

12: Independent RVs

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Quick slide reference

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Independent Discrete RVs

Independent discrete RVs

Recall the definition of independent events E and F :

$$P(EF) = P(E)P(F)$$

Two discrete random variables X and Y are **independent** if:

for all x, y :

$$P(X = x, Y = y) = P(X = x)P(Y = y)$$

Different notation,
same idea:

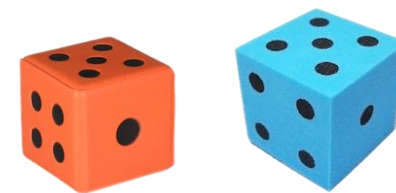
$$p_{X,Y}(x, y) = p_X(x)p_Y(y)$$

- Intuitively: knowing value of X tells us nothing about the distribution of Y (and vice versa)
- If two variables are not independent, they are called **dependent**.

Dice (after all this time, still our friends)

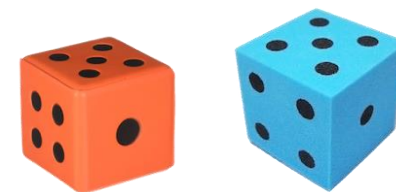
Let: D_1 and D_2 be the outcomes of two rolls
 $S = D_1 + D_2$, the sum of two rolls

- Each roll of a 6-sided die is an independent trial.
 - Random variables D_1 and D_2 are independent.
1. Are events $(D_1 = 1)$ and $(S = 7)$ independent?
 2. Are events $(D_1 = 1)$ and $(S = 5)$ independent?
 3. Are random variables D_1 and S independent?



Dice (after all this time, still our friends)

Let: D_1 and D_2 be the outcomes of two rolls
 $S = D_1 + D_2$, the sum of two rolls



- Each roll of a 6-sided die is an independent trial.
- Random variables D_1 and D_2 are independent.

1. Are events $(D_1 = 1)$ and $(S = 7)$ independent?

2. Are events $(D_1 = 1)$ and $(S = 5)$ independent?

3. Are random variables D_1 and S independent?

All events $(X = x, Y = y)$ must be independent for X, Y to be independent RVs.

What about continuous random variables?

Continuous random variables can also be independent! We'll see this later.

Today's goal:

How can we model sums of discrete random variables?

Big motivation: Model total successes observed over multiple experiments

Sums of independent Binomial RVs

Sum of independent Binomials

$$\begin{array}{l} X \sim \text{Bin}(n_1, p) \\ Y \sim \text{Bin}(n_2, p) \\ X, Y \text{ independent} \end{array} \quad \Rightarrow \quad X + Y \sim \text{Bin}(n_1 + n_2, p)$$

Intuition:

- Each trial in X and Y is independent and has same success probability p
- Define $Z = \#$ successes in $n_1 + n_2$ independent trials, each with success probability p . $Z \sim \text{Bin}(n_1 + n_2, p)$, and also $Z = X + Y$

Holds in general case:

$$\begin{array}{l} X_i \sim \text{Bin}(n_i, p) \\ X_i \text{ independent for } i = 1, \dots, n \end{array}$$

$$\Rightarrow \sum_{i=1}^n X_i \sim \text{Bin}\left(\sum_{i=1}^n n_i, p\right)$$

If only it were
always so
simple...

Convolution: Sum of independent Poisson RVs

Convolution: Sum of independent random variables

For any discrete random variables X and Y :

$$P(X + Y = n) = \sum_k P(X = k, Y = n - k)$$

In particular, for **independent** discrete random variables X and Y :

$$P(X + Y = n) = \sum_k P(X = k)P(Y = n - k)$$

the **convolution** of p_X and p_Y

Insight into convolution

For **independent** discrete random variables X and Y :

$$P(X + Y = n) = \sum_k P(X = k)P(Y = n - k)$$

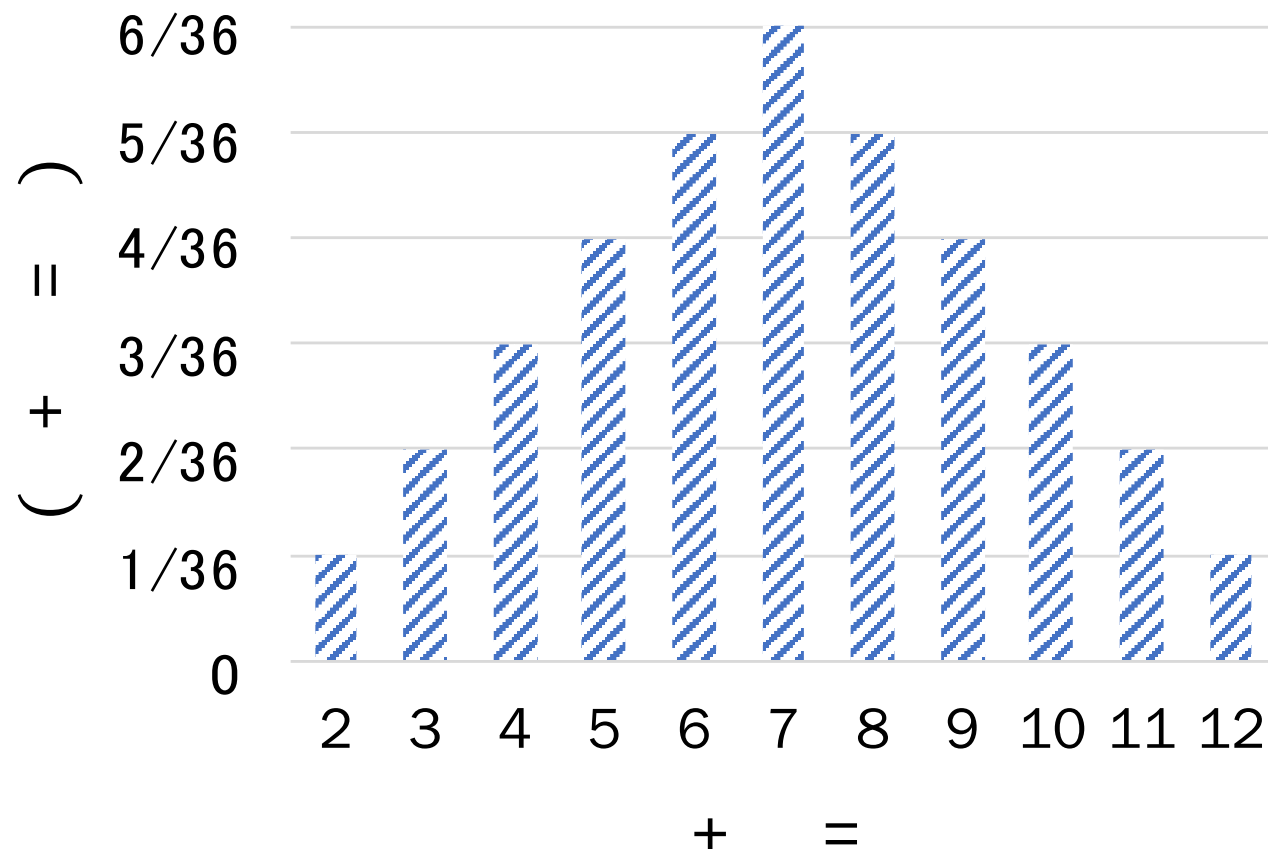
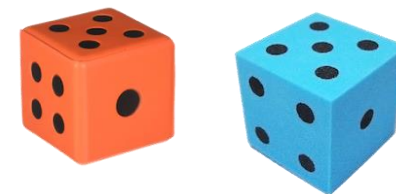
the **convolution** of p_X and p_Y

Suppose X and Y are independent, both with support $\{0, 1, \dots, n, \dots\}$:

		X						
		0	1	2	...	n	$n + 1$...
Y	0						✓	
			
	$n - 2$			✓				
	$n - 1$		✓					
	n	✓						
	$n + 1$							
	...							

- ✓ : event where $X + Y = n$
- Each event has probability:
 $P(X = k, Y = n - k)$
 $= P(X = k)P(Y = n - k)$
 (because X, Y are independent)
- $P(X + Y = n) =$ sum of mutually exclusive events

Sum of 2 dice rolls

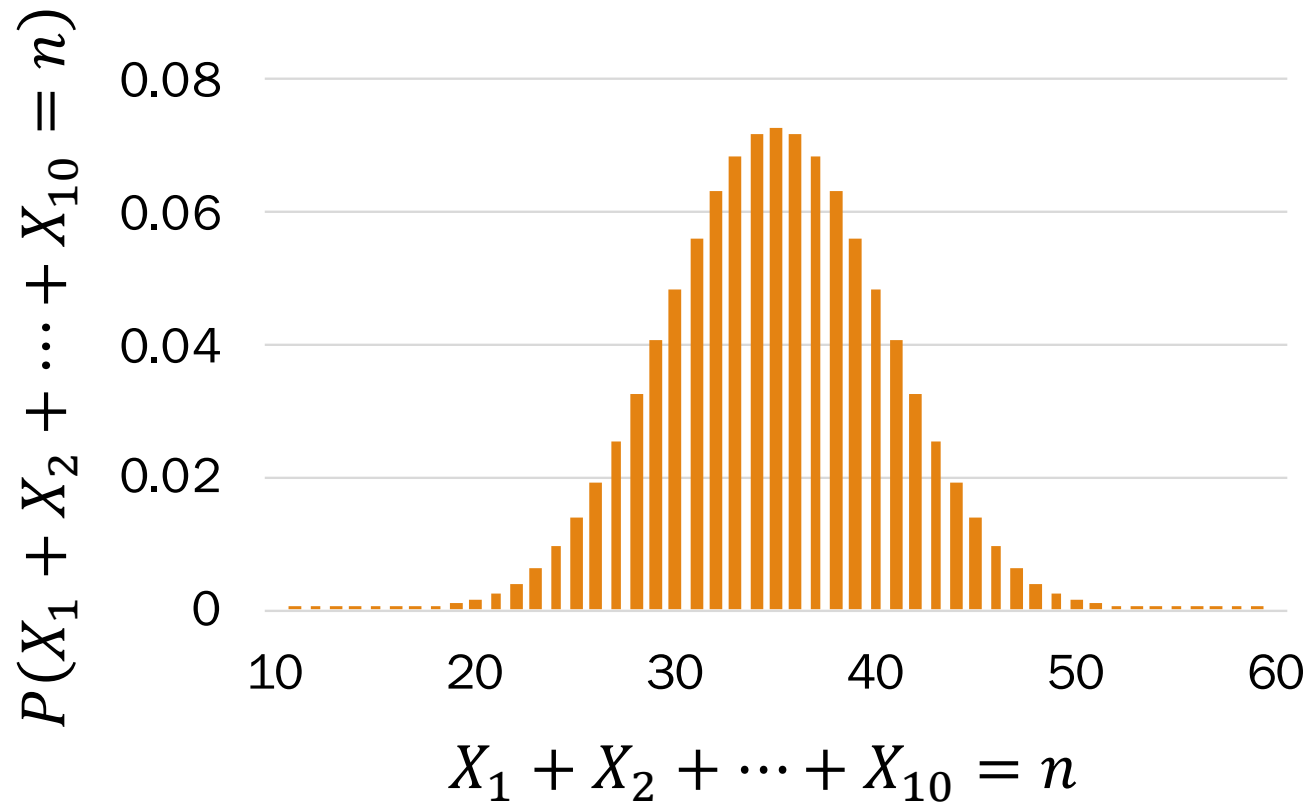
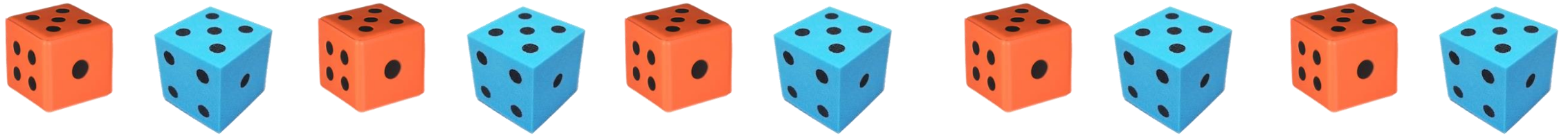


The distribution of a sum of 2 dice rolls is a convolution of 2 PMFs.

Example:

$$\begin{aligned} P(X + Y = 4) = & \\ & P(X = 1)P(Y = 3) \\ & + P(X = 2)P(Y = 2) \\ & + P(X = 3)P(Y = 1) \end{aligned}$$

Sum of 10 dice rolls (fun preview)



The distribution of a sum of 10 dice rolls is a convolution 10 PMFs.

Looks kinda Normal...???
(more on this in Week 7)

Sum of independent Poissons

$X \sim \text{Poi}(\lambda_1), Y \sim \text{Poi}(\lambda_2)$
 X, Y independent



$X + Y \sim \text{Poi}(\lambda_1 + \lambda_2)$

Proof (just for reference):

$$\begin{aligned} P(X + Y = n) &= \sum_k P(X = k)P(Y = n - k) \\ &= \sum_{k=0}^n e^{-\lambda_1} \frac{\lambda_1^k}{k!} e^{-\lambda_2} \frac{\lambda_2^{n-k}}{(n-k)!} = e^{-(\lambda_1 + \lambda_2)} \sum_{k=0}^n \frac{\lambda_1^k \lambda_2^{n-k}}{k! (n-k)!} \\ &= \frac{e^{-(\lambda_1 + \lambda_2)}}{n!} \sum_{k=0}^n \frac{n!}{k! (n-k)!} \lambda_1^k \lambda_2^{n-k} = \underbrace{\frac{e^{-(\lambda_1 + \lambda_2)}}{n!} (\lambda_1 + \lambda_2)^n}_{\text{Poi}(\lambda_1 + \lambda_2)} \end{aligned}$$

X and Y independent,
convolution

PMF of Poisson RVs

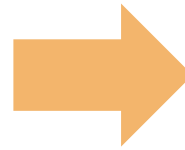
Binomial Theorem:

$$(a + b)^n = \sum_{k=0}^n \binom{n}{k} a^k b^{n-k}$$

General sum of independent Poissons

Holds in general case:

$X_i \sim \text{Poi}(\lambda_i)$
 X_i independent for $i = 1, \dots, n$



$$\sum_{i=1}^n X_i \sim \text{Poi}\left(\sum_{i=1}^n \lambda_i\right)$$



12: Independent RVs (live)

Lisa Yan

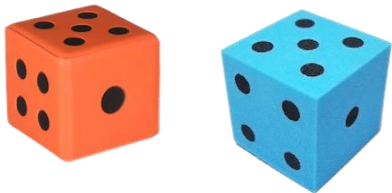
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Two discrete random variables X and Y are **independent** if:

for all x, y :

$$P(X = x, Y = y) = P(X = x)P(Y = y)$$

$$p_{X,Y}(x, y) = p_X(x)p_Y(y)$$



The sum of 2 dice and the outcome of 1st die are **dependent RVs**.

Important: Joint PMF must decompose into product of marginal PMFs for ALL values of X and Y for X, Y to be independent RVs.

Think

Slide 22 has a question to go over by yourself.

Post any clarifications here!

<https://us.edstem.org/courses/667/discussion/91824>

Think by yourself: 2 min



Coin flips

Flip a coin with probability p of “heads” a total of $n + m$ times.

Let $X =$ number of heads in first n flips. $X \sim \text{Bin}(n, p)$

$Y =$ number of heads in next m flips. $Y \sim \text{Bin}(m, p)$

$Z =$ total number of heads in $n + m$ flips.

1. Are X and Z independent?
2. Are X and Y independent?



(by yourself)

Coin flips

Flip a coin with probability p of “heads” a total of $n + m$ times.

Let $X =$ number of heads in first n flips. $X \sim \text{Bin}(n, p)$

$Y =$ number of heads in next m flips. $Y \sim \text{Bin}(m, p)$

$Z =$ total number of heads in $n + m$ flips.

1. Are X and Z independent? ✘

Counterexample: What if $Z = 0$?

2. Are X and Y independent? ✔

$$P(X = x, Y = y) = P\left(\begin{array}{l} \text{first } n \text{ flips have } x \text{ heads} \\ \text{and next } m \text{ flips have } y \text{ heads} \end{array}\right)$$

$$= \binom{n}{x} p^x (1-p)^{n-x} \binom{m}{y} p^y (1-p)^{m-y}$$

$$= P(X = x)P(Y = y)$$

of mutually exclusive outcomes in event : $\binom{n}{x} \binom{m}{y}$
 $P(\text{each outcome})$
 $= p^x (1-p)^{n-x} p^y (1-p)^{m-y}$

This probability (found through counting) is the product of the marginal PMFs.

Sum of independent Poissons

$$X \sim \text{Poi}(\lambda_1), Y \sim \text{Poi}(\lambda_2)$$

X, Y independent



$$X + Y \sim \text{Poi}(\lambda_1 + \lambda_2)$$

- n servers with independent number of requests/minute
- Server i 's requests each minute can be modeled as $X_i \sim \text{Poi}(\lambda_i)$

What is the probability that none of the servers receive any requests in the next minute?

Breakout Rooms

Slide 26 has two questions to go over in groups.

ODD breakout rooms: Try question 1

EVEN breakout rooms: Try question 2

Post any clarifications here!

<https://us.edstem.org/courses/667/discussion/91824>

Breakout rooms: 5 min. Introduce yourself!



Independent questions

1. Let $X \sim \text{Bin}(30, 0.01)$ and $Y \sim \text{Bin}(50, 0.02)$ be independent RVs.
 - How do we compute $P(X + Y = 2)$ using a Poisson approximation?
 - How do we compute $P(X + Y = 2)$ exactly?

2. Let $N = \#$ of requests to a web server per day. Suppose $N \sim \text{Poi}(\lambda)$.
 - Each request independently comes from a human (prob. p), or bot ($1 - p$).
 - Let X be $\#$ of human requests/day, and Y be $\#$ of bot requests/day.Are X and Y independent? What are their marginal PMFs?



1. Approximating the sum of independent Binomial RVs

Let $X \sim \text{Bin}(30, 0.01)$ and $Y \sim \text{Bin}(50, 0.02)$ be independent RVs.

- How do we compute $P(X + Y = 2)$ using a Poisson approximation?

- How do we compute $P(X + Y = 2)$ exactly?

$$\begin{aligned} P(X + Y = 2) &= \sum_{k=0}^2 P(X = k)P(Y = 2 - k) \\ &= \sum_{k=0}^2 \binom{30}{k} 0.01^k (0.99)^{30-k} \binom{50}{2-k} 0.02^{2-k} 0.98^{50-(2-k)} \approx \mathbf{0.2327} \end{aligned}$$

2. Web server requests

Let $N = \#$ of requests to a web server per day. Suppose $N \sim \text{Poi}(\lambda)$.

- Each request independently comes from a human (prob. p), or bot ($1 - p$).
- Let X be $\#$ of human requests/day, and Y be $\#$ of bot requests/day.

Are X and Y independent? What are their marginal PMFs?

$$P(X = n, Y = m) = P(X = n, Y = m | N = n + m)P(N = n + m) \quad \text{Law of Total Probability} \\ + P(X = n, Y = m | N \neq n + m)P(N \neq n + m)$$

$$= P(X = n | N = n + m)P(Y = m | X = n, N = n + m)P(N = n + m) \quad \text{Chain Rule}$$

$$= \binom{n + m}{n} p^n (1 - p)^m \cdot 1 \cdot e^{-\lambda} \frac{\lambda^{n+m}}{(n + m)!} \quad \text{Given } N = n + m \text{ indep. trials, } X|N = n + m \sim \text{Bin}(n + m, p)$$

$$= \frac{(n + m)!}{n! m!} e^{-\lambda} \frac{(\lambda p)^n (\lambda(1 - p))^m}{(n + m)!} = e^{-\lambda p} \frac{(\lambda p)^n}{n!} \cdot e^{-\lambda(1-p)} \frac{(\lambda(1 - p))^m}{m!}$$

$$= P(X = n)P(Y = m) \quad \text{where } X \sim \text{Poi}(\lambda p), Y \sim \text{Poi}(\lambda(1 - p))$$

Yes, X and Y are independent!

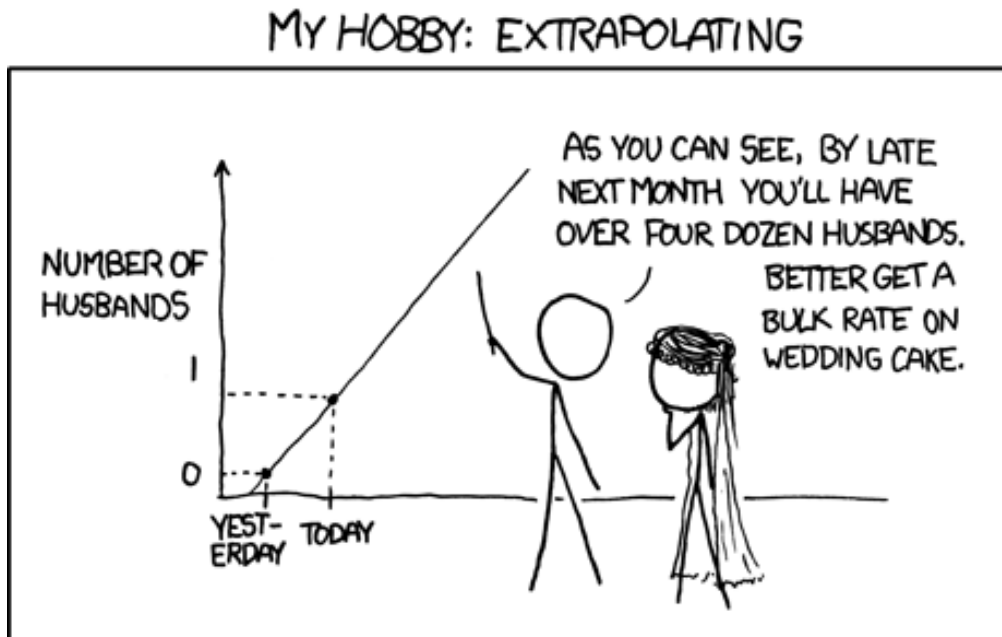
Interlude for jokes/announcements

Announcements

Quiz #1

24 hours, starting on Monday, 5pm PST

See Ed and the OH calendar for study help!



Credit: <https://xkcd.com/605/>

Interesting probability news

Column: Did Astros beat the Dodgers by cheating? The numbers say no



“...new analyses of the Astros’ 2017 season by baseball’s corps of unofficial statisticians — “sabermetricians,” to the sport — indicate that the Astros didn’t gain anything from their cheating; in fact, it may have hurt them.”

<https://www.latimes.com/business/story/2020-02-27/astros-cheating-analysis>

Independence of multiple random variables

Recall independence of n events E_1, E_2, \dots, E_n :

for $r = 1, \dots, n$:

for every subset E_1, E_2, \dots, E_r :

$$P(E_1, E_2, \dots, E_r) = P(E_1)P(E_2) \cdots P(E_r)$$

We have independence of n discrete random variables X_1, X_2, \dots, X_n if for all x_1, x_2, \dots, x_n :

$$P(X_1 = x_1, X_2 = x_2, \dots, X_n = x_n) = \prod_{i=1}^n P(X_i = x_i)$$

Independence is symmetric

If X and Y are independent random variables, then
 X is independent of Y , and Y is independent of X

Let N be the number of times you roll 2 dice repeatedly until a 4 is rolled (the player wins), or a 7 is rolled (the player loses).

Let X be the value (4 or 7) of the final throw's sum.

- Is N independent of X ?
 $P(N = n|X = 7) = P(N = n)?$
 $P(N = n|X = 4) = P(N = n)?$
- Is X independent of N ?
 $P(X = 4|N = n) = P(X = 4)?$
 $P(X = 7|N = n) = P(X = 7)?$ } (yes, easier to intuit)

In short: Independence is not always intuitive, but it is symmetric.

Statistics of Two RVs

Expectation and Covariance

In real life, we often have many RVs interacting at once.

- We've seen some simpler cases (e.g., sum of independent Poissons).
- Computing joint PMFs in general is hard!
- But **often you don't need to model** joint RVs completely.

Instead, we'll focus next on reporting **statistics** of multiple RVs:

- Expectation of sums (you've seen some of this)
- **Covariance**: a measure of how two RVs vary with *each other*

Properties of Expectation, extended to two RVs

1. Linearity:

$$E[aX + bY + c] = aE[X] + bE[Y] + c$$

2. Expectation of a sum = sum of expectation:

$$E[X + Y] = E[X] + E[Y]$$

(we've seen this;
we'll prove this next)

3. Unconscious statistician:

$$E[g(X, Y)] = \sum_x \sum_y g(x, y) p_{X, Y}(x, y)$$

True for both independent
and dependent random
variables!

Proof of expectation of a sum of RVs

$$E[X + Y] = E[X] + E[Y]$$

$$E[X + Y] = \sum_x \sum_y (x + y)p_{X,Y}(x, y)$$

LOTUS,
 $g(X, Y) = X + Y$

$$= \sum_x \sum_y xp_{X,Y}(x, y) + \sum_x \sum_y yp_{X,Y}(x, y)$$

$$= \sum_x x \sum_y p_{X,Y}(x, y) + \sum_y y \sum_x p_{X,Y}(x, y)$$

$$= \sum_x xp_X(x) + \sum_y yp_Y(y)$$

$$= E[X] + E[Y]$$

Linearity of summations
(cont. case: linearity of integrals)

Marginal PMFs for X and Y

Expectations of common RVs: Binomial

$$X \sim \text{Bin}(n, p) \quad E[X] = np$$

of successes in n independent trials with probability of success p

Recall: $\text{Bin}(1, p) = \text{Ber}(p)$

$$X = \sum_{i=1}^n X_i$$

Let $X_i = i$ th trial is heads
 $X_i \sim \text{Ber}(p), E[X_i] = p$



$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = \sum_{i=1}^n p = np$$

Think

Slide 40 has a question to go over by yourself.

Post any clarifications here!

<https://us.edstem.org/courses/667/discussion/91824>

Think by yourself: 2 min



Expectations of common RVs: Negative Binomial

$$Y \sim \text{NegBin}(r, p) \quad E[Y] = \frac{r}{p}$$

of independent trials with probability of success p until r successes

Recall: $\text{NegBin}(1, p) = \text{Geo}(p)$

$$Y = \sum_{i=1}^? Y_i$$

1. How should we define Y_i ?
2. How many terms are in our summation?



Expectations of common RVs: Negative Binomial

$$Y \sim \text{NegBin}(r, p) \quad E[Y] = \frac{r}{p}$$

of independent trials with probability of success p until r successes

Recall: $\text{NegBin}(1, p) = \text{Geo}(p)$

$$Y = \sum_{i=1}^? Y_i$$

Let $Y_i = \#$ trials to get i th success (after $(i - 1)$ th success)

$$Y_i \sim \text{Geo}(p), E[Y_i] = \frac{1}{p}$$



$$E[Y] = E\left[\sum_{i=1}^r Y_i\right] = \sum_{i=1}^r E[Y_i] = \sum_{i=1}^r \frac{1}{p} = \frac{r}{p}$$

Good luck!



(by yourself)