

13: Statistics of Multiple RVs

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Quick slide reference

3	Expectation of Common RVs	13a_expectation_sum
8	Coupon Collecting Problems	13b_coupon_collecting
14	Covariance	13c_covariance
20	Independence and Variance	13d_variance_sum
27	Exercises	LIVE
48	Correlation	LIVE

Expectation of Common RVs

Linearity of Expectation is useful

Expectation is a linear mathematical operation. If $X = \sum_{i=1}^n X_i$:

$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

- Even if you *don't know* the distribution of X (e.g., because the joint distribution of (X_1, \dots, X_n) is unknown), you can still compute *expectation* of the sum!!

- Problem-solving key:
Define X_i such that

$$X = \sum_{i=1}^n X_i$$

Most common use cases:

- $E[X_i]$ easy to calculate
- **Sum of dependent RVs**

Expectations of common RVs: Binomial

$$X \sim \text{Bin}(n, p) \quad E[X] = np$$

of successes in n independent trials with probability of success p

Recall: $\text{Bin}(1, p) = \text{Ber}(p)$

$$X = \sum_{i=1}^n X_i$$

Let $X_i = i$ th trial is heads
 $X_i \sim \text{Ber}(p), E[X_i] = p$



$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = \sum_{i=1}^n p = np$$

Expectations of common RVs: Negative Binomial

$$Y \sim \text{NegBin}(r, p) \quad E[Y] = \frac{r}{p}$$

of independent trials with probability of success p until r successes

Recall: $\text{NegBin}(1, p) = \text{Geo}(p)$

$$Y = \sum_{i=1}^? Y_i$$

1. How should we define Y_i ?
2. How many terms are in our summation?



Expectations of common RVs: Negative Binomial

$$Y \sim \text{NegBin}(r, p) \quad E[Y] = \frac{r}{p}$$

of independent trials with probability of success p until r successes

Recall: $\text{NegBin}(1, p) = \text{Geo}(p)$

$$Y = \sum_{i=1}^? Y_i$$

Let $Y_i = \#$ trials to get i th success (after $(i - 1)$ th success)

$$Y_i \sim \text{Geo}(p), E[Y_i] = \frac{1}{p}$$



$$E[Y] = E\left[\sum_{i=1}^r Y_i\right] = \sum_{i=1}^r E[Y_i] = \sum_{i=1}^r \frac{1}{p} = \frac{r}{p}$$

Coupon Collecting Problems

Linearity of Expectation is useful

Expectation is a linear mathematical operation. If $X = \sum_{i=1}^n X_i$:

$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

- Even if you *don't know* the distribution of X (e.g., because the joint distribution of (X_1, \dots, X_n) is unknown), you can still compute *expectation* of the sum!!

- Problem-solving key:
Define X_i such that

$$X = \sum_{i=1}^n X_i$$

Most common use cases:

- $E[X_i]$ easy to calculate
- **Sum of dependent RVs**

Coupon collecting problems: Server requests

The **coupon collector's problem** in probability theory:

- You buy boxes of cereal.
- There are k different types of coupons
- For each box you buy, you "collect" a coupon of type i .

1. How many coupons do you expect after buying n boxes of cereal?



What is the expected number of utilized servers after n requests?

Servers
requests
 k servers
request to
server i



- * 52% of Amazon profits
- ** more profitable than Amazon's North America commerce operations

[source](#)

Computer cluster utilization

$$E \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n E[X_i]$$

Consider a computer cluster with k servers. We send n requests.

- Requests independently go to server i with probability p_i
- Let $X = \#$ servers that receive ≥ 1 request.

What is $E[X]$?



Computer cluster utilization

$$E \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n E[X_i]$$

Consider a computer cluster with k servers. We send n requests.

- Requests independently go to server i with probability p_i
- Let $X = \#$ servers that receive ≥ 1 request.

What is $E[X]$?

1. Define additional random variables.

2. Solve.

Let: $A_i =$ event that server i
receives ≥ 1 request
 $X_i =$ indicator for A_i

$$\begin{aligned} E[X_i] &= P(A_i) = 1 - (1 - p_i)^n \\ E[X] &= E \left[\sum_{i=1}^k X_i \right] = \sum_{i=1}^k E[X_i] = \sum_{i=1}^k (1 - (1 - p_i)^n) \\ &= \sum_{i=1}^k 1 - \sum_{i=1}^k (1 - p_i)^n = k - \sum_{i=1}^k (1 - p_i)^n \end{aligned}$$

$$\begin{aligned} P(A_i) &= 1 - P(\text{no requests to } i) \\ &= 1 - (1 - p_i)^n \end{aligned}$$

Note: A_i are dependent!

Coupon collecting problems: Hash tables

The **coupon collector's problem** in probability theory:

- You buy boxes of cereal.
- There are k different types of coupons
- For each box you buy, you "collect" a coupon of type i .

<u>Servers</u>	<u>Hash Tables</u>
requests	strings
k servers	k buckets
request to server i	hashed to bucket i

1. How many coupons do you expect after buying n boxes of cereal?



What is the expected number of utilized servers after n requests?

2. How many boxes do you expect to buy until you have one of each coupon?



What is the expected number of strings to hash until each bucket has ≥ 1 string?

Stay tuned for live lecture!

Covariance

Statistics of sums of RVs

For any random variables X and Y ,

$$E[X + Y] = E[X] + E[Y]$$

$$\text{Var}(X + Y) = ?$$

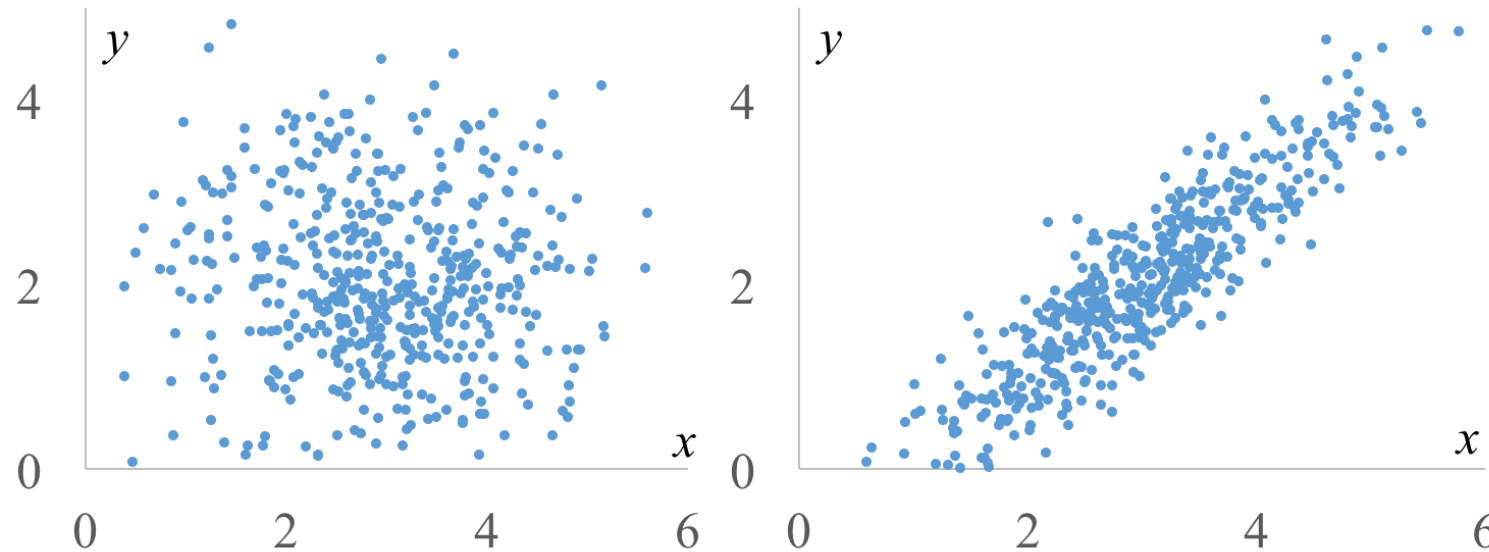
But first...
a new statistic!

Spot the difference

Compare/contrast the following two distributions:

Assume all points are equally likely.

$$P(X = x, Y = y) = \frac{1}{N}$$



Both distributions have the same $E[X]$, $E[Y]$, $\text{Var}(X)$, and $\text{Var}(Y)$

Difference: how the two variables vary with *each other*.

Covariance

The **covariance** of two variables X and Y is:

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

Proof of second part:

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY - XE[Y] - E[X]Y + E[X]E[Y]] \\ &= E[XY] - E[XE[Y]] - E[E[X]Y] + E[E[X]E[Y]] \\ &= E[XY] - E[X]E[Y] - E[X]E[Y] + E[X]E[Y] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

(linearity of
expectation)
($E[X]$, $E[Y]$ are
scalars)

Covarying humans

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

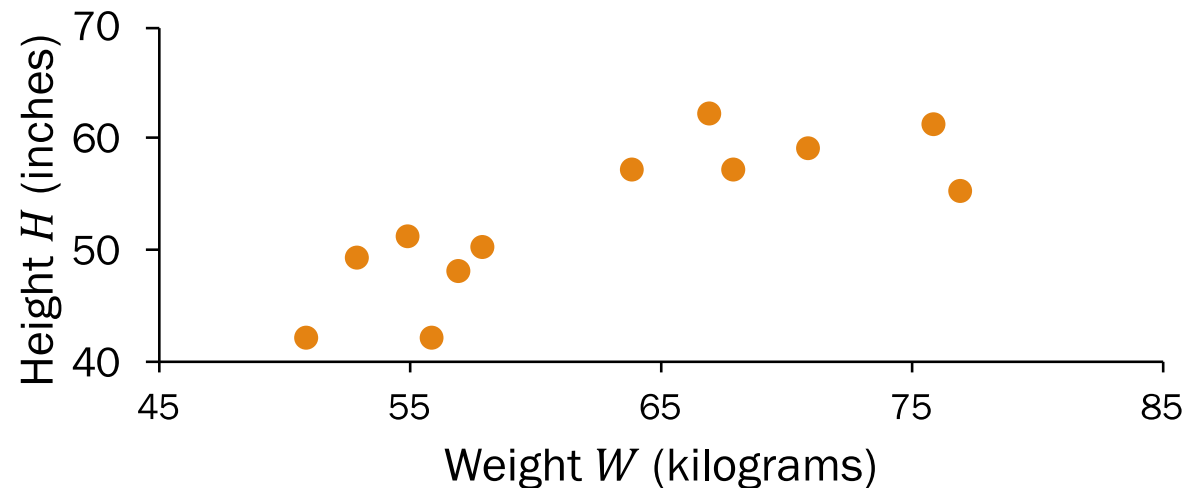
Weight (kg)	Height (in)	W · H
64	57	3648
71	59	4189
53	49	2597
67	62	4154
55	51	2805
58	50	2900
77	55	4235
57	48	2736
56	42	2352
51	42	2142
76	61	4636
68	57	3876

$$\begin{aligned}E[W] &= 62.75 \\ E[H] &= 52.75 \\ E[WH] &= 3355.83\end{aligned}$$

What is the covariance of weight W and height H ?

$$\begin{aligned}\text{Cov}(W, H) &= E[WH] - E[W]E[H] \\ &= 3355.83 - (62.75)(52.75) \\ &= 45.77\end{aligned}$$

(positive)



Covariance > 0 : one variable \uparrow , other variable \uparrow

Properties of Covariance

The **covariance** of two variables X and Y is:

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

Properties:

1. $\text{Var}(X) = E[X^2] - (E[X])^2 = \text{Cov}(X, X)$
 2. Symmetry
 3. Non-linearity
 4. Covariance of sums
- (to be discussed in live lecture)

Variance of sums of RVs

Statistics of sums of RVs

For any random variables X and Y ,

$$E[X + Y] = E[X] + E[Y]$$

$$\text{Var}(X + Y) = \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y)$$

Variance of general sum of RVs

For any random variables X and Y ,

$$\text{Var}(X + Y) = \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y)$$

Proof:

$$\text{Var}(X + Y) = \text{Cov}(X + Y, X + Y)$$

$$= \text{Cov}(X, X) + \text{Cov}(X, Y) + \text{Cov}(Y, X) + \text{Cov}(Y, Y)$$

$$= \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y)$$

$$\text{Var}(X) = \text{Cov}(X, X)$$

covariance of
all pairs

Symmetry of covariance +
 $\text{Cov}(X, X) = \text{Var}(X)$

More generally:

$$\text{Var}\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n \text{Var}(X_i) + 2 \sum_{i=1}^n \sum_{j=i+1}^n \text{Cov}(X_i, X_j) \quad (\text{proof in extra slides})$$

Statistics of sums of RVs

For any random variables X and Y ,

$$E[X + Y] = E[X] + E[Y]$$

$$\text{Var}(X + Y) = \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y)$$

For **independent** X and Y ,

$$E[XY] = E[X]E[Y]$$

(Lemma: proof in extra slides)

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Variance of sum of independent RVs

For **independent** X and Y ,

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

Proof:

$$\begin{aligned} 1. \quad \text{Cov}(X, Y) &= E[XY] - E[X]E[Y] \\ &= E[X]E[Y] - E[X]E[Y] \\ &= 0 \end{aligned}$$

def. of covariance

X and Y are **independent**

$$\begin{aligned} 2. \quad \text{Var}(X + Y) &= \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y) \\ &= \text{Var}(X) + \text{Var}(Y) \end{aligned}$$

NOT bidirectional:

$\text{Cov}(X, Y) = 0$ does NOT imply independence of X and Y !

Proving Variance of the Binomial

$$X \sim \text{Bin}(n, p) \quad \text{Var}(X) = np(1 - p)$$

To simplify the algebra a bit, let $q = 1 - p$, so $p + q = 1$.

So:

$$\begin{aligned} E(X^2) &= \sum_{k \geq 0} k^2 \binom{n}{k} p^k q^{n-k} \\ &= \sum_{k=0}^n kn \binom{n-1}{k-1} p^k q^{n-k} \\ &= np \sum_{k=1}^n k \binom{n-1}{k-1} p^{k-1} q^{(n-1)-(k-1)} \\ &= np \sum_{j=0}^{n-1} (j+1) \binom{n-1}{j} p^j q^{n-1-j} \\ &= np \left(\sum_{j=0}^{n-1} j \binom{n-1}{j} p^j q^{n-1-j} + \sum_{j=0}^{n-1} \binom{n-1}{j} p^j q^{n-1-j} \right) \\ &= np \left(\sum_{j=0}^{n-1} m \binom{m-1}{j-1} p^j q^{m-1-j} + \sum_{j=0}^{m-1} \binom{m-1}{j} p^j q^{m-1-j} \right) \\ &= np \left((n-1)p \sum_{j=1}^{n-1} \binom{m-1}{j-1} p^{j-1} q^{(m-1)-(j-1)} + \sum_{j=0}^{m-1} \binom{m-1}{j} p^j q^{m-1-j} \right) \\ &= np((n-1)p(p+q)^{m-1} + (p+q)^m) \\ &= np(n-1)p + 1 \\ &= n^2 p^2 + np(1-p) \end{aligned}$$

Then:

$$\begin{aligned} \text{var}(X) &= E(X^2) - (E(X))^2 \\ &= np(1-p) + n^2 p^2 - (np)^2 \quad \text{Expectation of Binomial Distribution: } E(X) = np \\ &= np(1-p) \end{aligned}$$

as required.

Definition of Binomial Distribution: $p + q = 1$

Factors of Binomial Coefficient: $k \binom{n}{k} = n \binom{n-1}{k-1}$

Change of limit: term is zero when $k - 1 = 0$

putting $j = k - 1, m = n - 1$

splitting sum up into two

Factors of Binomial Coefficient: $j \binom{m}{j} = m \binom{m-1}{j-1}$

Change of limit: term is zero when $j - 1 = 0$

Binomial Theorem

as $p + q = 1$

by algebra



Let's instead prove this using independence and variance!

proofwiki.org

Proving Variance of the Binomial

$$X \sim \text{Bin}(n, p) \quad \text{Var}(X) = np(1 - p)$$

Let
$$X = \sum_{i=1}^n X_i$$

Let $X_i = i$ th trial is heads

$$X_i \sim \text{Ber}(p)$$

$$\text{Var}(X_i) = p(1 - p)$$

X_i are **independent**
(by definition)

$$\text{Var}(X) = \text{Var}\left(\sum_{i=1}^n X_i\right)$$

$$= \sum_{i=1}^n \text{Var}(X_i)$$

$$= \sum_{i=1}^n p(1 - p)$$

$$= np(1 - p)$$

X_i are **independent**,
therefore variance of sum
= sum of variance

Variance of Bernoulli



(live)

13: Statistics of Multiple RVs

Slides by Lisa Yan
July 20, 2020

Where are we now? A roadmap of CS109

Today: Statistics of multiple RVs!

$$\text{Var}(X + Y)$$

$$E[X + Y]$$

$$\text{Cov}(X, Y)$$

$$\rho(X, Y)$$

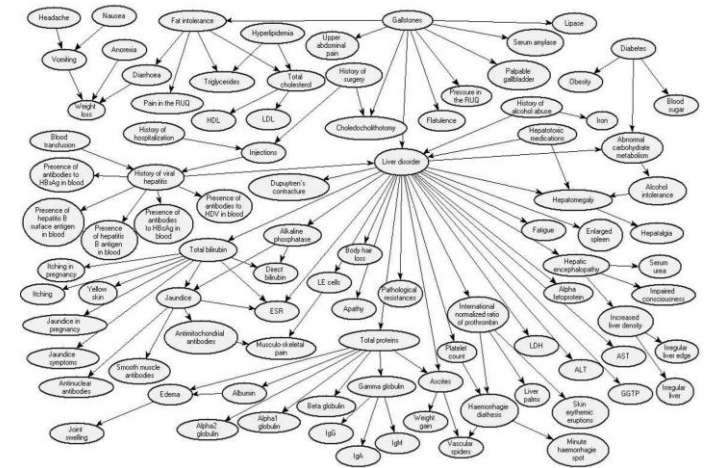
Wednesday:

Conditional distributions

$$p_{X|Y}(x|y)$$

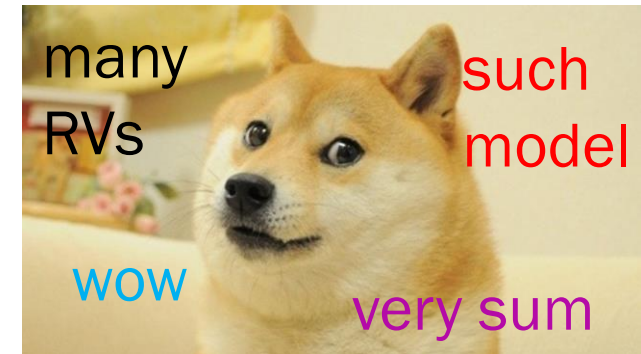
$$E[X|Y]$$

Also Wednesday: Modeling with Bayesian Networks



Last week: Joint distributions

$$p_{X,Y}(x, y)$$



Don't we already know linearity of expectation?

Expectation is a linear mathematical operation. If $X = \sum_{i=1}^n X_i$:

$$E[X] = E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i]$$

We covered this back in Lecture 6 (when we first learned expectation)!

- Proved binomial: sum of 1s or 0s
- Hat check (section): sum of 1s or 0s
- We ignored (in)dependence of **events**.

Why are we learning this again???

- Now we can prove it!
- We can now ignore (in)dependence of **random variables**.
- Our approach is still the same!

Coupon collecting problems: Hash tables

The **coupon collector's problem** in probability theory:

- You buy boxes of cereal.
- There are k different types of coupons
- For each box you buy, you "collect" a coupon of type i .

<u>Servers</u>	<u>Hash Tables</u>
requests	strings
k servers	k buckets
request to server i	hashed to bucket i

1. How many coupons do you expect after buying n boxes of cereal?



What is the expected number of utilized servers after n requests?

2. How many boxes do you expect to buy until you have one of each coupon?



What is the expected number of strings to hash until each bucket has ≥ 1 string?

Breakout Rooms

Check out the properties on the next slide.
Post any clarifications here!

<https://us.edstem.org/courses/667/discussion/93095>

Breakout rooms: 4 min. Introduce yourself!



Hash Tables

$$E \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n E[X_i]$$

Consider a hash table with k buckets.

- Strings are equally likely to get hashed into any bucket (independently).
- Let $Y = \#$ strings to hash until each bucket ≥ 1 string.

What is $E[Y]$?

1. Define additional random variables.

How should we define Y_i such that $Y = \sum_i Y_i$?

2. Solve.



Hash Tables

$$E \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n E[X_i]$$

Consider a hash table with k buckets.

- Strings are equally likely to get hashed into any bucket (independently).
- Let $Y = \#$ strings to hash until each bucket ≥ 1 string.

What is $E[Y]$?

1. Define additional random variables.

Let: $Y_i = \#$ of trials to get success after i -th success

- Success: hash string to previously empty bucket
- If i non-empty buckets: $P(\text{success}) = \frac{k-i}{k}$

2. Solve.

$$P(Y_i = n) = \left(\frac{i}{k}\right)^{n-1} \left(\frac{k-i}{k}\right)$$

$$\text{Equivalently, } Y_i \sim \text{Geo} \left(p = \frac{k-i}{k} \right) \quad E[Y_i] = \frac{1}{p} = \frac{k}{k-i}$$

Hash Tables

$$E \left[\sum_{i=1}^n X_i \right] = \sum_{i=1}^n E[X_i]$$

Consider a hash table with k buckets.

- Strings are equally likely to get hashed into any bucket (independently).
- Let $Y = \#$ strings to hash until each bucket ≥ 1 string.

What is $E[Y]$?

1. Define additional random variables.

Let: $Y_i = \#$ of trials to get success after i -th success

$$Y_i \sim \text{Geo} \left(p = \frac{k-i}{k} \right), \quad E[Y_i] = \frac{1}{p} = \frac{k}{k-i}$$

2. **Solve.** $Y = Y_0 + Y_1 + \dots + Y_{k-1}$

$$E[Y] = E[Y_0] + E[Y_1] + \dots + E[Y_{k-1}]$$

Even if Y_i dependent,
it wouldn't affect expectation!

$$= \frac{k}{k} + \frac{k}{k-1} + \frac{k}{k-2} + \dots + \frac{k}{1} = k \left[\frac{1}{k} + \frac{1}{k-1} + \dots + 1 \right] = O(k \log k)$$

The **covariance** of two variables X and Y is:

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

Think

The next slide has a question to go over by yourself.

Post any clarifications here!

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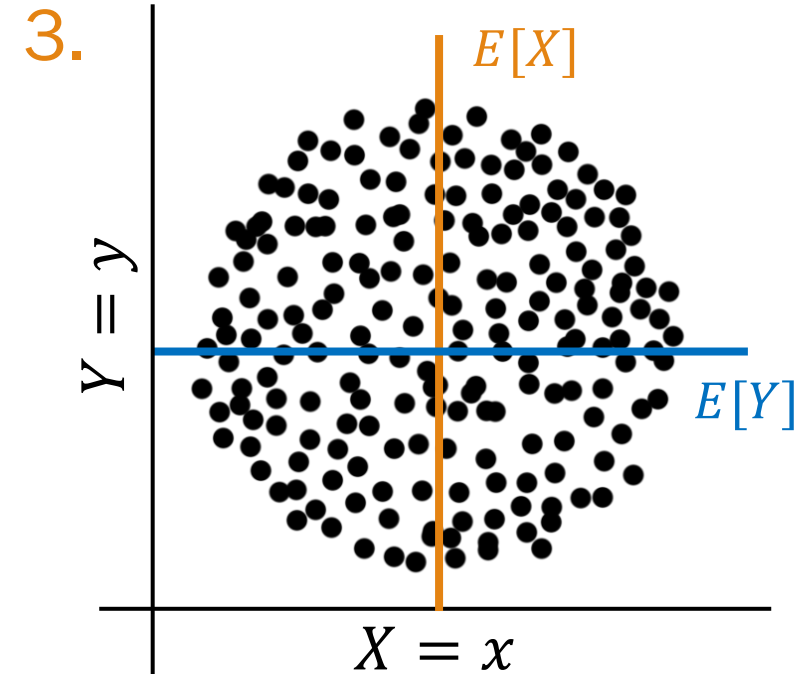
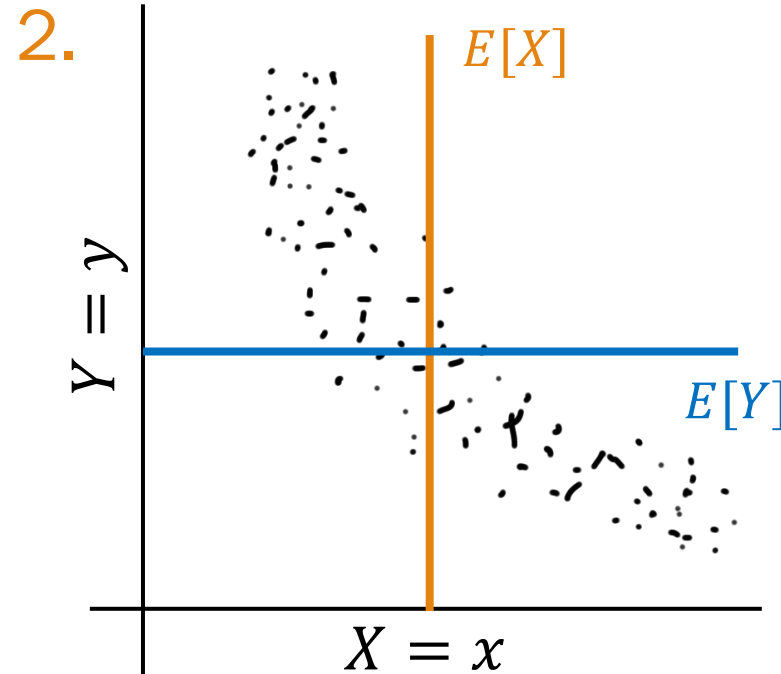
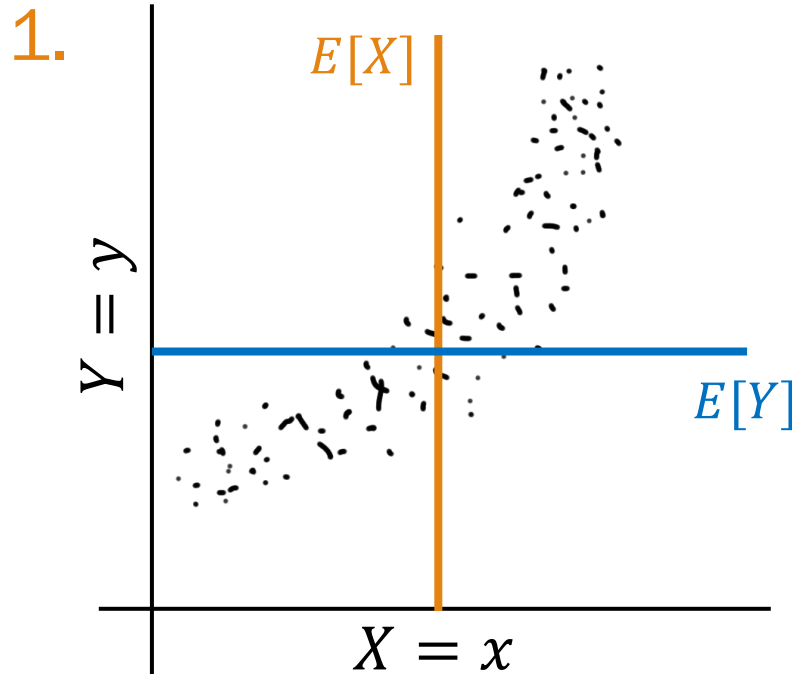
Think by yourself: 1 min



Feel the covariance

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

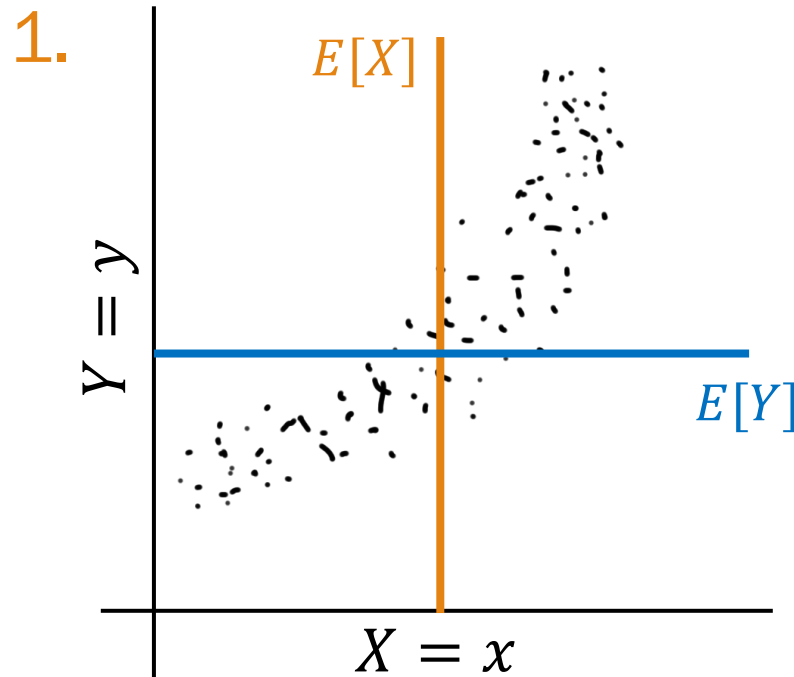
Is the covariance positive, negative, or zero?



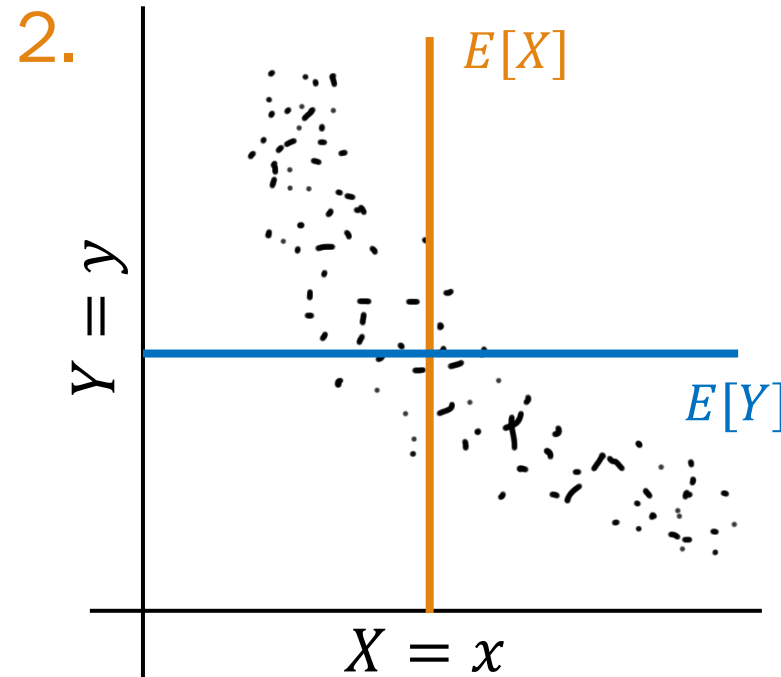
Feel the covariance

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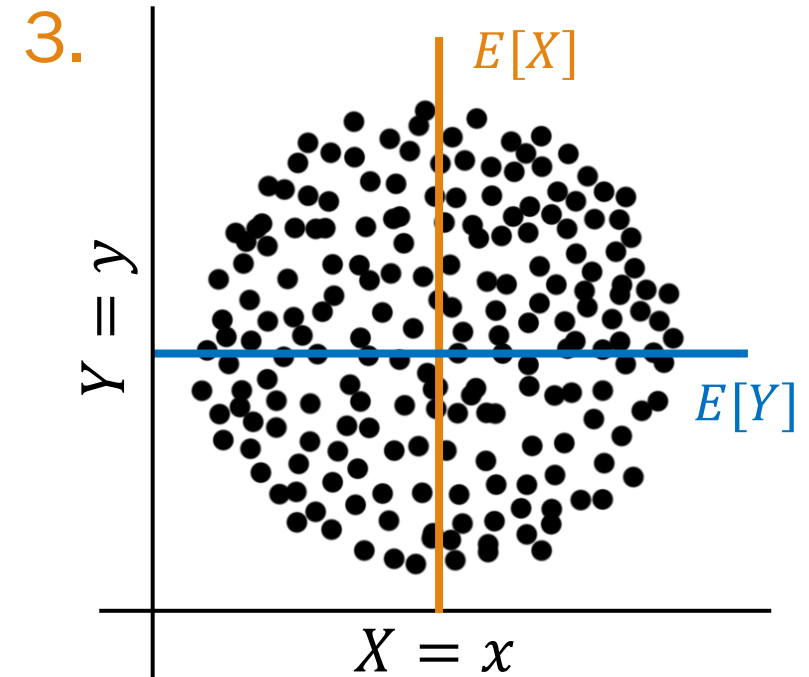
Is the covariance positive, negative, or zero?



positive



negative



zero

Properties of Covariance

The covariance of two variables X and Y is:

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

Properties:

1. $\text{Cov}(X, Y) = \text{Cov}(Y, X)$
2. $\text{Var}(X) = \text{Cov}(X, X)$
3. $\text{Cov}(\sum_i X_i, \sum_j Y_j) = \sum_i \sum_j \text{Cov}(X_i, Y_j)$
- ✗ 4. ~~$\text{Cov}(aX + b, Y) = a\text{Cov}(X, Y) + b$?~~
Covariance is non-linear: $\text{Cov}(aX + b, Y) = a\text{Cov}(X, Y)$

For any random variables X and Y ,

$$E[X + Y] = E[X] + E[Y]$$

$$\text{Var}(X + Y) = \text{Var}(X) + 2 \cdot \text{Cov}(X, Y) + \text{Var}(Y)$$

For **independent** X and Y ,

$$E[XY] = E[X]E[Y]$$

(Lemma: proof in extra slides)

$$\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$$

$\text{Cov}(X, Y) = 0$ does NOT imply independence of X and Y !

Zero covariance does not imply independence

Let X take on values $\{-1,0,1\}$
with equal probability $1/3$.

Define $Y = \begin{cases} 1 & \text{if } X = 0 \\ 0 & \text{otherwise} \end{cases}$

What is the joint PMF of X and Y ?

Breakout Rooms

Check out the properties on the next slide.
Post any clarifications here!

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Breakout rooms: 4 min. Introduce yourself!



Zero covariance does not imply independence

Let X take on values $\{-1, 0, 1\}$ with equal probability $1/3$.

Define $Y = \begin{cases} 1 & \text{if } X = 0 \\ 0 & \text{otherwise} \end{cases}$

		X			
		-1	0	1	
Y	0	1/3	0	1/3	2/3
	1	0	1/3	0	1/3
		1/3	1/3	1/3	

Marginal PMF of Y , $p_Y(y)$

Marginal PMF of X , $p_X(x)$

1. $E[X] =$ $E[Y] =$

2. $E[XY] =$

3. $\text{Cov}(X, Y) =$

4. Are X and Y independent?



Zero covariance does not imply independence

Let X take on values $\{-1, 0, 1\}$ with equal probability $1/3$.

Define $Y = \begin{cases} 1 & \text{if } X = 0 \\ 0 & \text{otherwise} \end{cases}$

		X			
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		1/3	1/3	1/3	

Marginal PMF of Y , $p_Y(y)$

Marginal PMF of X , $p_X(x)$

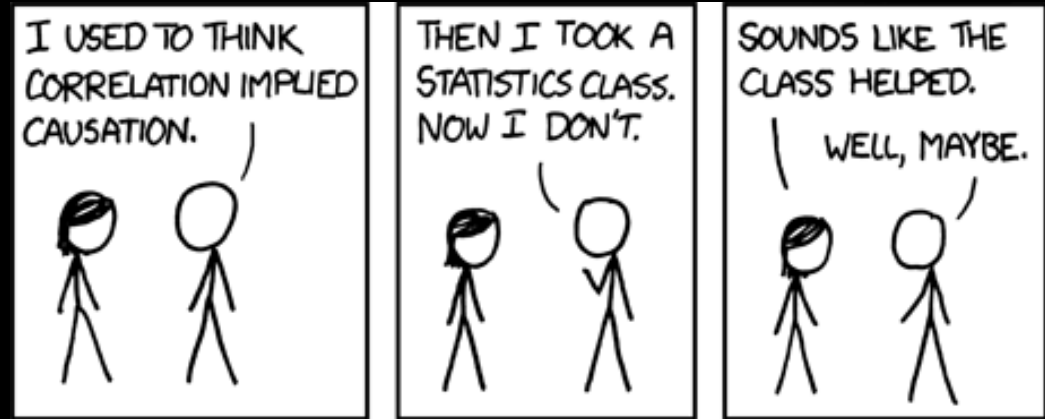
$$1. \quad E[X] = -1 \left(\frac{1}{3}\right) + 0 \left(\frac{1}{3}\right) + 1 \left(\frac{1}{3}\right) = 0 \quad E[Y] = 0 \left(\frac{2}{3}\right) + 1 \left(\frac{1}{3}\right) = 1/3$$

$$2. \quad E[XY] = (-1 \cdot 0) \left(\frac{1}{3}\right) + (0 \cdot 1) \left(\frac{1}{3}\right) + (1 \cdot 0) \left(\frac{1}{3}\right) = 0$$

$$3. \quad \text{Cov}(X, Y) = E[XY] - E[X]E[Y] = 0 - 0(1/3) = 0 \quad \triangle \text{ does not imply independence!}$$

4. Are X and Y independent? **✗**

$$P(Y = 0 | X = 1) = 1 \neq P(Y = 0) = 2/3$$



Interlude for fun/announcements

Announcements

Midterm Quiz

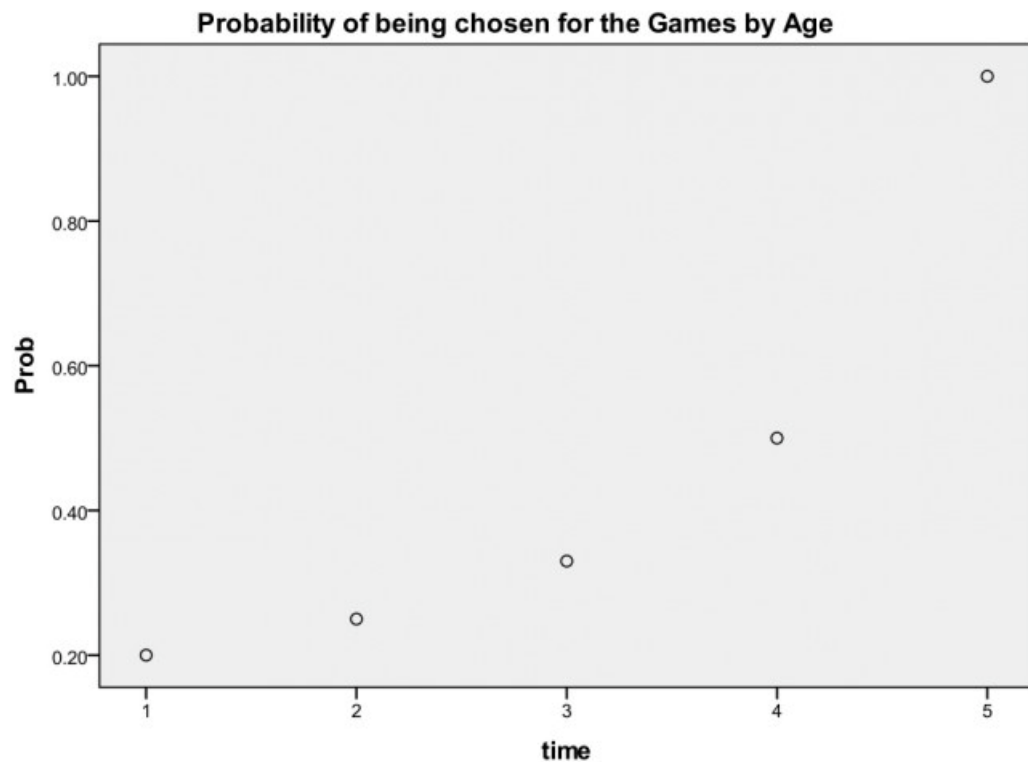
Start: Today (Mon) 5PM PDT – find on Website

Due: Tomorrow (Tue) 5PM PDT – submit to Gradescope

More notes: (no office hours tomorrow, Ed will be set to private-questions-only mode, we'll make clarifications via Ed)

Interesting probability news

Probability and Game Theory in *The Hunger Games*



1 = 12yrs. old; 2 = 13yrs. old; 3 = 14yrs. old; 4 = 15yrs. old; 5 = 16yrs. old

“Suppose the parents in a given district gave birth to only...five girls, and that all of these kids were born at the same time.”

- Not a probability mass function
- Also duh? ($P(\text{you get chosen if you're the only person}) = 1$)
- You now know enough Python/ probability to write a better simulation to model the Reaping!!!!
- (game theory part of the article is good)

<https://www.wired.com/2012/04/probability-and-game-theory-in-the-hunger-games/>

Ethics in Probability: Smoking and Cancer

Correlation does not imply causation

Does lung cancer cause smoking?

<https://towardsdatascience.com/correlation-does-not-imply-causation-92e4832a6713>

“Is it possible then, that lung cancer — that is to say, the pre-cancerous condition which must exist and is known to exist for years in those who are going to show over lung cancer — is one of the causes of smoking cigarettes? I don’t think it can be excluded.”

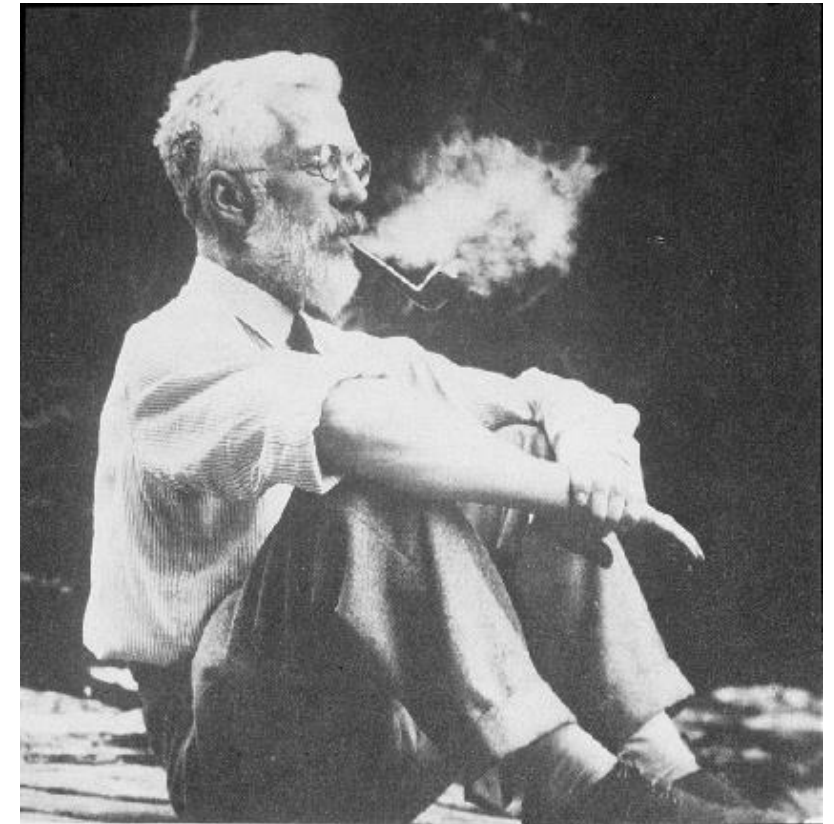
- Statistician R.A. Fisher

How, then, do we think about correlation and causation?

Fisher’s Paper (1958): <https://www.nature.com/articles/182596a0>

A reference paper from Judea Pearl on causal inference:

<https://www.ncbi.nlm.nih.gov/pmc/articles/PMC2836213/>



<http://www.economics.soton.ac.uk/staff/aldrich/fisherguide/Doc1.htm>

Correlation

Covarying humans

$$\begin{aligned}\text{Cov}(X, Y) &= E[(X - E[X])(Y - E[Y])] \\ &= E[XY] - E[X]E[Y]\end{aligned}$$

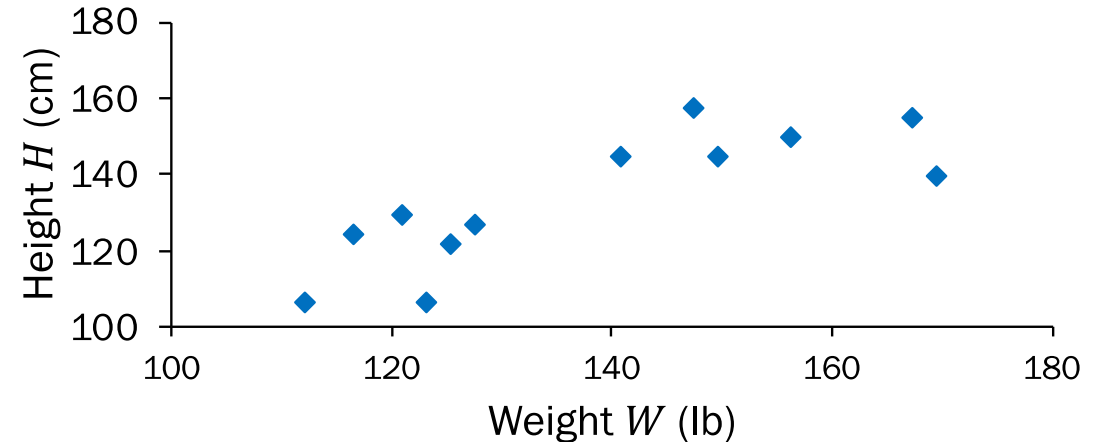
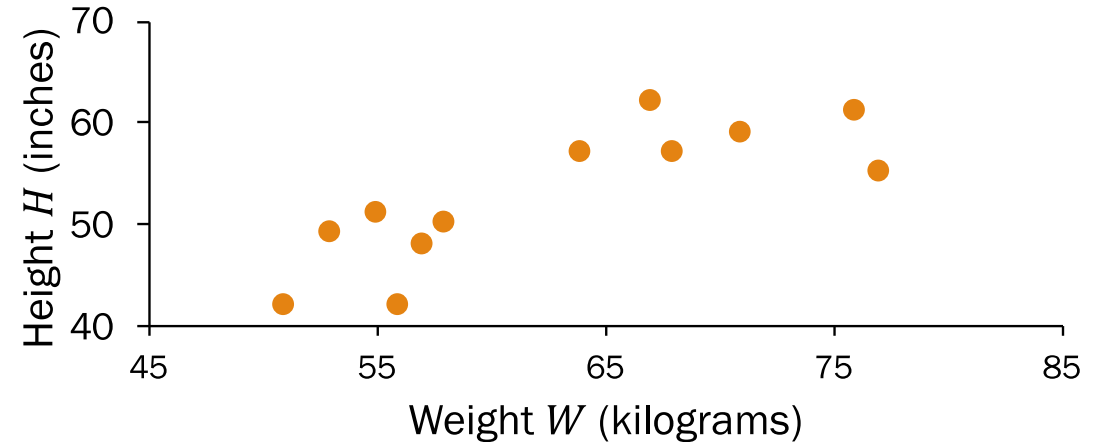
What is the covariance of weight W and height H ?

$$\begin{aligned}\text{Cov}(W, H) &= E[WH] - E[W]E[H] \\ &= 3355.83 - (62.75)(52.75) \\ &= 45.77 \text{ (positive)}\end{aligned}$$

What about weight (lb) and height (cm)?

$$\begin{aligned}\text{Cov}(2.20W, 2.54H) &= E[2.20W \cdot 2.54H] - E[2.20W]E[2.54H] \\ &= 18752.38 - (138.05)(133.99) \\ &= 255.06 \text{ (positive)}\end{aligned}$$

⚠ Covariance depends on units!



Sign of covariance (+/-) more meaningful than magnitude

Correlation

The **correlation** of two variables X and Y is:

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

$$\begin{aligned}\sigma_X^2 &= \text{Var}(X), \\ \sigma_Y^2 &= \text{Var}(Y)\end{aligned}$$

- Note: $-1 \leq \rho(X, Y) \leq 1$
- Correlation measures the **linear relationship** between X and Y :

$$\rho(X, Y) = 1 \quad \Rightarrow Y = aX + b, \text{ where } a = \sigma_Y / \sigma_X$$

$$\rho(X, Y) = -1 \quad \Rightarrow Y = -aX + b, \text{ where } a = \sigma_Y / \sigma_X$$

$$\rho(X, Y) = 0 \quad \Rightarrow \text{“uncorrelated” (absence of linear relationship)}$$

Think

The next slide has a question to go over by yourself.

Post any clarifications here!

<https://us.edstem.org/courses/667/discussion/93095>

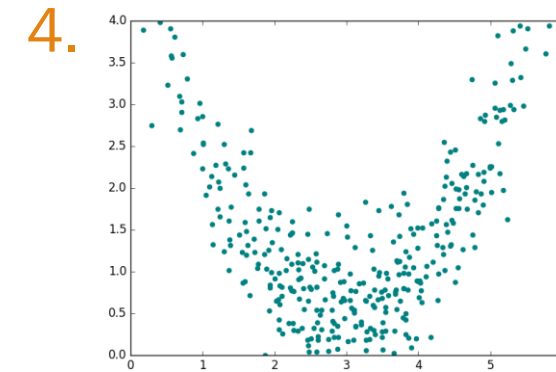
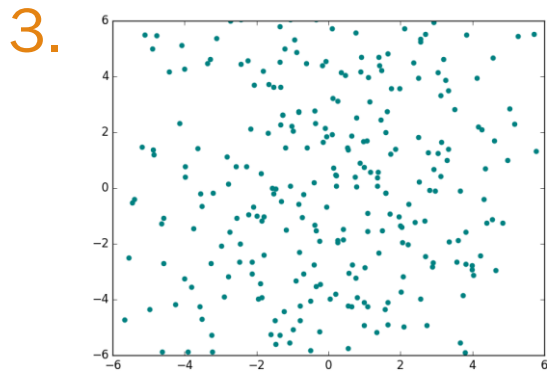
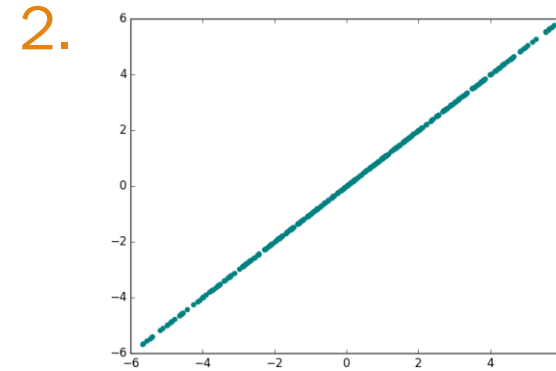
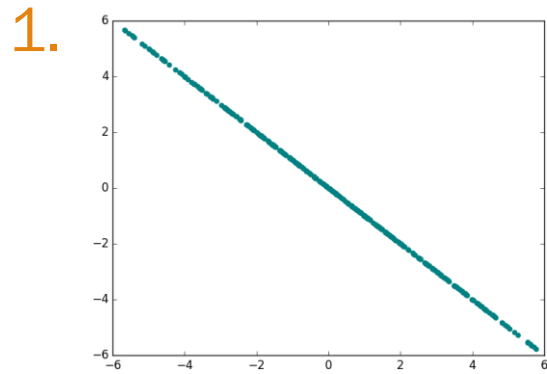
Think by yourself: 1 min



Correlation reps

- A. $\rho(X, Y) = 1$
- B. $\rho(X, Y) = -1$
- C. $\rho(X, Y) = 0$
- D. Other

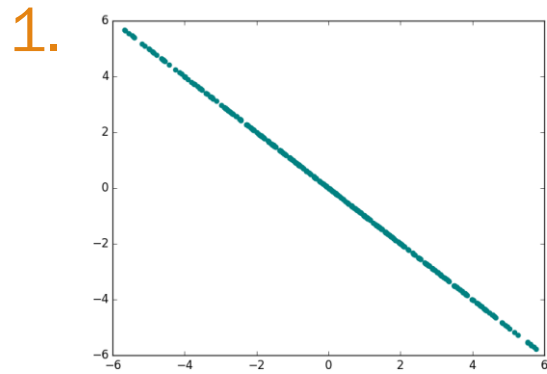
What is the correlation coefficient $\rho(X, Y)$?



Correlation reps

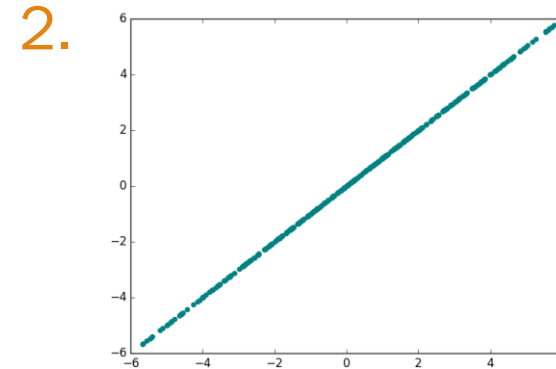
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What is the correlation coefficient $\rho(X, Y)$?



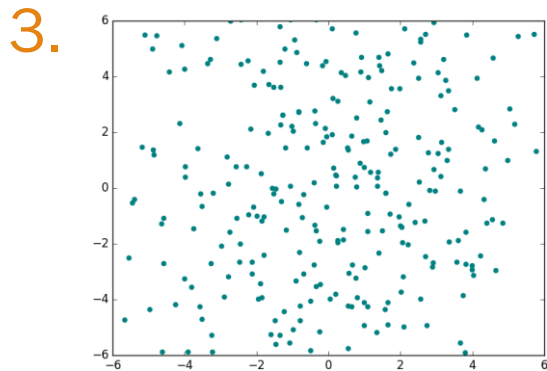
B. $\rho(X, Y) = -1$

$$Y = -aX + b$$
$$a > 0$$



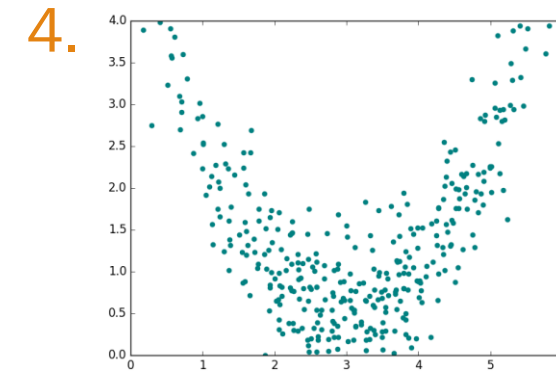
A. $\rho(X, Y) = 1$

$$Y = aX + b$$
$$a > 0$$



C. $\rho(X, Y) = 0$

“uncorrelated”

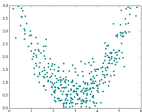


C. $\rho(X, Y) = 0$

$$Y = X^2$$

X and Y can be nonlinearly related even if $\rho(X, Y) = 0$.

CS103: Conditional statements

Statement $P \rightarrow Q$:	Independence \rightarrow No correlation	✓
Contrapositive $\neg Q \rightarrow \neg P$:	Correlation \rightarrow Dependence	✓ (logically equivalent)
Inverse $\neg P \rightarrow \neg Q$:	Dependence \rightarrow Correlation?	✗ (not always) $Y = X^2$ $\rho(X, Y) = 0$ 
Converse $Q \rightarrow P$:	No correlation \rightarrow Independence?	✗ (not always) Slide 46

“Correlation does not imply causation”

Spurious Correlations

$\rho(X, Y)$ is used a lot to statistically quantify the relationship b/t X and Y.

Correlation:
0.947091

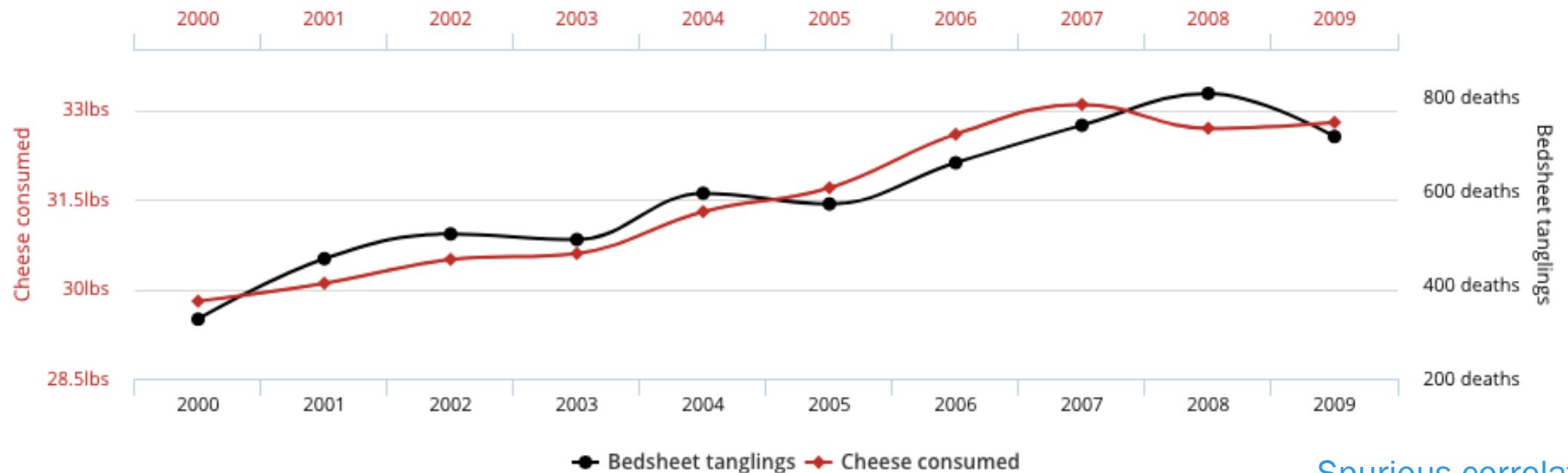


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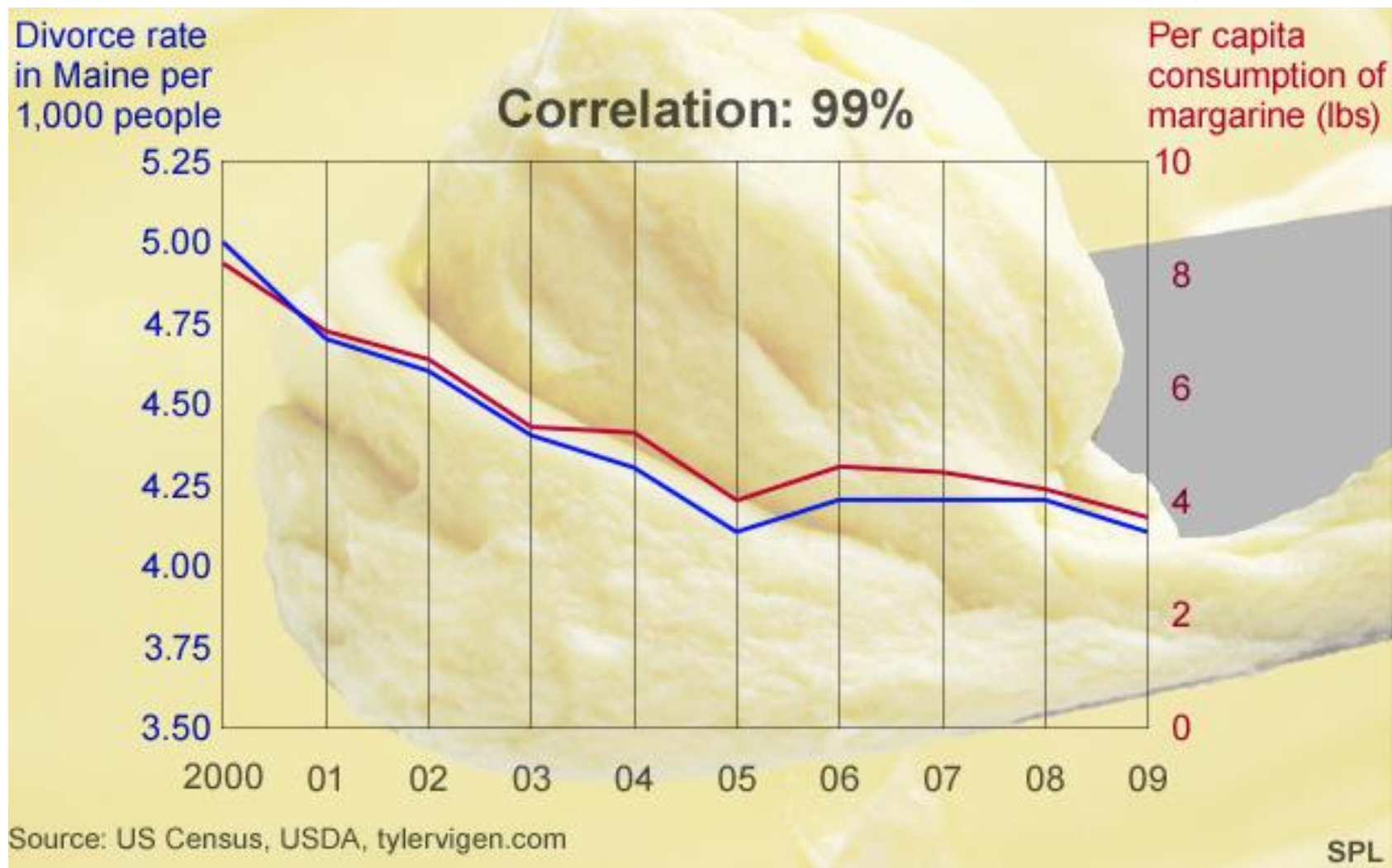
Per capita cheese consumption
correlates with
Number of people who died by becoming tangled in their bedsheets



[Spurious correlations](#)

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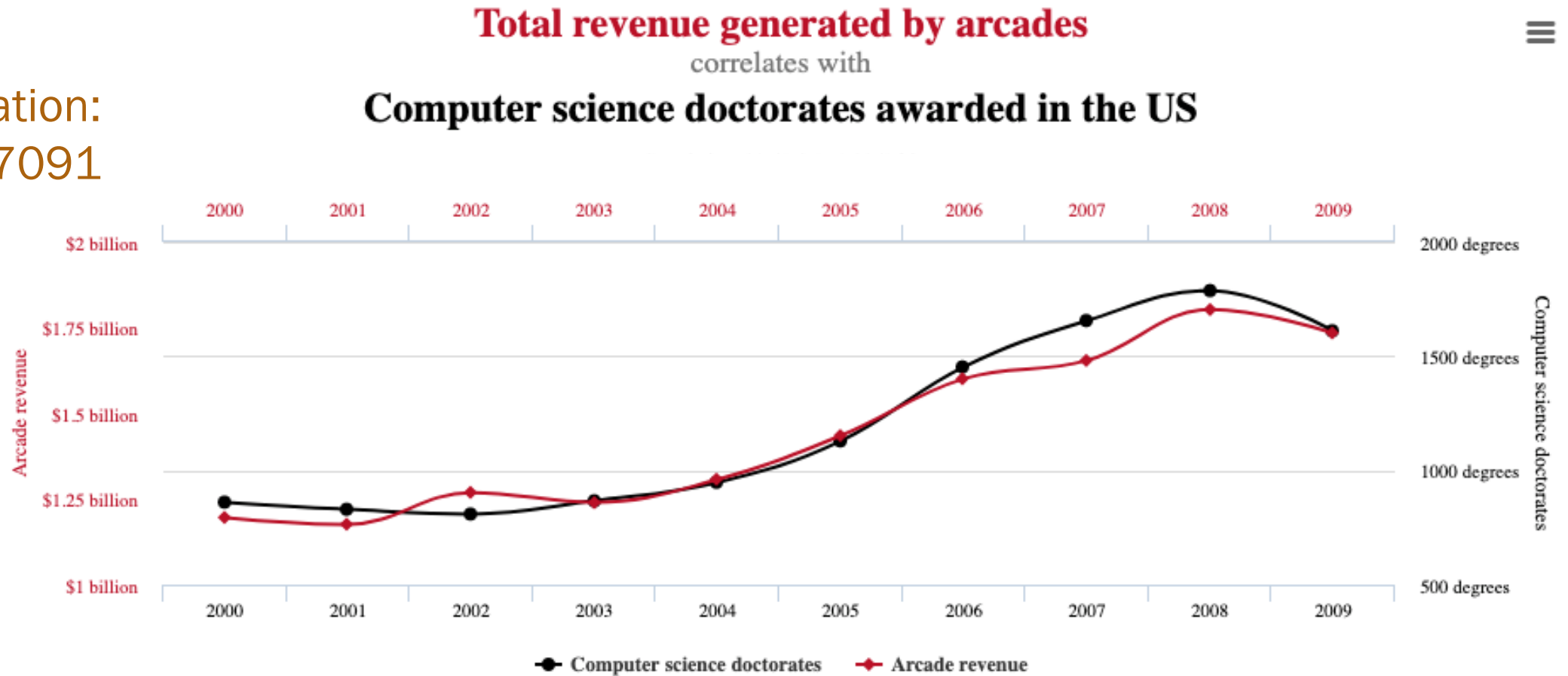
Divorce vs. Butter



<http://www.bbc.com/news/magazine-27537142>

Arcade revenue vs. CS PhDs

Correlation:
0.947091



Data sources: U.S. Census Bureau and National Science Foundation

tylervigen.com

[Spurious correlations](#)

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Extra

Expectation of product of independent RVs

If X and Y are
independent, then

$$E[XY] = E[X]E[Y]$$
$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

Proof: $E[g(X)h(Y)] = \sum_y \sum_x g(x)h(y)p_{X,Y}(x,y)$ (for continuous proof, replace summations with integrals)

$$= \sum_y \sum_x g(x)h(y)p_X(x)p_Y(y)$$
 X and Y are independent
$$= \sum_y \left(h(y)p_Y(y) \sum_x g(x)p_X(x) \right)$$
 Terms dependent on y are constant in integral of x
$$= \left(\sum_x g(x)p_X(x) \right) \left(\sum_y h(y)p_Y(y) \right)$$
 Summations separate
$$= E[g(X)]E[h(Y)]$$

Variance of Sums of Variables

$$\text{Var}\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n \text{Var}(X_i) + 2 \sum_{i=1}^n \sum_{j=i+1}^n \text{Cov}(X_i, X_j)$$

Proof:

$$\begin{aligned} \text{Var}\left(\sum_{i=1}^n X_i\right) & \stackrel{\text{Var}(X) = \text{Cov}(X, X)}{=} \text{Cov}\left(\sum_{i=1}^n X_i, \sum_{i=1}^n X_i\right) \stackrel{\text{covariance of all pairs}}{=} \sum_{i=1}^n \sum_{j=1}^n \text{Cov}(X_i, X_j) \\ & = \sum_{i=1}^n \text{Var}(X_i) + \sum_{i=1}^n \sum_{j=1, j \neq i}^n \text{Cov}(X_i, X_j) \\ & = \sum_{i=1}^n \text{Var}(X_i) + 2 \sum_{i=1}^n \sum_{j=i+1}^n \text{Cov}(X_i, X_j) \end{aligned}$$

Symmetry of covariance
 $\text{Cov}(X, X) = \text{Var}(X)$

Adjust summation bounds