

21:

Parameters and MLE

David Varodayan

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Adapted from slides by Lisa Yan

Rejection sampling algorithm

Inference question: What is $P(F_{lu} = 1 | U = 1, T = 1)$?

```
def rejection_sampling(event, observation):  
    samples = sample_a_ton()  
    samples_observation =  
        reject_inconsistent(samples, observation)  
    samples_event =  
        reject_inconsistent(samples_observation, event)  
    return len(samples_event)/len(samples_observation)
```

[flu, und, fev, tir]

```
Sampling...  
[0, 1, 0, 1]  
[0, 1, 0, 1]  
[0, 1, 0, 1]  
[0, 0, 0, 0]  
[0, 1, 0, 1]  
[0, 1, 1, 1]  
[0, 1, 0, 0]  
[1, 1, 1, 1]  
[0, 0, 1, 1]  
...  
[0, 1, 0, 1]  
Finished sampling
```

Rejection sampling

If you can sample enough from the joint distribution, you can answer any probability inference question.

With enough samples, you can correctly compute:

- Probability estimates
- Conditional probability estimates
- Expectation estimates

Because your samples are a representation of the joint distribution!

[flu, und, fev, tir]

```
Sampling...
[0, 1, 0, 1]
[0, 1, 0, 1]
[0, 1, 0, 1]
[0, 0, 0, 0]
[0, 1, 0, 1]
[0, 1, 1, 1]
[0, 1, 0, 0]
[1, 1, 1, 1]
[0, 0, 1, 1]
...
[0, 1, 0, 1]
Finished sampling
```

$$P(\text{has flu} \mid \text{undergrad and is tired}) = 0.122$$

Disadvantages of rejection sampling

$$P(F_{lu} = 1 | F_{ev} = 1)?$$

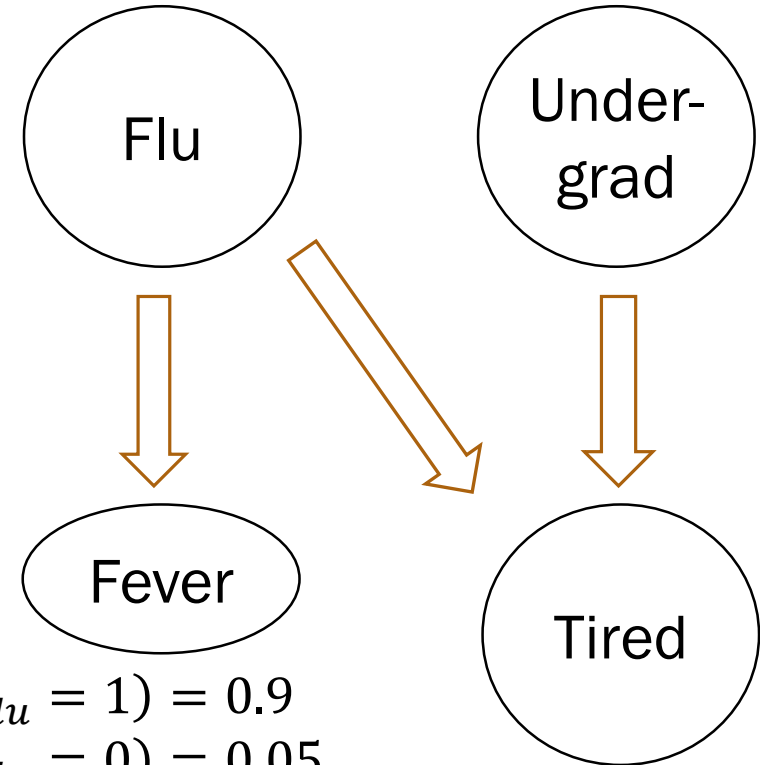
What if we never encounter some samples?

[flu=0, und, fev=1, tir]



$$P(F_{lu} = 1) = 0.1$$

$$P(U = 1) = 0.8$$



$$P(F_{ev} = 1 | F_{lu} = 1) = 0.9$$

$$P(F_{ev} = 1 | F_{lu} = 0) = 0.05$$

$$P(T = 1 | F_{lu} = 0, U = 0) = 0.1$$

$$P(T = 1 | F_{lu} = 0, U = 1) = 0.8$$

$$P(T = 1 | F_{lu} = 1, U = 0) = 0.9$$

$$P(T = 1 | F_{lu} = 1, U = 1) = 1.0$$

Disadvantages of rejection sampling

$$P(F_{lu} = 1 | F_{ev} = 99.4)?$$

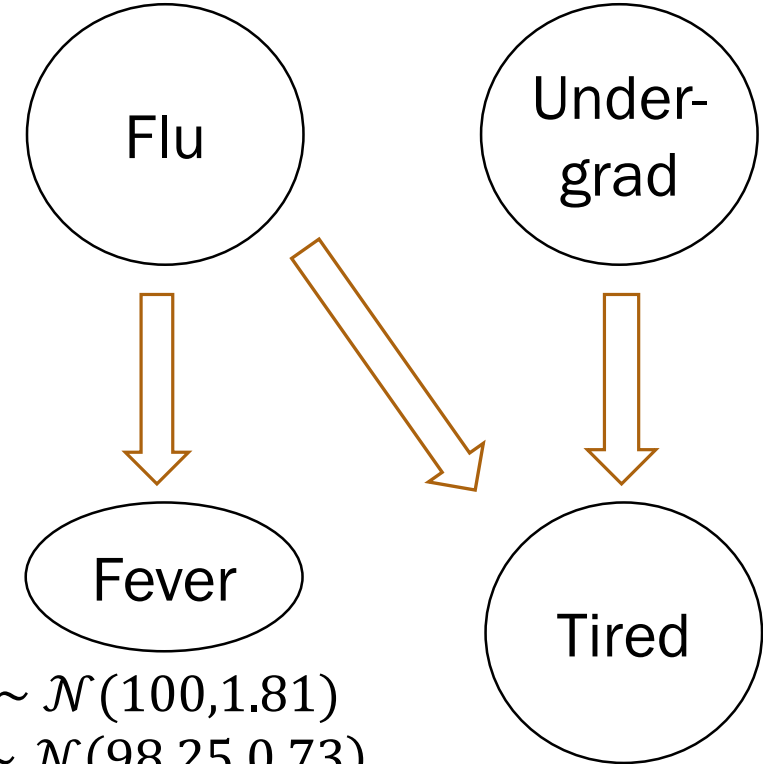
What if we never encounter some samples?

What if random variables are continuous?



$$P(F_{lu} = 1) = 0.1$$

$$P(U = 1) = 0.8$$



$$F_{ev} | F_{lu} = 1 \sim \mathcal{N}(100, 1.81)$$

$$F_{ev} | F_{lu} = 0 \sim \mathcal{N}(98.25, 0.73)$$

$$P(T = 1 | F_{lu} = 0, U = 0) = 0.1$$

$$P(T = 1 | F_{lu} = 0, U = 1) = 0.8$$

$$P(T = 1 | F_{lu} = 1, U = 0) = 0.9$$

$$P(T = 1 | F_{lu} = 1, U = 1) = 1.0$$

Gibbs Sampling (not covered)

Basic idea:

- Fix all observed events
- Incrementally sample a new value for each random variable
- Difficulty: More coding for computing different posterior probabilities

Learn in extra notebook!

(or by taking CS228/CS238)

Announcements

Problem Set 5

Due: Friday 2/28
Covers: Up to Lecture 19

Late Day Reminder

No late days permitted past
last day of the quarter, 3/13

Autograded Coding Problems

Run your code in the command line,
not just in a Jupyter notebook cell

CS109 Contest

Due: Monday 3/9 11:59pm

Today's plan

Inference:

1. Math
2. Rejection sampling (“joint” sampling)
3. Optional: Gibbs sampling (MCMC algorithm) (extra notebook)

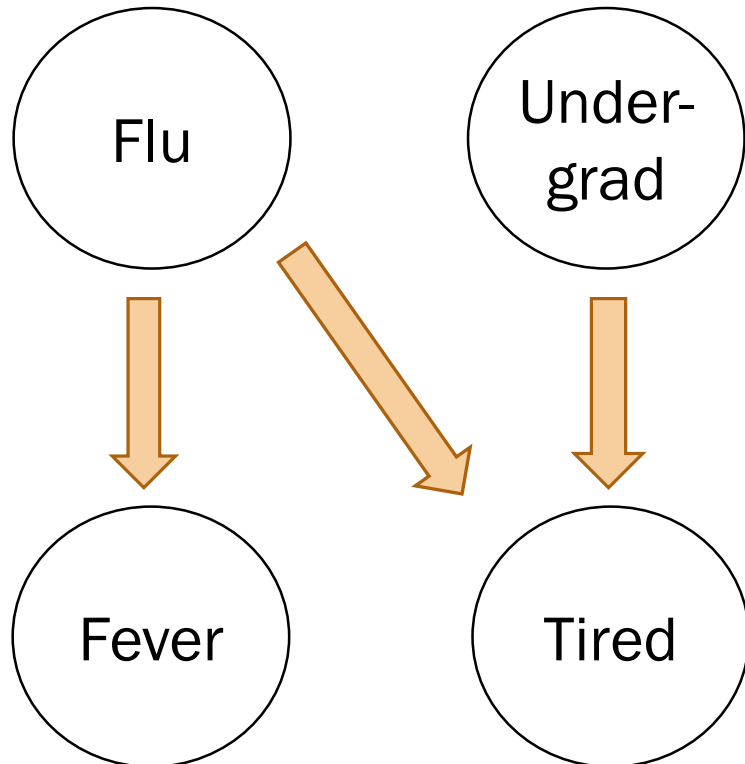
 Intro to Parameter Estimation

Maximum Likelihood Estimation (MLE)

Where do the numbers come from?

$$P(F_{lu} = 1) = 0.1$$

$$P(U = 1) = 0.8$$



Given experiment data,
how do we come up
with a reasonable
probabilistic **model**?

$$P(F_{ev} = 1 | F_{lu} = 1) = 0.9$$

$$P(F_{ev} = 1 | F_{lu} = 0) = 0.05$$

$$P(T = 1 | F_{lu} = 0, U = 0) = 0.1$$

$$P(T = 1 | F_{lu} = 0, U = 1) = 0.8$$

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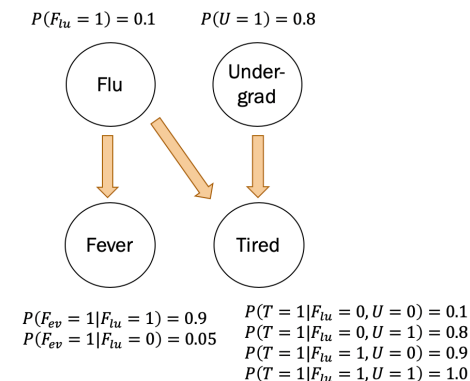
Story so far

At this point:

If you are given a **model** with all the necessary probabilities, you can make predictions.

$$Y \sim \text{Poi}(5)$$

$$X_1, \dots, X_n \text{ i.i.d.}$$
$$X \sim \text{Ber}(0.2),$$
$$X = \sum_{i=1}^n X_i$$

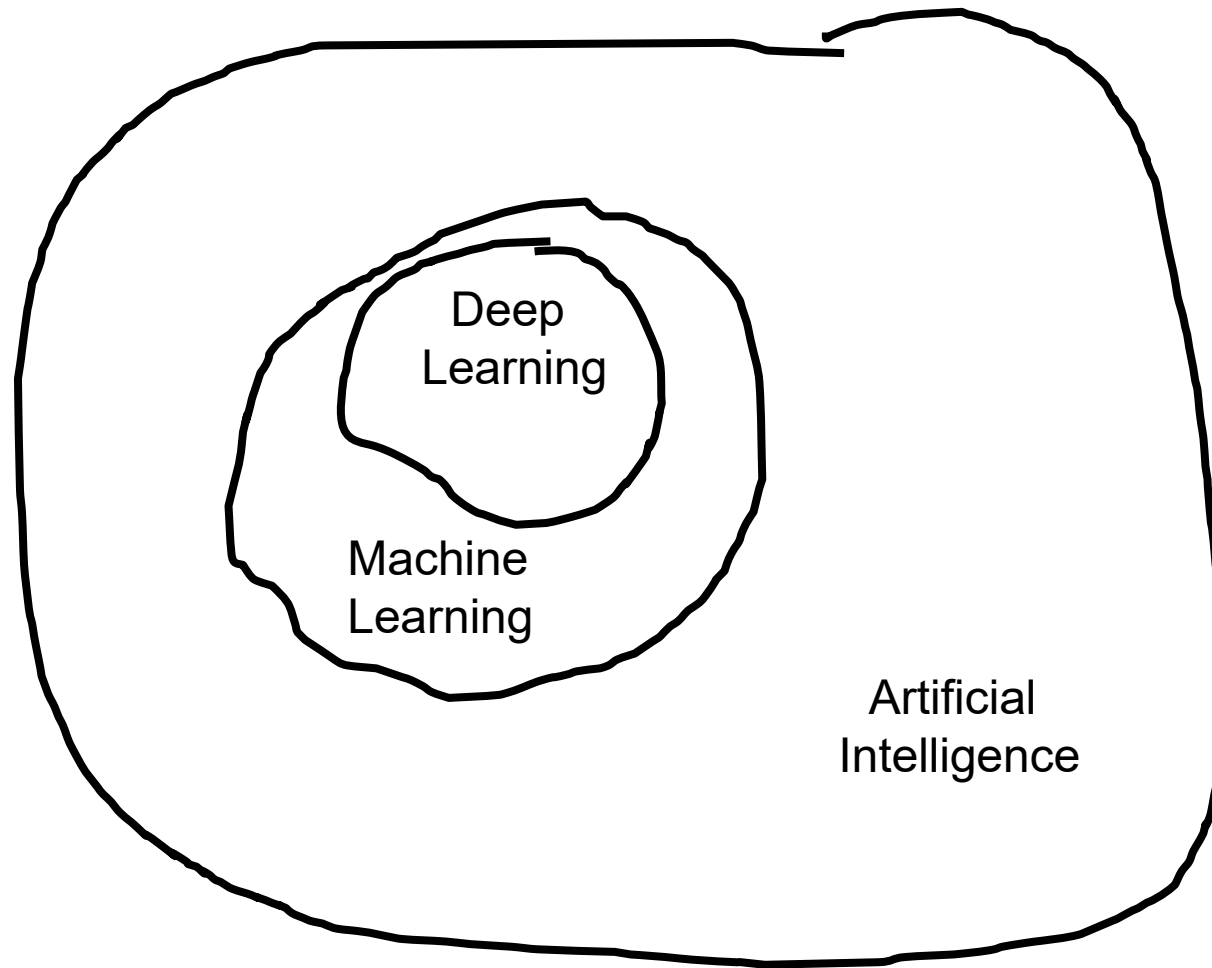


But what if you want to **learn** the probabilities in the model?

What if you want to learn the **structure** of the model, too?

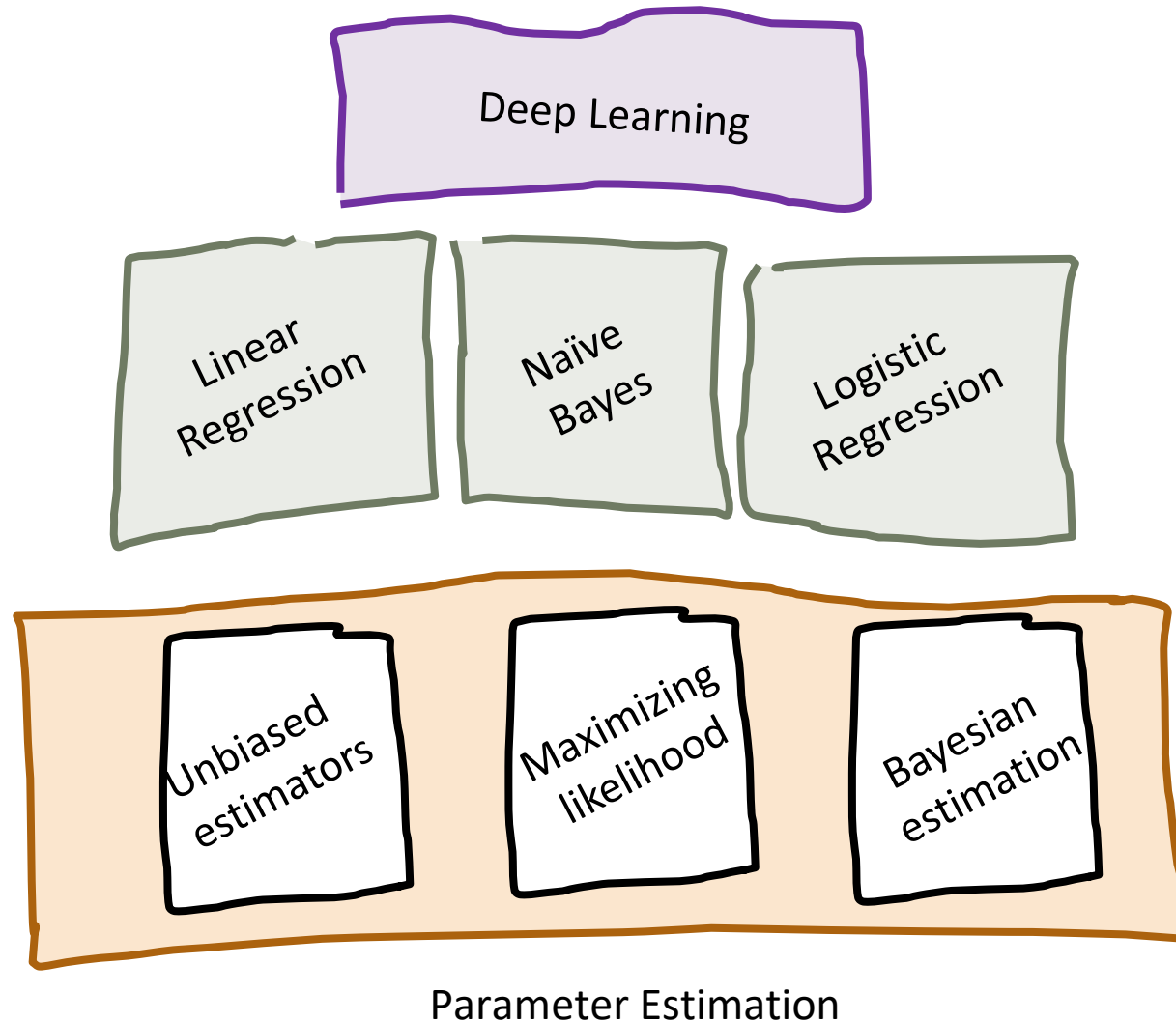
Machine Learning

AI and Machine Learning



ML: Rooted in probability theory

Our path from here



- Understand the theory to help you debug.
- Understand the theory to push on the grander challenges.

What are parameters?

def Many random variables we have learned so far are **parametric models**:

Distribution = model + parameter θ

ex The distribution $\text{Ber}(0.2)$ = Bernoulli model, parameter $\theta = 0.2$.

For each of the distributions below, what is the parameter θ ?

1. $\text{Ber}(p)$ $\theta = p$

2. $\text{Poi}(\lambda)$

3. $\text{Uni}(\alpha, \beta)$

4. $\mathcal{N}(\mu, \sigma^2)$

5. $Y = mX + b$

What are parameters?

def Many random variables we have learned so far are **parametric models**:

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For each of the distributions below, what is the parameter θ ?

1. $\text{Ber}(p)$ $\theta = p$
2. $\text{Poi}(\lambda)$ $\theta = \lambda$
3. $\text{Uni}(\alpha, \beta)$ $\theta = (\alpha, \beta)$
4. $\mathcal{N}(\mu, \sigma^2)$ $\theta = (\mu, \sigma^2)$
5. $Y = mX + b$ $\theta = (m, b)$

θ is the parameter of a distribution.
 θ can be a vector of parameters!

Why do we care?

In real world, we don't know the “true” parameters.

- But we do get to **observe data**: (# times coin comes up heads, lifetimes of disk drives produced, # visitors to website per day, etc.)

def **estimator** $\hat{\theta}$: random variable estimating parameter θ from data.

In parameter estimation,

We use the **point estimate** of parameter estimate (best single value):

- Better understanding of the process producing data
- Future **predictions** based on model
- Simulation of future processes

Today's plan

Inference:

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Intro to Parameter Estimation

 Maximum Likelihood Estimation (MLE)

Recall some estimators

Consider n i.i.d. random variables X_1, X_2, \dots, X_n .

- The sequence X_1, X_2, \dots, X_n is a **sample** from distribution F .
- X_i have distribution F with $E[X_i] = \mu, \text{Var}(X_i) = \sigma^2$.

Sample mean:

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

unbiased **estimate** of μ

$$E[\bar{X}] = \mu$$

Sample variance:

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

unbiased estimate of σ^2

$$E[S^2] = \sigma^2$$

Estimating a Bernoulli parameter

Consider n i.i.d. random variables X_1, X_2, \dots, X_n .

- The sequence X_1, X_2, \dots, X_n is a **sample** from distribution F .
- X_i have distribution F with $E[X_i] = \mu, \text{Var}(X_i) = \sigma^2$.
- Suppose distribution $F = \text{Ber}(\theta)$ with unknown parameter θ .
- Say you have three estimates $\hat{\theta}$: $\hat{\theta} = 0.5$, $\hat{\theta} = 0.8$, or $\hat{\theta} = 1$

Which estimate is most likely to give you the following sample ($n = 10$)?

$[0, 0, 1, 1, 1, 1, 1, 1, 1, 1]$

Estimating a Bernoulli parameter

Consider n i.i.d. random variables X_1, X_2, \dots, X_n .

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Which estimate is most likely to give you the following sample ($n = 10$)?

$[0, 0, 1, 1, 1, 1, 1, 1, 1, 1]$

$$P(\text{sample} | \theta = 0.5) = (0.5)^2 (0.5)^8 = 0.00097$$

$$P(\text{sample} | \theta = 0.8) = (0.2)^2 (0.8)^8 = 0.00671 \leftarrow \text{Estimate } \hat{\theta} = 0.8$$

$$P(\text{sample} | \theta = 1.0) = (0)^2 (1.0)^8 = 0$$

Defining the likelihood of data

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- X_i was drawn from a distribution with density function $f(X_i|\theta)$.
- Observed data: (x_1, x_2, \dots, x_n)

Note: now explicitly specify parameter θ of distribution

Likelihood question:

How likely is the observed data (x_1, x_2, \dots, x_n) given parameter θ ?

Likelihood function, $L(\theta)$:

$$L(\theta) = \prod_{i=1}^n f(X_i|\theta)$$

This is just a product, since X_i are i.i.d.

Maximum Likelihood Estimator

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

def The **Maximum Likelihood Estimator (MLE)** of θ is the value of θ that maximizes $L(\theta)$.

$$\theta_{MLE} = \arg \max_{\theta} L(\theta)$$

Maximum Likelihood Estimator

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Likelihood Function

$$L(\theta) = \prod_{i=1}^n f(X_i|\theta)$$

For continuous X_i , $f(X_i|\theta)$ is PDF; for discrete X_i , $f(X_i|\theta)$ is PMF

Maximum Likelihood Estimator

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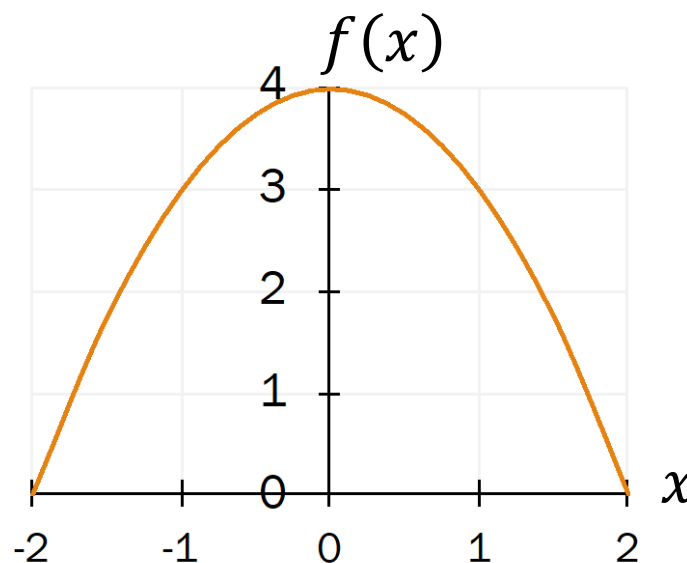
The argument θ
that maximizes $L(\theta)$

New function: arg max

$$\arg \max_x f(x)$$

The x that maximizes the function $f(x)$.

Let $f(x) = -x^2 + 4$,
where $-2 < x < 2$.



1. $\max_x f(x) ?$

2. $\arg \max_x f(x) ?$

Argmax properties

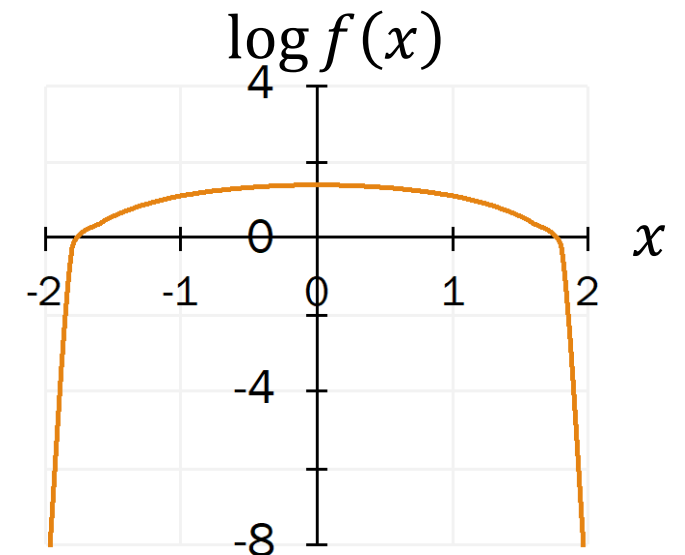
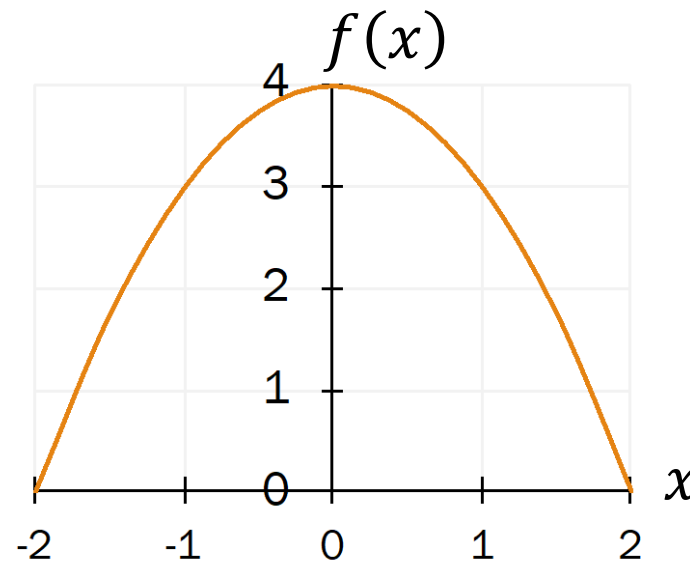
$$\arg \max_x f(x)$$

The x that maximizes the function $f(x)$.

$$= \arg \max_x \log f(x)$$

Let $f(x) = -x^2 + 4$,
where $-2 < x < 2$.

$$\arg \max_x f(x) = 0$$



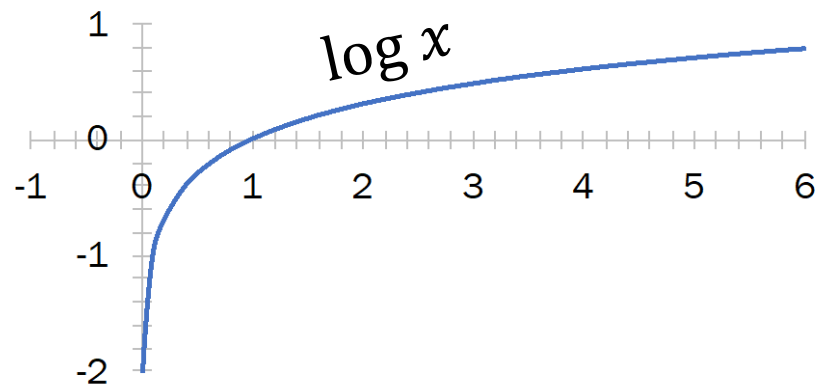
Argmax properties

$$\arg \max_x f(x)$$

The x that maximizes the function $f(x)$.

$$= \arg \max_x \log f(x)$$

- Log is **monotonic**:
 $x \leq y \Leftrightarrow \log x \leq \log y$



- Log of product = sum of logs:
 $\log(ab) = \log a + \log b$

Argmax properties

$$\arg \max_x f(x)$$

The x that maximizes the function $f(x)$.

$$= \arg \max_x \log f(x)$$

(log is monotonic:
 $x \leq y \Leftrightarrow \log x \leq \log y$)

$$= \arg \max_x (c \log f(x))$$

($x \leq y \Leftrightarrow c \log x \leq c \log y$)

for any positive constant c

Maximum Likelihood Estimator

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

def The Maximum Likelihood Estimator (MLE) of θ is the value of θ that maximizes $L(\theta)$.

$$\theta_{MLE} = \arg \max_{\theta} L(\theta)$$

θ_{MLE} also maximizes the **log-likelihood function** $LL(\theta)$:

$$LL(\theta) = \log L(\theta) = \log \left(\prod_{i=1}^n f(X_i|\theta) \right) = \sum_{i=1}^n \log f(X_i|\theta)$$

$$\theta_{MLE} = \arg \max_{\theta} LL(\theta)$$

(log is monotonic)

Story so far

- We want to estimate a parameter θ for a density $f(X_i|\theta)$.
- Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

$$\text{Likelihood } L(\theta) = \prod_{i=1}^n f(X_i|\theta) \quad \text{Log-likelihood } LL(\theta) = \sum_{i=1}^n \log f(X_i|\theta)$$

- We can choose θ by finding the argmax of the log-likelihood of data:

$$\theta_{MLE} = \arg \max_{\theta} LL(\theta) = \arg \max_{\theta} \sum_{i=1}^n \log f(X_i|\theta)$$

Computing the MLE

$$\theta_{MLE} = \arg \max_{\theta} LL(\theta)$$

General approach for finding θ_{MLE} , the MLE of θ :

1. Determine formula for $LL(\theta)$

$$LL(\theta) = \sum_{i=1}^n \log f(X_i|\theta)$$

2. Differentiate $LL(\theta)$ w.r.t. (each) θ

$$\frac{\partial LL(\theta)}{\partial \theta}$$

3. Solve resulting (simultaneous) equations

To maximize:
$$\frac{\partial LL(\theta)}{\partial \theta} = 0$$

(algebra or computer)

4. Make sure derived $\hat{\theta}_{MLE}$ is a maximum

- Check $LL(\theta_{MLE} \pm \epsilon) < LL(\theta_{MLE})$
- Often ignored in expository derivations
- We'll ignore it here too (and won't require it in class)

Maximum Likelihood with Bernoulli

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- Let $X_i \sim \text{Ber}(p)$.

What is $\theta_{MLE} = p_{MLE}$?

1. Determine formula for $LL(\theta)$

$$LL(\theta) = \sum_{i=1}^n \log f(X_i|p)$$

2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0

3. Solve resulting (simultaneous) equations

What is the PMF $f(X_i|p)$?

A. p

B. $1 - p$

C. $\begin{cases} p & \text{if } X_i = 1 \\ 1 - p & \text{if } X_i = 0 \end{cases}$

D. $p^{X_i}(1 - p)^{1 - X_i}$ where $X_i \in \{0, 1\}$

Maximum Likelihood with Bernoulli

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- Let $X_i \sim \text{Ber}(p)$.

What is $\theta_{MLE} = p_{MLE}$?

1. Determine formula for $LL(\theta)$

$$LL(\theta) = \sum_{i=1}^n \log f(X_i|p)$$

- Is differentiable
- Valid PMF over discrete domain



2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0

3. Solve resulting equations

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Maximum Likelihood with Bernoulli

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- Let $X_i \sim \text{Ber}(p)$.
- $f(X_i|p) = p^{X_i}(1-p)^{1-X_i}$ where $X_i \in \{0,1\}$

What is $\theta_{MLE} = p_{MLE}$?

1. Determine formula for $LL(\theta)$

2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0

3. Solve resulting equations

$$\begin{aligned} LL(\theta) &= \sum_{i=1}^n \log f(X_i|p) \\ &= \sum_{i=1}^n \log(p^{X_i}(1-p)^{1-X_i}) = \sum_{i=1}^n [X_i \log p + (1-X_i) \log(1-p)] \\ &= Y(\log p) + (n-Y) \log(1-p), \text{ where } Y = \sum_{i=1}^n X_i \end{aligned}$$

Maximum Likelihood with Bernoulli

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- Let $X_i \sim \text{Ber}(p)$.
- $f(X_i|p) = p^{X_i}(1-p)^{1-X_i}$ where $X_i \in \{0,1\}$

What is $\theta_{MLE} = p_{MLE}$?

1. Determine formula for $LL(\theta)$
2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0
3. Solve resulting equations

$$LL(\theta) = \sum_{i=1}^n [X_i \log p + (1 - X_i) \log(1 - p)] = Y(\log p) + (n - Y) \log(1 - p) \quad \text{where } Y = \sum_{i=1}^n X_i$$

$$\frac{\partial LL(\theta)}{\partial p} = Y \frac{1}{p} + (n - Y) \frac{-1}{1 - p} = 0$$

Maximum Likelihood with Bernoulli

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$$\frac{\partial LL(\theta)}{\partial p} = Y \frac{1}{p} + (n - Y) \frac{-1}{1 - p} = 0$$

$$p_{MLE} = \frac{1}{n} Y = \frac{1}{n} \sum_{i=1}^n X_i$$

MLE of the Bernoulli parameter, p_{MLE} , is the unbiased estimate of the mean, \bar{X} (sample mean)

Quick check

- You draw n i.i.d. random variables X_1, X_2, \dots, X_n from the distribution F , yielding the following sample:

$[0, 0, 1, 1, 1, 1, 1, 1, 1, 1]$ $(n = 10)$

- Suppose distribution $F = \text{Ber}(p)$ with unknown parameter p .

1. What is p_{MLE} , the MLE of the parameter p ?

- A. 1.0
- B. 0.5
- C. 0.8
- D. 0.2
- E. None/other

Quick check

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- Suppose distribution $F = \text{Ber}(p)$ with unknown parameter p .
1. What is p_{MLE} , the MLE of the parameter p ?
 2. What is the likelihood $L(\theta)$ of this particular sample?

Quick check

- You draw n i.i.d. random variables X_1, X_2, \dots, X_n from the distribution F , yielding the following sample:

$$[0, 0, 1, 1, 1, 1, 1, 1, 1, 1] \quad (n = 10)$$

- Suppose distribution $F = \text{Ber}(p)$ with unknown parameter p .

1. What is p_{MLE} , the MLE of the parameter p ?
2. What is the likelihood $L(\theta)$ of this particular sample?

$$f(X_i|p) = p^{X_i}(1-p)^{1-X_i} \text{ where } X_i \in \{0,1\}$$

$$L(\theta) = \prod_{i=1}^n f(X_i|p) \quad \text{where } \theta = p$$

$$= p^8(1-p)^2$$

Maximum Likelihood Algorithm

1. Decide on a model for the distribution of your samples.
Define the PMF/PDF for the distribution.

$$f(X_i|p)$$

2. Write out the log-likelihood function.

$$LL(\theta) = \sum_{i=1}^n \log f(X_i|p)$$

3. State that the optimal parameters are the argmax of the log-likelihood function.

$$\theta_{MLE} = \arg \max_{\theta} LL(\theta)$$

4. Use an optimization algorithm to calculate argmax:
 - Differentiate $LL(\theta)$ w.r.t (each) θ , set to 0
 - Solve resulting (simultaneous) equations

Maximum Likelihood with Poisson

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- Let $X_i \sim \text{Poi}(\lambda)$.
- PMF: $f(X_i|\lambda) = \frac{e^{-\lambda} \lambda^{X_i}}{X_i!}$

What is $\theta_{MLE} = \lambda_{MLE}$?

Maximum Likelihood with Poisson

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What is $\theta_{MLE} = \lambda_{MLE}$?

1. Determine
formula for $LL(\theta)$

2. Differentiate $LL(\theta)$
w.r.t. (each) θ , set to 0

3. Solve resulting
equations

$$\begin{aligned} LL(\theta) &= \sum_{i=1}^n \log \left(\frac{e^{-\lambda} \lambda^{X_i}}{X_i!} \right) = \sum_{i=1}^n -\lambda \log e + X_i \log \lambda - \log X_i! \\ &= -n\lambda + \log(\lambda) \sum_{i=1}^n X_i - \sum_{i=1}^n \log(X_i!) \quad (\text{using natural log, } \ln e = 1) \end{aligned}$$

Maximum Likelihood with Poisson

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- Let $X_i \sim \text{Poi}(\lambda)$.
- PMF: $f(X_i|\lambda) = \frac{e^{-\lambda} \lambda^{X_i}}{X_i!}$

What is $\theta_{MLE} = \lambda_{MLE}$?

1. Determine formula for $LL(\theta)$
2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0
3. Solve resulting equations

$$LL(\theta) = -n\lambda + \log(\lambda) \sum_{i=1}^n X_i - \sum_{i=1}^n \log(X_i!)$$

$$\frac{\partial LL(\theta)}{\partial \lambda} = -n + \frac{1}{\lambda} \sum_{i=1}^n X_i = 0 \quad (\sum_{i=1}^n \log(X_i!) \text{ is a constant w.r.t } \lambda)$$

Maximum Likelihood with Poisson

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

- Let $X_i \sim \text{Poi}(\lambda)$.
- PMF: $f(X_i|\lambda) = \frac{e^{-\lambda} \lambda^{X_i}}{X_i!}$

What is $\theta_{MLE} = \lambda_{MLE}$?

1. Determine formula for $LL(\theta)$
2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0
3. Solve resulting equations

$$LL(\theta) = -n\lambda + \log(\lambda) \sum_{i=1}^n X_i - \sum_{i=1}^n \log(X_i!)$$

$$\frac{\partial LL(\theta)}{\partial \lambda} = -n + \frac{1}{\lambda} \sum_{i=1}^n X_i = 0 \quad \Rightarrow \quad \lambda_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i$$

MLE of the Poisson parameter, λ_{MLE} , is the unbiased estimate of the mean, \bar{X} (sample mean)

Maximum Likelihood with Uniform

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

Let $X_i \sim \text{Uni}(\alpha, \beta)$.

$$f(X_i | \alpha, \beta) = \begin{cases} \frac{1}{\beta - \alpha} & \text{if } \alpha \leq X_i \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

1. Determine formula for $L(\theta)$

2. Differentiate $LL(\theta)$ w.r.t. (each) θ , set to 0

Likelihood:

$$L(\theta) = \begin{cases} \left(\frac{1}{\beta - \alpha}\right)^n & \text{if } \alpha \leq X_1, X_2, \dots, X_n \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

A. Great, let's do it

B. Differentiation is hard

C. Constraint

$$\alpha \leq X_1, X_2, \dots, X_n \leq \beta$$

makes differentiation hard

Example sample from a Uniform

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

Let $X_i \sim \text{Uni}(\alpha, \beta)$.

$$L(\theta) = \begin{cases} \left(\frac{1}{\beta - \alpha}\right)^n & \text{if } \alpha \leq X_1, X_2, \dots, X_n \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

Suppose $X_i \sim \text{Uni}(0, 1)$. [0.15, 0.20, 0.30, 0.40, 0.65, 0.70, 0.75]

You observe data:

- Which parameters would give you maximum $L(\theta)$?
- A. $\text{Uni}(\alpha = 0, \beta = 1)$
 - B. $\text{Uni}(\alpha = 0.15, \beta = 0.75)$
 - C. $\text{Uni}(\alpha = 0.15, \beta = 0.70)$

Example sample from a Uniform

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

Let $X_i \sim \text{Uni}(\alpha, \beta)$.

$$L(\theta) = \begin{cases} \left(\frac{1}{\beta - \alpha}\right)^n & \text{if } \alpha \leq X_1, X_2, \dots, X_n \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

Suppose $X_i \sim \text{Uni}(0, 1)$. [0.15, 0.20, 0.30, 0.40, 0.65, 0.70, 0.75]

You observe data:

- Which parameters would give you maximum $L(\theta)$?
- A. $\text{Uni}(\alpha = 0, \beta = 1)$ $(1)^7 = 1$
 - B.** $\text{Uni}(\alpha = 0.15, \beta = 0.75)$ $\left(\frac{1}{0.6}\right)^7 = 35.7$
 - C. $\text{Uni}(\alpha = 0.15, \beta = 0.70)$ $\left(\frac{1}{0.55}\right)^6 \cdot 0 = 0$

Original parameters may not yield maximum likelihood.

Maximum Likelihood with Uniform

Consider a sample of n i.i.d. random variables X_1, X_2, \dots, X_n .

Let $X_i \sim \text{Uni}(\alpha, \beta)$.

$$L(\theta) = \begin{cases} \left(\frac{1}{\beta - \alpha}\right)^n & \text{if } \alpha \leq X_1, X_2, \dots, X_n \leq \beta \\ 0 & \text{otherwise} \end{cases}$$

$$\theta_{MLE}: \alpha_{MLE} = \min(x_1, x_2, \dots, x_n) \quad \beta_{MLE} = \max(x_1, x_2, \dots, x_n)$$

Intuition:

- Want interval size $(\beta - \alpha)$ to be as small as possible to maximize likelihood function per datapoint
- Need to make sure all observed data is in interval (if not, then $L(\theta) = 0$)

(demo)

Small samples = problems with MLE

Maximum Likelihood Estimator θ_{MLE} :

- Best explains data we have seen
- Does not attempt to generalize to unseen data.

$$\theta_{MLE} = \arg \max_{\theta} L(\theta)$$

$$= \arg \max_{\theta} LL(\theta)$$

In many cases, $\mu_{MLE} = \frac{1}{n} \sum_{i=1}^n X_i$ Sample mean

(MLE for Bernoulli p ,
Poisson λ , Normal μ)

- Unbiased ($E[\mu_{MLE}] = \mu$ regardless of size of sample, n)

For some cases, like Uniform: $\alpha_{MLE} \geq \alpha$, $\beta_{MLE} \leq \beta$

- Biased. Problematic for small sample size
- Example: If $n = 1$ then $\alpha = \beta$, yielding an invalid distribution

Properties of MLE

Maximum Likelihood Estimator:

- Best explains data we have seen
- Does not attempt to generalize to unseen data.

$$\theta_{MLE} = \arg \max_{\theta} L(\theta)$$

$$= \arg \max_{\theta} LL(\theta)$$

-
- Often used when sample size n is large relative to parameter space
 - Potentially **biased** (though asymptotically less so, as $n \rightarrow \infty$)
 - **Consistent:** $\lim_{n \rightarrow \infty} P(|\hat{\theta} - \theta| < \varepsilon) = 1$ where $\varepsilon > 0$

As $n \rightarrow \infty$ (i.e., more data), probability that $\hat{\theta}$ significantly differs from θ is zero