# 21: Beta

Lisa Yan and Jerry Cain October 30, 2020

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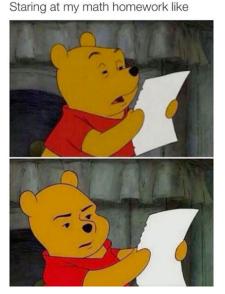
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21a\_mle\_multinomial

# MLE: Multinomial

Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes.  $P(\text{outcome } i) = p_i$ , where  $\sum_{i=1}^{m} p_i = 1$
- $X_i = #$  of trials with outcome *i*, where  $\sum_{i=1}^m X_i = n$



Let's give an example!

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Consider a sample of n i.i.d. random variables where

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Example: Suppose each RV is outcome of 6-sided die.

$$m = 6, \sum_{i=1}^{6} p_i = 1$$

- Roll the dice n = 12 times.
- Observe data: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

$$X_1 = 3, X_2 = 2, X_3 = 0,$$
  
 $X_4 = 3, X_5 = 1, X_6 = 3$ 

Check: 
$$X_1 + X_2 + \dots + X_6 = 12$$

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes.  $P(\text{outcome } i) = p_i$ , where  $\sum_{i=1}^m p_i = 1$
- $X_i = #$  of trials with outcome *i*, where  $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing the sample( $X_1, X_2, ..., X_m$ ), given the probabilities  $p_1, p_2, ..., p_m$ ?

A. 
$$\frac{n!}{X_{1}! X_{2}! \cdots X_{m}!} p_{1}^{X_{1}} p_{2}^{X_{2}} \cdots p_{m}^{X_{m}}$$
  
B. 
$$p_{1}^{X_{1}} p_{2}^{X_{2}} \cdots p_{m}^{X_{m}}$$
  
C. 
$$\frac{n!}{X_{1}! X_{2}! \cdots X_{m}!} X_{1}^{p_{1}} X_{2}^{p_{2}} \cdots X_{m}^{p_{m}}$$

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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes.  $P(\text{outcome } i) = p_i$ , where  $\sum_{i=1}^m p_i = 1$
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(A) 
$$\frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$$
  
B. 
$$p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$$
  
C. 
$$\frac{n!}{X_1! X_2! \cdots X_m!} X_1^{p_1} X_2^{p_2} \cdots X_m^{p_m}$$
  
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Consider a sample of n i.i.d. random variables where

- Each element is drawn from one of m outcomes.  $P(\text{outcome } i) = p_i$ , where  $\sum_{i=1}^m p_i = 1$
- $X_i = #$  of trials with outcome *i*, where  $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing the sample( $X_1, X_2, ..., X_m$ ), given the probabilities  $p_1, p_2, ..., p_m$ ?  $L(\theta) = \frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$
- 2. What is  $\theta_{MLE}$ ?

$$LL(\theta) = \log(n!) - \sum_{i=1}^{m} \log(X_i!) + \sum_{i=1}^{m} X_i \log(p_i)$$
, such that  $\sum_{i=1}^{m} p_i = 1$ 

Optimize with Lagrange multipliers in extra slides

$$\rightarrow \theta_{MLE}: p_i = \frac{X_i}{n}$$
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Intuitively, probability  $p_i = \text{proportion of outcomes}$ <u>Stanford University</u> 8

#### When MLEs attack!

MLE for  $p_i = \frac{X_i}{n}$ 

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

What is  $\theta_{MLE}$ ?



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#### When MLEs attack!

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

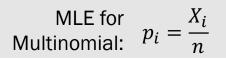
 $\theta_{MLE}$ :

$$p_1 = 3/12$$
  
 $p_2 = 2/12$   
 $p_3 = 0/12$   
 $p_4 = 3/12$   
 $p_5 = 1/12$   
 $p_6 = 3/12$ 

- MLE: you'll never...<u>EVER</u>... roll a three.
- Do you really believe that?

Today: A new definition of probability!

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21b\_bayesian

# Bayesian Statistics

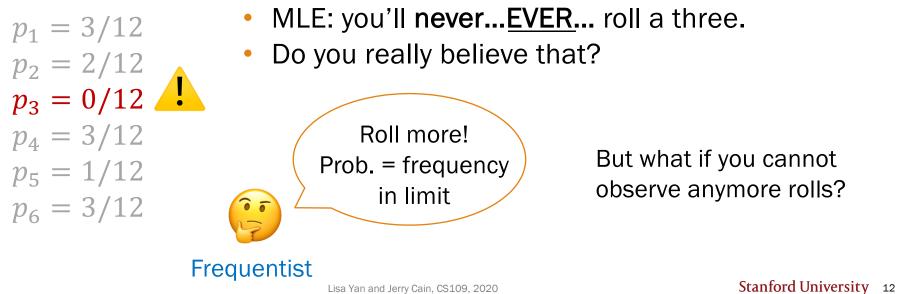
## When MLEs attack!

**Review** 

Consider a 6-sided die.

- Roll the dice n = 12 times.
- Observe: 3 ones, 2 twos, 0 threes, 3 fours, 1 fives, 3 sixes

 $\theta_{MLE}$ :



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#### Today's plan

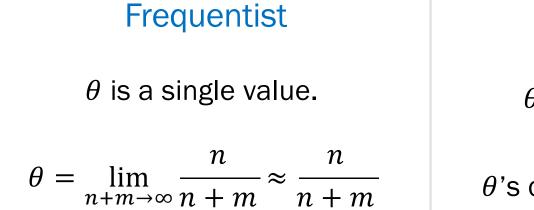
# Today we are going to learn something unintuitive, beautiful, and useful!

# We are going to think of probabilities as random variables.

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# A new definition of probability

Flip a coin n + m times, come up with n heads. We don't know the probability  $\theta$  that the coin comes up heads.





```
The world's first coin
```

Bayesian

 $\theta$  is a random variable.

 $\theta$ 's continuous support: (0, 1)

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### Let's play a game

Roll 2 dice. If *neither* roll is a 6, you win (event W). Else, I win (event  $W^C$ ).



- Before you play, what's the probability that you win?
- Play once. What's the probability that you win?
- Play three more times. What's the probability that you win?



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Bayesian probability

Bayesian statistics: Probability is a reasonable expectation representing a state of knowledge.

Mixing discrete and continuous random variables, combined with Bayes' Theorem, allows us to reason about probabilities as random variables.

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### Mixing discrete and continuous

Let X be a continuous random variable, and N be a discrete random variable.

Bayes' Theorem:

$$f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$$

Intuition: 
$$P(X = x | N = n) = \frac{P(N = n | X = x)P(X = x)}{P(N = n)}$$
$$f_{X|N}(x|n)\varepsilon_X = \frac{P(N = n | X = x)f_X(x)\varepsilon_X}{P(N = n)} \implies f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$$

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#### All your Bayes are belong to us

Let *X*, *Y* be continuous and *M*, *N* be discrete random variables.

OG Bayes:
$$p_{M|N}(m|n) = \frac{p_{N|M}(n|m)p_M(m)}{p_N(n)}$$
Mix Bayes #1: $f_{X|N}(x|n) = \frac{p_{N|X}(n|x)f_X(x)}{p_N(n)}$ Mix Bayes #2: $p_{N|X}(n|x) = \frac{f_{X|N}(x|n)p_N(n)}{f_X(x)}$ All continuous: $f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)}$ 



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# Mixing discrete and continuous

Let  $\theta$  be a random variable for the probability your coin comes up heads, and N be the number of heads you observe in an experiment.

posterior  

$$f_{\theta|N}(x|n) = \frac{\substack{\text{likelihood prior}}{p_N|_{\theta}(n|x)f_{\theta}(x)}}{p_N(n)}$$

normalization constant

- Prior belief of parameter  $\theta$
- Likelihood of N = n heads, given parameter  $\theta = x$ .
- Posterior updated belief of parameter  $\theta$ .

 $f_{\theta}(x)$  $p_{N|\theta}(n|x)$  $f_{\theta|N}(x|n)$ 

More in live lecture! Stanford University 19

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21c\_beta

# Beta RV

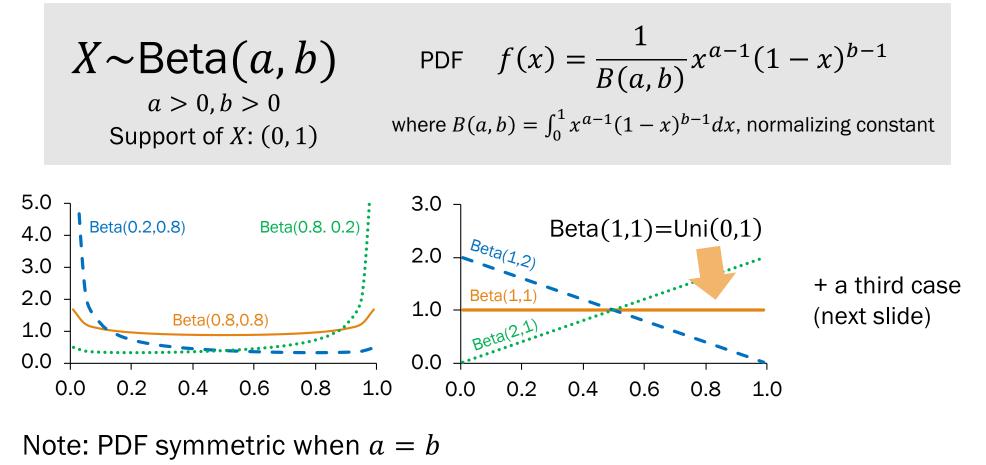
#### Beta random variable

<u>def</u> A Beta random variable *X* is defined as follows:

$$\begin{aligned} X \sim \text{Beta}(a, b) \\ a > 0, b > 0 \\ \text{Support of } X: (0, 1) \end{aligned} \qquad \text{PDF} \quad f(x) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} \\ \text{where } B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx, \text{ normalizing constant} \end{aligned}$$
$$\qquad \text{Where } B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx, \text{ normalizing constant} \end{aligned}$$
$$\qquad \text{Variance } Var(X) = \frac{ab}{(a+b)^2(a+b+1)}$$

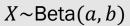
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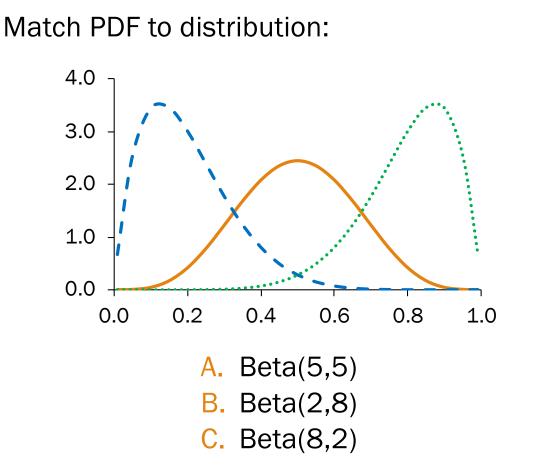
#### Beta RV with different *a*, *b*

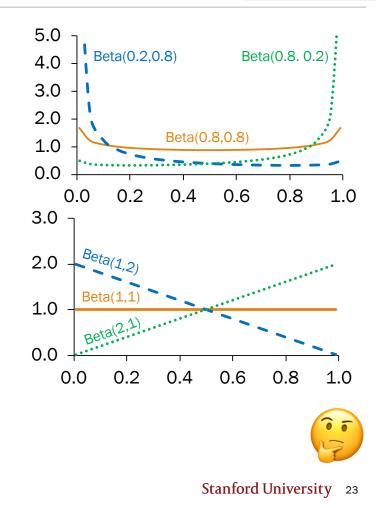


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#### Beta RV with different *a*, *b*



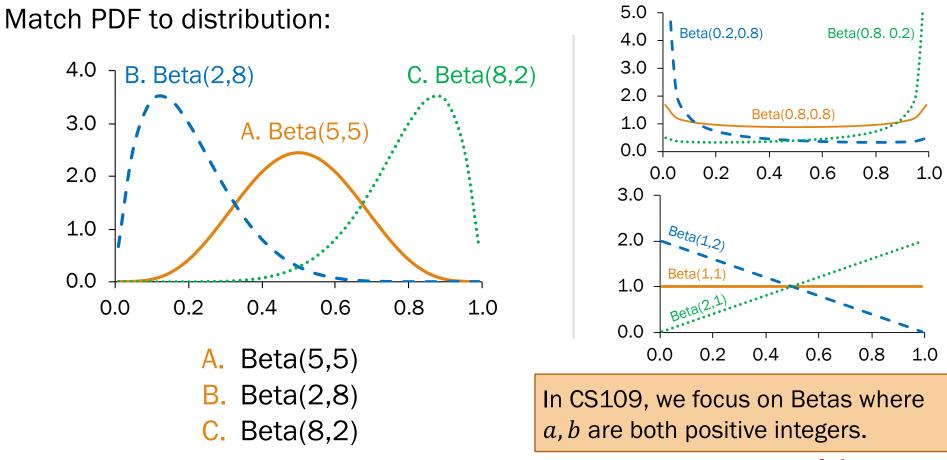




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#### Beta RV with different *a*, *b*





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#### Beta random variable

<u>def</u> A Beta random variable *X* is defined as follows:

$$X \sim \text{Beta}(a, b)$$
  

$$a > 0, b > 0$$

$$\text{PDF} \quad f(x) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1}$$
  
where  $B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$ , normalizing constant  

$$\text{Expectation} \quad E[X] = \frac{a}{a+b}$$

$$\text{Variance} \quad \text{Var}(X) = \frac{ab}{(a+b)^2(a+b+1)}$$

#### Beta can be a distribution of probabilities.

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## Beta can be a distribution of probabilities.

4.0 3.0 2.0 1.0 0.0 0.2 0.4 0.6 0.8 1.0Beta(8,2) Beta(8,2) 0.8 0.81.0

Beta parameters a, b <u>could</u> come from an experiment...

But which one? Stay tuned...

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 $X \sim \text{Beta}(a, b)$ 



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# Flipping a coin with unknown probability

# A new definition of probability

#### Review

Flip a coin n + m times, comes up with n heads. We don't know the probability  $\theta$  that the coin comes up heads.



 $\boldsymbol{\theta}$  is a single value.

 $\theta = \lim_{n+m \to \infty} \frac{n}{n+m} \approx \frac{n}{n+m}$ 



The world's first coin

Bayesian

 $\theta$  is a random variable.

 $\theta$ 's continuous support: (0, 1)

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# Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

- Before our experiment,  $\theta$  (the probability that the coin comes up heads) can be any probability.
- Let *N* = number of heads.
- Given  $\theta = x$ , coin flips are independent.

What is our updated belief of  $\theta$  after we observe N = n?

#### What are reasonable distributions of the following?

- **1.** *θ*
- 2.  $N|\theta = x$
- $3. \quad \theta | N = n$



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# Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

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What is our updated belief of  $\theta$  after we observe N = n?

#### What are reasonable distributions of the following?

- **1.**  $\theta$  Bayesian prior  $\theta \sim \text{Uni}(0,1)$
- 2.  $N|\theta = x$  Likelihood  $N|\theta = x \sim Bin(n + m, x)$
- **3.**  $\theta | N = n$  Bayesian posterior. Use Bayes'!

## Flip a coin with unknown probability

Flip a coin n + m times, observe n heads.

- Before our experiment,  $\theta$  (the probability that the coin comes up heads) can be any probability.
- Let N = number of heads.
- Given  $\theta = x$ , coin flips are independent.

What is our updated belief of  $\theta$  after we observe N = n?

$$\theta \sim \text{Uni}(0,1)$$

Likelihood:  

$$N|\theta = x \sim Bin(n + m, x)$$
  
? Posterior:  $f_{\theta|N}(\theta|n)$ 

$$\begin{aligned} f_{\theta|N}(x|n) &= \frac{p_{N|\theta}(n|x)f_{\theta}(x)}{p_{N}(n)} = \frac{\binom{n+m}{n}x^{n}(1-x)^{m} \cdot 1}{p_{N}(n)} \\ &= \frac{\binom{n+m}{n}}{p_{N}(n)}x^{n}(1-x)^{m} = \frac{1}{c}x^{n}(1-x)^{m}, \text{ where } c = \int_{0}^{1}x^{n}(1-x)^{m}dx \end{aligned}$$

constant with respect to x, doesn't depend on x

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#### Let's try it out

**1.** Start with a  $\theta \sim \text{Uni}(0,1)$  over probability that a coin lands heads.

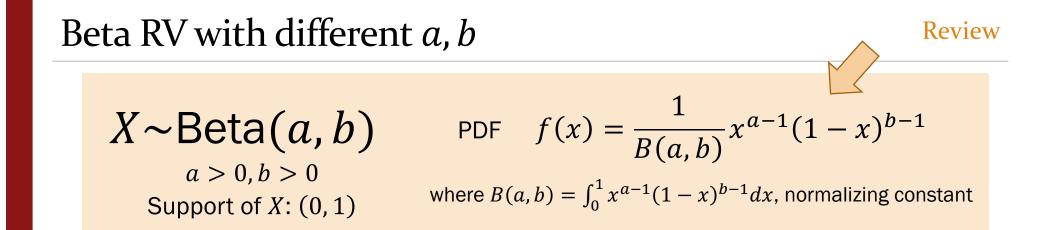
- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability  $\theta$ ?

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

c normalizes to valid PDF

#### Wait a minute! #tbplv

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$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$
 is the PDF for Beta(8,2)!

c normalizes to valid PDF

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#### Let's try it out

1. Start with a  $\theta \sim \text{Uni}(0,1)$  over probability that a coin lands heads.

- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability  $\theta$ ?

 $\begin{array}{c} \begin{array}{c} \begin{array}{c} \\ \\ \\ \\ \\ \end{array} \end{array} \begin{array}{c} 2.0 \\ 1.0 \\ 0.0 \\ 0.0 \\ 0.0 \\ 0.2 \\ 0.4 \\ 0.6 \\ 0.8 \\ 1.0 \end{array}$ 

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

4.0

3.0

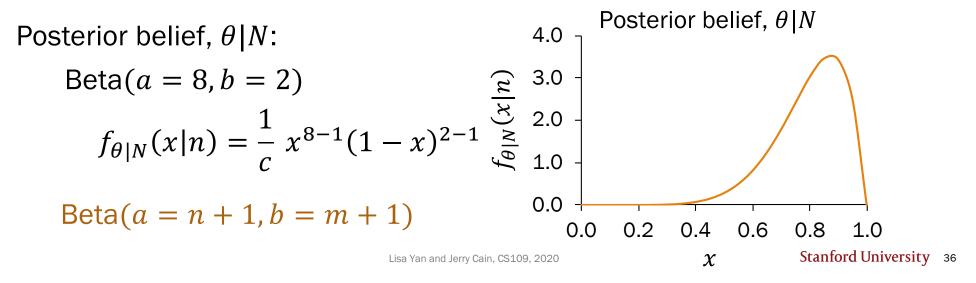
c normalizes to valid PDF

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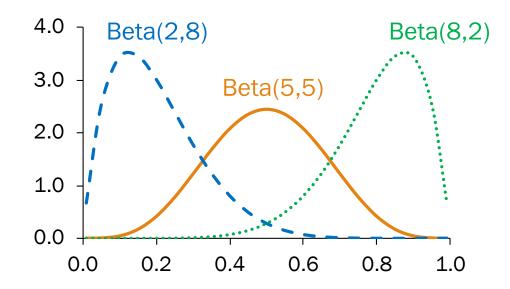
### **3**. What is our posterior belief of the probability $\theta$ ?

- Start with a  $\theta \sim \text{Uni}(0,1)$  over probability
- Observe n = 7 successes and m = 1 failures
- Your new belief about the probability of  $\theta$  is:

$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$
, where  $c = \int_0^1 x^7 (1-x)^1 dx$ 



#### CS109 focus: Beta where a, b both positive integers $X \sim Beta(a, b)$



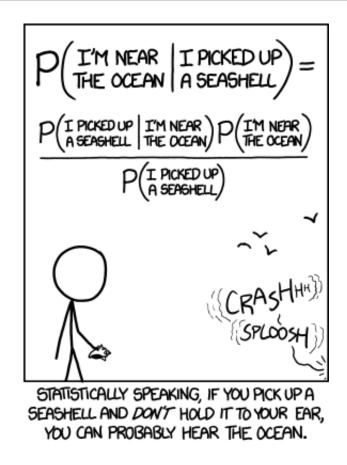
If *a*, *b* are positive integers, Beta parameters *a*, *b* could come from an experiment:

$$a =$$
 "successes" + 1  
 $b =$  "failures" + 1

- Beta (in CS109) models the randomness of the probability of experiment success.
- Beta parameters depend our data and our prior.

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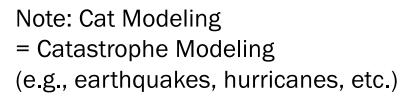
#### Bayes' on the waves



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#### Interesting probability news

Why Rejection Sampling Is Useful in Cat Modeling



4b. Compute 2. Scale q(x) by 4a. Compute p(x) anc q(x) : a ы 5a. Dra fro Unif(0,C x2 🛀 3b. Another draw x2 from q(x) x1 3a. Draw x1 🗖 K 6a. Accept x1 X from q(x)6b. Reject x2 since  $u1 \le p(x1)$ since u2>p(x2)

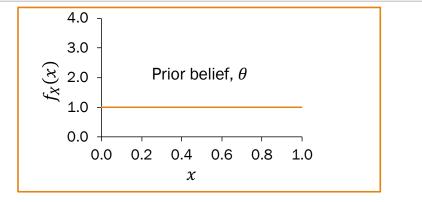
https://www.air-worldwide.com/blog/posts/2018/9/why-rejection-sampling-is-useful-in-cat-modeling/

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# Conjugate distributions

#### A note about our prior

1. Start with a  $\theta \sim \text{Uni}(0,1)$  over probability that a coin lands heads.



- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability  $\theta$ ?

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$$f_{\theta|N}(x|n) = \frac{1}{c} x^7 (1-x)^1$$

 $\boldsymbol{c}$  normalizes to valid PDF

Wait another minute!

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okay

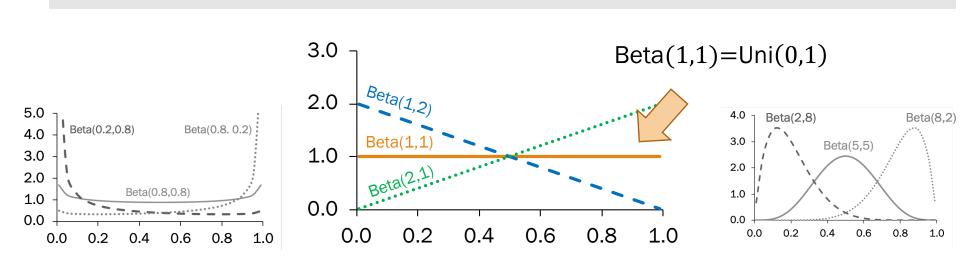
#### Beta RV with different *a*, *b*

 $X \sim \text{Beta}(a, b)$ 

a > 0, b > 0

Support of *X*: (0, 1)

#### Review



PDF  $f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$ 

where  $B(a,b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$ , normalizing constant

Note: PDF symmetric when a = b

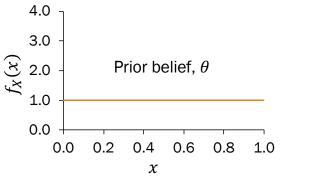
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#### A note about our prior

1. Start with a  $\theta \sim \text{Uni}(0,1)$  over probability that a coin lands heads.

Beta(1,1)

- 2. Flip a coin 8 times. Observe n = 7 heads and m = 1 tail
- 3. What is our posterior belief of the probability  $\theta$ ?



Check this out. Beta(a = 1, b = 1):

$$f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1}$$
$$= \frac{1}{\int_0^1 1 dx}$$

 $= 1 \qquad \text{where } 0 < x < 1$ 



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## Beta is a conjugate distribution for Bernoulli

Beta is a **conjugate distribution** for Bernoulli, meaning:

• Prior and posterior parametric forms are the same

(proof on next slide)

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#### Beta is a conjugate distribution for Bernoulli

Beta is a conjugate distribution for Bernoulli, meaning:

- 1. If our prior belief of the parameter is Beta, and
- 2. Our experiment is Bernoulli, then
- 3. Our posterior is also Beta.

Proof:  $\theta \sim \text{Beta}(a, b)$   $N | \theta \sim \text{Bin}(n + m, x)$  $f_{\theta|N}(x|n) = \frac{p_{N|\theta}(n|x)f_{\theta}(x)}{p_{N}(n)} = \frac{\binom{n+m}{m}x^{n}(1-x)^{m} \cdot \frac{1}{B(a,b)}x^{a-1}(1-x)^{b-1}}{p_{N}(n)}$   $\stackrel{\text{constants that}}{= C \cdot x^{n}(1-x)^{m} \cdot x^{a-1}(1-x)^{b-1}} = C \cdot x^{n+a-1}(1-x)^{m+b-1} \checkmark$   $= C \cdot x^{n+a-1}(1-x)^{m+b-1} \checkmark$  Example 2 Stanford University 45

(observe *n* successes, *m* failures)

## Beta is a conjugate distribution for Bernoulli

This is the main takeaway of Beta.

Beta is a **conjugate distribution** for Bernoulli, meaning:

- Prior and posterior parametric forms are the same
- Practically, conjugate means easy update: Add number of "heads" and "tails" seen to Beta parameters.

You can set the prior to reflect how biased you think the coin is a priori:

- $\theta \sim \text{Beta}(a, b)$ : have seen (a + b 2) imaginary trials, where (a 1) are heads, (b 1) tails
- Then Beta(1, 1) = Uni(0, 1) means we haven't seen any imaginary trials

Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Experiment Observe *n* successes and *m* failures Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$ 

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## The enchanted die

 $Beta(a = n_{imag} + 1, b = m_{imag} + 1)$ Prior Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$ 

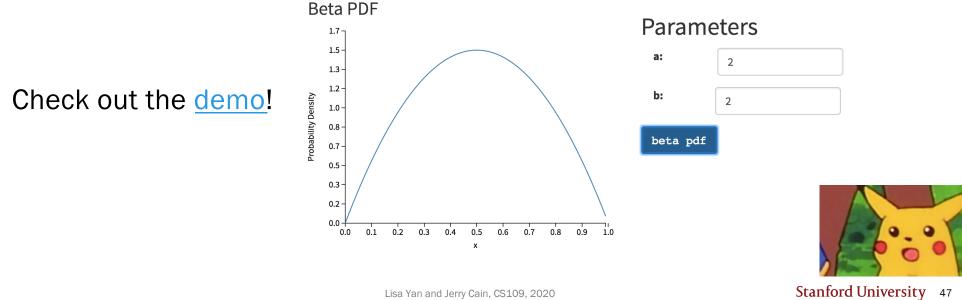
Let  $\theta$  be the probability of rolling a 6 on Lisa's die.

Prior: Imagine 1 out of 6 die rolls where only 6 showed up •



Observation: roll it a few times...

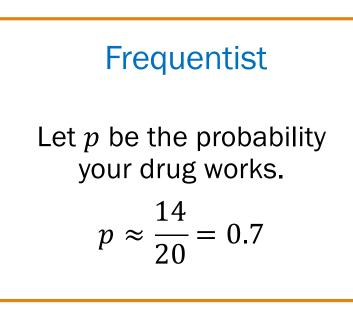
What is the updated distribution of  $\theta$  after our observation?



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- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?



#### Bayesian

A frequentist view will not incorporate prior/expert belief about probability.

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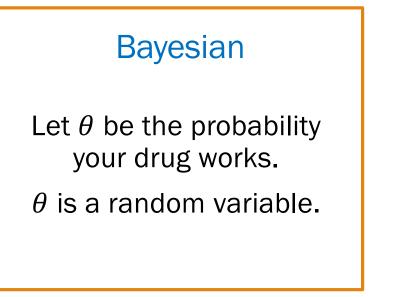
- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

Frequentist

Let *p* be the probability your drug works.

$$p \approx \frac{14}{20} = 0.7$$



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Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$ 

- Before being tested, a medicine is believed to "work" 80% of the time.
- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

(Bayesian interpretation)

What is the prior distribution of  $\theta$ ? (select all that apply)

- A.  $\theta \sim \text{Beta}(1, 1) = \text{Uni}(0, 1)$
- B.  $\theta \sim \text{Beta}(81, 101)$
- C.  $\theta \sim \text{Beta}(80, 20)$
- D.  $\theta \sim \text{Beta}(81, 21)$
- E.  $\theta \sim \text{Beta}(5,2)$



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 $Beta(a = n_{imag} + 1, b = m_{imag} + 1)$ Prior Posterior  $Beta(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$ 

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 $\theta \sim \text{Beta}(81, 21)$ Interpretation: 80 successes / 100 imaginary trials

 $\theta \sim \text{Beta}(5,2)$ 

(you can choose either based on how strong your belief is (an engineering choice). We choose E on next slide)

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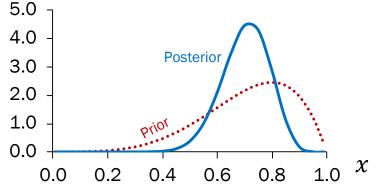
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- The medicine is tried on 20 patients.
- It "works" for 14, "doesn't work" for 6.

What is your new belief that the drug "works"?

- Prior:  $\theta \sim \text{Beta}(a = 5, b = 2)$
- Posterior:  $\theta \sim \text{Beta}(a = 5 + 14, b = 2 + 6)$ ~Beta(a = 19, b = 8)

(Bayesian interpretation)



Prior Beta $(a = n_{imag} + 1, b = m_{imag} + 1)$ Posterior Beta $(a = n_{imag} + n + 1, b = m_{imag} + m + 1)$ 

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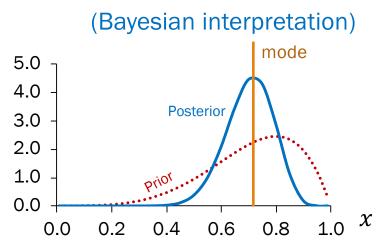
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#### What do you report to pharmacists?

- A. Expectation of posterior
- B. Mode of posterior
- C. Distribution of posterior
- D. Nothing





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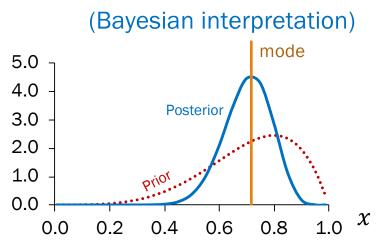
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What do you report to pharmacists?

$$E[\theta] = \frac{a}{a+b} = \frac{19}{19+8} \approx 0.70$$
$$mode(\theta) = \frac{a-1}{a+b-2} = \frac{18}{18+7} \approx 0.72$$



In CS109, we report the **mode**: The "most likely" parameter given the data.

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## Food for thought

In this lecture:



If we don't know the parameter p, Bayesian statisticians will:

- Treat the parameter as a random variable  $\theta$  with a Beta prior distribution
- Perform an experiment
- Based on experiment outcomes, update the posterior distribution of  $\theta$



Food for thought:

Any parameter for a "parameterized" random variable can be thought of as a random variable.

 $Y \sim \mathcal{N}(\mu, \sigma^2)$ 

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## Estimating our parameter directly

#### (our focus so far)

Maximum Likelihood Estimator (MLE) What is the parameter  $\theta$ that maximizes the likelihood of our observed data  $(x_1, x_2, \dots, x_n)$ ?

$$L(\theta) = f(X_1, X_2, \dots, X_n | \theta)$$
$$= \prod_{i=1}^n f(X_i | \theta)$$
$$= \arg \max f(X_i | X_i = X_i | \theta)$$

 $\theta_{MLE} = \arg \max_{\theta} f(X_1, X_2, \dots, X_n | \theta)$ likelihood of data

#### Observations:

- MLE maximizes probability of observing data given a parameter  $\theta$ .
- If we are estimating  $\theta$ , shouldn't we maximize the probability of  $\theta$  directly?

See you next time!

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(extra)

## Extra: MLE: Multinomial derivation

## Okay, just one more MLE with the Multinomial

Consider a sample of *n* i.i.d. random variables where

- Each element is drawn from one of *m* outcomes.  $P(\text{outcome } i) = p_i$ , where  $\sum_{i=1}^m p_i = 1$
- $X_i = \#$  of trials with outcome *i*, where  $\sum_{i=1}^m X_i = n$
- 1. What is the likelihood of observing  $L(\theta) = \frac{n!}{X_1! X_2! \cdots X_m!} p_1^{X_1} p_2^{X_2} \cdots p_m^{X_m}$ the sample( $X_1, X_2, \dots, X_m$ ), given the probabilities  $p_1, p_2, \dots, p_m$ ?
- 2. What is  $\theta_{MLE}$ ?

extra slides

$$LL(\theta) = \log(n!) - \sum_{i=1}^{m} \log(X_i!) + \sum_{i=1}^{m} X_i \log(p_i), \text{ such that } \sum_{i=1}^{m} p_i = 1$$
  
Optimize with  
Lagrange multipliers in  $\theta_{MLE}$ :  $p_i = \frac{X_i}{n}$  Intuitively, probability  
 $p_i = \text{proportion of outcomes}$ 

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n

## Optimizing MLE for Multinomial

$$\begin{array}{ll} \theta = (p_1, p_2, \dots, p_m) \\ \theta_{MLE} = \arg\max_{\theta} LL(\theta), \text{ where } \sum_{i=1}^m p_i = 1 & \text{Use Lagrange multipliers} \\ \text{to account for constraint} \\ \end{array}$$

$$\begin{array}{ll} \text{Lagrange} \\ \text{multipliers: } & A(\theta) = LL(\theta) + \lambda \left(\sum_{i=1}^m p_i - 1\right) = \sum_{i=1}^m X_i \log(p_i) + \lambda \left(\sum_{i=1}^m p_i - 1\right) \begin{array}{l} (\text{drop non-}p_i \\ \text{non-}p_i \\ \text{terms}) \end{array}$$

$$\begin{array}{ll} \text{Differentiate w.r.t.} \\ \text{each } p_i, \text{ in turn: } & \frac{\partial A(\theta)}{\partial p_i} = X_i \frac{1}{p_i} + \lambda = 0 \Rightarrow p_i = -\frac{X_i}{\lambda} \\ \text{Solve for } \lambda, \text{ noting } \\ \sum_{i=1}^m X_i = n, \sum_{i=1}^m p_i = 1: \\ \sum_{i=1}^m p_i = \sum_{i=1}^m -\frac{X_i}{\lambda} = 1 \quad \Rightarrow 1 = -\frac{n}{\lambda} \quad \Rightarrow \lambda = -n \\ \text{Substitute } \lambda \text{ into } p_i \quad p_i = \frac{X_i}{n} \end{array}$$

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