## 25: Logistic Regression

Lisa Yan June 3, 2020

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## Background

#### 1. Weighted sum

If 
$$X = (X_1, X_2, ..., X_m)$$
:

$$Z = \theta_1 X_1 + \theta_2 X_2 + \dots + \theta_m X_m$$

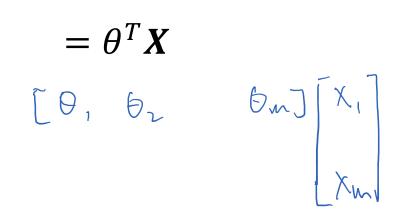
$$=\sum_{j=1}^m \theta_j X_j$$

weighted sum

$$= \theta^T X$$

dot product

$$[0, b_{2}]$$



#### 1. Weighted sum

Dot product/ weighted sum  $\theta^T X = \sum_{j=1}^m \theta_j X_j$ 

Recall the linear regression model, where  $X = (X_1, X_2, ..., X_m)$  and  $Y \in \mathbb{R}$ :

$$\widehat{Y} = g(X) = \theta_0 + \sum_{j=1}^m \theta_j X_j$$

How would you rewrite this expression as a single dot product?



#### 1. Weighted sum

Dot product/ weighted sum  $\theta^T X = \sum_{j=1}^m \theta_j X_j$ 

Recall the linear regression model, where  $X = (X_1, X_2, ..., X_m)$  and  $Y \in \mathbb{R}$ :

$$g(\mathbf{X}) = \theta_0 + \sum_{j=1}^m \theta_j X_j$$

How would you rewrite this expression as a single dot product?

$$g(\mathbf{X}) = \theta_0 X_0 + \theta_1 X_1 + \theta_2 X_2 + \dots + \theta_m X_m \qquad \text{Define } X_0 = 1$$

$$= \theta^T \mathbf{X} \qquad \text{New } \mathbf{X} = (1, X_1, X_2, \dots, X_m) \quad \theta^T \left( \mathbf{Q}_0, \mathbf{Q}_1, \dots, \mathbf{Q}_m \right)$$

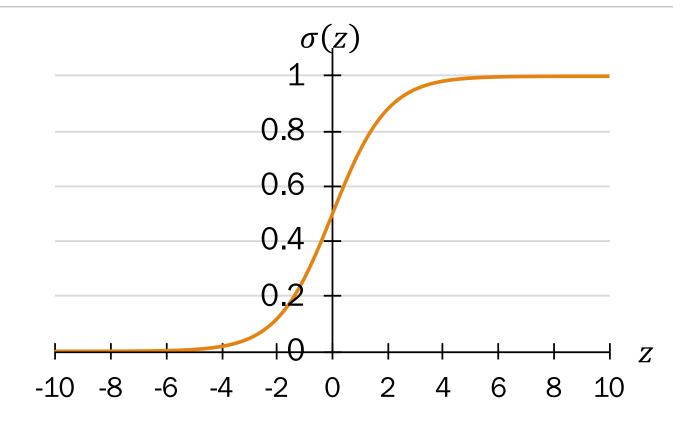
Prepending  $X_0 = 1$  to each feature vector X makes matrix operators more accessible.

#### 2. Sigmoid function $\sigma(z)$

The sigmoid function:

$$\sigma(z) = \frac{1}{1 + e^{-z}}$$

 Sigmoid squashes z to a number between 0 and 1.



Recall definition of probability:
 A number between 0 and 1

 $\sigma(z)$  can represent a probability.

#### 3. Conditional likelihood function

#### Training data (*n* datapoints):

•  $(x^{(i)}, y^{(i)})$  drawn i.i.d. from a distribution  $f(X = x^{(i)}, Y = y^{(i)}|\theta) = f(x^{(i)}, y^{(i)}|\theta)$ 

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta)$$

$$= \arg\max_{\theta} \sum_{i=1}^{n} \log f(y^{(i)} | \mathbf{x}^{(i)}, \theta)$$

$$= \arg\max_{\theta} LL(\theta)$$

### conditional likelihood of training data

log conditional likelihood

- MLE in this lecture is estimator that maximizes <u>conditional likelihood</u>
- Confusingly, log conditional likelihood is also written as  $LL(\theta)$

## Logistic Regression

#### Linear Regression (Regression)



$$\theta_0 + \sum_{j=1}^m \theta_j X_j$$

 $\widehat{Y} = \theta_0 + \sum_{i=1}^m \theta_i X_i$ 



 $\bigvee X$  can be dependent

 $\Re$  Regression model ( $\widehat{Y} \in \mathbb{R}$ , not discrete)

#### Naïve Bayes (Classification)





$$\widehat{Y} = \underset{y=\{0,1\}}{\operatorname{arg max}} P(Y \mid X)$$

$$= \underset{y=\{0,1\}}{\operatorname{arg max}} P(X \mid Y) P(Y)$$

- ✓ Tractable with NB assumption, but...
- $\triangle$  Realistically,  $X_i$  features not necessarily conditionally independent
- Actually models P(X, Y), not P(Y|X)?

#### Introducing Logistic Regression!



Linear Regression ideas

Classification models

+ compute power

#### Logistic Regression

$$\theta_0 + \sum_{j=1}^m \theta_j X_j$$



sigmoid function 
$$\sigma(z) = \frac{1}{1 + e^{-z}}$$



Logistic Regression Model:

$$P(Y = 1 | X = x) = \sigma \left(\theta_0 + \sum_{j=1}^m \theta_j x_j\right)$$

Predict  $\hat{Y}$  as the most likely Ygiven our observation X = x:

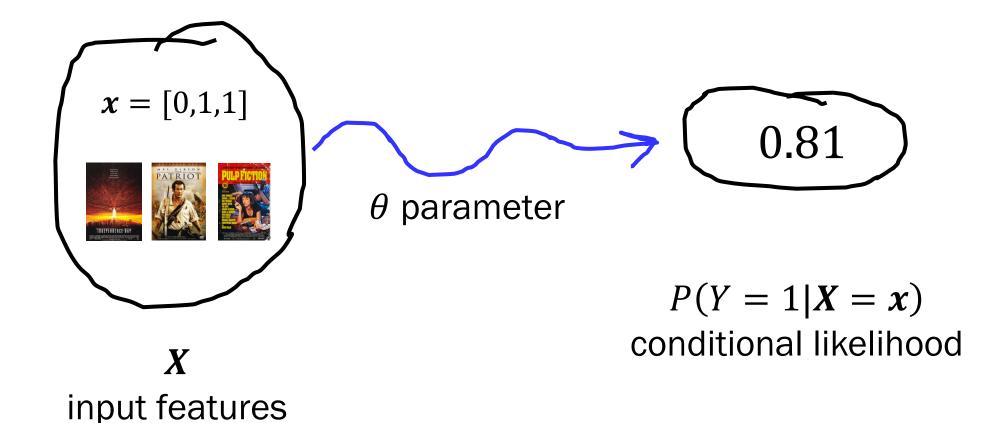
$$\widehat{Y} = \arg \max_{y = \{0,1\}} P(Y \mid X)$$

• Since 
$$Y \in \{0,1\}$$
,

$$P(Y = 0 | \mathbf{X} = \mathbf{x}) = 1 - \sigma(\theta_0 + \sum_{j=1}^m \theta_j x_j)$$

Sigmoid function also known as "logit" function

#### Logistic Regression



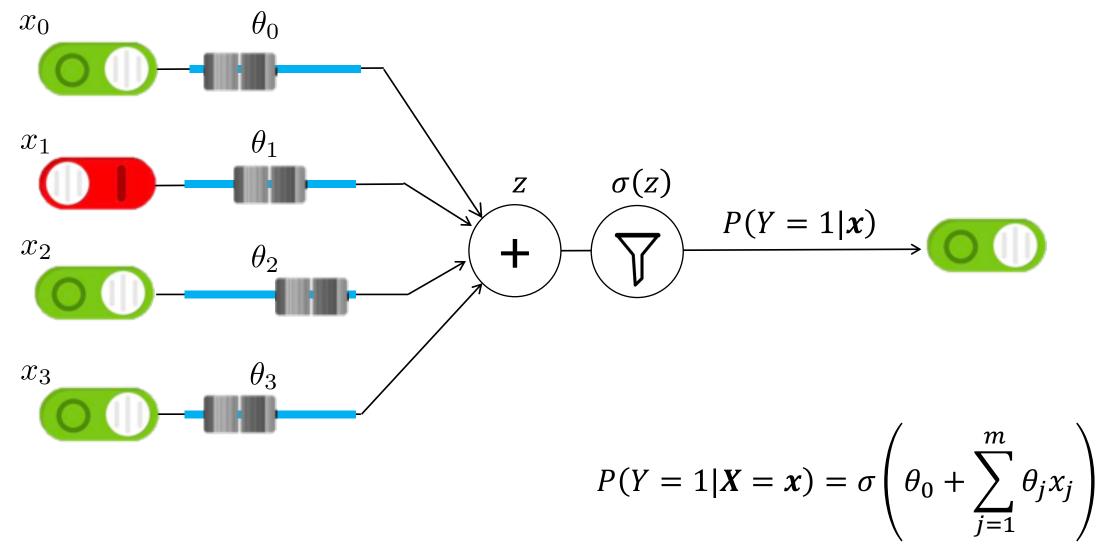
$$P(Y = 1 | X = x) = \sigma \left(\theta_0 + \sum_{j=1}^m \theta_j x_j\right)$$

#### Logistic Regression cartoon

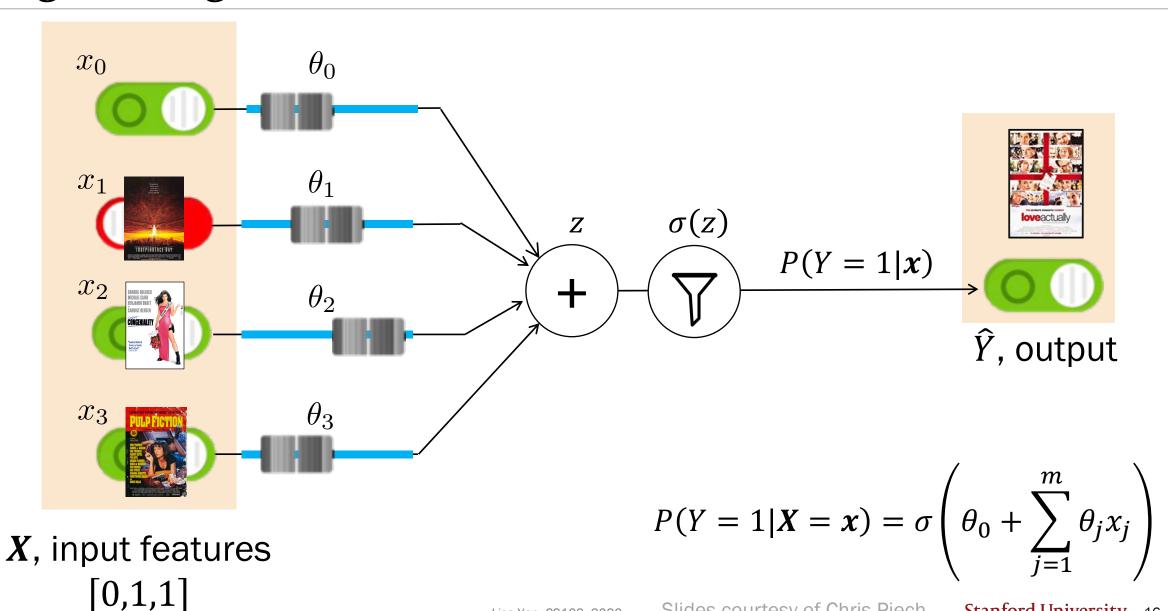


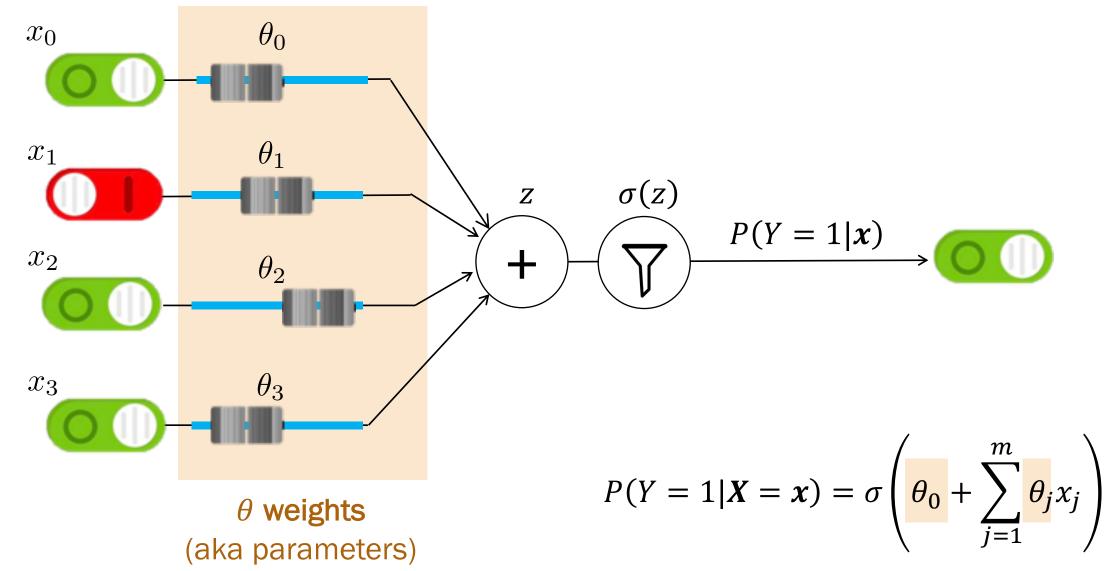
 $\theta$  parameter

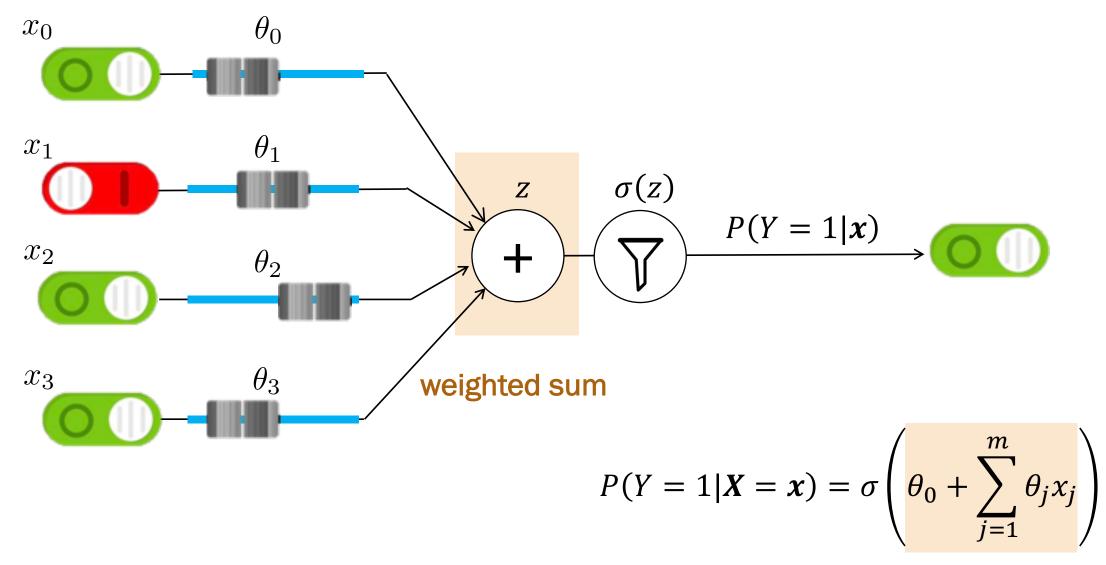
#### Logistic Regression cartoon

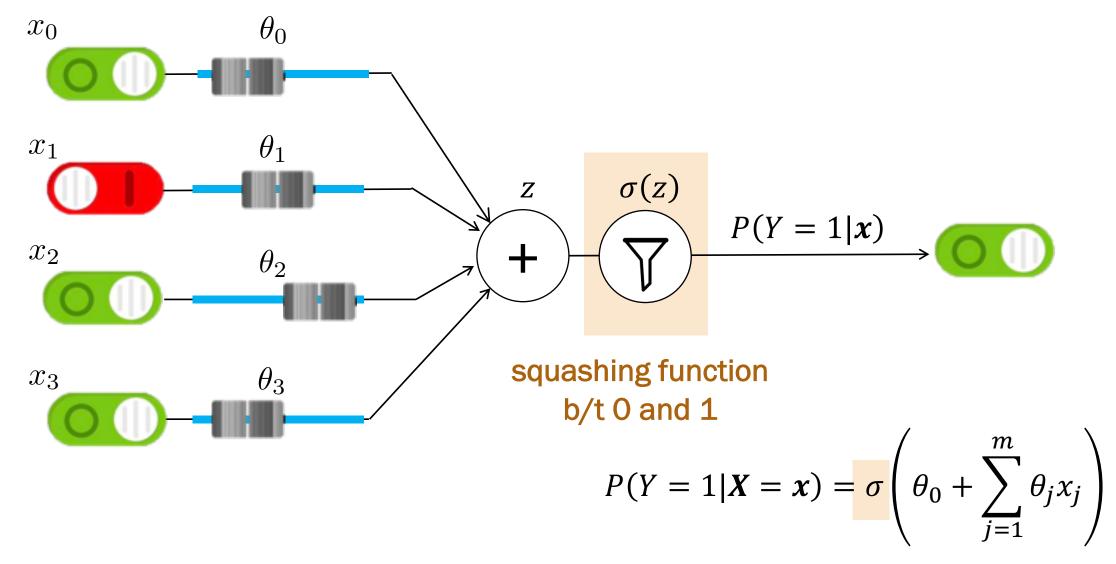


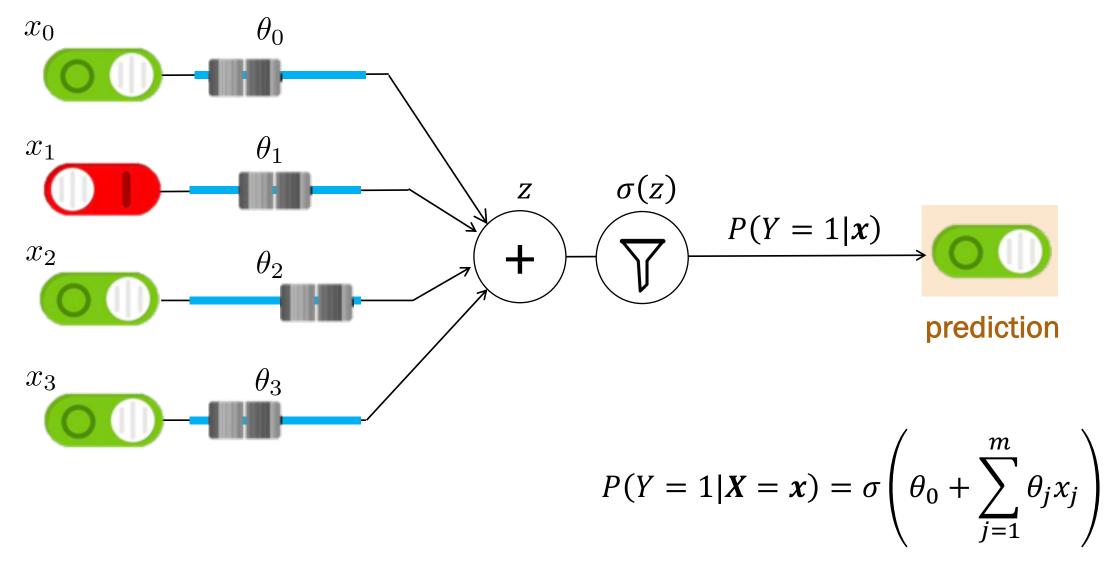
#### Logistic Regression cartoon



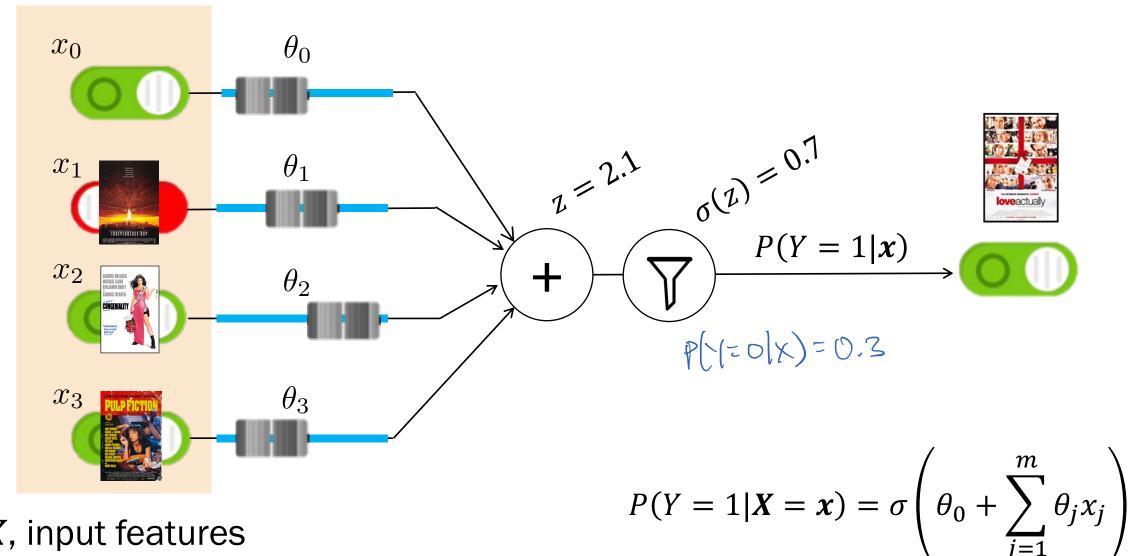






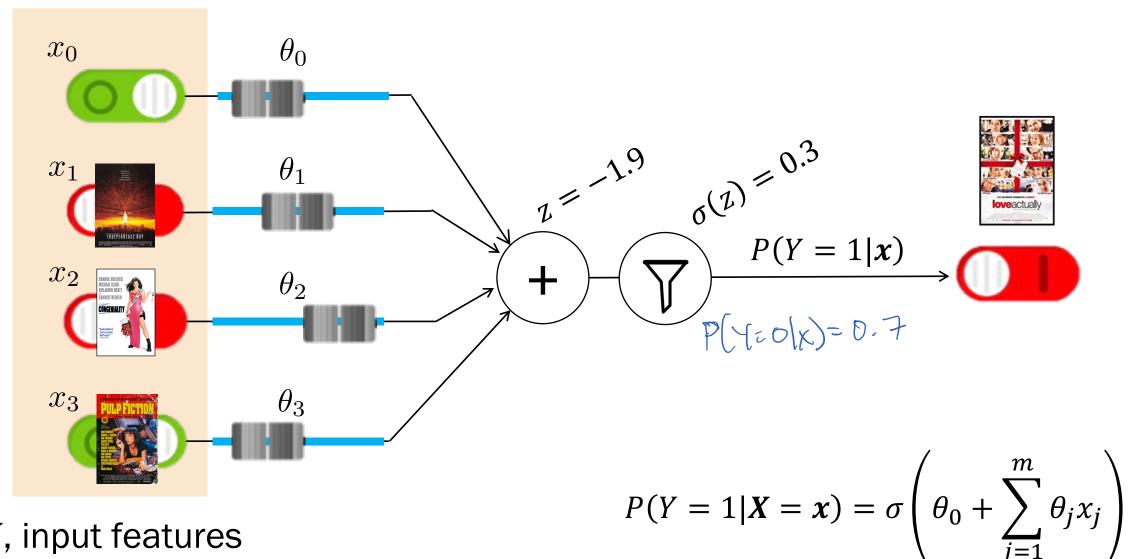


#### Different predictions for different inputs



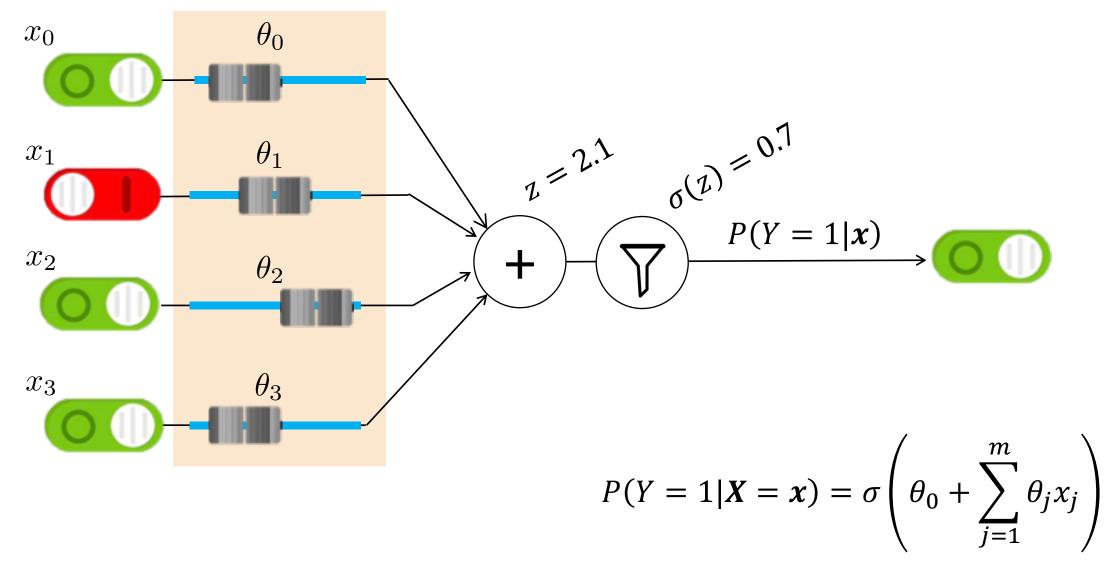
X, input features [0,1,1]

#### Different predictions for different inputs

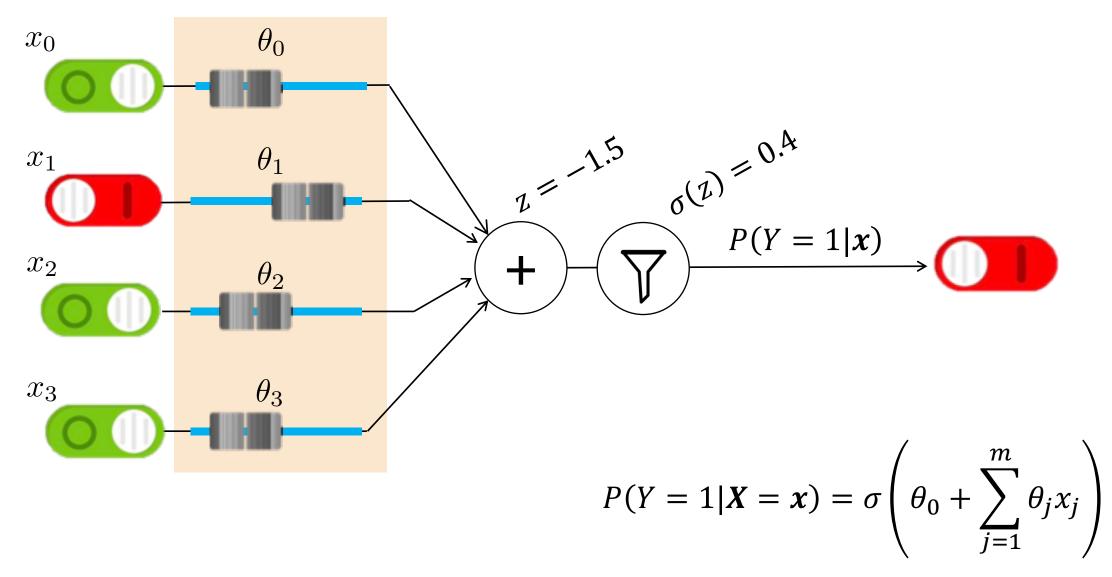


X, input features [0,0,1]

#### Parameters affect prediction



#### Parameters affect prediction



#### For simplicity

$$P(Y = 1 | X = x) = \sigma \left(\theta_0 + \sum_{j=1}^m \theta_j x_j\right)$$

$$P(Y = 1 | X = x) = \sigma \left( \sum_{j=0}^{m} \theta_{j} x_{j} \right) = \overline{\sigma(\theta^{T} x)} \text{ where } x_{0} = 1$$

$$\sum_{j=0}^{m} \overline{\sigma(\theta^{T} x)} \text{ where } x_{0} = 1$$

#### Logistic regression classifier

$$\widehat{Y} = \underset{y=\{0,1\}}{\arg \max} P(Y|X)$$

$$P(Y = 1|X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j}) = \sigma(\theta^{T} x)$$

**Training** 

Estimate parameters from training data

$$\theta = (\theta_0, \theta_1, \theta_2, \dots, \theta_m)$$

**Testing** 

Given an observation  $X = (X_1, X_2, ..., X_m)$ , predict  $\hat{Y} = \arg \max P(Y|X)$  $y = \{0,1\}$ 

# Training: The big picture

#### Logistic regression classifier

$$\hat{Y} = \arg \max_{y = \{0,1\}} P(Y|X)$$

$$P(Y = 1|X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j}) = \sigma(\theta^{T} x)$$

#### Training

Estimate parameters from training data

$$(x^{(i)},y^{(i)})$$
  $i=1,\ldots,n$ 

$$\theta = (\theta_0, \theta_1, \theta_2, \dots, \theta_m)$$

#### Choose $\theta$ that optimizes some objective:

- Determine objective function
- Find gradient with respect to  $\theta$
- Solve analytically by setting to 0, or computationally with gradient ascent

We are modeling P(Y|X)directly, so we maximize the conditional likelihood of training data.

#### Estimating $\theta$

1. Determine objective function

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta)$$

2. Gradient w.r.t.  $\theta_i$ , for j = 0, 1, ..., m

#### 3. Solve

- No analytical derivation of  $\theta_{MLE}$ ...
- ...but can still compute  $\theta_{MLE}$ with gradient ascent!

```
initialize x
repeat many times:
  compute gradient
  x += \eta * gradient
```

#### 1. Determine objective function

$$\theta_{MLE} = \arg \max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg \max_{\theta} LL(\theta)$$

$$P(Y = 1 | \mathbf{X} = \mathbf{x}) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j})$$

$$= \sigma(\theta^{T} \mathbf{x})$$

First: Interpret conditional likelihood with Logistic Regression Second: Write a differentiable expression for log conditional likelihood

#### 1. Determine objective function (interpret)

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)}|\mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta) \qquad P(Y = 1|\mathbf{X} = \mathbf{x}) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j}) \\ = \sigma(\theta^{T} \mathbf{x})$$
 Suppose you have  $n = 2$  training datapoints:  $(\mathbf{x}^{(1)}, 1), (\mathbf{x}^{(2)}, 0)$ 

Consider the following expressions for a given  $\theta$ :

A. 
$$\sigma(\theta^T \mathbf{x}^{(1)}) \sigma(\theta^T \mathbf{x}^{(2)})$$

C. 
$$\sigma(\theta^T \mathbf{x}^{(1)}) \left(1 - \sigma(\theta^T \mathbf{x}^{(2)})\right)$$

B. 
$$\left(1 - \sigma(\theta^T \boldsymbol{x}^{(1)})\right) \sigma(\theta^T \boldsymbol{x}^{(2)})$$

D. 
$$(1 - \sigma(\theta^T \mathbf{x}^{(1)})) (1 - \sigma(\theta^T \mathbf{x}^{(2)}))$$

- Interpret the above expressions as probabilities.
- If we let  $\theta = \theta_{MLE}$ , which probability should be highest?



#### 1. Determine objective function (interpret)

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$P(Y = 1 | \mathbf{X} = \mathbf{x}) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j})$$

$$= \sigma(\theta^{T} \mathbf{x})$$

Suppose you have n=2 training datapoints:

Consider the following expressions for a given  $\theta$ :

A. 
$$\sigma(\theta^T x^{(1)}) \sigma(\theta^T x^{(2)})$$

P(Y=1 |  $\chi = \chi^{(1)}$ ) P(Y=1 |  $\chi = \chi^{(2)}$ )

B.  $(1 - \sigma(\theta^T x^{(1)})) \sigma(\theta^T x^{(2)})$ 

P(Y=0 |  $\chi = \chi^{(1)}$ ) P(Y=1 |  $\chi = \chi^{(2)}$ )

$$(x^{(1)}, 1), (x^{(2)}, 0)$$
 $P(Y=5^{(2)}|X=x^{(2)})$ 
 $P(Y=5^{(2)}|X=x^{(2)})$ 

- C.  $\sigma(\theta^T x^{(1)}) \left(1 \sigma(\theta^T x^{(2)})\right)$   $P(\langle z | | \chi = \chi^{(1)}) P(\langle z | \chi = \chi^{(2)}) \rangle$
- D.  $(1 \sigma(\theta^T \mathbf{x}^{(1)})) (1 \sigma(\theta^T \mathbf{x}^{(2)}))$   $P((-\sigma(\mathbf{x} = \mathbf{x}^{(1)}))) P((-\sigma(\mathbf{x} = \mathbf{x}^{(1)})))$
- 1. Interpret the above expressions as probabilities.
- 2. If we let  $\theta = \theta_{MLE}$ , which probability should be highest?

#### 1. Determine objective function (write)

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | x^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$P(Y = 1 | X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j})$$

$$= \sigma(\theta^{T} x)$$

What is a differentiable expression for P(Y = y | X = x)?

$$P(Y = y | \mathbf{X} = \mathbf{x}) = \begin{cases} \sigma(\theta^T \mathbf{x}) & \text{if } y = 1\\ 1 - \sigma(\theta^T \mathbf{x}) & \text{if } y = 0 \end{cases}$$

2. What is a differentiable expression for  $LL(\theta)$ , log conditional likelihood?

$$LL(\theta) = \log \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta)$$



#### 1. Determine objective function (write)

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$P(Y = 1 | \mathbf{X} = \mathbf{x}) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j})$$

$$= \sigma(\theta^{T} \mathbf{x})$$

1. What is a differentiable expression for P(Y = y | X = x)?

$$P(Y = y | \mathbf{X} = \mathbf{x}) = \begin{cases} \sigma(\theta^T \mathbf{x}) & \text{if } y = 1\\ 1 - \sigma(\theta^T \mathbf{x}) & \text{if } y = 0 \end{cases}$$

Recall 
$$P(B=b)=\begin{cases} P & b=1 \\ P & b=0 \end{cases}$$
Bernoulli MLE!

$$\sigma(\Theta^T X)^{\delta} \left(1 - \sigma(\Theta^T X)\right)^{1-\delta}$$

$$\sum_{i=1}^{n} \log \left[ \sigma(\mathbf{p}^{r} \mathbf{x})^{i} \right] \left( 1 - \sigma(\mathbf{p}^{r} \mathbf{x}^{n})^{1 - \mathbf{y}^{(n)}} \right]$$

#### 1. Determine objective function (write)

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | x^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$P(Y = 1 | X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j})$$

$$= \sigma(\theta^{T} x)$$

1. What is a differentiable expression for P(Y = y | X = x)?

$$P(Y = y | X = x) = (\sigma(\theta^T x))^y (1 - \sigma(\theta^T x))^{1-y}$$

2. What is a differentiable expression for  $LL(\theta)$ , log conditional likelihood?

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^T \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log \left(1 - \sigma(\theta^T \mathbf{x}^{(i)})\right)$$

#### 2. Find gradient with respect to $\theta$

**Optimization** problem:

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log(1 - \sigma(\theta^{T} \mathbf{x}^{(i)}))$$

Gradient w.r.t.  $\theta_i$ , for j = 0, 1, ..., m:

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T \mathbf{x}^{(i)}) \right] x_j^{(i)}$$

How do we interpret the gradient contribution of the i-th training datapoint?



#### **2.** Find gradient with respect to $\theta$

Optimization problem:

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log(1 - \sigma(\theta^{T} \mathbf{x}^{(i)}))$$

Gradient w.r.t.  $\theta_j$ , for j = 0, 1, ..., m:

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T \boldsymbol{x}^{(i)}) \right] x_j^{(i)}$$

(derived later)

らてX(i) = ころり(x)

scale by j-th feature

#### 2. Find gradient with respect to $\theta$

**Optimization** problem:

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log(1 - \sigma(\theta^{T} \mathbf{x}^{(i)}))$$

Gradient w.r.t.  $\theta_i$ , for j = 0, 1, ..., m:

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \begin{bmatrix} y^{(i)} - \sigma(\theta^T x^{(i)}) \end{bmatrix} x_j^{(i)} \qquad \text{(derived later)}$$

$$1 \text{ or } 0 \quad P(Y = 1 | X = x^{(i)})$$

#### **2.** Find gradient with respect to $\theta$

Optimization problem:

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log(1 - \sigma(\theta^{T} \mathbf{x}^{(i)}))$$

Gradient w.r.t.  $\theta_i$ , for j = 0, 1, ..., m:

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T x^{(i)}) \right] x_j^{(i)}$$
 (derived later)

Suppose  $y^{(i)} = 1$  (the true class label for *i*-th datapoint):

- If  $\sigma(\theta^T x^{(i)}) \ge 0.5$ , correct
- If  $\sigma(\theta^T x^{(i)}) < 0.5$ , incorrect  $\rightarrow$  change  $\theta_i$  more

#### 3. Solve

1. Optimization problem:

$$\theta_{MLE} = \arg\max_{\theta} \prod_{i=1}^{n} f(y^{(i)} | \mathbf{x}^{(i)}, \theta) = \arg\max_{\theta} LL(\theta)$$

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log(1 - \sigma(\theta^{T} \mathbf{x}^{(i)}))$$

2. Gradient w.r.t.  $\theta_j$ , for j = 0, 1, ..., m:

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T \mathbf{x}^{(i)}) \right] x_j^{(i)}$$

3. Solve

Stay tuned!

# (live) 25: Logistic Regression

Lisa Yan June 3, 2020

#### Logistic Regression Model

$$\widehat{Y} = \arg \max_{y = \{0,1\}} P(Y|X)$$

$$P(Y = 1 | \mathbf{X} = \mathbf{x}) = \sigma \left( \sum_{j=0}^{m} \theta_{j} x_{j} \right) = \sigma(\theta^{T} \mathbf{x})$$

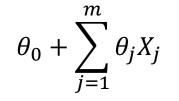
$$\hat{Y}$$
 is prediction of  $Y$ 

$$\chi = (\chi_1, \dots, \chi_m) + \chi_b = 1$$

$$\theta = (\theta_0, \theta_1, \dots, \theta_m)$$

where  $x_0 = 1$ 

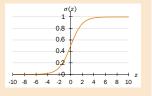






#### sigmoid function

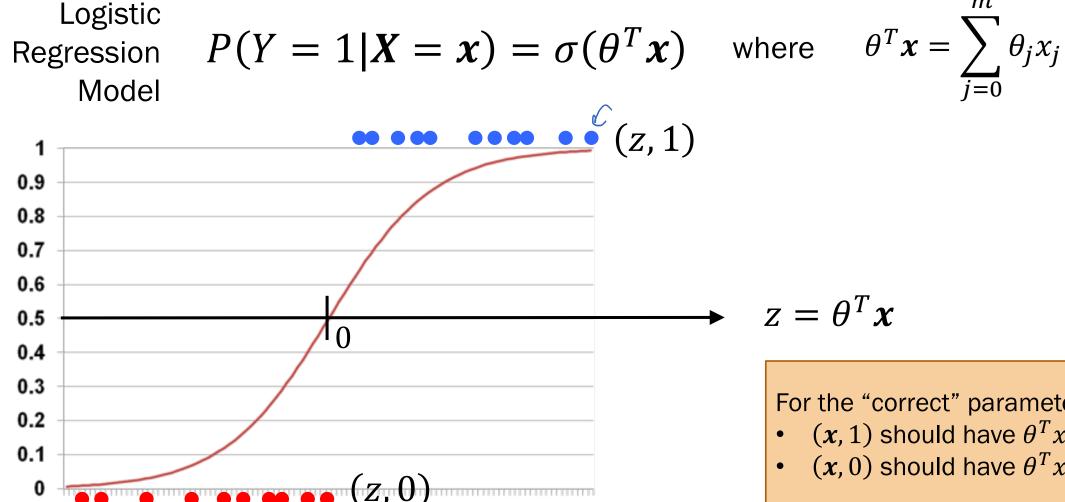
$$\sigma(z) = \frac{1}{1 + e^{-z}}$$





$$\widehat{P}(Y=1|\boldsymbol{X})$$

### Another view of Logistic Regression



$$e \quad \theta^T \mathbf{x} = \sum_{j=0}^m \theta_j x_j$$

$$z = \theta^T \mathbf{x}$$

For the "correct" parameters  $\theta$ :

- (x, 1) should have  $\theta^T x > 0$
- (x,0) should have  $\theta^T x \leq 0$

#### Learning parameters

#### **Training**

desiration)

Learn parameters  $\theta = (\theta_0, \theta_1, ..., \theta_m)$ 

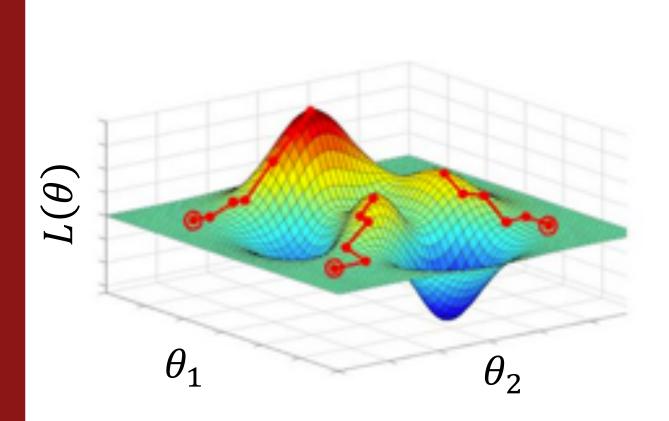
$$\theta_{MLE} = \arg\max_{\theta} LL(\theta)$$

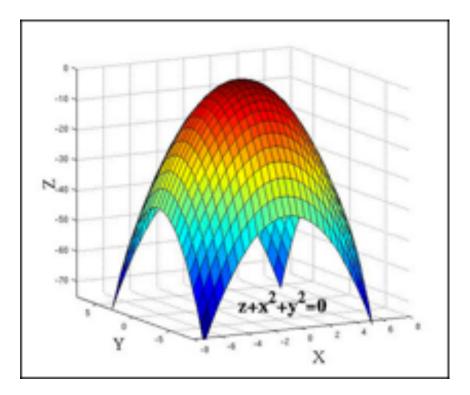
$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \boldsymbol{x}^{(i)}) + (1 - y^{(i)}) \log \left(1 - \sigma(\theta^{T} \boldsymbol{x}^{(i)})\right)$$

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T \boldsymbol{x}^{(i)}) \right] x_j^{(i)} \qquad \text{for } j = 0, 1, ..., m$$

- No analytical derivation of  $\theta_{MLE}$ ...
- ...but can still compute  $\theta_{MLE}$  with gradient ascent!

Walk uphill and you will find a local maxima (if your step is small enough).





Logistic regression  $LL(\theta)$ is concave

# Training: The details

## Training: Gradient ascent step

3. Optimize.

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \left[ y^{(i)} - \sigma(\theta^T \boldsymbol{x}^{(i)}) \right] x_j^{(i)}$$

#### repeat many times:

for all thetas:

$$\theta_{j}^{\text{new}} = \theta_{j}^{\text{old}} + \eta \cdot \frac{\partial LL(\theta^{\text{old}})}{\partial \theta_{j}^{\text{old}}}$$

$$= \theta_{j}^{\text{old}} + \eta \cdot \sum_{i=1}^{n} \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^{T}} \boldsymbol{x}^{(i)} \right) \right] x_{j}^{(i)}$$
What does this

What does this look like in code?

Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x}^{(i)} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   // compute all gradient[j]'s
   // based on n training examples
   \theta_i += \eta * gradient[j] for all 0 \le j \le m
```

Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x^{(i)}} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
     for each 0 \le j \le m:
        // update gradient[j] for
        // current (x,y) example
   \theta_i += \eta * gradient[j] for all 0 \le j \le m
```

Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x^{(i)}} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
                                                       j=[, ..., N
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
       for each 0 \le j \le m:
          gradient[j] += \left[y - \frac{1}{1 + \rho^{-\theta^T x}}\right] x_j
    \theta_i += \eta * gradient[j] for all 0 \le j \le m
```

What are the important details?



Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x^{(i)}} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
      for each 0 \le j \le m:
          gradient[j] += \left[y - \frac{1}{1 + e^{-\theta^T x}}\right] x_j
    \theta_i += \eta * gradient[j] for all 0 \le j \le m
```

•  $x_j$  is j-th feature of input  $\mathbf{x} = (x_1, ..., x_m)$ 

$$\begin{array}{l} \text{Gradient} \\ \text{Ascent Step} \ \theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \pmb{x}^{(i)} \right) \right] \, x_j^{(i)} \end{array}$$

Step D: insert 
$$X_0 = 1$$
  $(X_1, ..., X_n) \rightarrow (J_1 X_1, ..., X_n)$   $\Theta_0 + \sum_{j=1}^{n} \Theta_j X_j^*$ 

initialize  $\theta_j = 0$  for  $0 \le j \le m$  repeat many times:

```
gradient[j] = 0 for 0 ≤ j ≤ m
for each training example (x, y):
  for each 0 ≤ j ≤ m:
```

gradient[j] += 
$$\left[y - \frac{1}{1 + e^{-\theta^T x}}\right] x_j$$

$$\theta_i$$
 +=  $\eta$  \* gradient[j] for all  $0 \le j \le m$ 

- $x_j$  is j-th feature of input  $\mathbf{x} = (x_1, ..., x_m)$
- Insert  $x_0 = 1$  before training

$$\begin{array}{l} \text{Gradient} \\ \text{Ascent Step} \ \theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \pmb{x^{(i)}} \right) \right] \, x_j^{(i)} \end{array}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
      for each 0 \le j \le m:
         gradient[j] += \left[y-\frac{1}{1+\rho^{-\theta^T}x}\right]x_j
              * gradient[j] for all 0 ≤ j ≤ m
```

- $x_j$  is j-th feature of input  $\mathbf{x} = (x_1, ..., x_m)$
- Insert  $x_0 = 1$  before training
- Finish computing gradient before updating any part of  $\theta$

Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x^{(i)}} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
      for each 0 \le j \le m:
          gradient[j] += \left[y-\frac{1}{1+\rho^{-\theta^T}x}\right]x_j
   \theta_j += n * gradient[j] for all 0 \le j \le m
```

- $x_i$  is j-th feature of input  $\mathbf{x} = (x_1, \dots, x_m)$
- Insert  $x_0 = 1$  before training
- Finish computing gradient before updating any part of  $\theta$
- Learning rate  $\eta$  is a constant you set before training

Ascent Step 
$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x^{(i)}} \right) \right] x_j^{(i)}$$

```
initialize \theta_i = \emptyset for \emptyset \le j \le m
repeat many times:
   gradient[j] = 0 for 0 \le j \le m
   for each training example (x, y):
       for each 0 \le j \le m:
          gradient[j] += \left[y - \frac{1}{1 + \rho^{-\theta^T}x}\right] x_j
    \theta_i += \eta * gradient[j] for all 0 \le j \le m
```

- $x_i$  is j-th feature of input  $x = (x_1, ..., x_m)$
- Insert  $x_0 = 1$  before training
- Finish computing gradient before updating any part of  $\theta$
- Learning rate  $\eta$  is a constant you set before training

# Testing

### Introducing notation $\hat{y}$

$$\widehat{Y} = \underset{y=\{0,1\}}{\arg \max} P(Y|X)$$

$$P(Y = 1|X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j}) = \sigma(\theta^{T} x)$$

$$\hat{y} = P(Y = 1 | X = x) = \sigma(\theta^T x)$$

Small  $\hat{y}$  is conditional probability

$$P(Y = y | \mathbf{X} = \mathbf{x}) = \begin{cases} \hat{y} & \text{if } y = 1\\ 1 - \hat{y} & \text{if } y = 0 \end{cases}$$

# Testing: Classification with Logistic Regression

Training

Learn parameters 
$$\theta = (\theta_0, \theta_1, \dots, \theta_m)$$

via gradient ascent:

$$\theta_j^{\text{new}} = \theta_j^{\text{old}} + \eta \cdot \sum_{i=1}^n \left[ y^{(i)} - \sigma \left( \theta^{\text{old}^T} \boldsymbol{x}^{(i)} \right) \right] x_j^{(i)}$$

Testing

• Compute 
$$\hat{y} = P(Y = 1 | X = x) = \sigma(\theta^T x) = \frac{1}{1 + e^{-\theta^T x}}$$

Classify instance as:

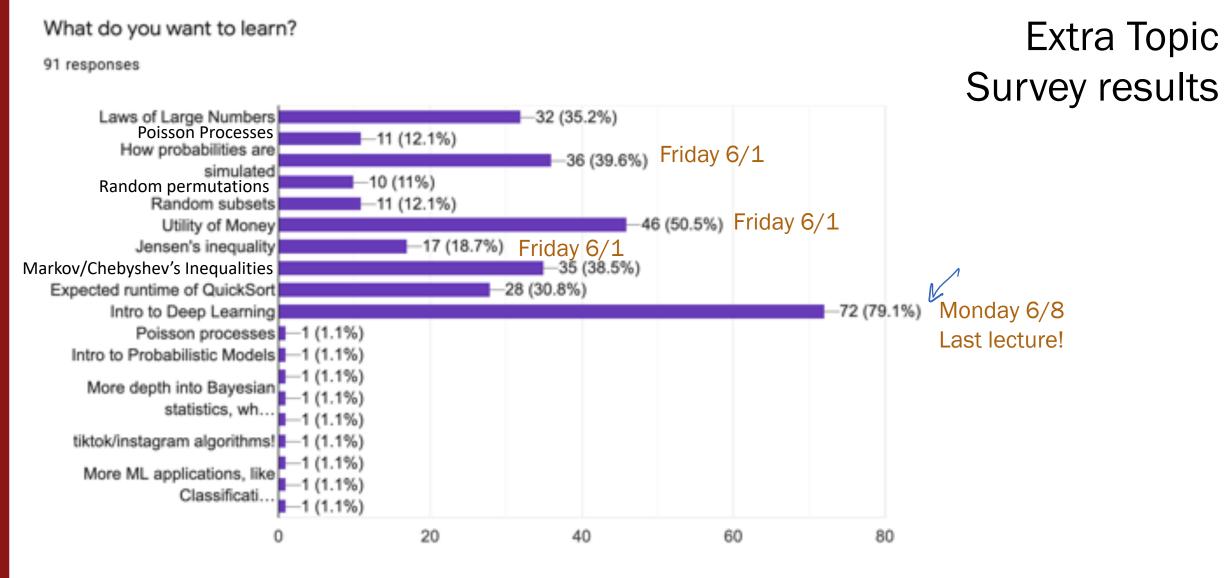
$$\begin{cases} 1 & \hat{y} > 0.5, \text{ equivalently } \theta^T x > 0 \\ 0 & \text{otherwise} \end{cases}$$



Parameters  $\theta_i$  are <u>not</u> updated during testing phase



#### Announcements



#### Interesting probability news

# The Time Everyone "Corrected" the World's Smartest Woman





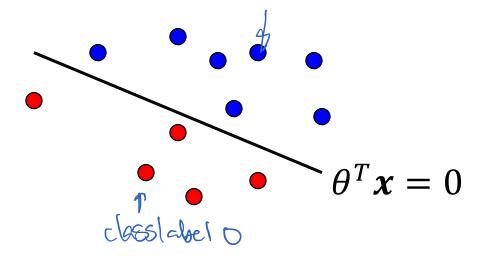
https://priceonomics.com/the-time-everyone-corrected-the-worlds-smartest/

# Philosophy

### Intuition about Logistic Regression

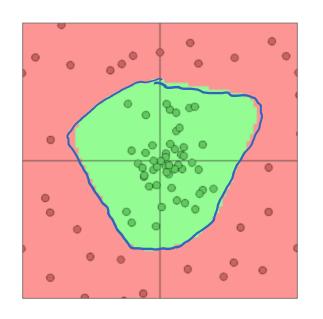
Regression 
$$P(Y=1|\boldsymbol{X}=\boldsymbol{x})=\sigma(\theta^T\boldsymbol{x})$$
 where  $\theta^T\boldsymbol{x}=\sum_{j=0}^m\theta_j\boldsymbol{x}$ 

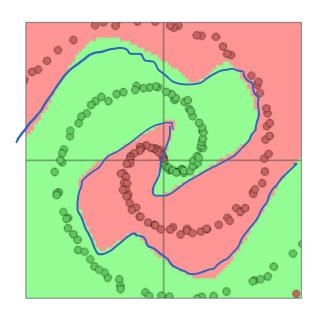
Logistic Regression is trying to fit a line that separates data instances where y = 1 from those where y = 0:



- We call such data (or functions generating the data <u>linearly separable</u>.
- Naïve Bayes is linear too, because there is no interaction between different features.

#### Data is often not linearly separable





- Not possible to draw a line that successfully separates all the y = 1 points (green) from the y = 0 points (red)
- Despite this fact, Logistic Regression and Naive Bayes still often work well in practice

## Many tradeoffs in choosing an algorithm

#### **Naïve Bayes**

Modeling goal

P(X,Y)

**Generative** or discriminative? Generative: could use joint distribution to generate new points ( but you might not need this extra effort)

Continuous input features

Needs parametric form (e.g., Gaussian) or discretized buckets (for multinomial features)

Discrete input features

Yes, multi-value discrete data = multinomial  $P(X_i|Y)$  Logistic Regression

P(Y|X)

**Discriminative**: just tries to discriminate y = 0 vs y = 1(X cannot generate new points b/c no P(X,Y)

Yes, easily forange, apple, Lavava?

Multi-valued discrete data hard (e.g., if  $X_i \in \{A, B, C\}$ , not necessarily good to encode as  $\{1, 2, 3\}$ Stanford University 65



# Gradient Derivation

20's

#### Background: Calculus

#### Calculus refresher #1:

Derivative(sum) = sum(derivative)

$$\frac{\partial}{\partial x} \sum_{i=1}^{n} f_i(x) = \sum_{i=1}^{n} \frac{\partial f_i(x)}{\partial x}$$

Calculus refresher #2:

Chain rule 😿 😿 😿

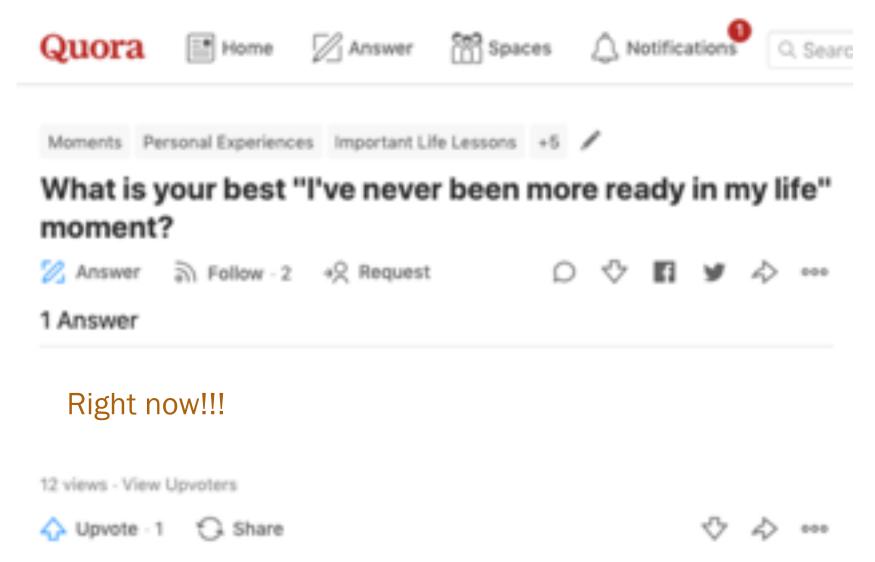
$$\frac{\partial f(x)}{\partial x} = \frac{\partial f(z)}{\partial z} \frac{\partial z}{\partial x}$$

Calculus Chain Rule

$$f(x) = f(z(x))$$

aka decomposition of composed functions

#### Are you ready?



Find: 
$$\frac{\partial LL(\theta)}{\partial \theta_j}$$
 where

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \boldsymbol{x}^{(i)}) + (1 - y^{(i)}) \log \left(1 - \sigma(\theta^{T} \boldsymbol{x}^{(i)})\right)$$

$$Couple denote the constant (2) simplifies ?$$

log conditional likelihood

### Aside: Sigmoid has a beautiful derivative

#### Sigmoid function:



$$\sigma(z) = \frac{1}{1 + e^{-z}}$$

#### Derivative:

$$\frac{d}{dz}\sigma(z) = \sigma(z)[1 - \sigma(z)]$$

What is 
$$\frac{\partial}{\partial \theta_j} \sigma(\theta^T \mathbf{x})$$
?

- A.  $\sigma(x_i)[1-\sigma(x_i)]x_i$
- B.  $\sigma(\theta^T x)[1 \sigma(\theta^T x)]x$
- C.  $\sigma(\theta^T x)[1 \sigma(\theta^T x)]x_i$
- D.  $\sigma(\theta^T \mathbf{x}) x_i [1 \sigma(\theta^T \mathbf{x}) x_i]$
- None/other



#### Aside: Sigmoid has a beautiful derivative

#### Sigmoid function:

$$\sigma(z) = \frac{1}{1 + e^{-z}}$$

Derivative:

$$\frac{d}{dz}\sigma(z) = \sigma(z)[1 - \sigma(z)]$$

What is 
$$\frac{\partial}{\partial \theta_j} \sigma(\theta^T \mathbf{x})$$
?

A. 
$$\sigma(x_i)[1-\sigma(x_i)]x_i$$

B. 
$$\sigma(\theta^T x)[1 - \sigma(\theta^T x)]x$$

C. 
$$\sigma(\theta^T \mathbf{x})[1 - \sigma(\theta^T \mathbf{x})]x_i$$

D. 
$$\sigma(\theta^T x) x_j [1 - \sigma(\theta^T x) x_j]$$

None/other

Let 
$$z = \theta^T \mathbf{x} = \sum_{k=0}^m \theta_k x_k$$
.  $\frac{\partial}{\partial Q} \sum_{k=0}^m \nabla_k \chi_k = \frac{\partial}{\partial Q} \nabla_k \chi_k$ 

$$\frac{\partial}{\partial \theta_j} \sigma(\theta^T \mathbf{x}) = \frac{\partial}{\partial z} \sigma(z) \cdot \frac{\partial z}{\partial \theta_j}$$
 (Chain Rule)
$$= \sigma(\theta^T \mathbf{x}) [1 - \sigma(\theta^T \mathbf{x})] x_j$$

#### Re-introducing notation $\hat{y}$

$$\widehat{Y} = \underset{y=\{0,1\}}{\arg \max} P(Y|X)$$

$$P(Y = 1|X = x) = \sigma(\sum_{j=0}^{m} \theta_{j} x_{j}) = \sigma(\theta^{T} x)$$

$$\hat{\mathbf{y}} = P(Y = 1 | \mathbf{X} = \mathbf{x}) = \sigma(\theta^T \mathbf{x})$$

$$P(Y = y | \mathbf{X} = \mathbf{x}) = \begin{cases} \hat{y} & \text{if } y = 1\\ 1 - \hat{y} & \text{if } y = 0 \end{cases}$$

$$P(Y = y | X = x) = (\hat{y})^{y} (1 - \hat{y})^{1-y}$$

Find: 
$$\frac{\partial LL(\theta)}{\partial \theta_j}$$
 where

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \sigma(\theta^{T} \mathbf{x}^{(i)}) + (1 - y^{(i)}) \log \left(1 - \sigma(\theta^{T} \mathbf{x}^{(i)})\right)$$

log conditional likelihood

$$LL(\theta) = \sum_{i=1}^{n} y^{(i)} \log \hat{y}^{(i)} + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)})$$

$$\frac{\partial LL(\theta)}{\partial \theta_{j}} = \sum_{i=1}^{n} \frac{\partial}{\partial \theta_{j}} \left[ y^{(i)} \log(\hat{y}^{(i)}) + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)}) \right] \qquad \text{Let } \hat{y}^{(i)} = \sigma(\theta^{T} \boldsymbol{x}^{(i)})$$

$$= \sum_{i=1}^{n} \frac{\partial}{\partial \hat{y}^{(i)}} \left[ y^{(i)} \log(\hat{y}^{(i)}) + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)}) \right] \cdot \frac{\partial \hat{y}^{(i)}}{\partial \theta_{j}} \qquad \text{(Chain Rule)}$$

$$= \sum_{i=1}^{n} \left[ y^{(i)} \frac{1}{\hat{y}^{(i)}} - (1 - y^{(i)}) \frac{1}{1 - \hat{y}^{(i)}} \right] \cdot \hat{y}^{(i)} (1 - \hat{y}^{(i)}) x_{j}^{(i)} \qquad \text{(calculus)}$$

$$= \sum_{i=1}^{n} \left[ y^{(i)} - \hat{y}^{(i)} \right] x_{j}^{(i)} \qquad = \sum_{i=1}^{n} \left[ y^{(i)} - \sigma(\theta^{T} \boldsymbol{x}^{(i)}) \right] x_{j}^{(i)} \qquad \text{(simplify)}$$

$$\frac{\partial LL(\theta)}{\partial \theta_{j}} = \sum_{i=1}^{n} \frac{\partial}{\partial \theta_{j}} \left[ y^{(i)} \log(\hat{y}^{(i)}) + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)}) \right] \qquad \text{Let } \hat{y}^{(i)} = \sigma(\theta^{T} \boldsymbol{x}^{(i)})$$

$$= \sum_{i=1}^{n} \frac{\partial}{\partial \hat{y}^{(i)}} \left[ y^{(i)} \log(\hat{y}^{(i)}) + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)}) \right] \cdot \frac{\partial \hat{y}^{(i)}}{\partial \theta_{j}} \qquad \text{(Chain Rule)}$$

$$= \sum_{i=1}^{n} \left[ y^{(i)} \frac{1}{\hat{y}^{(i)}} - (1 - y^{(i)}) \frac{1}{1 - \hat{y}^{(i)}} \right] \cdot \hat{y}^{(i)} (1 - \hat{y}^{(i)}) x_{j}^{(i)} \qquad \text{(calculus)}$$

$$= \sum_{i=1}^{n} \left[ y^{(i)} - \hat{y}^{(i)} \right] x_{j}^{(i)} \qquad = \sum_{i=1}^{n} \left[ y^{(i)} - \sigma(\theta^{T} \boldsymbol{x}^{(i)}) \right] x_{j}^{(i)} \qquad \text{(simplify)}$$

$$\frac{\partial LL(\theta)}{\partial \theta_j} = \sum_{i=1}^n \frac{\partial}{\partial \theta_j} \left[ y^{(i)} \log(\hat{y}^{(i)}) + (1 - y^{(i)}) \log(1 - \hat{y}^{(i)}) \right]$$
 Let  $\hat{y}^{(i)} = \sigma(\theta^T x^{(i)})$ 

$$= \sum_{i=1}^{n} \frac{\partial}{\partial \hat{y}^{(i)}} \left[ y^{(i)} \log(\hat{y}^{(i)}) + \left(1 - y^{(i)}\right) \log\left(1 - \hat{y}^{(i)}\right) \right] \cdot \frac{\partial \hat{y}^{(i)}}{\partial \theta_{j}}$$
 (Chain Rule)

$$= \sum_{i=1}^{n} \left[ y^{(i)} \frac{1}{\hat{y}^{(i)}} - (1 - y^{(i)}) \frac{1}{1 - \hat{y}^{(i)}} \right] \cdot \hat{y}^{(i)} (1 - \hat{y}^{(i)}) x_j^{(i)}$$
 (calculus)

$$= \sum_{i=1}^{n} [y^{(i)} - \hat{y}^{(i)}] x_j^{(i)} = \sum_{i=1}^{n} [y^{(i)} - \sigma(\theta^T x^{(i)})] x_j^{(i)}$$



(simplify)

modelium  $P(X_{3}=X|Y_{2})$   $P(X_{3}=X|Y_{2})$  EP(X|Y=1)P(Y=1) = P(X, Y=1) P(X|Y=0)P(Y=0) = P(X, Y=0)P(Y=1|X) + P(Y=0|X) = 1

22: P(Y=0|X) P(Y=1|X)>0.5

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