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## Martingale convergence

$(Y_n)_{n \geq 1}$  a MG (wrt seq  $(X_n)_{n \geq 1}$ )

In general  $\lim_{n \rightarrow \infty} Y_n$  does not exist

example  $(X_i)$  iid  $\mathbb{P}(X_i = +1) = \mathbb{P}(X_i = -1) = \frac{1}{2}$

$Y_n = \sum_{i=1}^n X_i$  does not converge!

However if  $Y_n$  has some "boundedness" properties then it converges.

- Will state a general thm + special cases

Thm Let  $(Y_n)$  be a sub MG and assume  
 $\exists M < \infty$  st  $\mathbb{E}(Y_n^+) \leq M \forall n$ .

Then ~~the~~

(1)  $\exists Y_\infty$  such that  $Y_n \xrightarrow{a.s.} Y_\infty$  as  $n \rightarrow \infty$ .

(2) If  $\mathbb{E}|Y_0| < \infty$  then  $\mathbb{E}|Y_\infty| < \infty$

(3) If  $(Y_n)$  Uniformly integrable  
then  $\mathbb{E}|Y_n - Y_\infty| \rightarrow 0$  ( $Y_n \xrightarrow{L^1} Y_\infty$ )

$(Y_n)$  UI if

$$\lim_{B \rightarrow \infty} \sup_n \mathbb{E}\{ |Y_n| \mathbb{1}_{|Y_n| \geq B} \} = 0$$

Sufficient condition

$$\sup_n \mathbb{E}(|Y_n|^\alpha) \leq C$$

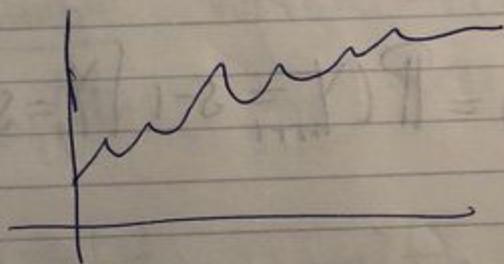
for some  $\alpha > 1$ ,  $C < \infty$ . ]

## Intuition

- A sub MG is something that grows ave

$$\mathbb{E}(Y_{n+1} | X_1^n) \geq Y_n$$

- Condition  $\mathbb{E}(Y_n^+) \leq M$  means the ~~average~~ cannot diverge to  $+\infty$



Corollary Let  $Y_n$  be a MG with

$\sup_n \mathbb{E}|Y_n| \leq M$ . Then  $\exists Y_\infty$  such that

$$\lim_n Y_n \xrightarrow{\text{d.s.}} Y_\infty$$

If  $\sup_n \mathbb{E}|Y_n|^q \leq M'$ , we also have

$$Y_n \xrightarrow{L^q} Y_\infty$$

~~Will prove MC convergence theorem~~

Will prove MC convergence theorem  
for martingales  $(Y_n)_{n \geq 1}$  under the stronger  
assumption

$$\sup_{n \geq 1} \mathbb{E}(Y_n^2) \leq M < \infty \quad (\text{bdd in } L^2)$$

For compactness

Note that  $\forall k \geq 0, n \geq 1$

$$0 \leq \mathbb{E}[(Y_{n+k} - Y_n)^2] =$$

$$= \mathbb{E}(Y_{n+k}^2) - 2\mathbb{E}(Y_{n+k}Y_n) + \mathbb{E}(Y_n^2) =$$

$$= \mathbb{E}(Y_{n+k}^2) - \mathbb{E}(Y_n^2)$$

$\Rightarrow \mathbb{E}(Y_n^2)$  is monotone non-decr bdd

$$\lim_{n \rightarrow \infty} \mathbb{E}(Y_n^2) = B \leq M$$

$$\lim_{n \rightarrow \infty} \sup_{k \geq 0} \mathbb{E}[(Y_{n+k} - Y_n)^2] = 0$$

Would like to use this to prove

$Z_n \xrightarrow{\text{a.s.}} Z_\infty$

$Y_n(\omega) \rightarrow Y_\infty(\omega)$  for almost every  $\omega$ .

Thm  $C_0$  [Kolmogorov ineq]

Let  $(Z_n)_{n \geq 1}$  be a sub MC,  $Z_n \geq 0 \forall n$

Then  $\forall t > 0$

$$\mathbb{P}(\max(Z_1, \dots, Z_n) \geq t) \leq \frac{\mathbb{E}(Z_n)}{t}$$

~~Coroll~~ Note: this is ~~to~~ very interesting because it bounds the max by the last element!

## Corollary

$$\mathbb{P}\left(\max_{0 \leq k \leq m} |Y_{n+k} - Y_n| \geq t\right) \leq \frac{\mathbb{E}[(Y_{n+m} - Y_n)^2]}{t^2}$$

Indeed  $(Y_{n+k} - Y_n)^2 = Z_k$  is a non-neg sub

& getting back to MG convergence  $\epsilon_n \downarrow 0$

$$\limsup_{m \rightarrow \infty} \mathbb{P}\left(\max_{0 \leq k \leq m} |Y_{n+k} - Y_n| \geq t\right) \leq \frac{\epsilon_n}{t^2}$$

$$\mathbb{P}\left(\sup_{k, l \geq n} |Y_k - Y_l| \geq t\right) \leq \frac{4\epsilon_n}{t^2}$$

$$\mathbb{P}\left(\limsup_{n \rightarrow \infty} \sup_{k, l \geq n} |Y_k - Y_l| \geq t\right) = 0$$

$\Rightarrow (Y_k)$  is a.s. Cauchy

$$\Rightarrow Y_k \xrightarrow{\text{a.s.}} Y_\infty$$

□

Corollary Any non-negative MG  
 $Y_n$  converges a.s. to a limit.

Example  $\sum_{i=1}^n Y_n = x_0 + \sum_{i=1}^n X_i$

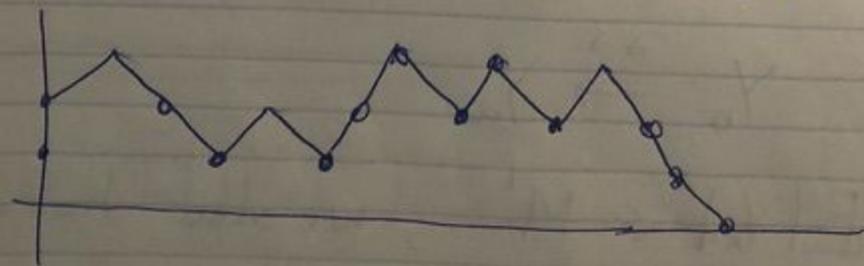
$x_0 \in \mathbb{N}$  Consider  $Y_n$  a RW on  $\mathbb{Z}_{\geq 0}$   
that stops at 0  $(X_n) = (Y_n)$

$$\mathbb{P}(Y_{n+1} = s+1 | Y_n = s) = \mathbb{P}(Y_{n+1} = s-1 | Y_n = s) = \frac{1}{2}$$

for  $s \geq 1$

$$\mathbb{P}(Y_{n+1} = 0 | Y_n = 0) = 1$$

$$Y_0 = x_0 > 0 \quad x_0 \in \mathbb{Z}_{>0}$$



$$Y_n \xrightarrow{d.s.} Y_\infty$$

$$\mathbb{P}(Y_\infty = k) = \mathbb{P}\left(\lim_{n \rightarrow \infty} Y_n = k\right) = 0 \quad \forall k > 0$$

$$\text{Hence } \mathbb{P}(Y_\infty = 0) = 1$$

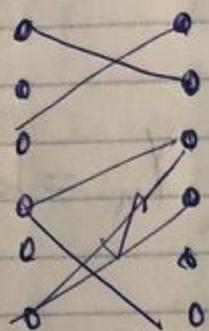
We proved  $Y_n \xrightarrow{d.s.} 0$  (sooner or later hits 0)

Note this applies to any Gambling scheme

Example Wright-Fisher model of genetic drift

- Population with  $M$  individuals
- ~~Each~~ Two alleles of a gene:  $A, a$ .
- Discrete time generations  $n, n+1, \dots$
- At each generation, each indiv. is equally likely to descend from any one in prev.

time  
→



$$Y_0 = k$$

$$Y_n \in \{0, \dots, m\}$$

Call  $Y_n = \#$  individuals with allele A at gen n.

$$\mathbb{P}(Y_{n+1} = k | Y_n = \frac{Y_n}{m}) = \text{Binom}\left(\frac{Y_n}{m}, m\right)$$

$$= \binom{m}{k} \left(\frac{Y_n}{m}\right)^k \left(1 - \frac{Y_n}{m}\right)^{m-k}$$

This is a MG

$$\begin{aligned} \mathbb{E}(Y_{n+1} | Y_1, \dots, Y_n) &= \mathbb{E}(Y_{n+1} | Y_n) = m \cdot \frac{Y_n}{m} \\ &= Y_n \end{aligned}$$

$$0 \leq Y_n \leq m$$

By DOM

$$\lim_{n \rightarrow \infty} Y_n = Y_\infty \quad \text{a.s.} \quad \mathbb{E}(|Y_n - Y_\infty|^q) \rightarrow 0 \quad \forall q$$

$$\begin{aligned} \mathbb{E}[(Y_{n+1} - Y_n)^2 | Y_n] &= \text{Var}(\text{Bin}(\frac{Y_n}{m}, m)) = \\ &= m \frac{Y_n}{m} (1 - \frac{Y_n}{m}) \end{aligned}$$

$$\mathbb{E}[(Y_{n+1} - Y_n)^2] = \frac{1}{m} \mathbb{E}[Y_n(m - Y_n)]$$

$$0 = \lim_{n \rightarrow \infty} \mathbb{E}[(Y_{n+1} - Y_n)^2] = \frac{1}{m} \mathbb{E}[Y_\infty(m - Y_\infty)]$$

$$\mathbb{P}(Y_\infty \in \{0, m\}) = 1$$

$$\mathbb{E}Y_\infty = k = \mathbb{E}Y_0 = \mathbb{E}Y_\infty = m \mathbb{P}(Y_\infty = m)$$

$$k = \boxed{\mathbb{P}(Y_\infty = m) = \frac{k}{m}}$$

The allele fixates to A with prob  $\frac{k}{m}$