

# Lecture 10

## Spatio-Temporal Point Processes

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Stats 253

July 23, 2014

# Outline of Lecture

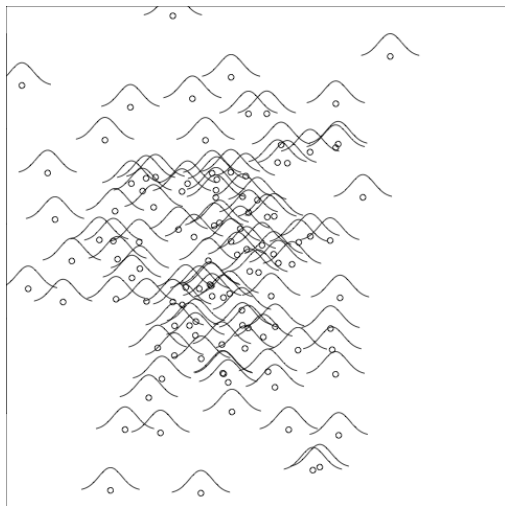
- 1 Review of Last Lecture
- 2 Spatio-temporal Point Processes
- 3 The Spatio-temporal Poisson Process
- 4 Modeling Interactions
- 5 Wrapping Up

# Where are we?

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# Intensity Estimation in Poisson Processes

$\lambda(\mathbf{s})$  can be estimated nonparametrically...



# Intensity Estimation in Poisson Processes

or parametrically, such as...

$$\log \lambda(\mathbf{s}) = \mathbf{x}(\mathbf{s})^T \boldsymbol{\beta}$$

To do this, note that the likelihood of  $(\mathbf{s}_1, \dots, \mathbf{s}_{N(D)}, N(D))$  is:

$$L(\lambda(\cdot)) = e^{-\int_D \lambda(\mathbf{s}) d\mathbf{s}} \frac{(\int_D \lambda(\mathbf{s}) d\mathbf{s})^{N(D)}}{N(D)!} \prod_{i=1}^{N(D)} \frac{\lambda(\mathbf{s}_i)}{\int_D \lambda(\mathbf{s}) d\mathbf{s}}$$

so the log-likelihood is

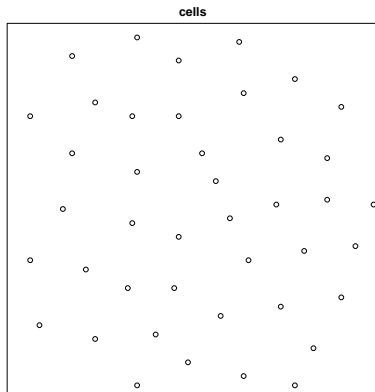
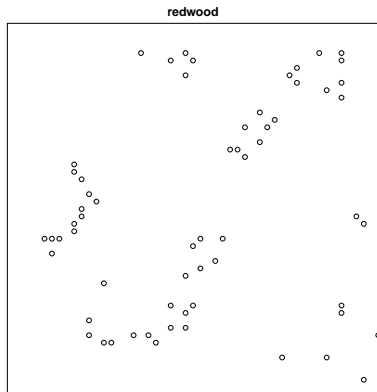
$$\ell(\lambda(\cdot)) = -\int_D \lambda(\mathbf{s}) d\mathbf{s} + \sum_{i=1}^{N(D)} \log \lambda(\mathbf{s}_i) + \text{constants},$$

or in terms of  $\boldsymbol{\beta}$ ...

$$\ell(\boldsymbol{\beta}) = -\int_D \exp\{\mathbf{x}(\mathbf{s})^T \boldsymbol{\beta}\} d\mathbf{s} + \sum_{i=1}^{N(D)} \mathbf{x}(\mathbf{s}_i)^T \boldsymbol{\beta} + \text{constants}$$

# Second-Order Properties

Processes may still exhibit **clustering** or **inhibition**.



## Second-Order Intensity

$$\lambda(\mathbf{s}) = \lim_{|d\mathbf{s}| \rightarrow 0} \frac{\mathbb{E}(N(d\mathbf{s}))}{|d\mathbf{s}|} \quad \mu(\mathbf{s}) = \mathbb{E}(y(\mathbf{s}))$$

$$\lambda_2(\mathbf{s}, \mathbf{s}') = \lim_{|d\mathbf{s}| \rightarrow 0} \frac{\mathbb{E}(N(d\mathbf{s})N(d\mathbf{s}'))}{|d\mathbf{s}||d\mathbf{s}'|} \quad \Sigma(\mathbf{s}, \mathbf{s}') = \mathbb{E}(y(\mathbf{s})y(\mathbf{s}')) - \mu(\mathbf{s})\mu(\mathbf{s}')$$

$\lambda_2$  is called the **second-order intensity**.

A process is stationary if  $\lambda(\mathbf{s}) \equiv \lambda$  and  $\lambda_2(\mathbf{s}, \mathbf{s}') = \lambda_2(\mathbf{s} - \mathbf{s}')$ .

# Ripley's $K$ -function

$$\begin{aligned}
 K(r) &= \frac{1}{\lambda} \mathbb{E} \# \{ \text{events within distance } r \text{ of a randomly chosen event} \} \\
 &= \frac{1}{\lambda} \mathbb{E} \left[ \frac{1}{N(D)} \sum_{i=1}^{N(D)} \sum_{j \neq i} 1 \{ d(\mathbf{s}_i, \mathbf{s}_j) \leq r \} \right]
 \end{aligned}$$

This has a natural estimator:

$$\hat{K}(r) = \frac{1}{\hat{\lambda}} \frac{\# \{ (i, j) : d(\mathbf{s}_i, \mathbf{s}_j) \leq r, i \neq j \}}{N(D)}$$

Provides us with a strategy for fitting models:

$$\text{minimize}_{\theta} \int_0^{r_0} w(r) (\hat{K}(r) - K_{\theta}(r))^2 dr$$



## Relationship between $K$ and $\lambda_2$

If the point process is stationary and isotropic,

$$K(r) = \frac{2\pi}{\lambda^2} \int_0^r \lambda_2(r') r' dr'$$

## Handling Inhomogeneity

- What if  $\lambda(\mathbf{s}) \neq \lambda$  but is known?
- Then we might hope  $\frac{\lambda_2(\mathbf{s}, \mathbf{s}')}{\lambda(\mathbf{s})\lambda(\mathbf{s}')} = \rho(\|\mathbf{s} - \mathbf{s}'\|)$  is stationary and isotropic.
- New definitions:

$$K_I(r) = \mathbb{E} \left[ \frac{1}{|D|} \sum_{i=1}^{N(D)} \sum_{j \neq i} \frac{1\{d(\mathbf{s}_i, \mathbf{s}_j) \leq r\}}{\lambda(\mathbf{s}_i)\lambda(\mathbf{s}_j)} \right] = 2\pi \int_0^r \rho(r')r' dr'$$

$$\hat{K}_I(r) = \frac{1}{|D|} \sum_{i=1}^{N(D)} \sum_{j \neq i} \frac{1\{d(\mathbf{s}_i, \mathbf{s}_j) \leq r\}}{\lambda(\mathbf{s}_i)\lambda(\mathbf{s}_j)}$$

- All computations proceed with  $K_I$  and  $\hat{K}_I$  instead of  $K$  and  $\hat{K}$ .

# Where are we?

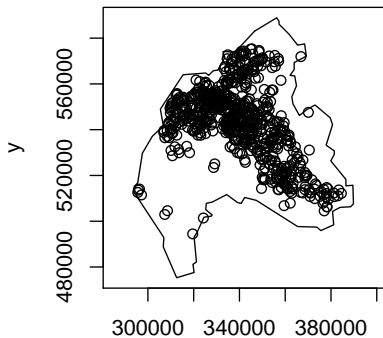
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# What is a Spatio-temporal Point Process?

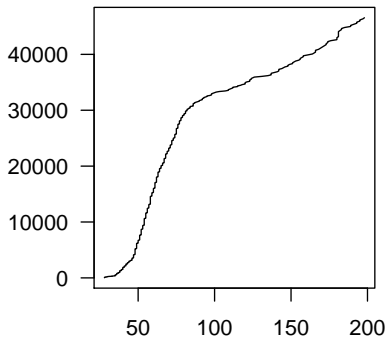
We now observe locations and times  $(s_i, t_i)$ .

```
library(stpp)
data(fmd, northcumbria)
fmd <- as.3dpoints(fmd)
plot(fmd, s.region=northcumbria)
```

**xy-locations**



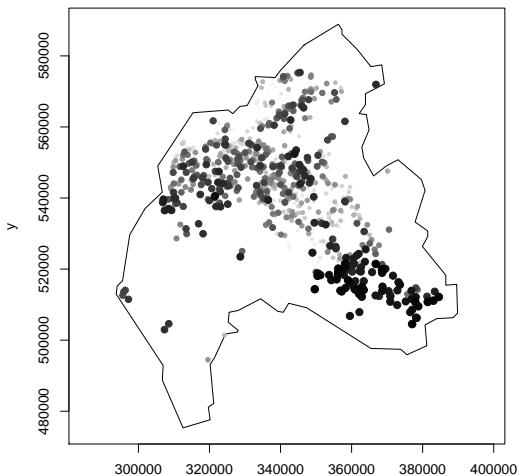
**cumulative number**



# Summarizing spatial and temporal information jointly

We now observe locations and times  $(s_i, t_i)$ .

```
plot(fmd, s.region=northcumbria, pch=19, mark=T)
```



## Animations tell the best story

We now observe locations and times  $(s_i, t_i)$ .

```
animation(fmd, s.region=northcumbria)
```

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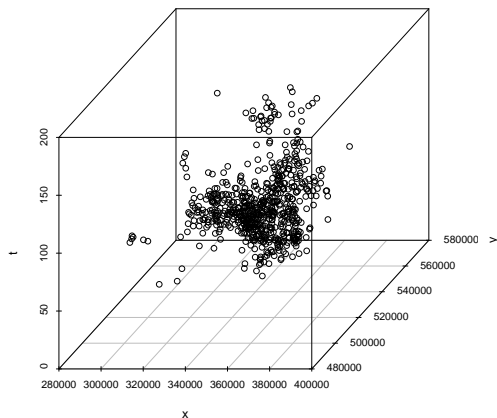
# Spatio-temporal Poisson Process

- Model events as occurring in space time with intensity  $\lambda(\mathbf{s}, t)$ .
- If  $A \subset D \times [0, T]$ , then  $N(A) \sim \text{Pois} \left( \int_A \lambda(\mathbf{s}, t) d\mathbf{s} dt \right)$ .
- How do we estimate  $\lambda(\cdot, \cdot)$ ?



# Estimating the Intensity Function

Exactly the same as before!



How can differences between space and time be captured when estimating  $\lambda(\cdot, \cdot)$ ?

What if  $\lambda(\cdot, \cdot)$  is **separable**?

$$\lambda(\mathbf{s}, t) = \lambda_s(\mathbf{s})\lambda_t(t)$$

The need for modeling interactions between observations becomes even more acute in the space-time setting.

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## Approach 1: $K$ -functions

The second-order spatio-temporal intensity function is:

$$\lambda_2((\mathbf{s}, t), (\mathbf{s}', t')) \stackrel{\text{def}}{=} \lim_{|\mathbf{ds}|, |\mathbf{ds}'|, dt, dt' \rightarrow 0} \frac{\mathbb{E}(N(\mathbf{ds} \times dt)N(\mathbf{ds}' \times dt'))}{|\mathbf{ds}||\mathbf{ds}'| dt dt'}$$

If the point process is stationary and isotropic in space and in time:

$$\begin{aligned} K(r, h) &= \frac{1}{\lambda} \mathbb{E} \left( \begin{array}{l} \# \text{ events within radius } r \text{ and time } h \\ \text{of randomly chosen event} \end{array} \right) \\ &= \frac{1}{\lambda} \mathbb{E} \left[ \frac{1}{N} \sum_{i=1}^N \sum_{j>i} 1\{d(\mathbf{s}_i, \mathbf{s}_j) \leq r\} 1\{t_j - t_i \leq h\} \right] \end{aligned}$$

which can be estimated in the usual way. We also have the relation:

$$K(r, h) = \frac{2\pi}{\lambda^2} \int_0^h \int_0^r \lambda_2(r', h') r' dr' dh'$$

## Approach 1: $K$ -functions

- Suppose we have a process that depends on some parameters  $\theta$ .
- e.g., Parents generate  $S \sim \text{Pois}(\mu)$  offspring. Each of the  $S$  offspring appear at times  $T_i$  according to a Poisson process with rate  $r$ .
- If we can calculate the theoretical  $K$ -function  $K_\theta$ , then we can estimate  $\theta$  by solving

$$\underset{\theta}{\text{minimize}} \int_0^{r_0} \int_0^{h_0} w(r, h) (\hat{K}(r, h) - K_\theta(r, h))^2$$

## Approach 2: Conditional Intensity Function

- In space-time, it's natural to model the conditional intensity given the past  $\mathcal{H}_t = \{(\mathbf{s}_i, t_i) : t_i < t\}$ .

$$\lambda_c(\mathbf{s}, t | \mathcal{H}_t) = \lim_{|d\mathbf{s}| \rightarrow 0, dt \rightarrow 0} \frac{\mathbb{E}(N(d\mathbf{s} \times dt) | \mathcal{H}_t)}{|d\mathbf{s}| dt}$$

- For a Poisson process,  $\lambda_c(\mathbf{s}, t | \mathcal{H}_t) = \lambda(\mathbf{s}, t)$ .
- e.g., pairwise interaction model:

$$\lambda_c(\mathbf{s}, t | \mathcal{H}_t) = \alpha(t) \prod_{i=1}^{|\mathcal{H}_t|} h_\theta(\mathbf{s}, \mathbf{s}_i)$$

- $h_\theta(\mathbf{s}, \mathbf{s}_i) = 1 - e^{-\theta \|\mathbf{s} - \mathbf{s}_i\|}$
- $h_\theta(\mathbf{s}, \mathbf{s}_i) = e^{-\theta \|\mathbf{s} - \mathbf{s}_i\|}$

## Approach 2: Conditional Intensity Function

We typically estimate the intensity function using maximum likelihood.  
The log-likelihood of  $(\mathbf{s}_i, t_i)$  is

$$\ell(\lambda_c(\cdot, \cdot | \mathcal{H}_\cdot)) = - \int_0^T \int_D \lambda_c(\mathbf{s}, t | \mathcal{H}_t) d\mathbf{s} dt + \sum_{i=1}^n \log \lambda_c(\mathbf{s}_i, t_i | \mathcal{H}_{t_i}) + \text{constants.}$$

**NB.** The integral will probably have to be approximated.

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# Summary

- Spatio-temporal point processes are a messy and emerging field.
- Poisson processes are not viable models for spatio-temporal processes; must take into account interactions across space-time.
- There are two main approaches for modeling and fitting interaction models:  $K$ -functions and conditional intensity.
- The `stpp` package contains many useful routines for visualizing, simulating, and (less so) fitting spatio-temporal point process models.



# Homeworks

- Homework 3 due Friday. Don't forget to upload your files to the website (should be in .csv format).
- Homework 4a will be posted tonight.