14: Conditional Expectation

Lisa Yan and Jerry Cain October 14, 2020

Quick slide reference

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14a_conditional_distributions

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LIVE

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14a_conditional_distributions

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Discrete conditional distributions

Discrete conditional distributions

Recall the definition of the conditional probability of event *E* given event *F*:

$$P(E|F) = \frac{P(EF)}{P(F)}$$

For discrete random variables X and Y, the conditional PMF of X given Y is

$$P(X = x | Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

Different notation, same idea:

$$p_{X|Y}(x|y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$$

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Discrete probabilities of CS109

Each student responds	with:
-----------------------	-------

Year Y

- 1: Frosh/Soph
- 2: Jr/Sr
- 3: Co-term/grad/NDO

Timezone T (12pm California time in my timezone is):

- -1: AM
- 0: noon
- 1: PM

Joint PMF				
	Y = 1	Y = 2	Y = 3	
T = -1	.06	.01	.01	
T = 0	.29	.14	.09	
T = 1	.30	.08	.02	
P(Y = 3, T = 1)				

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Joint PMFs sum to 1.

Discrete probabilities of CS109

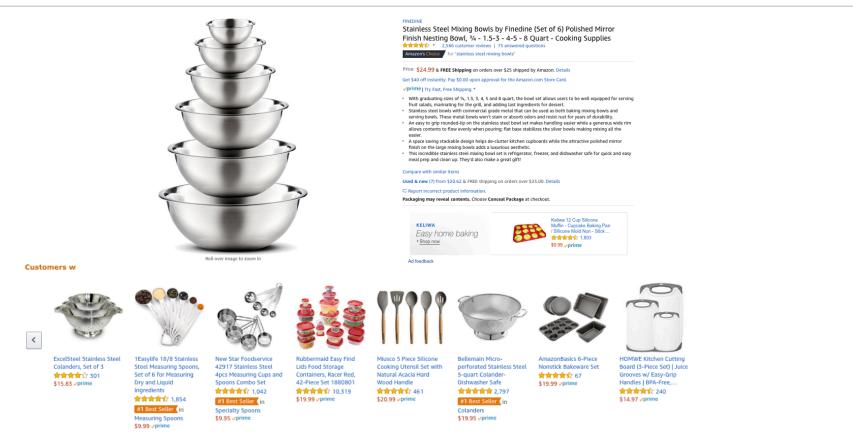
The below are conditional probability tables for conditional PMFs					$\begin{array}{c} \underline{\text{Joint PMF}} \\ Y = 1 Y = 2 Y = 3 \end{array}$				
			d (B) P	(T = t Y = y).	T = -1	.0	6.01	01	
		-			T = 0	.29	9.14	.09	
						.02			
2. What's the missing probability?									
I				1					
	Y = 1 Y	Y = 2 Y	f = 3	<u> </u>	= 1 Y =	= 2 Y =	= 3		
T = -1	.09	.04	.08	T = -1	.75 .1	.25			
T = 0	.45	.61	.75	T = 0	.56	.27	.17		
T = 1	.46	.35	.17	T = 1	.75	.2	.05		

Discrete probabilities of CS109

The below are conditional proba for conditional PMFs	bility tables		$\begin{array}{c c} \underline{\text{Joint}} \\ Y = 1 & Y \end{array}$	$\frac{PMF}{Y=2} Y=3$
(A) $P(Y = y T = t)$ and (B) $P(T)$	= t Y = v).	T = -1	.06	.01 .01
1. Which is which?		T = 0	.29	.14 .09
		T = 1	.30	.08 .02
2. What's the missing probabilit (B) $P(T = t Y = y)$	-	A) $P(Y =$	y T = t)	
Y = 1 Y = 2 Y = 3	Y	Y = 1 Y =	= 2 Y = 3	175125
T = -1 .09 .04 .08	T = -1	.75 .1	.25 .125	
T = 0 .45 .61 .75	T = 0	.56	.27 .17	,
T = 1 .46 .35 .17	T = 1	.75	.2 .05)
.30/(.06+.29+.30)	Conditional P different ever		sum to 1 cc	onditioned on

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Extended to Amazon



P(bought item X | bought item Y)

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Quick check

$$P(X = x | Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

Number or function?

1. P(X = 2|Y = 5)

2. P(X = x | Y = 5)

3. P(X = 2|Y = y)

 $4. \quad P(X = x | Y = y)$

True or false?

5.
$$\sum_{x} P(X = x | Y = 5) = 1$$

6. $\sum_{y} P(X = 2 | Y = y) = 1$
7. $\sum_{y} \sum_{y} P(X = x | Y = y) = 1$

8.
$$\sum_{x} \left(\sum_{y} P(X = x | Y = y) P(Y = y) \right) = 1$$

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 \overline{x} y

Quick check

$$P(X = x | Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)}$$

Number or function?

1.
$$P(X = 2|Y = 5)$$

number

- 2. P(X = x | Y = 5)1-D function
- 3. P(X = 2|Y = y)1-D function

4.
$$P(X = x | Y = y)$$

2-D function

True or false?

5.
$$\sum_{x} P(X = x | Y = 5) = 1 \quad \text{true}$$

6.
$$\sum_{y} P(X = 2 | Y = y) = 1 \quad \text{false}$$

7.
$$\sum_{x} \sum_{y} P(X = x | Y = y) = 1 \quad \text{false}$$

8.
$$\sum_{x} \left(\sum_{y} P(X = x | Y = y) P(Y = y) \right) = 1 \quad \text{true}$$

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14b_web_servers

Web server requests, redux

Web server requests (Lecture: Independent RVs)

Review

Let N = # of requests to a web server per day. Suppose $N \sim Poi(\lambda)$.

- Each request independently comes from a human (prob. p), or bot (1 p).
- Let *X* be *#* of human requests/day, and *Y* be *#* of bot requests/day.

Are X and Y independent? What are their marginal PMFs?

Our approach:

• Yes, independent Poisson random variables:

$$X \sim \mathsf{Poi}(\lambda p), Y \sim \mathsf{Poi}(\lambda(1-p))$$

- Two big parts of our derivation:
 - P(X = n, Y = m) = P(X = n | N = n + m)P(N = n)
 - $X|N = n + m \sim Bin(n + m, p)$

A conditional distribution, X | N!

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Web server requests, redux

(Note: this is a different problem setup from the previous slide)

Consider the number of requests to a web server per day.

- Let X = # requests from humans/day. $X \sim Poi(\lambda_1)$
- Let Y = # requests from bots/day.

What is P(X = k | X + Y = n)?

X and Y are independent.

$$Y \sim \text{Poi}(\lambda_2)$$

$$\rightarrow X + Y \sim \operatorname{Poi}(\lambda_1 + \lambda_2)$$

 $P(X = k | X + Y = n) = \frac{P(X = k, Y = n - k)}{P(X + Y = n)} = \frac{P(X = k)P(Y = n - k)}{P(X + Y = n)}$ (X,Y indep.) $= \frac{e^{-\lambda_1}\lambda_1^k}{k!} \cdot \frac{e^{-\lambda_2}\lambda_2^{n-k}}{(n-k)!} \cdot \frac{n!}{e^{-(\lambda_1 + \lambda_2)}(\lambda_1 + \lambda_2)^n} = \frac{n!}{k!(n-k)!} \cdot \frac{\lambda_1^k\lambda_2^{n-k}}{(\lambda_1 + \lambda_2)^n}$ $= \binom{n}{k} \left(\frac{\lambda_1}{\lambda_1 + \lambda_2}\right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2}\right)^{n-k} X | X + Y \sim \text{Bin}\left(X + Y, \frac{\lambda_1}{\lambda_1 + \lambda_2}\right)$

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14c_cond_expectation

Conditional Expectation

Conditional expectation

Recall the the conditional PMF of X given Y = y:

$$p_{X|Y}(x|y) = P(X = x|Y = y) = \frac{p_{X,Y}(x,y)}{p_Y(y)}$$

The **conditional expectation** of *X* given Y = y is

$$E[X|Y = y] = \sum_{x} xP(X = x|Y = y) = \sum_{x} xp_{X|Y}(x|y)$$

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It's been so long, our dice friends

- Roll two 6-sided dice.
- Let roll 1 be D_1 , roll 2 be D_2 .
- Let S = value of $D_1 + D_2$.



 $E[X|Y = y] = \sum x p_{X|Y}(x|y)$

1. What is
$$E[S|D_2 = 6]$$
? $E[S|D_2 = 6] = \sum_{x=7}^{12} xP(S = x|D_2 = 6)$
 $= \left(\frac{1}{6}\right)(7 + 8 + 9 + 10 + 11 + 12)$
 $= \frac{57}{6} = 9.5$

Intuitively: $6 + E[D_1] = 6 + 3.5 = 9.5$ Let's prove this!

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Properties of conditional expectation

1. LOTUS:

$$E[g(X)|Y = y] = \sum_{x} g(x)p_{X|Y}(x|y)$$

2. Linearity of conditional expectation:

$$E\left[\sum_{i=1}^{n} X_{i} \mid Y = y\right] = \sum_{i=1}^{n} E[X_{i}|Y = y]$$

3. Law of total expectation (next time)

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It's been so long, our dice friends

- Roll two 6-sided dice.
- Let roll 1 be D_1 , roll 2 be D_2 .
- Let S = value of $D_1 + D_2$.
- 1. What is $E[S|D_2 = 6]$?
- 2. What is $E[S|D_2]$?
 - A. A function of S
 - B. A function of D_2
 - C. A number
- 3. Give an expression for $E[S|D_2]$.



 $E[X|Y = y] = \sum x p_{X|Y}(x|y)$



 $\frac{57}{6} = 9.5$

It's been so long, our dice friends

- Roll two 6-sided dice.
- Let roll 1 be D_1 , roll 2 be D_2 .
- Let S = value of $D_1 + D_2$.
- 1. What is $E[S|D_2 = 6]$?
- 2. What is $E[S|D_2]$?
 - A. A function of S B A function of D_2 C. A number
- 3. Give an expression for $E[S|D_2]$.



 $E[X|Y = y] = \sum x p_{X|Y}(x|y)$

$$E[S|D_{2} = d_{2}] = E[D_{1} + d_{2}|D_{2} = d_{2}]$$

$$= \sum_{d_{1}} (d_{1} + d_{2})P(D_{1} = d_{1}|D_{2} = d_{2})$$

$$= \sum_{d_{1}} d_{1}P(D_{1} = d_{1}) + d_{2} \sum_{d_{1}} P(D_{1} = d_{1})$$

$$= E[D_{1}] + d_{2} = 3.5 + d_{2}$$

$$E[S|D_{2}] = 3.5 + D_{2}$$

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 $\frac{57}{6} = 9.5$

14d_law_of_total_expectation

Law of Total Expectation

Properties of conditional expectation

1. LOTUS:

$$E[g(X)|Y = y] = \sum_{x} g(x)p_{X|Y}(x|y)$$

2. Linearity of conditional expectation:

$$E\left[\sum_{i=1}^{n} X_{i} \mid Y = y\right] = \sum_{i=1}^{n} E[X_{i} \mid Y = y]$$

3. Law of total expectation:

$$E[X] = E[E[X|Y]] \quad \text{what?!}$$

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Proof of Law of Total Expectation

$$E[X] = E[E[X|Y]]$$

$$E[E[X|Y]] = E[g(Y)] = \sum_{y} P(Y = y)E[X|Y = y]$$
(LOTUS, $g(Y) = E[X|Y]$)

(def of conditional expectation)

$$= \sum_{y} P(Y = y) \sum_{x} xP(X = x | Y = y)$$

$$= \sum_{y} \left(\sum_{x} xP(X = x | Y = y)P(Y = y) \right) = \sum_{y} \left(\sum_{x} xP(X = x, Y = y) \right)$$
(chain rule)

$$= \sum_{x} \sum_{y} xP(X = x, Y = y) = \sum_{x} x \sum_{y} P(X = x, Y = y)$$
 (switch order of summations)

(marginalization)

= E[X]...what?

 $=\sum_{x}xP(X=x)$

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Another way to compute E[X]

$$E[E[X|Y]] = \sum_{y} P(Y=y)E[X|Y=y] = E[X]$$

If we only have a conditional PMF of X on some discrete variable Y, we can compute E[X] as follows:

- **1.** Compute expectation of *X* given some value of Y = y
- 2. Repeat step 1 for all values of Y
- 3. Compute a weighted sum (where weights are P(Y = y))

```
def recurse():
    if (random.random() < 0.5):
        return 3
    else: return (2 + recurse())</pre>
```

Useful for analyzing recursive code!!

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14: Conditional Expectation

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Where are we now? A roadmap of CS109

Monday: Statistics of
multiple RVs!Friday: Modeling with
Bayesian NetworksVar(X + Y)Var(X + Y)Last week: Joint
distributionsE[X + Y]
Cov(X, Y)
 $\rho(X, Y)$ Image: Cov(X, Y)
 $\rho(X, Y)$

Today: Conditional distributions $p_{X|Y}(x|y)$ E[X|Y]



Time to kick it up a notch! Lisa Yan and Jerry Cain, CS109, 2020

Conditional Expectation



Conditional Distributions



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Breakout Rooms

Check out the question on the next slide (Slide 28). Post any clarifications here!

https://us.edstem.org/courses/2678/discussion/146233

Breakout rooms: 4 min. Introduce yourself!



Quick check

- **1.** E[X]
- $2. \quad E[X,Y]$
- **3.** E[X + Y]
- $4. \quad E[X|Y]$
- 5. E[X|Y = 6]
- 6. E[X = 1]
- **7**^{*}. E[Y|X = x]

- A. value
- B. one RV, function on *Y*
- C. one RV, function on X
- D. two RVs, function on *X* and *Y*
- E. doesn't make sense



Conditional Expectation

Review

The conditional expectation of *X* given Y = y is

$$E[X|Y = y] = \sum_{x} xP(X = x|Y = y) = \sum_{x} xp_{X|Y}(x|y)$$

• Interpret: E[X|Y] is a random variable that takes on the value E[X|Y = y] with probability P(Y = y)

The Law of Total Expectation states that

$$E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y) = E[X]$$

• Apply: E[X] can be calculated as the expectation of E[X|Y]

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```
def recurse():
    # equally likely values 1,2,3
    x = np.random.choice([1,2,3])
    if (x == 1): return 3
    elif (x == 2): return (5 + recurse())
    else: return (7 + recurse())
```

$$E[X] = E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y)$$

Let Y =return value of recurse(). What is E[Y]?

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def recurse():
 # equally likely values 1,2,3
 x = np.random.choice([1,2,3])
 if (x == 1): return 3
 elif (x == 2): return (5 + recurse())
 else: return (7 + recurse())

 $E[X] = E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y)$

Let Y =return value of recurse(). What is E[Y]?

E[Y] = E[Y|X = 1]P(X = 1) + E[Y|X = 2]P(X = 2) + E[Y|X = 3]P(X = 3)

E[Y|X = 1] = 3When X = 1, return 3.

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Think

Slide 33 has a question to go over by yourself.

Post any clarifications here!

https://us.edstem.org/courses/2678/discussion/146233

Think by yourself: 2 min



32

def recurse(): # equally likely values 1,2,3 x = np.random.choice([1,2,3])**if** (x == 1): **return** 3 elif (x == 2): return (5 + recurse()) else: return (7 + recurse())

If Y discrete $E[X] = E[E[X|Y]] = \sum E[X|Y = y]P(Y = y)$

Let Y =return value of recurse(). What is E[Y]?

E[Y] = E[Y|X = 1]P(X = 1) + E[Y|X = 2]P(X = 2) + E[Y|X = 3]P(X = 3)

E[Y|X = 1] = 3

What is E[Y|X = 2]? A. E[5] + YB. E[Y + 5] = 5 + E[Y]C. 5 + E[Y|X = 2]

(by yourself



```
def recurse():
    # equally likely values 1,2,3
    x = np.random.choice([1,2,3])
    if (x == 1): return 3
    elif (x == 2): return (5 + recurse())
    else: return (7 + recurse())
```

 $E[X] = E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y)$

Let Y =return value of recurse(). What is E[Y]?

 $E[Y] = E[Y|X = 1]P(X = 1) + \frac{E[Y|X = 2]}{P(X = 2)} + \frac{E[Y|X = 3]P(X = 3)}{P(X = 3)}$

E[Y|X = 1] = 3 When X = 2, return 5 + a future return value of recurse(). What is E[Y|X = 2]? A. E[5] + YB. E[Y + 5] = 5 + E[Y]C. 5 + E[Y|X = 2]

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def recurse():
 # equally likely values 1,2,3
 x = np.random.choice([1,2,3])
 if (x == 1): return 3
 elif (x == 2): return (5 + recurse())
 else: return (7 + recurse())

 $E[X] = E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y)$

Let Y =return value of recurse(). What is E[Y]?

E[Y] = E[Y|X = 1]P(X = 1) + E[Y|X = 2]P(X = 2) + E[Y|X = 3]P(X = 3)

$$E[Y|X = 1] = 3$$
 $E[Y|X = 2] = E[5 + Y]$

When X = 3, return 7 + a future return value of recurse().

$$E[Y|X=3] = E[7+Y]$$

def recurse():
 # equally likely values 1,2,3
 x = np.random.choice([1,2,3])
 if (x == 1): return 3
 elif (x == 2): return (5 + recurse())
 else: return (7 + recurse())

 $E[X] = E[E[X|Y]] = \sum_{y} E[X|Y = y]P(Y = y)$

Let Y =return value of recurse(). What is E[Y]?

E[Y] = E[Y|X = 1]P(X = 1) + E[Y|X = 2]P(X = 2) + E[Y|X = 3]P(X = 3) $E[Y|X = 1] = 3 \qquad E[Y|X = 2] = E[5 + Y] \qquad E[Y|X = 3] = E[7 + Y]$ E[Y] = 3(1/3) + (5 + E[Y])(1/3) + (7 + E[Y])(1/3) E[Y] = (1/3)(15 + 2E[Y]) = 5 + (2/3)E[Y] E[Y] = 15On your own: What is Var(Y)?

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Interlude for jokes/announcements

Announcements

Quizzes Are Graded

Your custom solution available <u>here</u>! Regrade requests accepted through Monday, 11:59pm

Problem Set 3

Due:Friday 10/16 1pmCovers:Up to and including Lecture 11

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Interesting probability news

U.S. Recession Model at 100% Confirms Downturn Is Already Here

"Bloomberg Economics created a model last year to determine America's recession odds."

 I encourage you to read through and understand the parameters used to define this model!



Chance of Recession Within 12 Months

https://www.bloomberg.com/graphics/us-economicrecession-tracker/

Independent RVs, defined another way

If X and Y are independent discrete random variables, then $\forall x, y$:

$$P(X = x | Y = y) = \frac{P(X = x, Y = y)}{P(Y = y)} = \frac{P(X = x)P(Y = y)}{P(Y = y)} = P(X = x)$$
$$p_{X|Y}(x|y) = \frac{p_{X,Y}(x, y)}{p_Y(y)} = \frac{p_X(x)p_Y(y)}{p_Y(y)} = p_X(x)$$

Note for conditional expectation, independent X and Y implies

$$E[X|Y = y] = \sum_{x} x p_{X|Y}(x|y) = \sum_{x} x p_{X}(x) = E[X]$$

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Breakout Rooms

Check out the question on the next slide (Slide 42). Post any clarifications here!

https://us.edstem.org/courses/2678/discussion/146233

Breakout rooms: 4 min.



Random number of random variables

indep X, YE[X|Y = y] = E[X]

Say you have a website: BestJokesEver.com. Let:

- X = # of people per day who visit your site. $X \sim Bin(100, 0.5)$
- $Y_i = #$ of minutes spent per day by visitor $i = Y_i$
- X and all Y_i are independent.

The time spent by all visitors per day is W

$$W = \sum_{i=1}^{X} Y_i$$
. What is $E[W]$?



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Random number of random variables

Say you have a website: BestJokesEver.com. Let: X = # of people per day who visit your site. $X \sim Bin(100, 0.5)$ $Y_i = \#$ of minutes spent by visitor *i*. $Y_i \sim \text{Poi}(8)$ • X and all Y_i are independent. The time spent by all visitors per day is $W = \sum_{i=1}^{N} Y_i$. What is E[W]? $E[W] = E\left[\sum_{i=1}^{X} Y_i\right] = E\left[E\left[\sum_{i=1}^{X} Y_i | X\right]\right]$ Suppose X = x. $E\left[\sum_{i=1}^{x} Y_{i} | X = x\right] = \sum_{i=1}^{x} E[Y_{i} | X = x]$ (linearity) $= E[XE[Y_i]]$ $=\sum_{i=1}^{\infty} E[Y_i]$ (independence) $= E[Y_i]E[X] \qquad (\text{scalar } E[Y_i])$ $= xE[Y_i]$ $= 8 \cdot 50$

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See you next time!

Have a super Wednesday!



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(no video)

Extra

Your company has only one job opening for a software engineer.

- *n* candidates interview, in order (*n*! orderings equally likely)
- Must decide hire/no hire *immediately* after each interview
- Strategy: 1. Interview *k* (of *n*) candidates and reject all *k*
 - 2. Accept the next candidate better than all of first *k* candidates.

What is your target k that maximizes P(get best candidate)?

Fun fact:

- There is an α-to-1 factor difference in productivity b/t the "best" and "average" software engineer.
- Steve jobs said α =25, Mark Zuckerberg claims α =100, some even claim α =300

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Your company has only one job opening for a software engineer.

- *n* candidates interview, in order (*n*! orderings equally likely)
- Must decide hire/no hire *immediately* after each interview

Strategy: 1. Interview *k* (of *n*) candidates and reject all *k*

2. Accept the next candidate better than all of first *k* candidates.

What is your target k that maximizes P(get best candidate)?

Define: X = position of best engineer candidate (1, 2, ..., n) B = event that you hire the best engineer Want to maximize for k: $P_k(B)$ = probability of B when using strategy for a given k $P_k(B) = \sum_{i=1}^n P_k (B|X = i) P(X = i) = \frac{1}{n} \sum_{i=1}^n P_k(B|X = i)$ (law of total probability)

Your company has only one job opening for a software engineer. Strategy: 1. Interview k (of n) candidates and reject all k 2. Accept the next candidate better than all of first k candidates. What is your target k that maximizes P(get best candidate)? Define: X = position of best engineer candidate B = event that you hire the best engineer If $i \leq k$: $P_k(B|X=i) = 0$ (we fired best candidate already) Else: $P_k(B|X=i) = \frac{k}{i-1}$ We must not hire prior to the *i*-th candidate. \rightarrow We must have fired the best of the *i*-1 first candidates. \rightarrow The best of the i-1 needs to be our comparison point for positions k+1, ..., i-1. \rightarrow The best of the *i*-1 needs to be one of our first k comparison/auto-fire $P_k(B) = \frac{1}{n} \sum_{i=1}^{n} P_k(B|X=i) = \frac{1}{n} \sum_{i=1}^{n} \frac{k}{i-1}$ $= \frac{1}{n} \sum_{i=1}^{n} \frac{k}{i-1}$ Want to maximize over k Stanford University 48

Your company has only one job opening for a software engineer.

- Strategy: 1. Interview *k* (of *n*) candidates and reject all *k*
 - 2. Accept the next candidate better than all of first *k* candidates.

What is your target k that maximizes P(get best candidate)?

Want to maximize over k:

$$P_k(B) = \frac{1}{n} \sum_{i=k+1}^n \frac{k}{i-1} \approx \frac{k}{n} \int_{i=k+1}^n \frac{1}{i-1} di = \frac{k}{n} \ln(i-1) \Big|_{i=k+1}^n = \frac{k}{n} \ln \frac{n-1}{k} \approx \frac{k}{n} \ln \frac{n}{k}$$

Maximize by differentiating w.r.t k , set to 0, solve for k:

$$\frac{d}{dk} \left(\frac{k}{n} \ln \frac{n}{k}\right) = \frac{1}{n} \ln \frac{n}{k} + \frac{k}{n} \cdot \frac{k}{n} \cdot \frac{-n}{k^2} = 0$$

$$\ln \frac{n}{k} = 1$$

$$\lim_{k \to \infty} \frac{n}{k} \cdot \frac{n}{k} \cdot \frac{-n}{k^2} = 0$$

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$$\lim_{k \to \infty} \frac{n}{k^2} \cdot \frac{n}{k^2} \cdot$$