

## Homework 8 Solutions

Due on December 5, 2018

- Solutions should be complete and concisely written. Please, use a separate sheet (or set of sheets) for each problem.
- We will be using Gradescope (<https://www.gradescope.com>) for homework submission (you should have received an invitation) - no paper homework will be accepted. Handwritten solutions are still fine though, just make a good quality scan and upload it to Gradescope.
- You are welcome to discuss problems with your colleagues, but should write and submit your own solution.

## Problems on hypothesis testing

Solve problems 3.32, 4.2 and 4.19 from Lehmann, Romano, *Testing Statistical Hypotheses*.

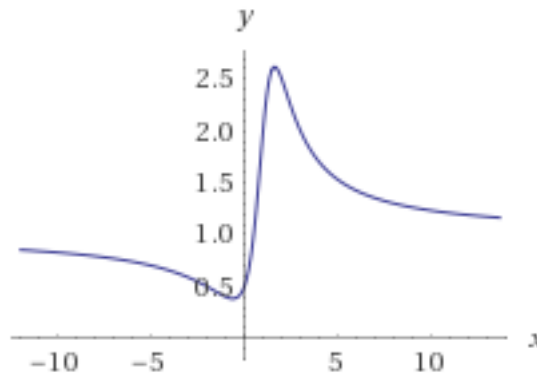
## TSH 3.32 solution (sketch)

The likelihood ratio is:

$$r(x) = \frac{p_{\theta_1}(x)}{p_{\theta_0}(x)} = \frac{1 + (x - \theta_0)^2}{1 + (x - \theta_1)^2}$$

by Neyman-Pearson, the MP test rejects when  $r(x) > k$  with  $k$  chosen such that  $P_{\theta_0}(r(x) > k) = \alpha$ . Since  $X$  has a continuous distribution, the behavior at  $k$  doesn't matter.

We plot an example with  $\theta_0 = 0$  and  $\theta_1 = 1$ .



Notice that  $r(x) \rightarrow 1$  as  $x \rightarrow \pm\infty$ , unlike for normal distributions and other MLR examples that we have seen. It is clear that as we vary  $\theta_1$ , the set  $r(x) \geq \alpha$  changes, so there is no UMP test for  $\theta = 0$  versus  $\theta > 0$ .

To characterize the set of distributions, we set  $r(x) = k$  for various values of  $\theta_0, \theta_1$  and  $k$  and look at the shape. We find that we can get any set of the form:

- $[a, b]$  for  $-\infty \leq a < b \leq \infty$
- $[-\infty, a] \cap [b, \infty]$  for  $-\infty < a < b < \infty$

## TSH 4.2 solution

$$\begin{aligned} P_\theta(\hat{\alpha} \leq \alpha) &= P_\theta(x \in S_{\alpha'} \text{ for some } \alpha' \leq \alpha) \\ &= P_\theta(x \in S_\alpha) \text{ by the nested sets assumption} \\ &\geq P_{\theta_0}(x \in X_\alpha) \text{ by unbiasedness} \\ &= \alpha \end{aligned}$$

## TSH 4.19 solution (sketch)

Let  $T' = \sum_{i=1}^N T(X_i)$  and notice that  $(N, T')$  are the sufficient statistics of a multi-parameter exponential family:

$$P(N, T) \propto e^{N \log(\lambda/C(\theta)) + Q(\theta)T'} h(X, N).$$

We now wish to apply theorem 4.4.1 from TSH, but the theorem doesn't apply because we don't know that the parameter space  $\{(\log(\lambda/C(\theta)), Q(\theta)) : \lambda > 0, \theta \in \Theta\}$  is convex. Looking at the proof of theorem 4.4.1, we find that we only need  $\omega_j = \{(\log(\lambda/C(\theta)), Q(\theta)) : \lambda > 0, \theta = Q^{-1}(c)\}$  to contain a rectangle, which it does because  $\lambda$  can be any positive number. Thus the proof of theorem 4.4.1 goes through in our situation, as desired.