21: Beta

Lisa Yan and Jerry Cain October 30, 2020

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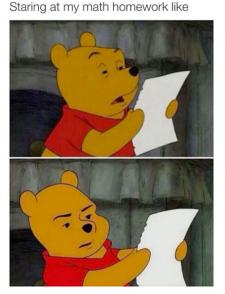
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21a_mle_multinomial

MLE: Multinomial

Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^{m} p_i = 1$
- $X_i = #$ of trials with outcome *i*, where $\sum_{i=1}^m X_i = n$



Let's give an example!

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^m p_i = 1$
- $X_i = #$ of trials with outcome *i*, where $\sum_{i=1}^m X_i = n$

Example: Suppose each RV is outcome of 6-sided die.

$$m = 6, \sum_{i=1}^{6} p_i = 1$$

- Roll the dice n = 12 times.
- Observe data: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

$$X_1 = 3, X_2 = 2, X_3 = 0,$$

 $X_4 = 3, X_5 = 1, X_6 = 3$

Check:
$$X_1 + X_2 + \dots + X_6 = 12$$

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^m p_i = 1$
- $X_i = #$ of trials with outcome *i*, where $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing the sample($X_1, X_2, ..., X_m$), given the probabilities $p_1, p_2, ..., p_m$?

A.
$$\frac{n!}{X_{1}! X_{2}! \cdots X_{m}!} p_{1}^{X_{1}} p_{2}^{X_{2}} \cdots p_{m}^{X_{m}}$$

B.
$$p_{1}^{X_{1}} p_{2}^{X_{2}} \cdots p_{m}^{X_{m}}$$

C.
$$\frac{n!}{X_{1}! X_{2}! \cdots X_{m}!} X_{1}^{p_{1}} X_{2}^{p_{2}} \cdots X_{m}^{p_{m}}$$

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^m p_i = 1$
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(A)
$$\frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$$

B.
$$p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$$

C.
$$\frac{n!}{X_1! X_2! \cdots X_m!} X_1^{p_1} X_2^{p_2} \cdots X_m^{p_m}$$

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^m p_i = 1$
- $X_i = #$ of trials with outcome *i*, where $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing the sample($X_1, X_2, ..., X_m$), given the probabilities $p_1, p_2, ..., p_m$? $L(\theta) = \frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$
- 2. What is θ_{MLE} ?

$$LL(\theta) = \log(n!) - \sum_{i=1}^{m} \log(X_i!) + \sum_{i=1}^{m} X_i \log(p_i)$$
, such that $\sum_{i=1}^{m} p_i = 1$

Optimize with Lagrange multipliers in extra slides

$$\rightarrow \theta_{MLE}: p_i = \frac{X_i}{n}$$
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Intuitively, probability $p_i = \text{proportion of outcomes}$ <u>Stanford University</u> 8

When MLEs attack!

MLE for $p_i = \frac{X_i}{n}$

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

What is θ_{MLE} ?



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When MLEs attack!

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

 θ_{MLE} :

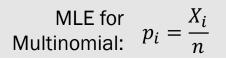
$$p_1 = 3/12$$

 $p_2 = 2/12$
 $p_3 = 0/12$
 $p_4 = 3/12$
 $p_5 = 1/12$
 $p_6 = 3/12$

- MLE: you'll never...<u>EVER</u>... roll a three.
- Do you really believe that?

Today: A new definition of probability!

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21b_bayesian

Bayesian Statistics

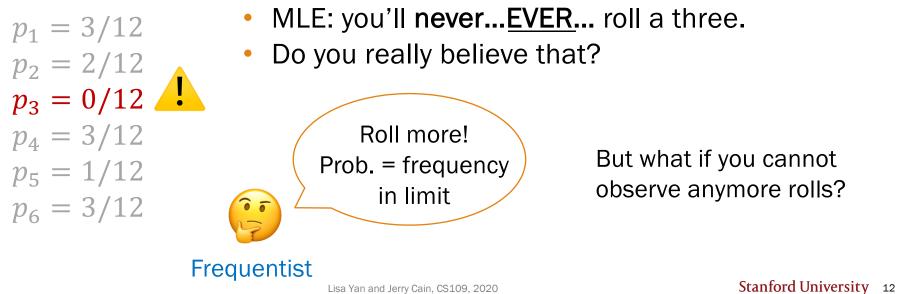
When MLEs attack!

Review

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

 θ_{MLE} :



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Today's plan

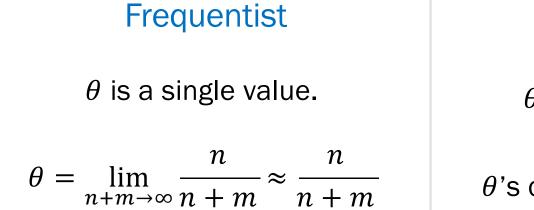
Today we are going to learn something unintuitive, beautiful, and useful!

We are going to think of probabilities as random variables.

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A new definition of probability

Flip a coin n + m times, come up with n heads. We don't know the probability θ that the coin comes up heads.





```
The world's first coin
```

Bayesian

 θ is a random variable.

 θ 's continuous support: (0, 1)

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Let's play a game

Roll 2 dice. If *neither* roll is a 6, you win (event W). Else, I win (event W^C).



- Before you play, what's the probability that you win?
- Play once. What's the probability that you win?
- Play three more times. What's the probability that you win?



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Bayesian probability

Bayesian statistics: Probability is a reasonable expectation representing a state of knowledge.

Mixing discrete and continuous random variables, combined with Bayes' Theorem, allows us to reason about probabilities as random variables.

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Mixing discrete and continuous

Let X be a continuous random variable, and N be a discrete random variable.

Bayes' Theorem:

$$f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$$

Intuition:
$$P(X = x | N = n) = \frac{P(N = n | X = x)P(X = x)}{P(N = n)}$$
$$f_{X|N}(x|n)\varepsilon_X = \frac{P(N = n | X = x)f_X(x)\varepsilon_X}{P(N = n)} \implies f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$$

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All your Bayes are belong to us

Let *X*, *Y* be continuous and *M*, *N* be discrete random variables.

OG Bayes:
$$p_{M|N}(m|n) = \frac{p_{N|M}(n|m)p_M(m)}{p_N(n)}$$
Mix Bayes #1: $f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$ Mix Bayes #2: $p_{N|X}(n|x) = \frac{f_{X|N}(x|n)p_N(n)}{f_X(x)}$ All continuous: $f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$



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Mixing discrete and continuous

Let θ be a random variable for the probability your coin comes up heads, and N be the number of heads you observe in an experiment.

posterior

$$f_{\theta|N}(x|n) = \frac{\substack{\text{likelihood prior}}{p_N|_{\theta}(n|x)f_{\theta}(x)}}{p_N(n)}$$

normalization constant

- Prior belief of parameter θ
- Likelihood of N = n heads, given parameter $\theta = x$.
- Posterior updated belief of parameter θ .

 $f_{\theta}(x)$ $p_{N|\theta}(n|x)$ $f_{\theta|N}(x|n)$

More in live lecture! Stanford University 19

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21c_beta

Beta RV

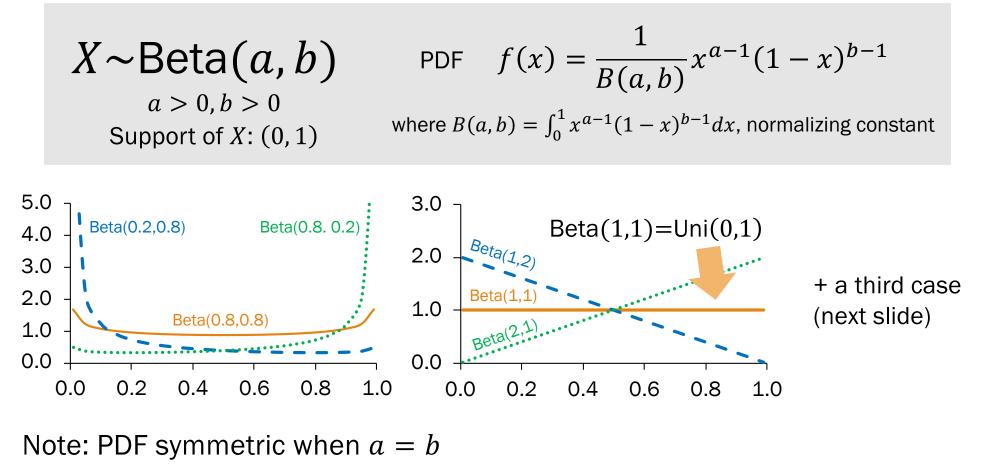
Beta random variable

<u>def</u> A Beta random variable *X* is defined as follows:

$$\begin{aligned} X \sim \text{Beta}(a, b) \\ a > 0, b > 0 \\ \text{Support of } X: (0, 1) \end{aligned} \qquad \text{PDF} \quad f(x) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} \\ \text{where } B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx, \text{ normalizing constant} \end{aligned}$$
$$\qquad \text{Where } B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx, \text{ normalizing constant} \end{aligned}$$
$$\qquad \text{Variance } Var(X) = \frac{ab}{(a+b)^2(a+b+1)}$$

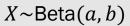
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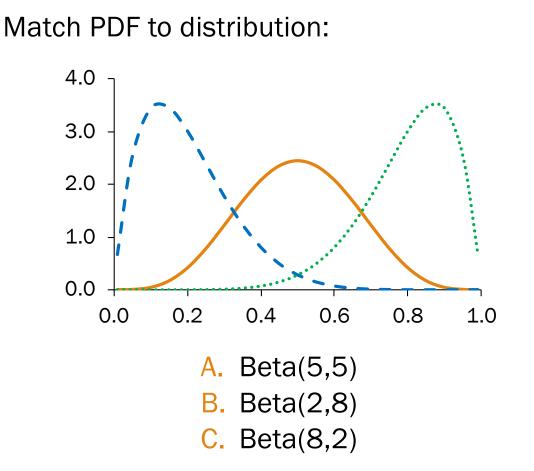
Beta RV with different *a*, *b*

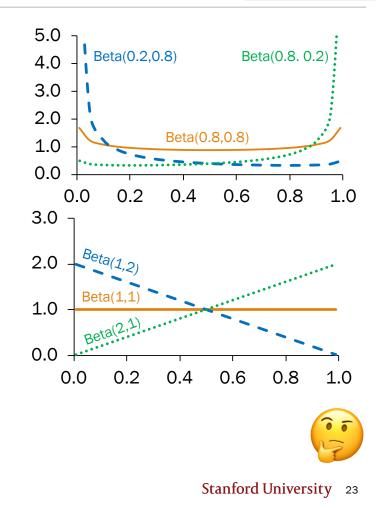


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Beta RV with different *a*, *b*



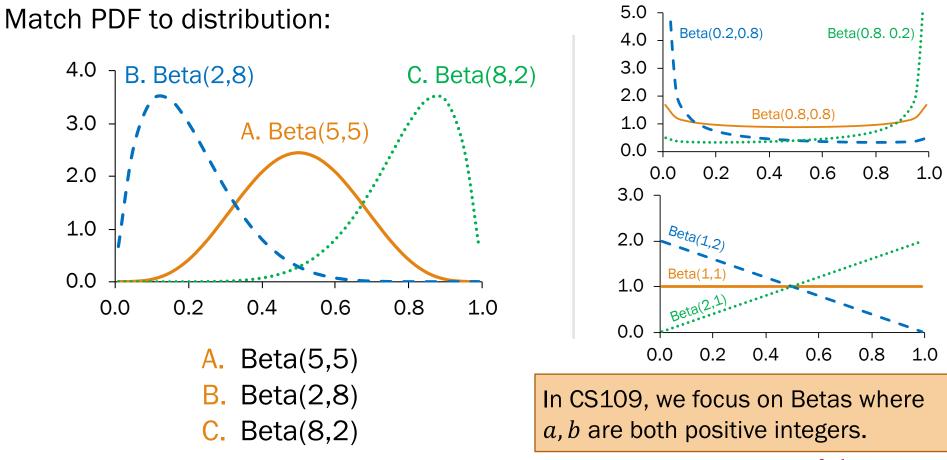




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Beta RV with different *a*, *b*





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Beta random variable

<u>def</u> A Beta random variable *X* is defined as follows:

$$X \sim \text{Beta}(a, b)$$

$$a > 0, b > 0$$

$$\text{PDF} \quad f(x) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1}$$

where $B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$, normalizing constant

$$\text{Expectation} \quad E[X] = \frac{a}{a+b}$$

$$\text{Variance} \quad \text{Var}(X) = \frac{ab}{(a+b)^2(a+b+1)}$$

Beta can be a distribution of probabilities.

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Beta can be a distribution of probabilities.

4.0 3.0 2.0 1.0 0.0 0.2 0.4 0.6 0.8 1.0Beta(8,2) Beta(8,2) 0.8 0.81.0

Beta parameters a, b <u>could</u> come from an experiment...

But which one? Stay tuned...

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 $X \sim \text{Beta}(a, b)$



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Flipping a coin with unknown probability

A new definition of probability

Review

Flip a coin n + m times, comes up with n heads. We don't know the probability θ that the coin comes up heads.



 $\boldsymbol{\theta}$ is a single value.

 $\theta = \lim_{n+m \to \infty} \frac{n}{n+m} \approx \frac{n}{n+m}$



The world's first coin

Bayesian

 θ is a random variable.

 θ 's continuous support: (0, 1)

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Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

- Before our experiment, θ (the probability that the coin comes up heads) can be any probability.
- Let *N* = number of heads.
- Given $\theta = x$, coin flips are independent.

What is our updated belief of θ after we observe N = n?

What are reasonable distributions of the following?

- **1.** *θ*
- 2. $N|\theta = x$
- $3. \quad \theta | N = n$



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Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

- Before our experiment, θ (the probability that the coin comes up heads) can be any probability.
- Let *N* = number of heads.
- Given $\theta = x$, coin flips are independent.

What is our updated belief of θ after we observe N = n?

What are reasonable distributions of the following?

- **1.** θ Bayesian prior $\theta \sim \text{Uni}(0,1)$
- 2. $N|\theta = x$ Likelihood $N|\theta = x \sim Bin(n + m, x)$
- **3.** $\theta | N = n$ Bayesian posterior. Use Bayes'!

Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

- Before our experiment, θ (the probability that the coin comes up heads) can be any probability.
- Let N = number of heads.
- Given $\theta = x$, coin flips are independent.

What is our updated belief of θ after we observe N = n?

$$\theta \sim \text{Uni}(0,1)$$

Likelihood:

$$N|\theta = x \sim Bin(n + m, x)$$

? Posterior: $f_{\theta|N}(\theta|n)$

$$\begin{aligned} f_{\theta|N}(x|n) &= \frac{p_{N|\theta}(n|x)f_{\theta}(x)}{p_{N}(n)} = \frac{\binom{n+m}{n}x^{n}(1-x)^{m} \cdot 1}{p_{N}(n)} \\ &= \frac{\binom{n+m}{n}}{p_{N}(n)}x^{n}(1-x)^{m} = \frac{1}{c}x^{n}(1-x)^{m}, \text{ where } c = \int_{0}^{1}x^{n}(1-x)^{m}dx \end{aligned}$$

constant with respect to x, doesn't depend on x

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Let's try it out

1. Start with a $\theta \sim \text{Uni}(0,1)$ over probability that a coin lands heads.

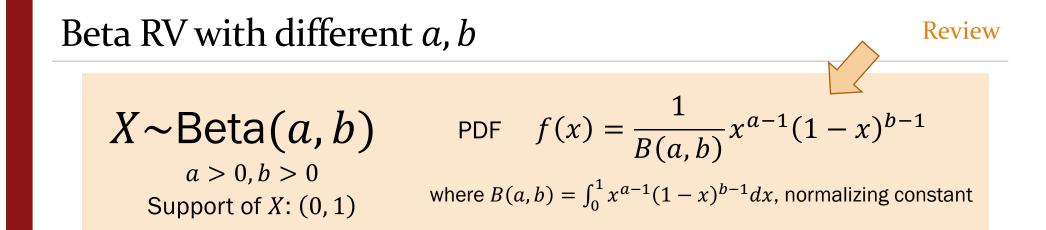
- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability θ ?

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

c normalizes to valid PDF

Wait a minute! #tbplv

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$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$
 is the PDF for Beta(8,2)!

c normalizes to valid PDF

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Let's try it out

1. Start with a $\theta \sim \text{Uni}(0,1)$ over probability that a coin lands heads.

- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability θ ?

 $\begin{array}{c} \begin{array}{c} \begin{array}{c} \\ \\ \\ \\ \\ \end{array} \end{array} \begin{array}{c} 2.0 \\ 1.0 \\ 0.0 \\ 0.0 \\ 0.0 \\ 0.2 \\ 0.4 \\ 0.6 \\ 0.8 \\ 1.0 \end{array}$

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

4.0

3.0

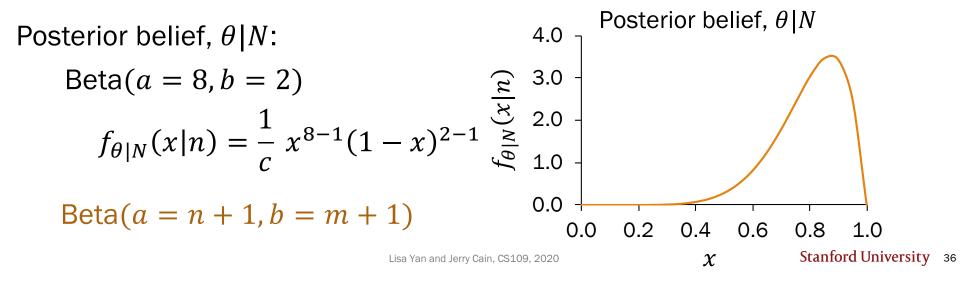
c normalizes to valid PDF

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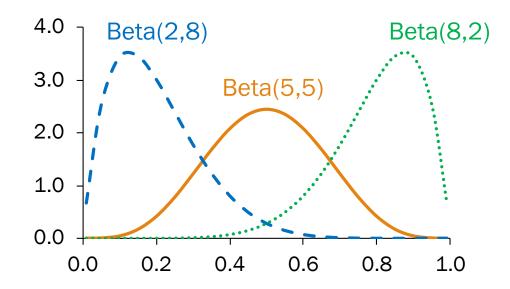
3. What is our posterior belief of the probability θ ?

- Start with a $\theta \sim \text{Uni}(0,1)$ over probability
- Observe n = 7 successes and m = 1 failures
- Your new belief about the probability of θ is:

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$
, where $c = \int_0^1 x^7 (1-x)^1 dx$



CS109 focus: Beta where a, b both positive integers $X \sim Beta(a, b)$



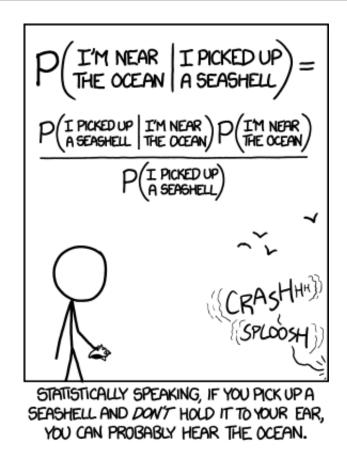
If *a*, *b* are positive integers, Beta parameters *a*, *b* could come from an experiment:

$$a =$$
 "successes" + 1
 $b =$ "failures" + 1

- Beta (in CS109) models the randomness of the probability of experiment success.
- Beta parameters depend our data and our prior.

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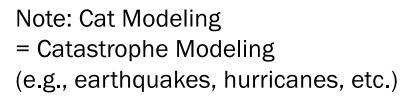
Bayes' on the waves



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Interesting probability news

Why Rejection Sampling Is Useful in Cat Modeling



4b. Compute 2. Scale q(x) by 4a. Compute p(x) anc q(x) : a ы 5a. Dra fro Unif(0,C x2 🛀 3b. Another draw x2 from q(x) x1 3a. Draw x1 🗖 K 6a. Accept x1 X from q(x)6b. Reject x2 since $u1 \le p(x1)$ since u2>p(x2)

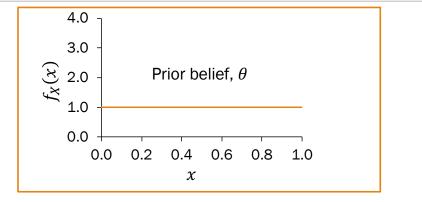
https://www.air-worldwide.com/blog/posts/2018/9/why-rejection-sampling-is-useful-in-cat-modeling/

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Conjugate distributions

A note about our prior

1. Start with a $\theta \sim \text{Uni}(0,1)$ over probability that a coin lands heads.



- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability θ ?

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$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

 \boldsymbol{c} normalizes to valid PDF

Wait another minute!

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okay

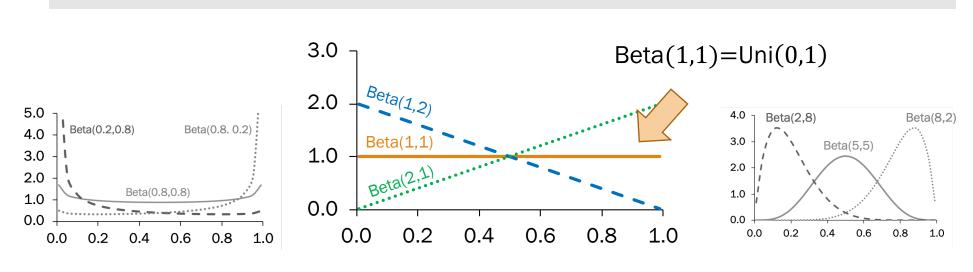
Beta RV with different *a*, *b*

 $X \sim \text{Beta}(a, b)$

a > 0, b > 0

Support of *X*: (0, 1)

Review



PDF $f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$

where $B(a,b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$, normalizing constant

Note: PDF symmetric when a = b

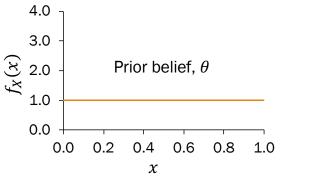
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A note about our prior

1. Start with a $\theta \sim \text{Uni}(0,1)$ over probability that a coin lands heads.

Beta(1,1)

- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability θ ?



Check this out. Beta(a = 1, b = 1):

$$f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$$
$$= \frac{1}{\int_0^1 1 dx}$$

 $= 1 \qquad \text{where } 0 < x < 1$



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Beta is a conjugate distribution for Bernoulli

Beta is a **conjugate distribution** for Bernoulli, meaning:

• Prior and posterior parametric forms are the same

(proof on next slide)

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Beta is a conjugate distribution for Bernoulli

Beta is a conjugate distribution for Bernoulli, meaning:

- 1. If our prior belief of the parameter is Beta, and
- 2. Our experiment is Bernoulli, then
- 3. Our posterior is also Beta.

Proof: $\theta \sim \text{Beta}(a, b)$ $N | \theta \sim \text{Bin}(n + m, x)$ $f_{\theta|N}(x|n) = \frac{p_{N|\theta}(n|x)f_{\theta}(x)}{p_{N}(n)} = \frac{\binom{n+m}{m}x^{n}(1-x)^{m} \cdot \frac{1}{B(a,b)}x^{a-1}(1-x)^{b-1}}{p_{N}(n)}$ $\stackrel{\text{constants that}}{= C \cdot x^{n}(1-x)^{m} \cdot x^{a-1}(1-x)^{b-1}} = C \cdot x^{n+a-1}(1-x)^{m+b-1} \checkmark$ $= C \cdot x^{n+a-1}(1-x)^{m+b-1} \checkmark$ Example 2 Stanford University 45

(observe *n* successes, *m* failures)

Beta is a conjugate distribution for Bernoulli

This is the main takeaway of Beta.

Beta is a **conjugate distribution** for Bernoulli, meaning:

- Prior and posterior parametric forms are the same
- Practically, conjugate means easy update: Add number of "heads" and "tails" seen to Beta parameters.

You can set the prior to reflect how biased you think the coin is a priori:

- $\theta \sim \text{Beta}(a, b)$: have seen (a + b 2) imaginary trials, where (a 1) are heads, (b 1) tails
- Then Beta(1, 1) = Uni(0, 1) means we haven't seen any imaginary trials

Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Experiment Observe *n* successes and *m* failures Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

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The enchanted die

 $Beta(a = n_{imag} + 1, b = m_{imag} + 1)$ Prior Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

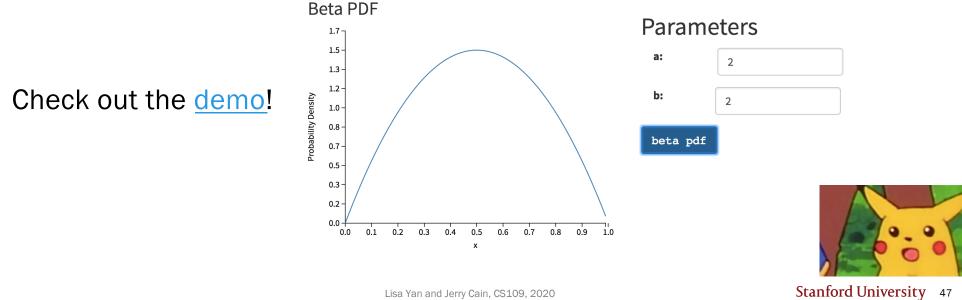
Let θ be the probability of rolling a 6 on Lisa's die.

Prior: Imagine 1 out of 6 die rolls where only 6 showed up •



Observation: roll it a few times...

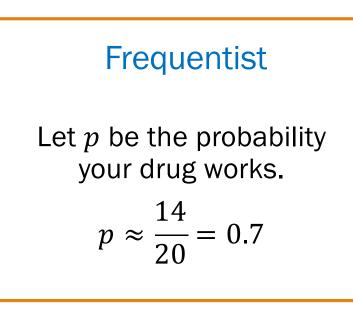
What is the updated distribution of θ after our observation?



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- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?



Bayesian

A frequentist view will not incorporate prior/expert belief about probability.

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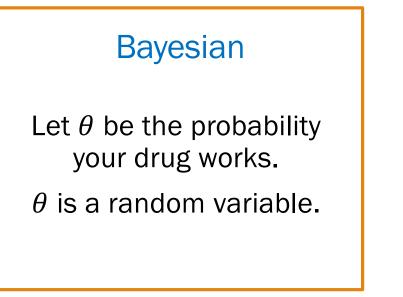
- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

Frequentist

Let *p* be the probability your drug works.

$$p \approx \frac{14}{20} = 0.7$$



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Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

(Bayesian interpretation)

What is the prior distribution of θ ? (select all that apply)

- A. $\theta \sim \text{Beta}(1, 1) = \text{Uni}(0, 1)$
- B. $\theta \sim \text{Beta}(81, 101)$
- C. $\theta \sim \text{Beta}(80, 20)$
- D. $\theta \sim \text{Beta}(81, 21)$
- E. $\theta \sim \text{Beta}(5,2)$



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 $Beta(a = n_{imag} + 1, b = m_{imag} + 1)$ Prior Posterior $Beta(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

(Bayesian interpretation)

What is the prior distribution of θ ? (select all that apply)

- $\theta \sim \text{Beta}(1,1) = \text{Uni}(0,1)$ Α.
- B. $\theta \sim \text{Beta}(81, 101)$
- C. $\theta \sim \text{Beta}(80, 20)$

 $\theta \sim \text{Beta}(81, 21)$ Interpretation: 80 successes / 100 imaginary trials

 $\theta \sim \text{Beta}(5,2)$

(you can choose either based on how strong your belief is (an engineering choice). We choose E on next slide)

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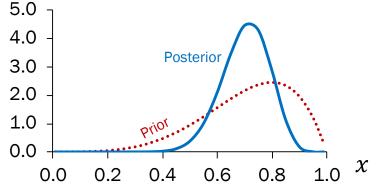
Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

- Prior: $\theta \sim \text{Beta}(a = 5, b = 2)$
- Posterior: $\theta \sim \text{Beta}(a = 5 + 14, b = 2 + 6)$ ~Beta(a = 19, b = 8)

(Bayesian interpretation)



Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

- Before being tested, a medicine is believed to "work" 80% of the time.
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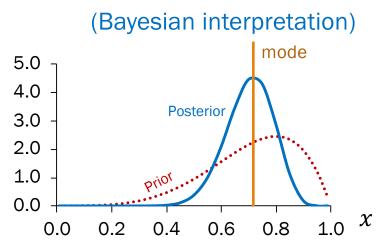
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Posterior: $\theta \sim \text{Beta}(a = 5 + 14, b = 2 + 6)$ ~Beta(a = 19, b = 8)

What do you report to pharmacists?

- A. Expectation of posterior
- B. Mode of posterior
- C. Distribution of posterior
- D. Nothing





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Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$

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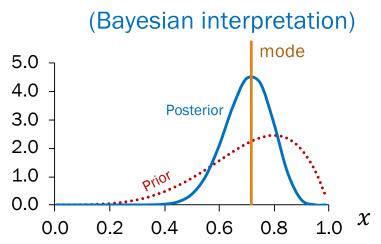
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Posterior: $\theta \sim \text{Beta}(a = 5 + 14, b = 2 + 6)$ ~Beta(a = 19, b = 8)

What do you report to pharmacists?

$$E[\theta] = \frac{a}{a+b} = \frac{19}{19+8} \approx 0.70$$
$$mode(\theta) = \frac{a-1}{a+b-2} = \frac{18}{18+7} \approx 0.72$$



In CS109, we report the **mode**: The "most likely" parameter given the data.

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Food for thought

In this lecture:



If we don't know the parameter p, Bayesian statisticians will:

- Treat the parameter as a random variable θ with a Beta prior distribution
- Perform an experiment
- Based on experiment outcomes, update the posterior distribution of θ



Food for thought:

Any parameter for a "parameterized" random variable can be thought of as a random variable.

 $Y \sim \mathcal{N}(\mu, \sigma^2)$

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Estimating our parameter directly

(our focus so far)

Maximum Likelihood Estimator (MLE) What is the parameter θ that maximizes the likelihood of our observed data (x_1, x_2, \dots, x_n) ?

$$L(\theta) = f(X_1, X_2, \dots, X_n | \theta)$$
$$= \prod_{i=1}^n f(X_i | \theta)$$
$$= \arg \max f(X_i | X_i = X_i | \theta)$$

 $\theta_{MLE} = \arg \max_{\theta} f(X_1, X_2, \dots, X_n | \theta)$ likelihood of data

Observations:

- MLE maximizes probability of observing data given a parameter θ .
- If we are estimating θ , shouldn't we maximize the probability of θ directly?

See you next time!

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(extra)

Extra: MLE: Multinomial derivation

Okay, just one more MLE with the Multinomial

Consider a sample of *n* i.i.d. random variables where

- Each element is drawn from one of *m* outcomes. $P(\text{outcome } i) = p_i$, where $\sum_{i=1}^m p_i = 1$
- $X_i = \#$ of trials with outcome *i*, where $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing $L(\theta) = \frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$ the sample(X_1, X_2, \dots, X_m), given the probabilities p_1, p_2, \dots, p_m ?
- 2. What is θ_{MLE} ?

extra slides

$$LL(\theta) = \log(n!) - \sum_{i=1}^{m} \log(X_i!) + \sum_{i=1}^{m} X_i \log(p_i), \text{ such that } \sum_{i=1}^{m} p_i = 1$$

Optimize with
Lagrange multipliers in θ_{MLE} : $p_i = \frac{X_i}{n}$ Intuitively, probability
 $p_i = \text{proportion of outcomes}$

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n

Optimizing MLE for Multinomial

$$\begin{array}{ll} \theta = (p_1, p_2, \dots, p_m) \\ \theta_{MLE} = \arg\max_{\theta} LL(\theta), \text{ where } \sum_{i=1}^m p_i = 1 & \text{Use Lagrange multipliers} \\ \text{to account for constraint} \\ \end{array}$$

$$\begin{array}{ll} \text{Lagrange} \\ \text{multipliers: } & A(\theta) = LL(\theta) + \lambda \left(\sum_{i=1}^m p_i - 1\right) = \sum_{i=1}^m X_i \log(p_i) + \lambda \left(\sum_{i=1}^m p_i - 1\right) \begin{array}{l} (\text{drop non-}p_i \\ \text{non-}p_i \\ \text{terms}) \end{array}$$

$$\begin{array}{ll} \text{Differentiate w.r.t.} \\ \text{each } p_i, \text{ in turn: } & \frac{\partial A(\theta)}{\partial p_i} = X_i \frac{1}{p_i} + \lambda = 0 \Rightarrow p_i = -\frac{X_i}{\lambda} \\ \text{Solve for } \lambda, \text{ noting } \\ \sum_{i=1}^m X_i = n, \sum_{i=1}^m p_i = 1: \\ \sum_{i=1}^m p_i = \sum_{i=1}^m -\frac{X_i}{\lambda} = 1 \quad \Rightarrow 1 = -\frac{n}{\lambda} \quad \Rightarrow \lambda = -n \\ \text{Substitute } \lambda \text{ into } p_i \quad p_i = \frac{X_i}{n} \end{array}$$

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