G) P(F | G)$$

Or, equivalently:  $P(E | F G) = P(E | G)$

**NETFLIX**

**And Learn**

# Netflix and Learn

What is the probability  
that a user will like  
Life is Beautiful?

$$P(E)$$



---

$$P(E) = \lim_{n \rightarrow \infty} \frac{n(E)}{n} \approx \frac{\# \text{ people who liked movie}}{\# \text{ people who watched movie}}$$

$$P(E) = 50,234,231 / 50,923,123 = 0.97$$

# Netflix and Learn

What is the probability that a user will like Life is Beautiful, given they liked Amelie?

$$P(E|F)$$



---

$$P(E|F) = \frac{P(EF)}{P(F)} = \frac{\text{people who liked both}}{\text{people who watched both}} = \frac{\text{people who liked amelie}}{\text{people who watched amelie}}$$

$$P(E|F) = 0.99$$

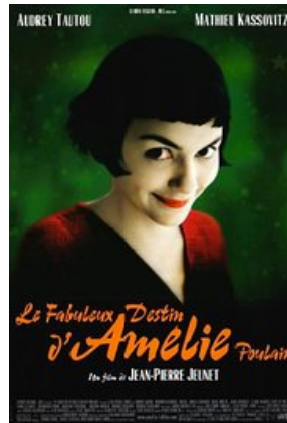
Conditioned on liking a set of movies?

# Netflix and Learn

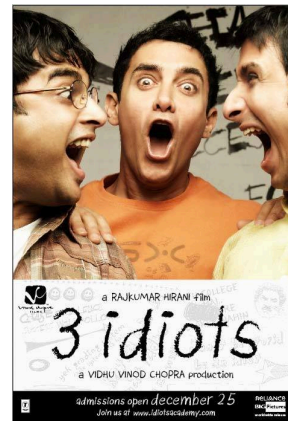
Each event corresponds to liking a particular movie



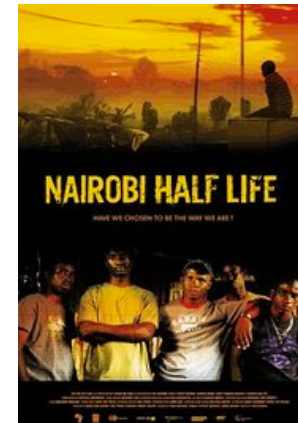
$E_1$



$E_2$



$E_3$



$E_4$

$$P(E_4 | E_1, E_2, E_3)?$$

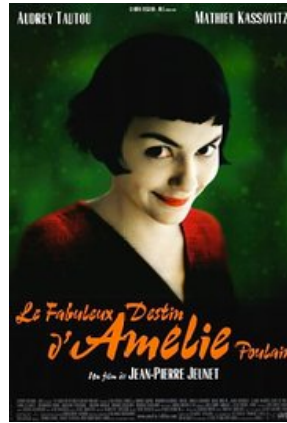
Is  $E_4$  independent of  $E_1, E_2, E_3$ ?

# Netflix and Learn

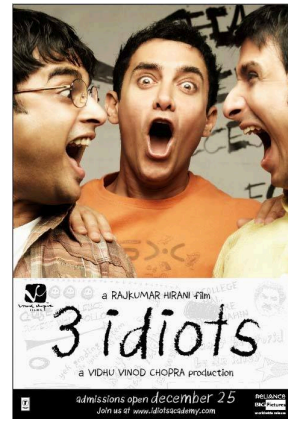
Is  $E_4$  independent of  $E_1, E_2, E_3$ ?



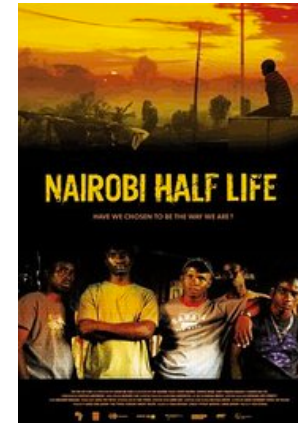
$E_1$



$E_2$



$E_3$



$E_4$

$$P(E_4|E_1, E_2, E_3) \stackrel{?}{=} P(E_4)$$



# Netflix and Learn

- What is the probability that a user watched four particular movies?
  - There are 13,000 titles on Netflix
  - The user watches 30 random titles
  - $E$  = movies watched include the given four.

- Solution:

$$P(E) = \frac{\binom{4}{4} \binom{12996}{24}}{\binom{13000}{30}} = 10^{-11}$$

Watch those four

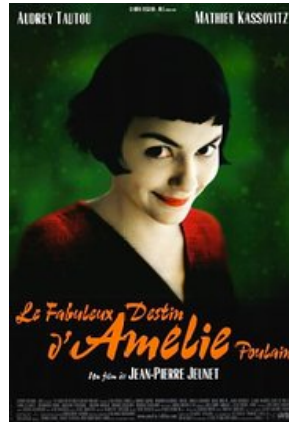
Choose 24 movies not in the set

Choose 30 movies from netflix

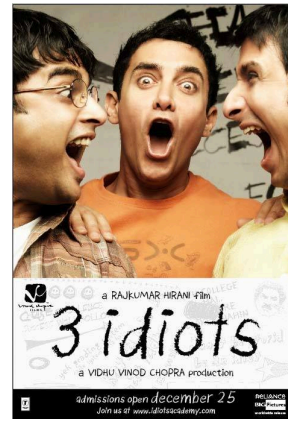
# Netflix and Learn



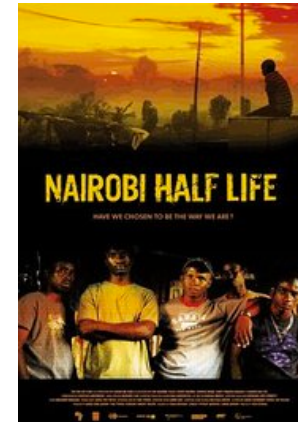
$E_1$



$E_2$



$E_3$

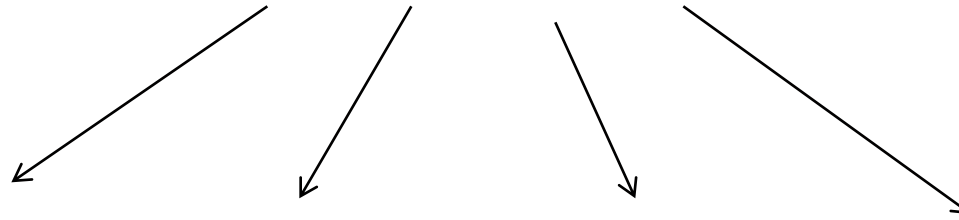


$E_4$

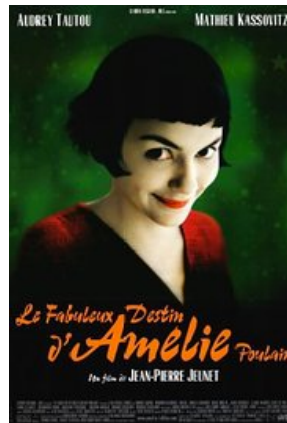
# Netflix and Learn

$K_1$

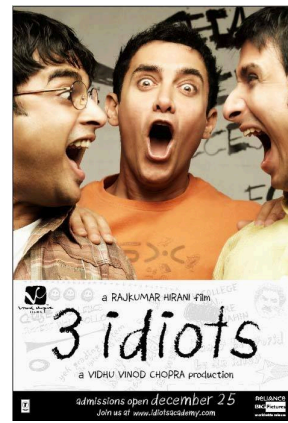
*Like foreign emotional comedies*



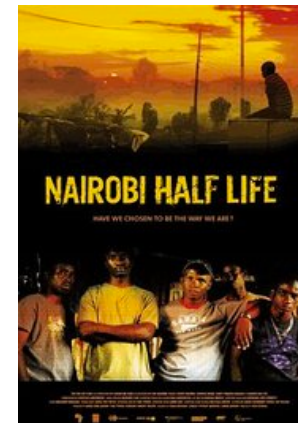
$E_1$



$E_2$



$E_3$



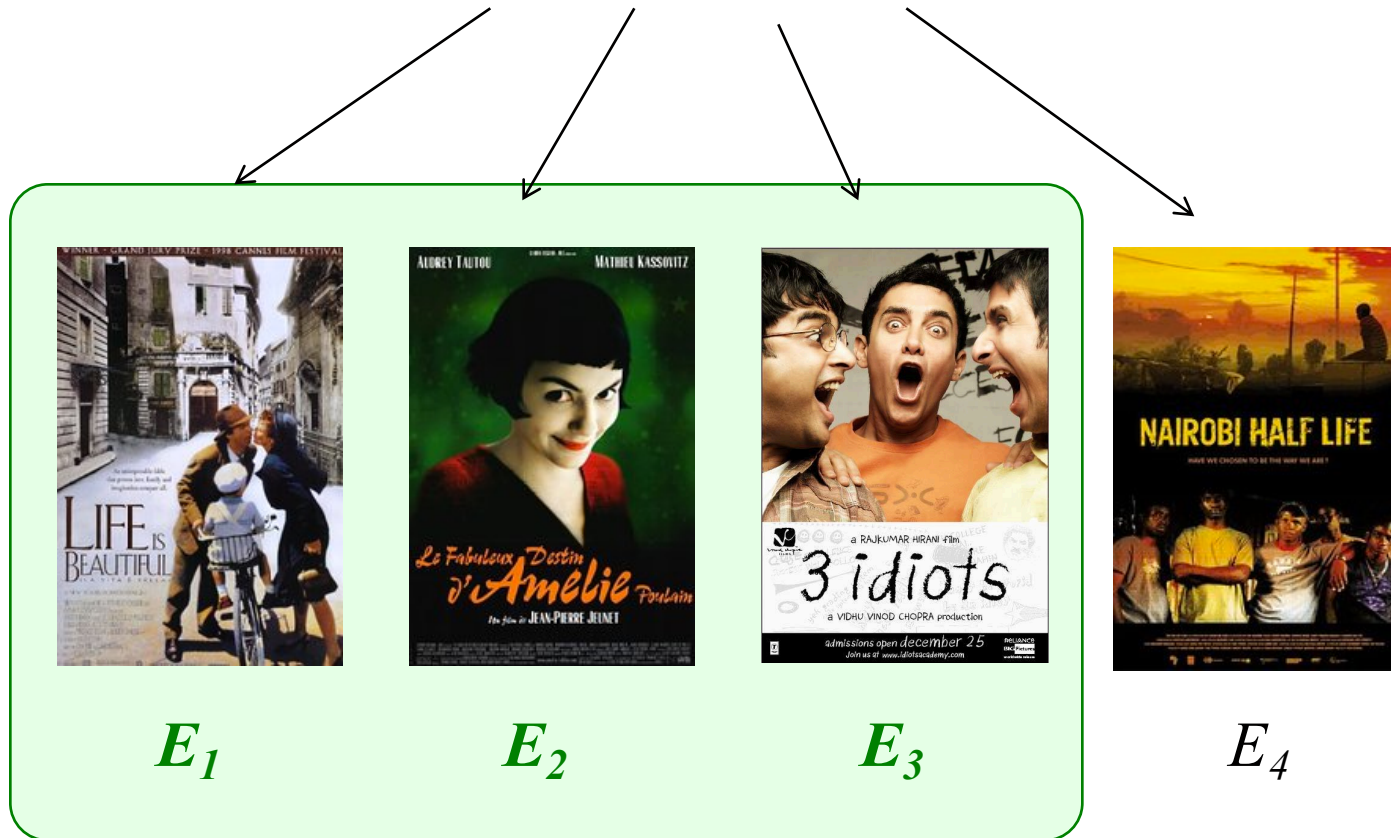
$E_4$

Assume  $E_1$ ,  $E_2$ ,  $E_3$  and  $E_4$  are conditionally independent given  $K_1$

# Netflix and Learn

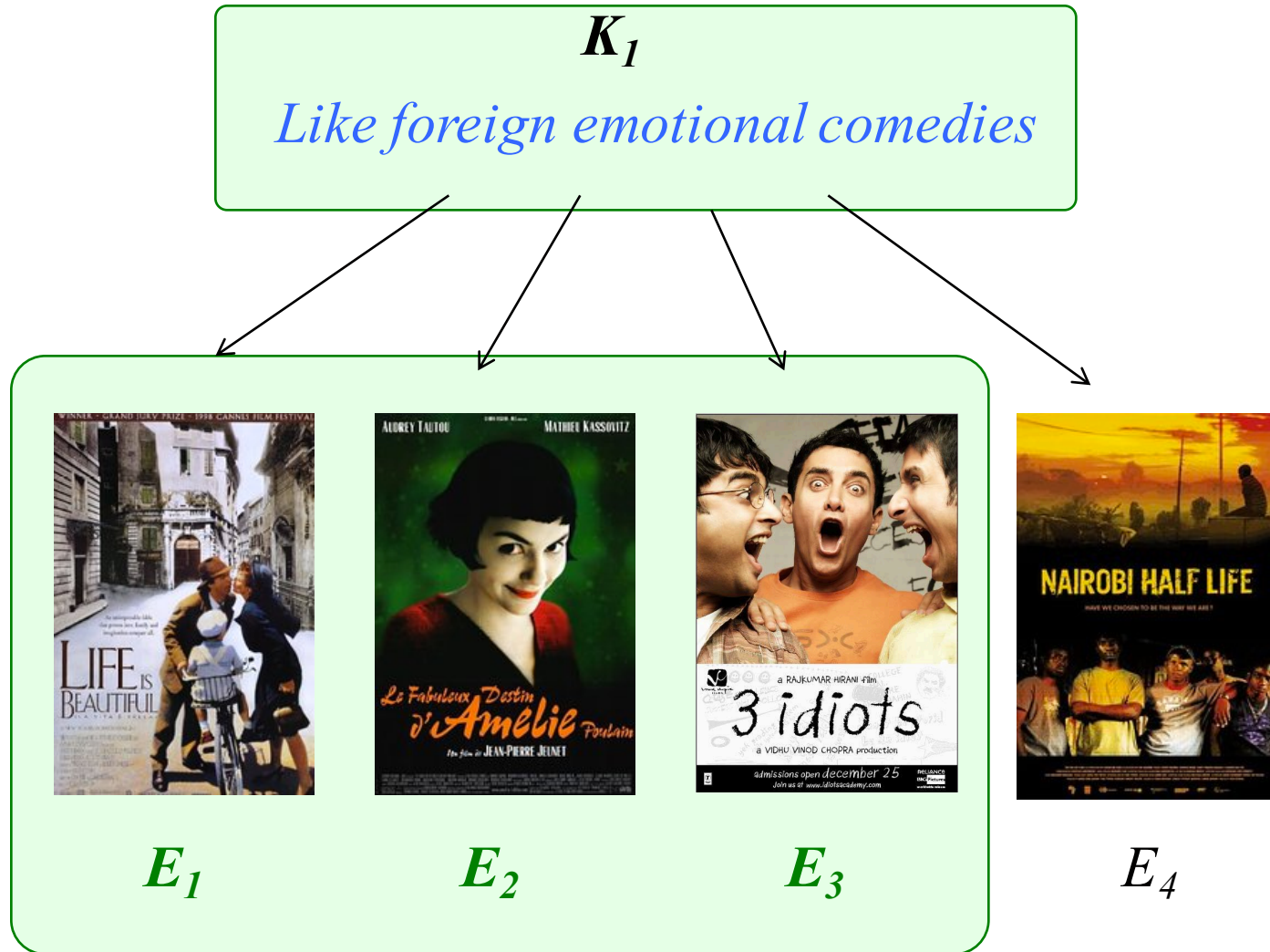
$K_1$

*Like foreign emotional comedies*



Assume  $E_1, E_2, E_3$  and  $E_4$  are conditionally independent given  $K_1$

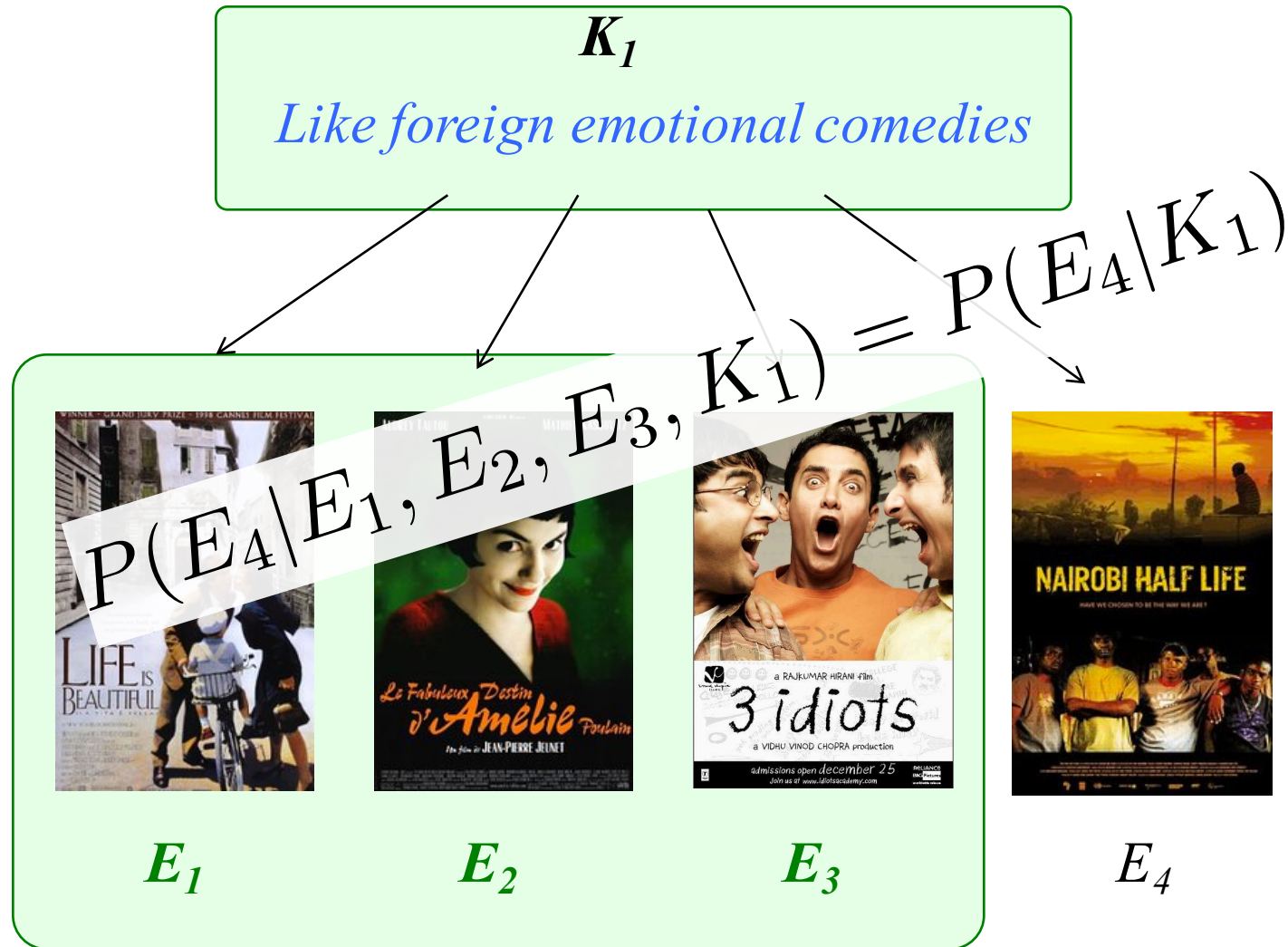
# Netflix and Learn



Assume  $E_1$ ,  $E_2$ ,  $E_3$  and  $E_4$  are conditionally independent given  $K_1$



# Netflix and Learn



Assume  $E_1, E_2, E_3$  and  $E_4$  are conditionally independent given  $K_1$

Conditional independence is a practical, real world way of decomposing hard probability questions.

# Big Deal

“Exploiting conditional independence to generate fast probabilistic computations is one of the main contributions CS has made to probability theory”

-Judea Pearl wins 2011 Turing Award, *“For fundamental contributions to artificial intelligence through the development of a calculus for probabilistic and causal reasoning”*



Extra problem given time

# Reminder: Geometric Series

- Geometric series:  $x^0 + x^1 + x^2 + x^3 + \dots + x^n = \sum_{i=0}^n x^i$
- If  $x$  is greater than 0 and less than 1:

$$\sum_{i=0}^{\infty} x^i = \frac{1}{1-x}$$

- From your “Calculation Reference” handout:

# Simplified Craps

- Two 6-sided dice repeatedly rolled (roll = ind. trial)
  - $E = 5$  is rolled before a 7 is rolled
  - What is  $P(E)$ ?
- Solution
  - $F_n =$  no 5 or 7 rolled in first  $n - 1$  trials, 5 rolled on  $n^{\text{th}}$  trial
  - $P(E) = P\left(\bigcup_{n=1}^{\infty} F_n\right) = \sum_{n=1}^{\infty} P(F_n)$
  - $P(5 \text{ on any trial}) = 4/36$        $P(7 \text{ on any trial}) = 6/36$
  - $P(F_n) = (1 - (10/36))^{n-1} (4/36) = (26/36)^{n-1} (4/36)$
  - $P(E) = \frac{4}{36} \sum_{n=1}^{\infty} \left(\frac{26}{36}\right)^{n-1} = \frac{4}{36} \sum_{n=0}^{\infty} \left(\frac{26}{36}\right)^n = \frac{4}{36} \frac{1}{\left(1 - \frac{26}{36}\right)} = \frac{2}{5}$