

11: Joint (Multivariate) Distributions

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Adapted from slides by Lisa Yan

$$X \sim \mathcal{N}(\overset{\text{mean}}{\mu}, \overset{\text{variance}}{\sigma^2})$$

$$P(X \leq x) = F(x) = \int_{-\infty}^x \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy$$

CDF has no closed form

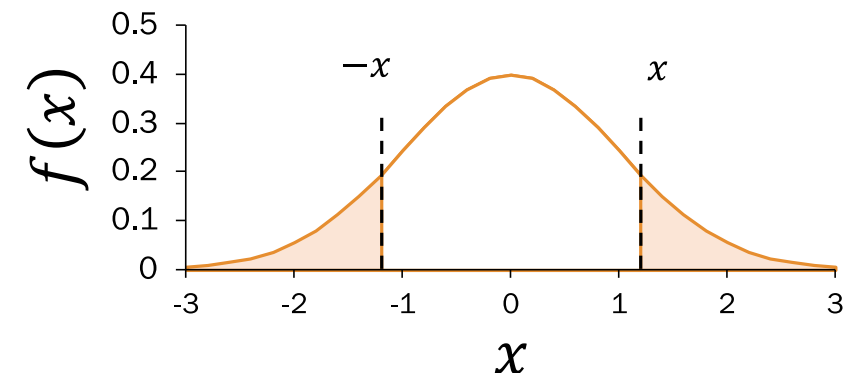
CDF of $X \sim \mathcal{N}(\mu, \sigma^2)$

If $X \sim \mathcal{N}(\mu, \sigma^2)$, then

$$F(x) = \Phi\left(\frac{x - \mu}{\sigma}\right)$$

- Symmetry of the PDF of Normal RV implies $\Phi(-x) = 1 - \Phi(x)$

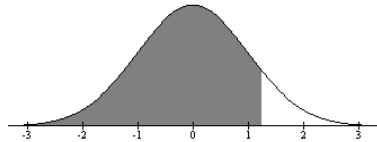
CDF of Standard Normal Z, solved for numerically



Standard Normal Table

Standard Normal Table

An entry in the table is the area under the curve to the left of z , $P(Z \leq z) = \Phi(z)$.



Z	0.00	0.01	0.02	0.03	0.04	0.05	0.06	0.07	0.08	0.09
0.0	0.5000	0.5040	0.5080	0.5120	0.5160	0.5199	0.5239	0.5279	0.5319	0.5359
0.1	0.5398	0.5438	0.5478	0.5517	0.5557	0.5596	0.5636	0.5675	0.5714	0.5753
0.2	0.5793	0.5832	0.5871	0.5910	0.5948	0.5987	0.6026	0.6064	0.6103	0.6141
0.3	0.6179	0.6217	0.6255	0.6293	0.6331	0.6368	0.6406	0.6443	0.6480	0.6517
0.4	0.6554	0.6591	0.6628	0.6664	0.6700	0.6736	0.6772	0.6808	0.6844	0.6879
0.5	0.6915	0.6950	0.6985	0.7019	0.7054	0.7088	0.7123	0.7157	0.7190	0.7224
0.6	0.7257	0.7291	0.7324	0.7357	0.7389	0.7422	0.7454	0.7486	0.7517	0.7549
0.7	0.7580	0.7611	0.7642	0.7673	0.7703	0.7734	0.7764	0.7793	0.7823	0.7852
0.8	0.7881	0.7910	0.7939	0.7967	0.7995	0.8023	0.8051	0.8078	0.8106	0.8133
0.9	0.8159	0.8186	0.8212	0.8238	0.8264	0.8289	0.8315	0.8340	0.8365	0.8389
1.0	0.8413	0.8438	0.8461	0.8485	0.8508	0.8531	0.8554	0.8577	0.8599	0.8621
1.1	0.8643	0.8665	0.8686	0.8708	0.8729	0.8749	0.8770	0.8790	0.8810	0.8830
1.2	0.8849	0.8869	0.8888	0.8906	0.8925	0.8943	0.8962	0.8980	0.8997	0.9015
1.3	0.9032	0.9049	0.9066	0.9082	0.9099	0.9115	0.9131	0.9147	0.9162	0.9177
1.4	0.9192	0.9207	0.9222	0.9236	0.9251	0.9265	0.9279	0.9292	0.9306	0.9319
1.5	0.9332	0.9345	0.9357	0.9370	0.9382	0.9394	0.9406	0.9418	0.9429	0.9441

- $Z \sim \mathcal{N}(0, 1)$ has a numeric lookup table for $\Phi(x)$, where $x \geq 0$.
- Computing implications: saving one lookup table for $\mathcal{N}(0, 1)$ enables you to quickly compute probabilities for general $\mathcal{N}(\mu, \sigma^2)$!

Today's plan

→ Normal approximation for Binomial

Joint distributions (discrete)

Multinomial Random Variable

Text analysis

Website testing

- 100 people are given a new website design.
- $X = \#$ people whose time on site increases
- CEO will endorse the new design if $X \geq 65$.
- The design actually has no effect, so $P(\text{time on site increases}) = 0.5$ independently.

What is $P(\text{CEO endorses change})$? *Give a numerical approximation.*

- Strategy:
- A. Poisson
 - B. Bayes' Theorem
 - C. Binomial
 - D. Normal (Gaussian)
 - E. Uniform

Website testing

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What is $P(\text{CEO endorses change})$? *Give a numerical approximation.*

Approach 1: Binomial

Define

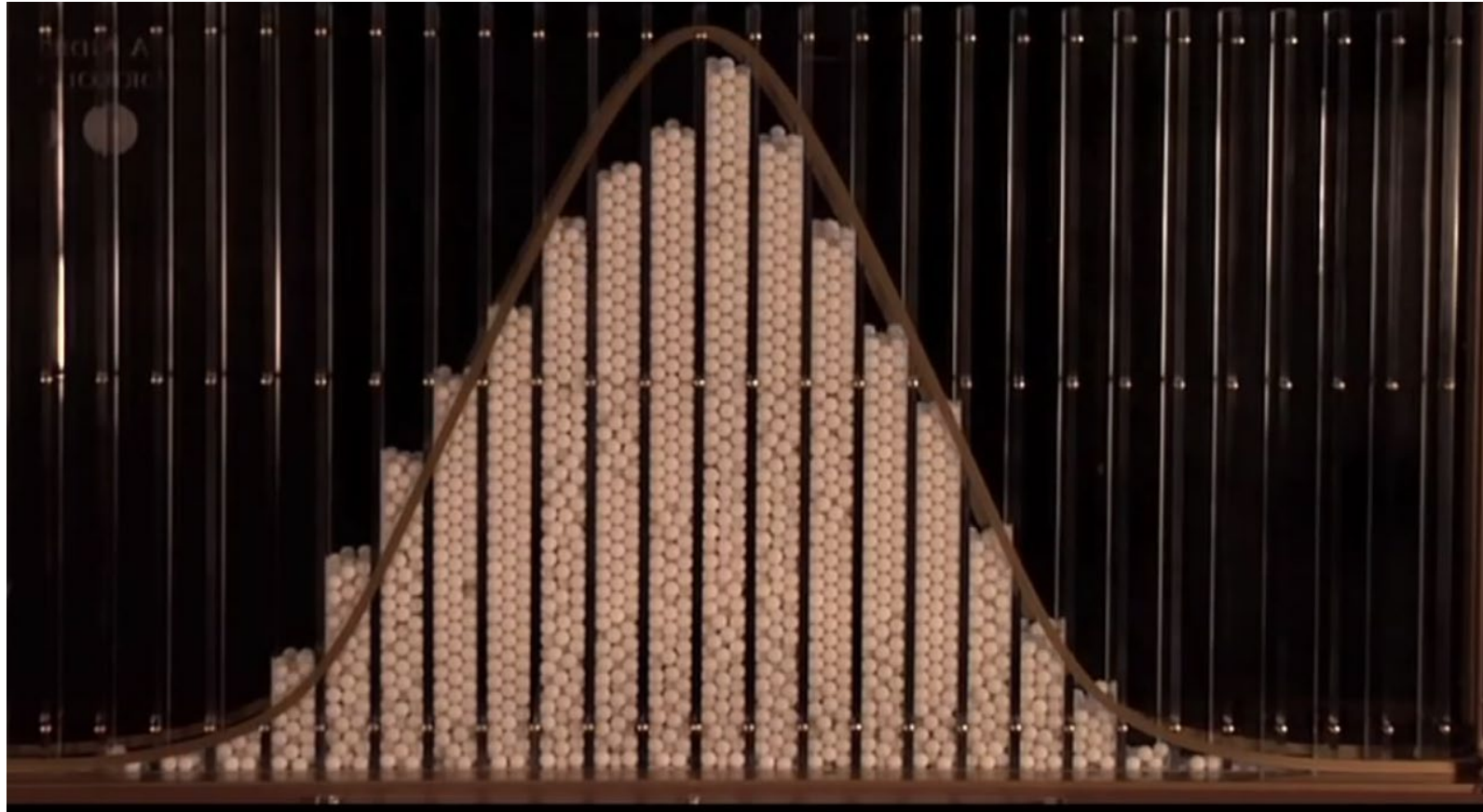
$$X \sim \text{Bin}(n = 100, p = 0.5)$$

Want: $P(X \geq 65)$

Solve

$$P(X \geq 65) = \sum_{i=65}^{100} \binom{100}{i} 0.5^i (1 - 0.5)^{100-i}$$

Don't worry, Normal approximates Binomial



Galton Board

(We'll explain where this approximation comes from in 2 weeks' time)

Website testing

- 100 people are given a new website design.
- $X = \#$ people whose time on site increases
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What is $P(\text{CEO endorses change})$? Give a numerical approximation.

Approach 1: Binomial

Define

$$X \sim \text{Bin}(n = 100, p = 0.5)$$

Want: $P(X \geq 65)$

Solve

$$P(X \geq 65) \approx \mathbf{0.0018}$$

Approach 2: approximate with Normal

Define

$$Y \sim \mathcal{N}(\mu, \sigma^2)$$

$$\mu = np = 50$$

$$\sigma^2 = np(1 - p) = 25$$

$$\sigma = \sqrt{25} = 5$$

Solve

$$P(X \geq 65) \approx P(Y \geq 65) = 1 - F_Y(65)$$

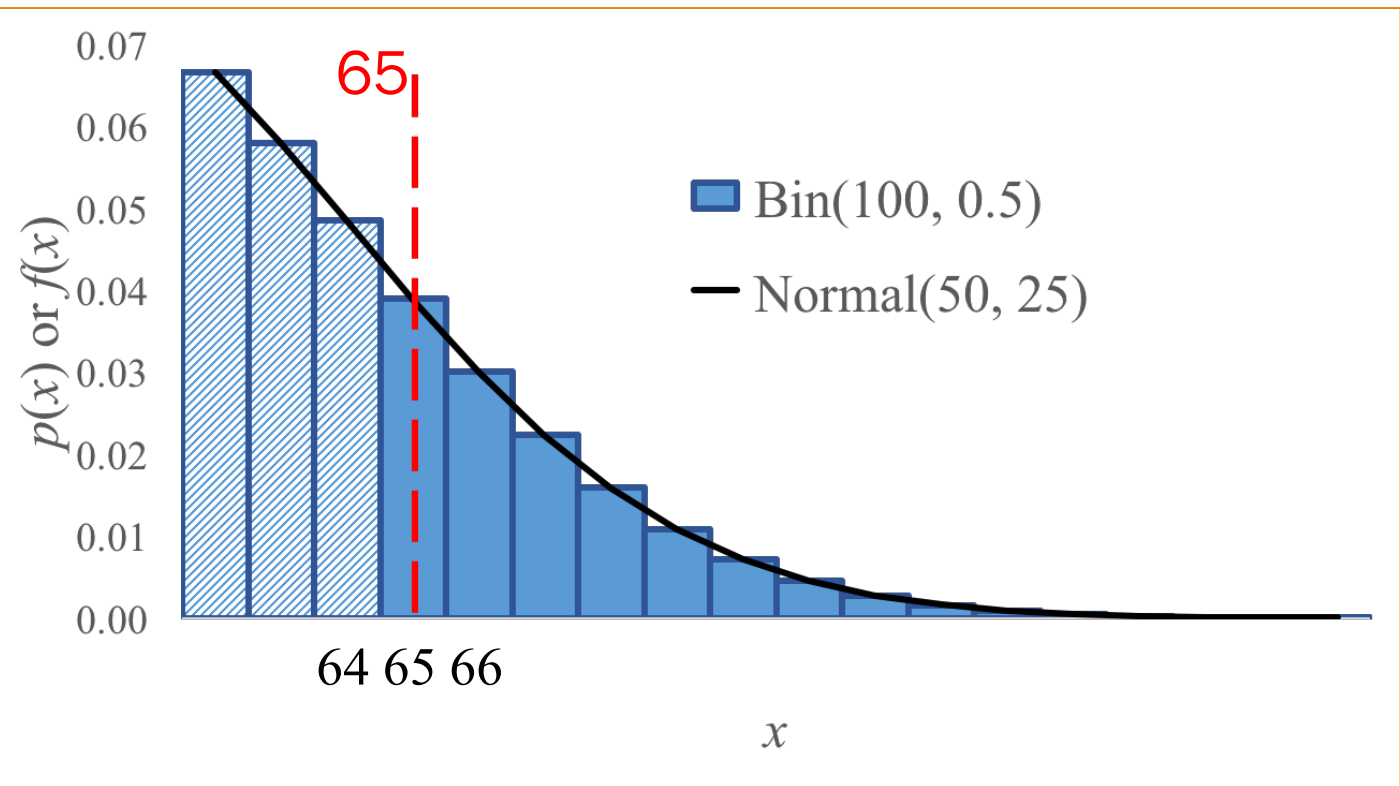
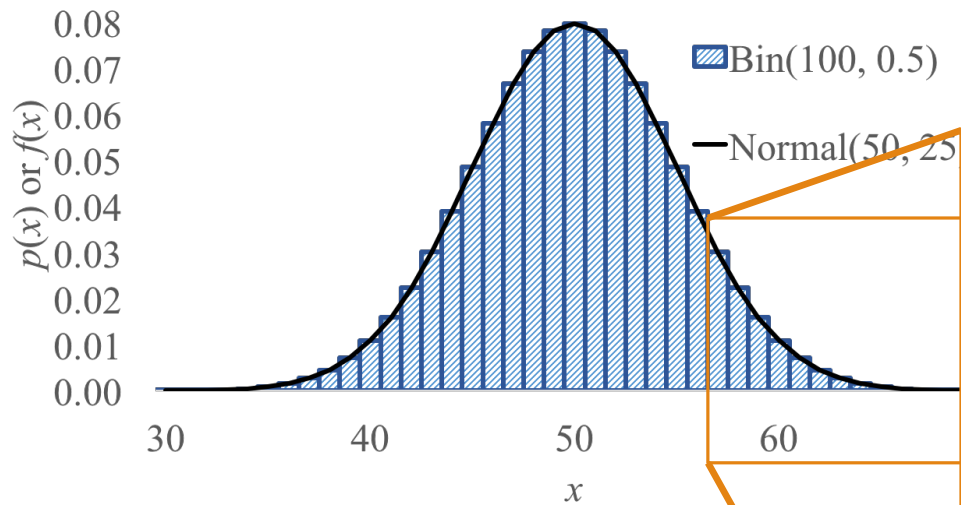
$$= 1 - \Phi\left(\frac{65-50}{5}\right) = 1 - \Phi(3) \approx \mathbf{0.0013?}$$

(this approach is actually missing something)

Website testing with continuity correction

You must perform a **continuity correction** when approximating a discrete RV with a continuous RV.

$Y \sim \mathcal{N}(50, 25)$ approximates $X \sim \text{Bin}(100, 0.5)$



$$P(X \geq 65) \text{ Binomial}$$

$$\approx P(Y \geq 64.5) \text{ Normal}$$

$$\approx 0.0018$$

Continuity correction

If $Y \sim \mathcal{N}(np, np(1 - p))$ approximates $X \sim \text{Bin}(n, p)$,
how do we approximate the following probabilities?

Discrete (e.g., Binomial)
probability question



Continuous (Normal)
probability question

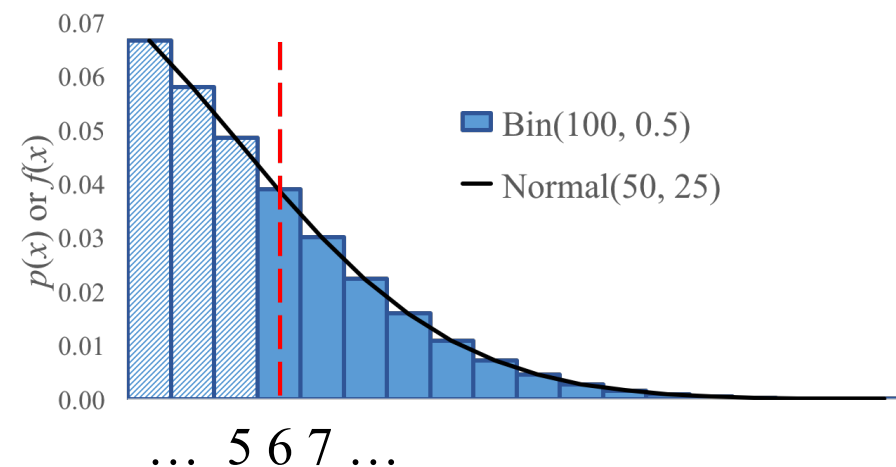
$$P(X = 6)$$

$$P(X \geq 6)$$

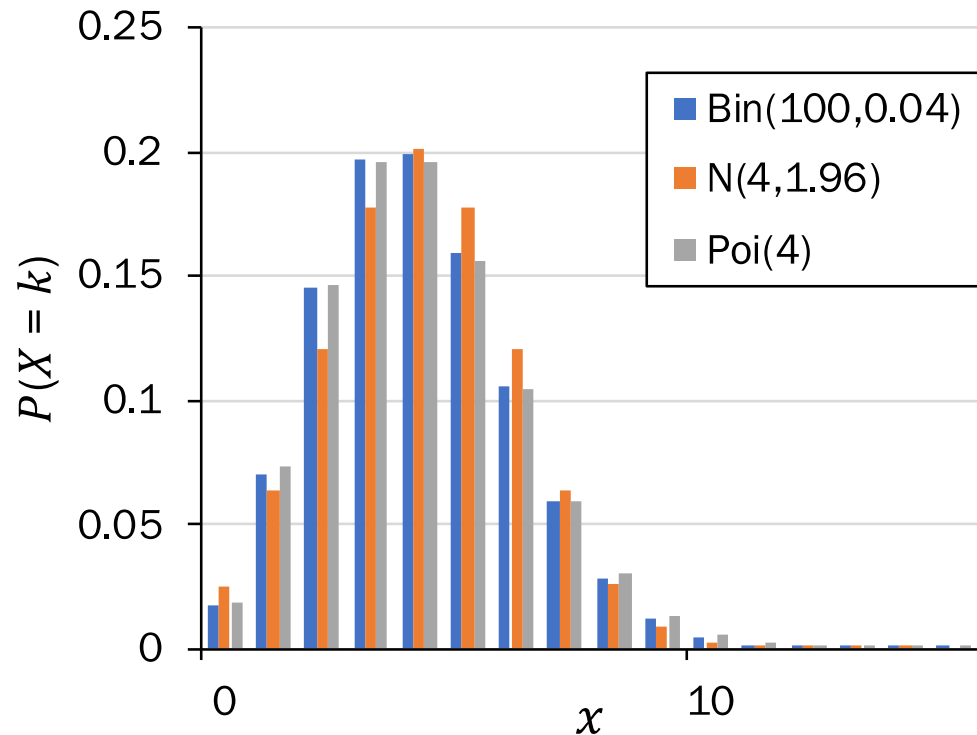
$$P(X > 6)$$

$$P(X < 6)$$

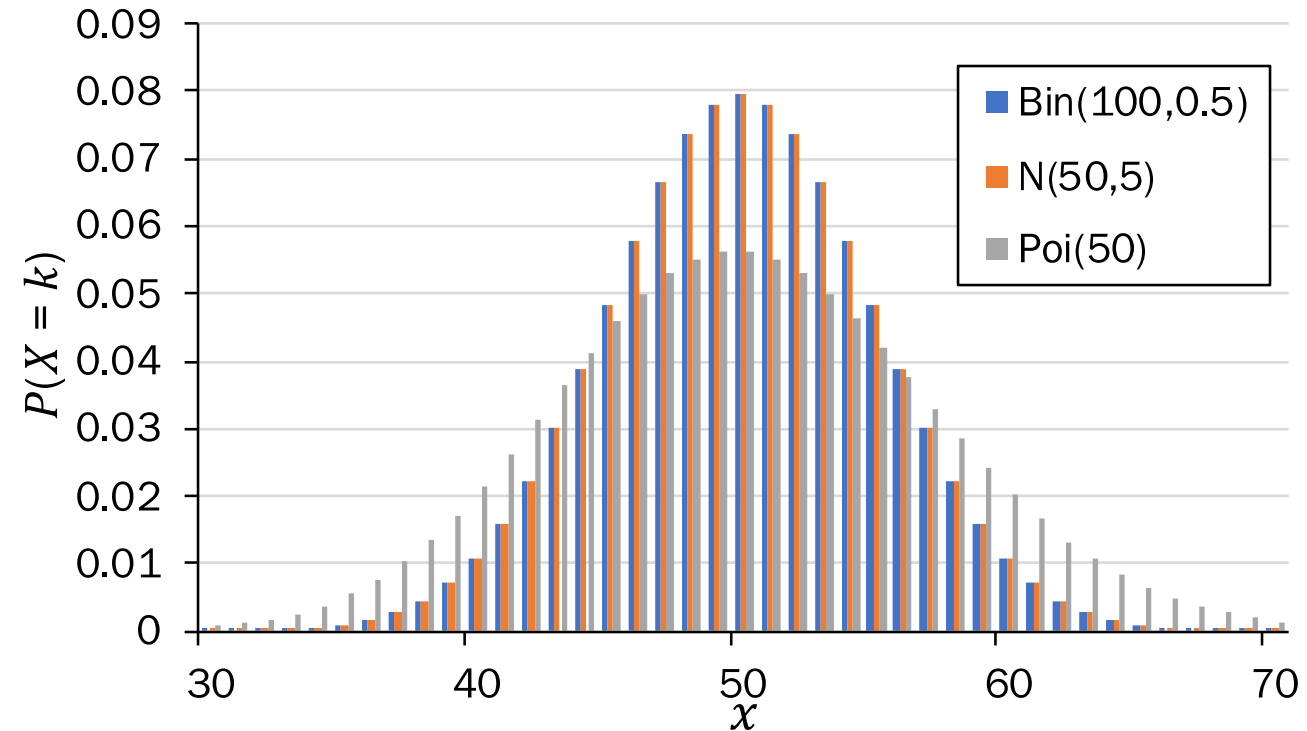
$$P(X \leq 6)$$



Who gets to approximate?



Poisson approximation
 n large (> 20), p small (< 0.05)
slight dependence okay



Normal approximation
 n large (> 20), p mid-ranged ($np(1 - p) > 10$)
independence

1. If there is a choice, use Normal to approx.
2. When using Normal to approximate a discrete RV, use a continuity correction.

Stanford Admissions (a while back)

Stanford accepts 2480 students.

- Each accepted student has 68% chance of attending (independent trials)
- Let $X = \#$ of students who will attend

What is $P(X > 1745)$? *Give a numerical approximation.*

- Strategy:
- A. Just Binomial
 - B. Poisson
 - C. Normal
 - D. None/other

Stanford Admissions (about 20 years ago)

Stanford accepts 2480 students.

- Each accepted student has 68% chance of attending (independent trials)
- Let $X = \#$ of students who will attend

What is $P(X > 1745)$? Give a numerical approximation.

Strategy:

A. Just Binomial	$n = 2480$, computationally expensive
B. Poisson	$p = 0.68$, not small enough
C. Normal	Variance $np(1 - p) = 540 > 10$
D. None/other	

Define an approximation

Let $Y \sim \mathcal{N}(E[X], \text{Var}(X))$

$$E[X] = np = 1686$$

$$\text{Var}(X) = np(1 - p) \approx 540 \rightarrow \sigma = 23.3$$

$$P(X > 1745) \approx P(Y \geq 1745.5)$$

Continuity
correction

Solve

$$P(Y \geq 1745.5) = 1 - F(1745.5)$$

$$= 1 - \Phi\left(\frac{1745.5 - 1686}{23.3}\right)$$

$$= 1 - \Phi(2.54) \approx \mathbf{0.0055}$$

Today's plan

Normal approximation for Binomial

→ Cool normal facts

Joint distributions (discrete)

Multinomial Random Variable

Text analysis

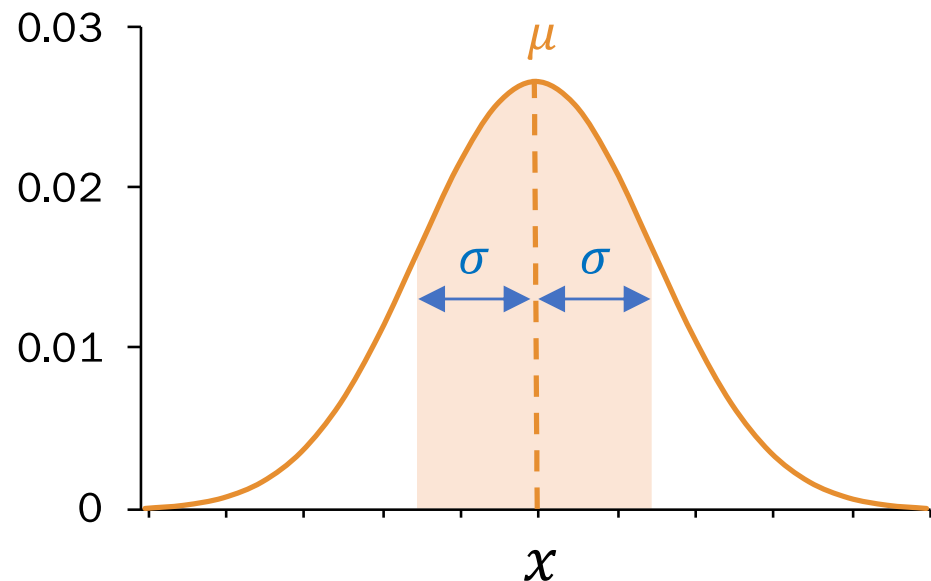
68% rule

You may have heard the statement:

“68% of the class will fall within 1 standard deviation of the exam average.”

This is only true of **normal distributions**:

Let $X \sim \mathcal{N}(\mu, \sigma^2)$ with CDF F .



$$\begin{aligned} P(|X - \mu| < \sigma) &= P(\mu - \sigma < X < \mu + \sigma) \\ &= F(\mu + \sigma) - F(\mu - \sigma) \\ &= \Phi\left(\frac{(\mu + \sigma) - \mu}{\sigma}\right) - \Phi\left(\frac{(\mu - \sigma) - \mu}{\sigma}\right) \\ &= \Phi(1) - \Phi(-1) = \Phi(1) - (1 - \Phi(1)) \\ &= 2\Phi(1) - 1 \approx 2(0.8413) - 1 = \mathbf{0.6826} \end{aligned}$$

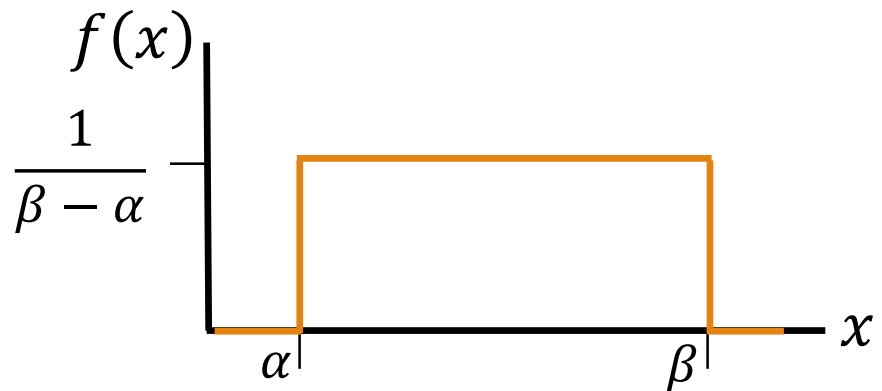
68% rule

You may have heard the statement:

“68% of the class will fall within 1 standard deviation of the exam average.”

This is only true of **normal distributions**:

Counterexample: Let $X \sim \text{Unif}(\alpha, \beta)$.



$$\mu = E[X] = \frac{\alpha + \beta}{2}$$

$$\text{Var}(X) = \frac{(\beta - \alpha)^2}{12} \Rightarrow \sigma = \text{SD}(X) = \frac{\beta - \alpha}{\sqrt{12}}$$

$$P(|X - \mu| < \sigma) = P(\mu - \sigma < X < \mu + \sigma)$$

$$= \frac{1}{\beta - \alpha} \cdot [(\mu + \sigma) - (\mu - \sigma)]$$

$$= \frac{1}{\beta - \alpha} [2\sigma] = \frac{1}{\beta - \alpha} \cdot \left[2 \cdot \frac{\beta - \alpha}{\sqrt{12}} \right]$$

$$= 2/\sqrt{12} \approx 0.58$$

Today's plan

Normal approximation for Binomial

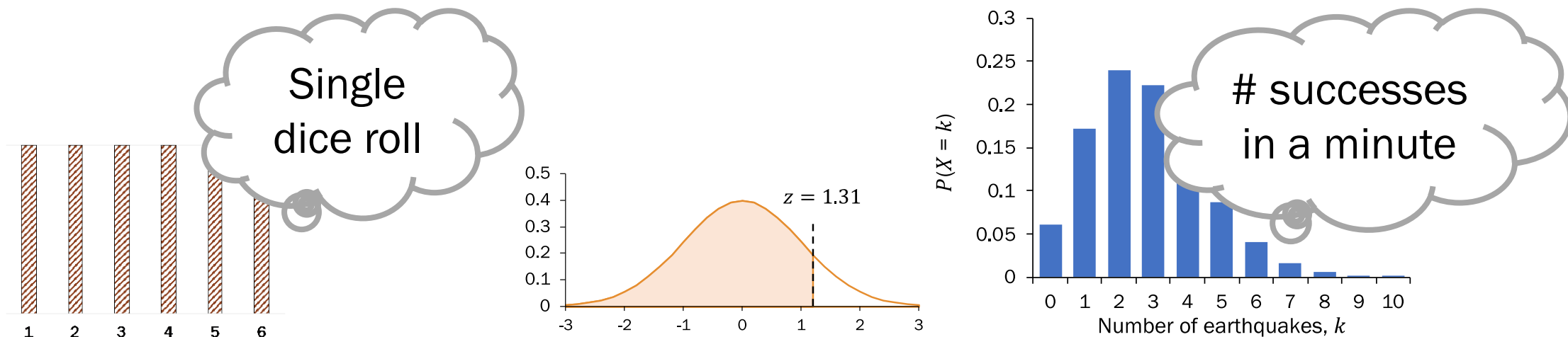
 Joint distributions (discrete)

Multinomial Random Variable

Text analysis

Joint distributions

So far, we have only worked with 1-dimensional random variables:



However, in the real world, events often occur with other events.

Outcomes on two dice rolls

2 successes in minute 1,
none in minutes 2-4,
3 successes in minute 5



What is the probability that the Warriors win?
How do you model zero-sum games?

ELO ratings

Want: $P(\text{Warriors win}) = P(A_W > A_B)$

```
from scipy import stats
WARRIORS_ELO = 1657
OPPONENT_ELO = 1470
STDEV = 200
NTRIALS = 10000
```

nSuccess = 0

for i in range(NTRIALS):

w = stats.norm.rvs(WARRIORS_ELO, STDEV)

b = stats.norm.rvs(OPPONENT_ELO, STDEV)

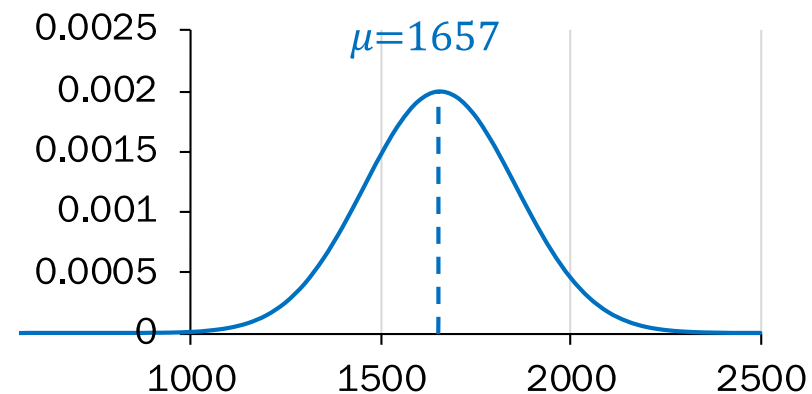
if w > b:

nSuccess += 1

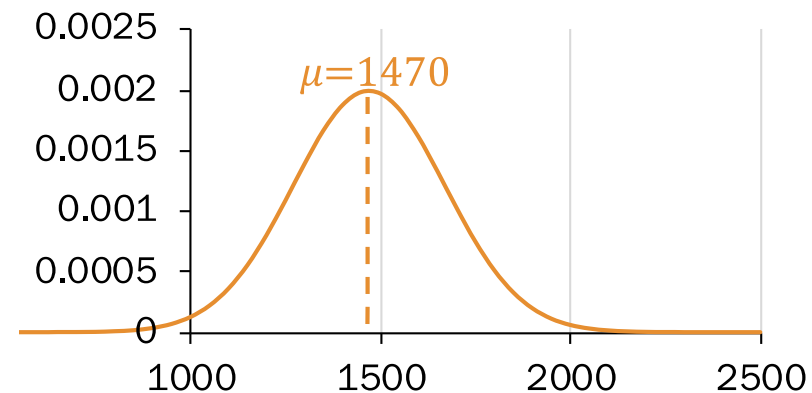
print("Warriors sampled win fraction", float(nSuccess) / NTRIALS)

≈ 0.7488, calculated by sampling

Warriors $A_W \sim \mathcal{N}(S = 1657, 200^2)$



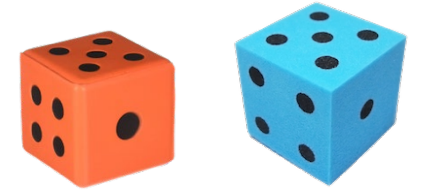
Opponents $A_B \sim \mathcal{N}(S = 1470, 200^2)$



CS109 Goal: Reason about probabilities involving multiple random variables.

Joint probability mass functions

Roll two 6-sided dice, yielding values X and Y .



X

random variable

$$P(X = 1)$$

probability of
an event

$$P(X = k)$$

probability mass function

X, Y

random variables

$$P(X = 1 \cap Y = 6)$$

$$P(X = 1, Y = 6)$$

new notation: the comma

probability of the intersection
of two events

$$P(X = a, Y = b)$$



joint probability mass function

Discrete joint distributions

For two discrete joint random variables X and Y , the **joint probability mass function** is defined as:

$$p_{X,Y}(a, b) = P(X = a, Y = b)$$

The **marginal distributions** of the joint PMF are defined as:

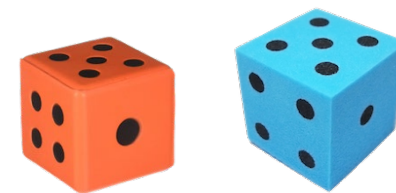
$$p_X(a) = P(X = a) = \sum_y p_{X,Y}(a, y) \quad p_Y(b) = P(Y = b) = \sum_x p_{X,Y}(x, b)$$

Use marginal distributions to get a 1-D RV from a joint PMF.

Two dice

Roll two 6-sided dice, yielding values X and Y .

1. What is the joint PMF of X and Y ?



$$p_{X,Y}(a, b) = 1/36 \quad (a, b) \in \{(1,1), \dots, (6,6)\}$$

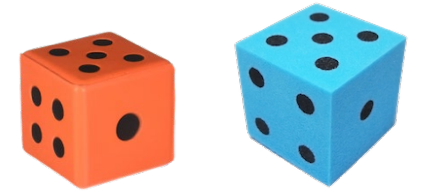
2. What is the marginal PMF of X ?

$$p_X(a) = P(X = a) = \sum_y p_{X,Y}(a, y) = \sum_{y=1}^6 \frac{1}{36} = \frac{1}{6} \quad a \in \{1, \dots, 6\}$$

Two dice

Roll two 6-sided dice, yielding values X and Y .

1. What is the joint PMF of X and Y ? $p_{X,Y}(a,b) = 1/36$



		X					
		1	2	3	4	5	6
Y	1	1/36	1/36
	2
	3
	4
	5
	6	1/36	1/36

An orange arrow points from the text $P(X = 4, Y = 2)$ to the cell in the table where $X=4$ and $Y=2$.

Probability table

- All possible outcomes for several discrete RVs
- Not parametric (e.g., parameter p in $\text{Ber}(p)$)

2. What is the marginal PMF of X ?

Announcements

Midterm exam

When: Monday, February 10, 7:00pm-9:00pm

Where: Cubberley Auditorium

Not permitted: book/computer/calculator

Permitted: **Three** 8.5"x11" double-sided sheets of notes

Covers: Up to (and including) week 4 + **Lecture Notes 11**

Practice: <http://web.stanford.edu/class/cs109/exams/midterm.html>

Review session: Saturday, 3-5pm, STLC 111
not recorded; materials will be posted though

A computer (or three) in every house.

Consider households in Silicon Valley.

- A household has C computers, where $C = X$ Macs + Y PCs.
- Each computer in a household is equally likely to be a Mac or PC.

$$P(C = c) = \begin{cases} 0.16, & c = 0 \\ 0.24, & c = 1 \\ 0.28, & c = 2 \\ 0.32, & c = 3 \end{cases}$$

What is $P(X = 1, Y = 0)$, the missing entry in the probability table?

- A. $1 - (.16 + .12 + .07 + \dots + .04) = 0.12$
- B. $.24 - (.12) = 0.12$
- C. $0.5(.24) = 0.12$
- D. All of the above
- E. None/other

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	?	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0

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Which entries in the probability table correspond to $P(C = 3)$?

A.

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0

B.

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0

C.

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0

A computer (or three) in every house.

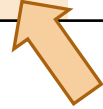
Consider households in Silicon Valley.

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- Each computer in a household is equally likely to be a Mac or PC.

$$P(C = c) = \begin{cases} 0.16, & c = 0 \\ 0.24, & c = 1 \\ 0.28, & c = 2 \\ 0.32, & c = 3 \end{cases}$$

How do you compute $P(X = 0, Y = 3)$?

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0



A computer (or three) in every house.

Consider households in Silicon Valley.

- A household has C computers, where $C = X$ Macs + Y PCs.
- Each computer in a household is equally likely to be a Mac or PC.

$$P(C = c) = \begin{cases} 0.16, & c = 0 \\ 0.24, & c = 1 \\ 0.28, & c = 2 \\ 0.32, & c = 3 \end{cases}$$

How do you compute $P(X = 0, Y = 3)$?


$$P(X = 0, Y = 3)$$

$$\begin{aligned} \text{Law of Total Probability} &= P(X = 0, Y = 3 | C = 3)P(C = 3) \\ &\quad + P(X = 0, Y = 3 | C \neq 3)P(C \neq 3) \end{aligned}$$

$$\text{Bin}(n=3, p=0.5) = \binom{3}{0} 0.5^0 0.5^3 \cdot (0.32) + 0$$

$$= 0.04$$

		X (# Macs)			
		0	1	2	3
Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0
	2	.07	.12	0	0
	3	.04	0	0	0



A computer (or three) in every house.

Consider households in Silicon Valley.

- A household has C computers, where $C = X$ Macs + Y PCs.
- Each computer in a household is equally likely to be a Mac or PC.

$$P(C = c) = \begin{cases} 0.16, & c = 0 \\ 0.24, & c = 1 \\ 0.28, & c = 2 \\ 0.32, & c = 3 \end{cases}$$

Which entries in the probability table correspond to the marginal PMF of X ?

A.		B.		C.							
		X (# Macs)		X (# Macs)		X (# Macs)					
		0	1	2	3	0	1	2	3		
Y (# PCs)	0	.16	.12	.07	.04	Y (# PCs)	0	.16	.12	.07	.04
	1	.12	.14	.12	0		1	.12	.14	.12	0
	2	.07	.12	0	0		2	.07	.12	0	0
	3	.04	0	0	0		3	.04	0	0	0
		Sum rows here						sum cols here			

A computer (or three) in every house.

Consider households in Silicon Valley.

- A household has C computers, where $C = X$ Macs + Y PCs.
- Each computer in a household is equally likely to be a Mac or PC.

$$P(C = c) = \begin{cases} 0.16, & c = 0 \\ 0.24, & c = 1 \\ 0.28, & c = 2 \\ 0.32, & c = 3 \end{cases}$$

		X (# Macs)				
		0	1	2	3	
Y (# PCs)	0	.16	.12	.07	.04	.39
	1	.12	.14	.12	0	.38
	2	.07	.12	0	0	.19
	3	.04	0	0	0	.04
		.39	.38	.19	.04	

Marginal PMF of Y,

$$p_Y(y) = \sum_x p_{X,Y}(x, y)$$

$$\text{Marginal PMF of } X, p_X(x) = \sum_y p_{X,Y}(x, y)$$

To find a marginal distribution over one variable, sum over all other variables.

Today's plan

Normal approximation for Binomial

Joint distributions (discrete)

 Multinomial Random Variable

Text analysis

Counting unordered objects

Binomial coefficient

How many ways are there to group n objects into **two** groups of size k and $n - k$, respectively?

$$\binom{n}{k} = \frac{n!}{k!(n-k)!}$$

Called the binomial coefficient because of something from Algebra

Multinomial coefficient

How many ways are there to group n objects into r groups of sizes n_1, n_2, \dots, n_r respectively?

$$\binom{n}{n_1, n_2, \dots, n_r} = \frac{n!}{n_1! n_2! \cdots n_r!}$$

Multinomial generalizes Binomial for counting.

Probability

Binomial RV

What is the probability of getting k successes and $n - k$ failures in n trials?

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}$$

Binomial # of ways of ordering the successes

Probability of each ordering of k successes is equal + mutually exclusive

Multinomial RV

What is the probability of getting c_1 of outcome 1, c_2 of outcome 2, ..., and c_m of outcome m in n trials?

Multinomial RVs also generalize Binomial RVs for probability!

Multinomial Random Variable

Consider an experiment:

- n independent trials
- Each trial results in one of m outcomes with respective probabilities p_1, p_2, \dots, p_m where $\sum_{i=1}^m p_i = 1$
- Let $X_i = \#$ of trials with outcome i .

def A **Multinomial** random variable X is defined as follows:

Joint PMF

$$P(X_1 = c_1, X_2 = c_2, \dots, X_m = c_m) = \binom{n}{c_1, c_2, \dots, c_m} p_1^{c_1} p_2^{c_2} \cdots p_m^{c_m}$$

Multinomial # of ways of ordering the outcomes

Probability of each ordering is equal + mutually exclusive

where $\sum_{i=1}^m c_i = n$ and $\binom{n}{c_1, c_2, \dots, c_m} = \frac{n!}{c_1! c_2! \cdots c_m!}$ is a multinomial coefficient

Hello dice rolls, my old friends

A 6-sided die is rolled 7 times.

What is the probability of getting:

- 1 one
- 0 threes
- 0 fives
- 1 two
- 2 fours
- 3 sixes

Strategy (choose all that apply):

- A. Law of total probability
- B. Counting
- C. Multinomial RV
- D. Binomial RV
- E. None/other

Hello dice rolls, my old friends

A 6-sided die is rolled 7 times.

What is the probability of getting:

- 1 one
- 1 two
- 0 threes
- 2 fours
- 0 fives
- 3 sixes

$$P(X_1 = 1, X_2 = 1, X_3 = 0, X_4 = 2, X_5 = 0, X_6 = 3)$$

$$= \binom{7}{1,1,0,2,0,3} \left(\frac{1}{6}\right)^1 \left(\frac{1}{6}\right)^1 \left(\frac{1}{6}\right)^0 \left(\frac{1}{6}\right)^2 \left(\frac{1}{6}\right)^0 \left(\frac{1}{6}\right)^3 = 420 \left(\frac{1}{6}\right)^7$$

Today's plan

Normal approximation for Binomial

Joint distributions (discrete)

Multinomial Random Variable

 Text analysis

Probabilistic text analysis

Ignoring the order of words...

What is the probability of any given word that you write in English?

- $P(\text{word} = \text{"the"}) > P(\text{word} = \text{"multinomial"})$
- $P(\text{word} = \text{"Stanford"}) > P(\text{word} = \text{"Cal"})$

Probabilities of *counts* of words = Multinomial distribution



A document is a large multinomial.

(according to the Global Language Monitor, there are 988,968 words in the English language used on the internet.)

Probabilistic text analysis

Probabilities of *counts* of words = Multinomial distribution

Example document:

#words: $n = 48$

“When my late husband was alive he deposited some amount of Money with china Bank in which the amount will be declared to you once you respond to this message indicating your interest in helping to receive the fund and use it for Gods work as my wish.”

$$P \left(\begin{array}{l} \text{bank} = 1 \\ \text{fund} = 1 \\ \text{money} = 1 \\ \text{wish} = 1 \\ \dots \\ \text{to} = 3 \end{array} \middle| \text{spam} \right) = \frac{n!}{1! 1! 1! 1! \dots 3!} p_{\text{bank}}^1 p_{\text{fund}}^1 \dots p_{\text{to}}^3$$

Note: $P(\text{bank} | \text{spam}) \gg P(\text{bank} | \text{writer=you})$

Probabilistic text analysis

Probabilities of *counts* of words = Multinomial distribution

What about probability of those same words in someone else's writing?

- $P(\text{word} = \text{"probability"} \mid \text{writer} = \text{you}) > P(\text{word} = \text{"probability"} \mid \text{non-CS109 student})$

To determine authorship:

1. Estimate $P(\text{word} \mid \text{writer})$ from known writings
2. Use Bayes' Theorem to determine $P(\text{writer} \mid \text{document})$ for a new writing!

Who wrote The Federalist Papers?

Old and New Analysis

Authorship of The Federalist Papers

- 85 essays advocating ratification of the US constitution
- Written under the pseudonym “Publius” (really, Alexander Hamilton, James Madison, John Jay)



Who wrote which essays?

- Analyze probability of words in each essay and compare against word distributions from known writings of three authors

Let's write a program!

<http://web.stanford.edu/class/cs109/demos/federalistpapers.html>