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Martingales

Refresher about conditional expectation
 $(X_1, \dots, X_k)$ r. variables, Y r. var.

$$\mathbb{E}(Y | \underline{X}) = \mathbb{E}(Y | X_1, \dots, X_k) \text{ is } \underline{\underline{\text{d.r.v.}}}$$

defined by 2 conditions

$$(1) \quad \mathbb{E}(Y | \underline{X}) = \hat{\mathbb{F}}_Y(\underline{X}) \text{ for some function } \hat{\mathbb{F}}_Y$$

$$(2) \quad \mathbb{E}[\mathbb{E}(Y | \underline{X}) g(\underline{X})] = \mathbb{E}[Y g(\underline{X})]$$

For any function g -

Explicit

$$(\mathbb{E}(Y | \emptyset) = \mathbb{E}(Y))$$

$$\mathbb{E}(Y | c) = \mathbb{E}(Y) \text{ for a const } c$$

Explicit formula for discrete r.v.

$$P_{X,Y}(a,b) = P(X=a, Y=b)$$

$$\hat{f}_Y(a) = E(Y | X=a)$$

$$= \frac{\sum_b b P_{X,Y}(a,b)}{\sum_{b'} P_{X,Y}(a,b')}$$

For r.v. with joint density

$$\hat{f}_Y(a) = \frac{\int b f_{X,Y}(a,b) db}{\int f_{X,Y}(a,b') db'}$$

Tower property

$$E[E(Y | X_1, X_2) | X_1] = E(Y | X_1)$$

Def Let $(X_n)_{n \geq 0}$ $(Y_n)_{n \geq 0}$ be two sequences of
Randomization / change of variables.

$$\phi_k: \mathbb{R}^k \rightarrow \mathbb{R}^k \text{ invertible}$$

$$\mathbb{E}(Y | X_1, \dots, X_k) = \mathbb{E}(Y | \phi_k(X_1, \dots, X_k)) \quad \square$$

(Intuition: ϕ_k does not change the amount of information.)

$$\text{Also } \forall h: \mathbb{R}^k \rightarrow \mathbb{R}^m$$

$$\mathbb{E}(Y | X_1, \dots, X_k, h(X_1, \dots, X_k)) = \mathbb{E}(Y | X_1, \dots, X_k)$$

(side remark)

For all of these reasons it is convenient to introduce σ -algebras $\mathcal{F}_k = \sigma(X_1, \dots, X_k)$

$$\begin{aligned} \mathbb{E}(Y | \mathcal{F}_k) &= \mathbb{E}(Y | X_1, \dots, X_k) \\ &= \mathbb{E}(Y | \phi_k(X_1, \dots, X_k)) \\ &= \dots \end{aligned}$$

Def Let $(X_n)_{n \geq 0}$ $(Y_n)_{n \geq 0}$ be two sequences of r.v.'s. We say that $(Y_n)_{n \geq 0}$ is a MG wrt $(X_n)_{n \geq 0}$ if $\forall n$:

$$(1) \quad \mathbb{E} |X_n| < \infty$$

$$(2) \quad \mathbb{E}(Y_{n+1} | X_0, \dots, X_n) = Y_n.$$

In particular $Y_n = f_n(X_0, \dots, X_n)$

The

Example 1 Random walk (symmetric)

$$(X_i)_{i \geq 1} \text{ iid } \mathbb{P}(X_i = +1) = \mathbb{P}(X_i = -1) = \frac{1}{2}$$

$$Y_n = \sum_{i=1}^n X_i = Y_{n-1} + X_n$$

$$\mathbb{E}(Y_{n+1} | X_1, \dots, X_n) = \mathbb{E}(Y_n + X_{n+1} | X_1, \dots, X_n)$$

generalizes to iid (X_i) $\mathbb{E}|X_i| < \infty$
 $\mathbb{E}(X_i) = 0.$

$$= Y_n + \mathbb{E}(X_{n+1} | X_1, \dots, X_n)$$

$$= Y_n$$

Example 2 $(X_i)_{i \geq 1}$ indep. $\mathbb{E}(e^{\theta X_1}) < \infty$
 $\forall \theta \in \mathbb{R}$
 $\phi(\theta)$

$$Y_n = C_n \cdot \exp\left(\theta \sum_{i=1}^n X_i\right)$$

$$\mathbb{E}[Y_{n+1} | X_1, \dots, X_n] = C_{n+1} e^{\theta \sum_{i=1}^n X_i} \mathbb{E}[e^{\theta X_{n+1}} | X_1, \dots, X_n]$$

$$= \frac{C_{n+1}}{C_n} Y_n e^{\phi(\theta)}$$

a MG if $C_n = e^{-n\phi(\theta)}$

$$Y_n = e^{\theta \sum_{i=1}^n X_i - n\phi(\theta)}$$

eg ~~X_i~~

example * $X_i = \begin{cases} +1 & \text{w prob } p \\ -1 & \text{w prob } q = 1-p \end{cases}$

~~$e^{\phi(\theta)}$~~

$$e^{\phi(\theta)} = p e^{\theta} + q e^{-\theta}$$

$$Y_n = (p e^{\theta} + q e^{-\theta})^{-n} e^{\theta \sum_{i=1}^n X_i}$$

* $X_i \sim N(\mu, \sigma^2) \quad \phi(\theta) = \mu\theta + \frac{1}{2} \sigma^2 \theta^2$

$$Y_n = \exp \left\{ \theta \sum_{i=1}^n (X_i - \mu) - \frac{n\sigma^2 \theta^2}{2} \right\}$$

Example $(X_n)_{n \geq 1}$ is a ^{discrete time} Markov Chain on countable space S

$$\mathbb{P}(X_{n+1} = j \mid X_n = i) = p_{ij} \quad \text{tr prob}$$

$$\mathbb{P}(X_{n+1} = j \mid X_1, \dots, X_n) = \phi_{X_n, j}$$

Want to construct a MG $Y_n = \psi(X_n)$

$$\mathbb{E}(Y_{n+1} | X_1, \dots, X_n) = \mathbb{E}[\psi(X_{n+1}) | X_n] =$$

$$= \sum_{j \in S} \psi(j) p_{ij}$$

$X_n = i$

So I need $\psi(x) = \sum_{y \in S} p_{xy} \psi(y)$

Harmonic functions

Example X_n ^{lazy} random walk in

$\mathbb{Z}^2 \cap [-L, L]^2$ stop at the bound

$$p_{xy} = \begin{cases} \frac{1}{4}(1-\epsilon) & \text{if } \|x-y\|_2 = 1 \\ \epsilon & \text{if } x=y \end{cases} \quad \text{for } |x|_\infty < L$$

on the bd $p_{xx} = 1$

$$\psi(x) = 1$$

$$\psi(x) = \frac{1}{4}(1-\epsilon) \sum_{y: \|x-y\|_2=1} \psi(y) + \epsilon \psi(x)$$

$y \in \partial x$

$$\sum_{y \in \partial x} [\psi(y) - \psi(x)] = 0 \quad \Delta \psi(x) = 0$$

Example 4 [Doob MG]

$$Z \geq r.v. \quad \mathbb{E}|Z| < \infty$$

$$(X_n)_{n \geq 1} \text{ r.v.}$$

$$\text{defin } Y_n = \mathbb{E}[Z | X_1, \dots, X_n]$$

$$\text{Indeed } X_1^n := (X_1, \dots, X_n)$$

$$\begin{aligned} \mathbb{E}[Y_{n+1} | X_1^n] &= \mathbb{E}[\mathbb{E}[Z | X_1^n, X_{n+1}] | X_1^n] \\ &= \mathbb{E}[Z | X_1^n] = Y_n \end{aligned}$$

Sub example

Color Chromatic

Define SUB-SUPER MG

$$\text{sub MG } \mathbb{E}(Y_n^+) < \infty$$

$$\mathbb{E}(Y_{n+1} | X_1, \dots, X_n) \geq Y_n$$