Estimating the Effect of Elite Communications on Public Opinion Using Instrumental Variables

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A central question in the study of democratic polities is the extent to which elite opinion about policy shapes public opinion. Estimating the impact of elites on mass opinion is difficult because of endogeneity, omitted variables, and measurement error. This article proposes an identification strategy for estimating the causal effect of elite messages on public support for European integration employing changes in political institutions as instrumental variables. We find that more negative elite messages about European integration do indeed decrease public support for Europe. Our analysis suggests that OLS estimates are biased, underestimating the magnitude of the effect of elite messages by 50%. We also find no evidence that this effect varies for more politically aware individuals, and our estimates are inconsistent with a mainstreaming effect in which political awareness increases support for Europe in those settings in which elites have a favorable consensus on the benefits of integration.

The relationship between mass and elite opinion is a central issue to the study of voting behavior, parties and elections, public opinion, and representation in democratic systems. For a variety of theoretical reasons, scholars expect elite opinion to affect mass attitudes and behavior. The literatures on priming, persuasion, and cue taking all offer theoretical accounts about how elite opinion shapes how voters approach public policy issues and what attitudes they adopt. In contrast, much of the theoretical literature on representation and electoral competition tells the opposite story: party and elite policy positions respond to voter policy preferences.

This reciprocal relationship raises fundamental methodological problems in attempting to isolate empirically the effect of elite communication on public opinion. While a great deal of research has recognized and attempted to address this problem (for further discussion and examples, see, e.g., Bartels 1993; Gerber and Jackson 1993; Kinder 1998), the extant solutions have serious limitations. Time-series analyses designed to exploit temporal dynamics neither solve the endogeneity problem if elites correctly anticipate mass opinion nor account for omitted variables that might jointly determine both public and elite opinion. Experimental methods often cannot examine the elite-mass linkages that theories of elite communication and persuasion imply, and the external validity of the findings is an open question.

The arguments tested in this article assess the effects of one particular characteristic of elite messages—the extent of elite consensus—on public opinion in the specific context of European integration. When an elite consensus exists in the member countries of the European Union, it is favorable. Consequently, decreasing consensus or increasing polarization implies less favorable elite messages about Europe. If these communications affect mass opinion, the simplest effect of increasing polarization is for those less favorable messages to reduce public support for European integration. The existing public opinion literature suggests, however, that this effect may depend on the characteristics of individuals. Specifically, it will be only those politically aware individuals who are likely to be exposed to and understand elite messages that are influenced...
by these communications. Finally, we examine the possibility of mainstreaming effects. Zaller (1992, 99) predicts that, when elites agree on a policy question, the level of support for the elite position will increase with levels of political awareness. This suggests that political awareness should increase support for European integration in those countries with a pro-integration elite consensus but that this relationship should weaken as elites become more polarized.

Evaluation of these arguments involves estimating the effect of elite consensus on individual opinion formation and thus faces the methodological problems characterizing the literature on the impact of elite communications on public opinion generally. Endogeneity is an obvious concern as changes in elite messages may be responses to changes in the public’s evaluations. Alternatively, elites and the public may simply be responding to similar stimuli consistent with the standard omitted variables concern. Measurement error in different operationalizations of elite consensus is also likely to be an important problem.

This article proposes an identification strategy for estimating the causal effect of elite consensus on public support for European integration employing observational data and instrumental variable estimation. Specifically, we test the three arguments outlined above about the impact of elite messages on public support for European integration by instrumenting for the relevant features of communication using changes in the salient characteristics of national electoral systems. The article presents three main empirical results. First, we find that more negative elite messages about European integration do indeed decrease public support for Europe. Our analysis suggests that OLS estimates that ignore the endogeneity, omitted variables, and measurement problems that typically occur in estimating the effects of elite communication are biased, underestimating the magnitude of the effect of elite messages by 50%. Second, we find no evidence that this effect of elite messages varies for more politically aware individuals. Third, our estimates are inconsistent with a mainstreaming effect in which political awareness increases support for Europe in those political settings in which elites have a favorable consensus on the benefits of integration. This result is in sharp contrast to the OLS analysis that incorrectly suggests a mainstreaming effect.

**Elite Influence on Public Opinion: The Problem**

Questions related to whether and how elites influence mass opinion about politics have been a central concern of social scientists since at least the beginning of the twentieth century. It is well beyond the scope of this article to review the development of the massive body of research that engages these questions (for reviews, see, e.g., Druckman and Lupia 2000; Kinder 1998). We do, however, want to highlight a set of widely recognized methodological problems in this literature that has significantly limited how much is actually known about how elites affect mass political opinions.

We begin by simply describing a general form shared by many theories of elite effects on mass opinion. The generic argument defines some form of elite communication that is hypothesized to affect some specific aspect of individual opinion. Thus, empirical studies generally attempt to estimate the effect of communication on opinion. For example, studies of priming are interested in whether media attention to certain issues (communication) leads citizens to evaluate candidates based on those issues (opinion) (e.g., Iyengar and Kinder 1987). Similarly, studies of persuasion often focus on whether variation in characteristics of the source of information (communication) affect the degree to which citizens are persuaded (opinion) (e.g., Lupia and McCubbins 1998). And, studies of political parties and electoral competition examine how the polarization of the party system (communication) activates voter partisanship and shapes voting behavior (opinion) (e.g., Hetherington 2001).

The natural starting point for estimating the effect of elite communication on mass political opinion is estimates of the contemporaneous (partial) correlation between some measure of the former and the latter. These correlations, however, may be poor estimates of the causal effect of elite communication on opinion for numerous reasons, including poor measurement of the theoretical concepts (measurement error), failure to account for other unobserved or unmeasured determinants of opinion (omitted variable bias), and ignoring the possibility that opinion may be an important determinant of communication (endogeneity). In short, estimating the effect of elite communication on mass political opinion using contemporaneous (partial) correlations is especially vulnerable to the usual limitations of cross-sectional analysis of observational data.

These problems are widely recognized by public opinion scholars and have been so from a very early stage in the development of the literature (see, e.g., Bartels 1993; Gerber and Jackson 1993). Setting aside, for the moment, the problem of measurement error, we focus on considering the methods used by researchers to address the problems of omitted variable bias and endogeneity for making reliable inferences about the impact of elites on mass opinion. The endogeneity problem is straightforward. Most
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The main limitations of the laboratory and survey experimental studies are that it is not clear whether the results generalize to the natural political environments in which elites and the public interact and further that the experiments actually test the most important observable implications of elite opinion leadership. The first point is simply the well-known problem of external validity in laboratory and survey experimentation. For example, a laboratory experiment might provide solid evidence that the content of an experimentally controlled news-cast changes the criteria subjects use in assessing political candidates (priming effects), but this is not direct evidence that the content of actual newscasts primes voters. Druckman (2004) shows that experimental results about elite effects on public opinion are highly sensitive to the specification of the character of elite competition and voter deliberation. Indeed, he finds that many experimental framing effects disappear when the experiment is modified to incorporate more realistic assumptions about the informational environment.

The second limitation of laboratory and survey experimentation is that some of the most direct observable implications of theories of elite influence involve patterns of elite communication that are difficult to replicate in a lab or survey questionnaire. In particular, predictions about the effect of general characteristics of the national political informational environment are difficult to test directly in a laboratory or survey setting.

Another alternative research design to estimate the effect of elite political communications on public opinion is field experiments. This approach has all the advantages discussed above for laboratory and survey experiments but fewer questions about the external validity of the findings (see, e.g., Gerber, Karlan, and Bergan 2006). While much more could be learned about the effects of elite communications on mass opinion from more field experimentation, questions of external validity remain, and it again seems likely that some observable implications of the relevant theories to be tested involve patterns of elite communication that are difficult to experimentally manipulate. The degree of elite consensus on issues of national importance—examined in this article—is arguably one such pattern of communication.

The general problem addressed in this article is how to generate more reliable estimates of the effect of elite political communications on public opinions. We are specifically interested in producing better estimates using observational data for which endogeneity, omitted variable bias, and measurement error are salient concerns. Our approach is to argue that theories of electoral systems and electoral competition suggest a number of potentially valid instrumental variables that are correlated

studies of political competition assume that candidates for political office take positions on policies in order to maximize their vote shares. In this common framework, any correlation between elite and mass policy opinions is determined by the influence of public opinion on elite policy positions rather than the reverse. For many problems investigated in the literature on elite communication, it would be surprising if the public was not exercising at least some influence on elites. To the extent that this is even partially true, estimates of the effect of elite communication on mass opinion from contemporaneous correlations are biased and inconsistent.

The omitted variable problem in the cross-sectional setting is also straightforward. Evidence of a contemporaneous (partial) correlation between elite messages and opinion on a policy issue may be simply due to the fact that the researcher has failed to measure or cannot observe how the objective costs and benefits of policy alternatives vary across cases. These factors surely influence elite and public assessments of policies, and thus their omission may often lead researchers to find elite effects where there are none.

The two most common solutions to these problems have been the use of time-series analyses and survey/laboratory experiments. Time-series analyses attempt to solve ambiguity about the direction of causation by looking at the timing of changes in elite messages and public opinion. Do elite communications lead public opinions? This strategy is a clear improvement on cross-sectional research designs but has at least two limitations. First, elites may appear to lead mass opinion simply because they correctly anticipate changes in opinion rather than cause them. Second, time-series analyses do little to address the omitted variables problem. Elite and mass opinions may change roughly at the same time in response to changes in the political environment. For example, approval of engagement in a foreign war may change among elites and the public as a consequence of changes in events and/or expectations about the costs and benefits of the conflict (see, e.g., Gelpi, Feaver, and Reifler 2006). Many of these time-varying omitted factors may be difficult or impossible to measure and control for.

The most common alternatives to cross-sectional or time-series observational studies are laboratory and survey experiments. These experiments are not subject to endogeneity and omitted variable bias problems because there is experimental control of political communications from elite sources. For the most part, the internal validity of these studies is quite sound, and this approach has been the source of most of what has been learned so far about the empirical effects of elite communications on opinion.
with variation in key dimensions of elite communication but are otherwise uncorrelated with the specific aspects of opinion thought to be affected by elites. Instrumental variable (IV) estimation purges the potentially endogenous regressor—in our case communication—of variation that is not a function of exogenous variables. IV estimates can also eliminate bias due to omitted variables and measurement error.

IV estimation is not, however, a costless solution to the methodological problems facing researchers interested in the causal effects of elite communications on public opinion. IV estimates are inefficient compared to ordinary least squares estimation and can be badly biased if the assumptions of the model are violated (see, e.g., Bartels 1991). Moreover, experimental approaches, despite the limitations discussed earlier, will typically constitute an important component of a compelling research design for many substantive questions in this literature. The claim of this article is simply that IV research designs provide a viable, but to date, underemployed strategy for estimating the effects of elite political communications on public opinion. The next section introduces the specific empirical problem addressed in the article.

**Elite Communication and Support for Europe**

**Existing Literature**

In this article, we estimate the effect of elite communications on public support for membership in the European Union (EU) among the national publics of the EU. A long research tradition characterizes European integration as an elite-driven enterprise executed before an uninterested and poorly informed public. In this context, scholars have often attributed variation in public opinion to cue taking from party elites. However, these studies suffer from the methodological problems highlighted above. For example, this research presents cross-sectional evidence that parties’ positions relate to their voters’ positions, but these studies do not demonstrate that changes in party cues affect changes in public opinion. In fact other studies have shown this not to be the case. Evans (1998, 2002) uses panel surveys of British voters to demonstrate that partisans did not change their positions on issues of European integration when they perceived their parties to have shifted on these issues.

The problem of reverse causality is also acute in the literature on cue taking and support for European integration. If voters’ opinions on integration affect the positions parties take, then findings of a linkage between party and voter positions cannot be clearly attributed to cue taking. Several studies have shown evidence that this methodological concern is pertinent. For instance, Carrubba (2001), using an instrumental variable for voters’ opinions, demonstrates that national parties have changed their positions on integration in response to the opinions of their electoral base.

Very few studies have investigated how the broader informational environment affects mass opinions toward the EU. Perhaps the most relevant to our analysis, Hooghe and Marks (2005) and Ray (2003) evaluated whether elite consensus in terms of party positions affects support for European integration. Hooghe and Marks found a strong negative effect of elite division on public support for integration. Ray tested an “elite consensus” hypothesis that the effect of party positions on their partisans’ opinions strengthens as the variation in party positions increases.

Neither of these studies, however, examined the same set of hypotheses at issue in our study. But their analyses raise relevant measurement and methodological questions for our study. First, both studies measure the level of elite consensus based on the same expert survey of party positions that we employ. Yet neither study examined the validity of their measures of elite consensus, which we do here. Second, both studies treat the level of elite consensus as exogenous to public opinion. As discussed above, this is probably unwarranted, as parties may become more polarized about this issue due to anticipation of changes in citizens’ opinions about integration or simply because they (and the public) are responding to new (and omitted) stimuli in the political environment. Consequently, we assume rather that the degree of consensus in party positions on European integration and public opinion on integration are jointly determined.

**Theoretical Framework**

Our analysis focuses on three specific elite effects, all of which are related to the impact of elite consensus or polarization about Europe on individual support for integration. Standard arguments about the effects of elite communication suggest that support for integration should

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1Studies proposing to identify media effects through natural experiments such as the introduction of new news sources (e.g., Della Vigna and Kaplan 2007) share many of the same strengths and potential weaknesses as our approach in this article.

2Note that in his estimates of cue taking, Ray (2003) recognizes the problem that party positions may be endogenous and employs an IV estimation to address the problem. The analysis does not, however, present a valid test of whether the instrument is correlated with the endogenous regressor, and the article makes implausible assumptions for the exclusion restrictions (in fact, replication of the analysis shows that it fails an overidentification test).
decline as the elite consensus supporting integration weakens. When elites send a consistent and positive message, we expect the mass public to express more favorable attitudes toward integration than when the public is exposed to a heterogeneous set of elite messages. In practice, the only elite consensus about Europe has been favorable. Thus, increasing polarization in this context means less favorable messages about Europe, which we hypothesize decrease public support for integration. This is the most direct form of elite influence on public opinion and has informed the work of public opinion scholars from at least Lippmann (1922) forward.

The work of Zaller (1992) and others suggests that the impact of elite messages on individual opinion formation may depend on various individual characteristics. Zaller, in particular, focuses on how the influence of elite communications depends on an individual’s level of political awareness. Political awareness means the ability of the survey respondent to receive and understand information, messages, and cues from the elite-driven informational environment. Zaller’s model is rich with implications for opinion formation about European integration, but we focus our attention on just two which constitute our second and third hypotheses.

The second hypothesis that we evaluate is based on Zaller’s argument that elite communications are likely to have little effect on an individual’s policy opinions if the individual does not receive or understand the message. Zaller states that “the greater a person’s level of cognitive engagement with an issue, the more likely he or she is to be exposed to and comprehend—in a word, to receive—political messages concerning the issue” (1992, 42). This implies that consistent elite messages supporting a policy position will only affect voters who are politically aware. For the case of support for European integration, this suggests that the negative effect of elite polarization on opinion discussed above is limited to or at least substantially larger for politically aware respondents who are likely to be exposed to and understand elite messages.

Finally, the third hypothesis that we evaluate is that citizens’ support for European integration should increase with their level of political awareness when elites have a supporting consensus. Zaller argues that the effect of elite consensus on mass opinion is stronger among politically aware than among politically unaware voters. In an informational setting where elites primarily send the same message, Zaller (1992) predicts a “mainstream effect” of elite communication. Specifically, Zaller predicts that “the greater a person’s level of political awareness, the greater the number of mainstream messages the person would internalize in the form of considerations and hence, all else equal, the greater the person’s level of expressed support for the mainstream policy” (1992, 98). Further, he suggests that “if the mainstream argument is correct, correlations between awareness and support for a policy should be strongest when elite consensus is strongest and less strong when elite consensus is less strong or nonexistent” (1992, 99). This has clear implications in the context of public support for integration. In those countries with broad support for integration among the major political parties and elites, we expect politically aware EU citizens to express greater support for integration than the less politically aware. This effect is likely to decline if not disappear altogether as elites become more polarized about integration.

To summarize, we evaluate three hypotheses about how elite communications affect public support for European integration:

**Hypothesis 1:** Elite polarization about European integration, because it implies more negative elite messages, has a negative effect on public support for Europe.

**Hypothesis 2:** The negative effect of increased elite polarization about European integration on support for Europe holds only for politically aware citizens who are likely to be exposed to and understand elite messages.

**Hypothesis 3:** Political awareness increases support for Europe when an elite consensus exists—a mainstreaming effect—but does not as polarization increases.

### Support for the European Union

#### Data and Model

Our analysis employs data at the individual and the national level. We use the Eurobarometer survey to measure public opinion toward the EU, political awareness, and other individual-level controls. We use expert survey data to measure the degree of elite consensus in each national political environment. While we expect that some Europeans may receive politically meaningful messages from subnational and supranational sources, we assume that the primary venue for political discourse on this topic in the EU member states is national. Since EU membership is a national policy, we believe this assumption is warranted.

We restrict our analysis to the two standard Eurobarometer surveys conducted in the years corresponding to the expert survey: 1984, 1988, 1992, 1996, and 1999. The dependent variable, *Europe*, is defined based on answers to the following question asked in all 10 surveys:
Generally speaking, do you think that (your country’s) membership of the European Community [common market] is: (1) a bad thing, (2) neither good nor bad, or (3) a good thing.

For the more recent surveys, “European Union” replaced “European Community [common market].” This variable measures support for the policy of interest here: membership in the European Union. While European integration involves more than just the EU—i.e., Europe has integrated politically and economically through other organizations—the EU is far and away the most profound and the broadest example of integration. Thus support for the policy of EU membership is support for European integration. Not surprisingly, previous research shows that responses to this question also reflect respondents’ support for integration generally and support for a variety of specific policy components of EU membership (Gabel 1998a, 1998b).

The main independent variables in this study are elite polarization and political awareness. We measure elite polarization with data from an expert survey conducted by Marks et al. (n.d.), which extends a survey by Ray (1999). The survey asked country experts to evaluate the nation’s political parties’ positions on European integration in 1984, 1988, 1992, 1996, and 1999. All EU members except Luxembourg were included. Based on the average expert evaluation, each party was assigned a score, ranging from 1 to 7, where 1 represents strong opposition to integration and 7 represents strong support.

We want to use these data to construct a valid measure of the level of elite—not simply political party—consensus regarding European integration in each country. We considered three possibilities: the range of party positions, the standard deviation of party positions, and the weighted standard deviation, which weights by the most recent vote percent of each party. Conceptually, we prefer the weighted standard deviation, as we would expect that all parties do not have an equal impact on the informational environment and that their impact should increase with their electoral success.4

To examine the validity of these different measures empirically, we estimated the correlation between the three possible measures and an independent measure of elite consensus. In 1996, the European Commission executed an elite survey called “The Top Decision-Makers” (TDM) Survey.5 This survey included elected politicians, senior civil servants, business and labor leaders, media leaders, and persons playing leading roles in the academic, cultural, or religious life of their countries. The study consisted of 3,500 interviews designed to provide a representative sample by type of elite within each country and with an equal number of interviews in each category of elites. That is, within a group (e.g., elected politicians) the study used a representative sample. Thus, the weighting scheme we used with the party data is unnecessary here.6 All elites interviewed were asked the same survey question as the one we use for our dependent variable. This elite survey, therefore, provides information about a broad range of elite opinion, not just political parties. If this survey had been administered over time, it would provide an attractive measure of elite consensus.

The single survey does, however, allow us to evaluate the validity of the different measures constructed from the expert survey of party positions. Specifically, we estimated the correlation between the aforementioned three measures and the standard deviation of a nation’s elite opinion about EU membership derived from the TDM survey for the 14 countries of the expert survey regarding party positions. The weighted standard deviation performed the best, with a 0.53 correlation. The correlation for the standard deviation of party positions was 0.13. The correlation for the range of party positions was 0.50.7 These results indicate that the weighted standard deviation, which we labelled Elite Polarization, serves as a reasonable proxy for the level of elite consensus.

A key assumption in our empirical analysis is that when a national informational environment approaches or arrives at a one-message setting, that one-message or mainstream position is pro-EU. Thus, increasing polarization implies more negative elite messages about Europe. Our data support this assumption. As the level of consensus increases (i.e., the weighted standard deviation decreases), the weighted mean level of party support for the EU increases.8 Indeed, for all national contexts with a level of elite polarization lower than one, the mean

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4Following Warwick (1994), we calculate the weighted standard deviation for each nation as \( \sqrt{\frac{\sum (x_i - \bar{x})^2}{n}} \) where \( x_i \) is the party’s vote share, \( \bar{x} \) is the party’s position on Europe, and \( \bar{x} \) is the mean position of all the parties in a given country.

5Flash Eurobarometer No. 39, February/March 1996.

6We could not weight by party in any event. The survey does not identify the elites interviewed or their party affiliation.

7We also examined the correlations between these three measures and the variation in party positions reflected in their election manifestos (Budge et al. 2001). We coded party positions on European integration according to Carrubba (2001). The weighted standard deviation measure from the expert survey was the most strongly related to variation in party positions from the most recent manifestos.

8The mean party support for integration is weighted by the party vote shares in the most recent national legislative elections.
position is never less than five. On the 1–7 scale of the expert survey, a four is neutral and higher values are positive appraisals of EU membership. Thus, all of the low polarization national environments have a high level of average party support. Furthermore, the correlation between the weighted standard deviation of party positions and the weighted mean of party positions is −0.63. That is, high support for the EU among parties in a nation is associated with low variation in party positions on the EU.

Zaller (1992) provides a thorough discussion of different measures of political awareness. Political awareness means the ability of the survey respondent to receive information, messages, and cues from the elite-driven informational environment. He considers four types of measures: factual political knowledge, education, media exposure, and interest or participation in politics. Factual political knowledge performed best, but he deemed education a good substitute. Unfortunately for our study, a measure of factual political knowledge is unavailable. Although two of the 10 surveys include some questions of factual knowledge regarding the EU (e.g., questions about the European Parliament), these questions are not repeated consistently. We should also note that we suspect respondents who are favorably predisposed toward the EU are more likely to acquire such specific information about the EU than respondents who are less supportive. Consequently, these neutral factual questions are probably poor measures of political awareness for our purposes. Furthermore, we cannot use education to measure political awareness. Although we do have a measure of education, we use it to control for human capital, which theoretically and empirically has been shown to affect attitudes toward European integration independent of political awareness (Gabel 1998a). This leaves us with two options: self-reported media use or interest in politics. Price and Zaller (1993) and Zaller (1992, 334) argue persuasively that media exposure, particularly self-reported media use, is a poor measure of political awareness. Krosnick and Brannon (1993) show that, when evaluated in the same model, these different aspects of political sophistication can have distinct effects on how voters receive and process elite communication. But they also show that political knowledge, political interest, and media exposure have fairly similar effects as proxies for political awareness when each is considered separately as we do here with political interest. As a result, we use political interest, which is measured with the following survey question:

When you get together with your friends, would you say you discuss political matters frequently, occasionally, or never?

For Political Awareness, we coded never as (1), occasionally as (2), and frequently as (3).9

In addition to our measures of elite polarization, political awareness and their interaction, we include several control variables for factors that influence individual support for Europe: education, gender, income, and occupational status as manual worker or as unemployed. We constructed the variable Education as a 9-point indicator of increasing educational attainment based on the individual’s age at the time he or she finished formal education. The variable Female is a dummy variable equal to 1 for women and 0 for men. Income indicates the income quartile of the respondent’s income relative to the distribution in his or her country and ranges from 1 to 4. The variables Manual Worker and Unemployed are dichotomous measures indicating the relevant occupational category of the respondent. We also included controls for country and year to account for various possible omitted factors that are either time constant features of each country or common shocks over time that may affect support for European integration.10

We model support for European integration in the following way:

\[
\text{Europe}_{i,c,t} = \beta_0 + \beta_1 \text{EP}_{i,c,t} + \beta_2 \text{PA}_{i,c,t} + \beta_3 \text{PA}_{i,c,t} \times \text{EP}_{i,c,t} + \Gamma \text{Z}_{i,c,t} + \alpha_c + \delta_t + \epsilon_{i,c,t},
\]

(1)

where \(i, c, \) and \(t\) index individuals, countries, and years, respectively; \(\text{EP}\) indicates Elite Polarization; \(\text{PA}\) indicates Political Awareness; \(Z\) is the vector of control variables described above; \(\beta_0, \beta_1, \beta_2, \beta_3, \Gamma, \alpha_c, \delta_t\) are parameters to be estimated with \(\alpha_c\) and \(\delta_t\) indicating country and year effects, respectively; and \(\epsilon_{i,c,t}\) is the error term.

As discussed above, we examine three hypotheses about the effects of elite polarization on public support for European integration. First, increasing polarization implies more negative messages about European integration and thus the marginal effect of Elite Polarization on Europe should be negative (Hypothesis 1). Formally,

\[
\beta_1 + \beta_3 \times \text{Political Awareness} < 0
\]

Second, we expect that elite messages only affect those politically aware individuals likely to receive and understand the content of the messages and so the marginal effect of Elite Polarization is negative only for respondents with relatively high levels of Political Awareness (Hypothesis 2).

9Responses of “don’t know” were coded as missing.

10The means (standard deviations) of the variables for the main sample of 85,012 respondents are Europe 2.452 (0.744), Political Awareness 1.876 (0.658), Elite Polarization 1.313 (0.388), Female 0.500 (0.500), Education 4.420 (2.923), Income 2.528 (1.109), Unemployed 0.057 (0.232), Manual Worker 0.149 (0.356), Upper 12.246 (16.678), and LME Ethnic 0.289 (0.324).
Formally,

\[ \beta_1 + \beta_3 \times \text{Political Awareness} < 0 \]

if Political Awareness is high

\[ \beta_1 + \beta_3 \times \text{Political Awareness} = 0 \]

if Political Awareness is low

Of course, this hypothesis could be restated to say simply that the absolute value of the effect was larger for politically aware respondents. Our estimates will allow us to evaluate both versions of this argument.

Finally, we look for evidence of a mainstream effect. Recall again that where there was elite consensus about Europe it was in support of European integration. Therefore, one test of Zaller’s mainstreaming hypothesis as applied to public support for European integration is that individuals who are more politically aware are more likely to support European integration in countries and times for which elite polarization is low (Hypothesis 3). In contrast, political awareness is unlikely to have a systematic effect on support for Europe in those countries and times for which elite messages are polarized. Formally,

\[ \beta_2 + \beta_3 \times \text{Elite Polarization} > 0 \]

if Elite Polarization is low

\[ \beta_2 + \beta_3 \times \text{Elite Polarization} = 0 \]

if Elite Polarization is high

### Instrumental Variables for Estimation of Elite Polarization Effects

To evaluate these hypotheses, we need consistent estimates of \( \beta_1, \beta_2, \) and \( \beta_3 \) in Equation (1). Estimation by ordinary least squares (OLS) is an obvious starting point. The concern highlighted above is that Elite Polarization and its interaction with political awareness are endogenous. Elites may converge or polarize in their views about Europe in response to changes in mass opinion or in anticipation of changes in mass opinion. Alternatively, both elites and the public may simply be responding to specific events that are unmeasured and omitted from the model. It is also possible, despite our evidence for the validity of the Elite Polarization measure, that the variable is measured with error. In the presence of all three of these problems, OLS is biased and inconsistent. The direction of the bias depends on the relative importance of each of these issues. If the main problem is that elites are responding to changes in mass opinion, this suggests that the OLS estimates will overestimate the effect of elites on opinion. In contrast, if random measurement error is in fact the chief issue, this suggests that the OLS estimates will underestimate the impact of elites.

The simplest alternative estimator for Equation (1) is to employ instrumental variables for the potentially endogenous regressors. To do so, we have to identify measures that are correlated with polarization and its interaction with political awareness but otherwise uncorrelated with individual support for European integration. More specifically, the instruments must have a significant partial correlation with Elite Polarization and Political Awareness* Elite Polarization controlling for all the other exogenous determinants of support for European integration (Europe)—including the country and year dummy variables—while being uncorrelated with the error term \( e \) in Equation (1).\(^{11}\) Note that because of the inclusion of the country and year indicator variables, identification comes from within country variation over time in the instruments.

Our approach for identifying valid instruments is to specify characteristics of national electoral systems that are theoretically expected to be correlated with policy polarization among competing political parties and the interaction of these institutional characteristics with our measure of political awareness. These variables are expected to be correlated with Elite Polarization and Political Awareness* Elite Polarization but uncorrelated with the error term in this equation. The overall strategy for forming the instruments is based on the idea that electoral institutions generate incentives for elite polarization/consensus about European integration that are exogenous to public opinion about European integration.

Electoral laws affect the degree to which parties and candidates differentiate their policy positions and the degree of ideological dispersion among political elites. One way this happens is through the effect of electoral laws on the number of parties in a party system. We would expect that electoral laws and social conditions that promote a relatively large number of parties will also produce a relatively heterogeneous set of party positions over various public policies. This provides a useful starting point for identifying instruments, since we have a now very well-developed body of research on how a variety of electoral laws affect party systems (see, e.g., Clark and Golder 2006; Cox 1997).\(^{12}\) It is important to note that some electoral laws do not change over time for the time period relevant

\(^{11}\)See Bertrand (2004) for an analogous IV analysis with two endogenous regressors, one of which is a component of an interaction term. This feature of the identification problem is essentially immaterial, and we estimate the IV specification using standard methods for any regression with two endogenous variables.

\(^{12}\)See also Orellana and Monroe (2004) for a study that shows that not only can changes in electoral laws lead to greater diversity in
for this study. Since we include controls for national fixed effects in our model, we must therefore focus on those electoral laws that demonstrate some intranational variation. With this in mind, we selected the following variables used to explain the number of electoral parties in the party system:

1. the logarithm of the median legislator’s district magnitude (lower tier) interacted with the level of ethnic heterogeneity ($LML^{Ethnic}$)$^{14}$
2. the percent of legislative seats awarded in the upper tier ($Upper$)$^{15}$

Cox (1997) demonstrates that $LML^{Ethnic}$ is positively related to the number of parties, as greater district magnitude in the first-tier allocation of seats combined with greater social heterogeneity provides opportunities for more parties to win seats. Cox (1997) also finds that the number of parties increases with the percent of legislative seats distributed at the upper tier. Since most upper-tier allocations are designed to compensate for disproportionality in the lower-tier allocation of seats, as the upper-tier portion increases we expect more parties to enter the elections as small parties have an increased chance of winning seats. Note that some electoral systems (e.g., Greece and Austria) permit changes in the percent of upper-tier seats from one election to the next, depending on the distribution of votes. When the percent of legislative seats awarded in the upper tier increases, it raises the potential for smaller parties to win seats and therefore increases the effective number of legislative parties. Given that, particularly between elections, legislative parties dominate discourse among political parties, we expect this intranational variation in $Upper$ to be positively related to the level of elite polarization.

Recall that one of the two endogenous variables for which we are devising instruments is the interaction between political awareness and the level of elite polarization on European integration. Thus, we interacted each of these electoral law variables with our measure of political awareness so that they would correlate with the endogenous interaction variable as well as with elite polarization itself.

The discussion so far has focused on why changes in electoral systems influence the distribution of elite policy positions, including views about European integration—i.e., why our instruments are relevant. For our identification strategy to be valid, it also must be the case that changes in electoral systems do not have a direct effect on individual opinions about Europe other than from the resulting changes in elite positions—i.e., that the exclusion restriction holds. We know of no theories of opinion formation about European integration that have proposed that electoral institutions directly influence individual evaluations of Europe and view such a link as implausible. Further, it is important to note that the sources of changes in electoral institutions within countries seem to be largely due to domestic factors such as corruption scandals that are unrelated to changes in individual opinions about European integration.

Although we draw on a substantial theoretical and empirical literature to make our instrumentation strategy plausible ex ante, it is still necessary for us to provide as much evidence as possible that the chosen instruments meet the assumptions for IV to provide consistent estimates. The assumption that the instruments are not correlated with the error term cannot be verified in the data directly. However, since our model is overidentified, we can provide some evidence that the overidentifying assumptions are valid. We will defer presenting this evidence until the next section where we present the main empirical results in the article. In the remainder of this section, we focus on the question of whether the instruments are in fact significantly correlated with the endogenous regressors in our data.

Variation in these electoral laws within countries appears to track changes in the level of disagreement among parties on European integration. For example, the value of $LML^{Ethnic}$ for Belgium increased monotonically by 12% from 1984 to 1999 and the weighted standard deviation of Belgian party positions on European integration increased from 0.78 to 1.45. In Italy, the adoption of new electoral laws in 1993 increased the percent of seats awarded at the upper tier, and this was associated with an increase in the level of party polarization on European integration. The average $Upper$ score was 11.89 in the pre-1993 period and 25 afterwards. The level of elite polarization rose from 1.23 pre-1993 to 1.73 afterwards.

To evaluate these relationships systematically, Table 1 reports the results of an ordinary least squares regression of Elite Polarization on our instruments and all the included exogenous variables specified in Equation (1). Our
Table 1  Ordinary Least Squares Estimates of Elite Polarization

<table>
<thead>
<tr>
<th>Regressor</th>
<th>Coefficient (SE)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Upper</td>
<td>0.061 (0.024)</td>
</tr>
<tr>
<td>LML* Ethnic</td>
<td>1.317 (0.785)</td>
</tr>
<tr>
<td>Upper* Political Awareness</td>
<td>0.00029 (0.00016)</td>
</tr>
<tr>
<td>LML* Ethnic* Political Awareness</td>
<td>0.009 (0.006)</td>
</tr>
<tr>
<td>Included Exogenous Individual-level Regressors</td>
<td>Yes</td>
</tr>
<tr>
<td>Country Fixed Effects</td>
<td>Yes</td>
</tr>
<tr>
<td>Year Fixed Effects</td>
<td>Yes</td>
</tr>
<tr>
<td>Standard Error of Regression</td>
<td>0.183</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.777</td>
</tr>
<tr>
<td>Observations</td>
<td>85,012</td>
</tr>
<tr>
<td>Partial R-squared for excluded instruments</td>
<td>0.070</td>
</tr>
<tr>
<td>F-statistic for test of excluded instruments</td>
<td>3.08</td>
</tr>
<tr>
<td>F p-value</td>
<td>0.023</td>
</tr>
</tbody>
</table>

Notes: This table reports the OLS coefficient estimates and their robust standard errors clustered on country and year (in parentheses). The dependent variable is Elite Polarization.

interest is whether the instruments are significantly correlated with Elite Polarization, controlling for the included exogenous variables in the model including country and year fixed effects, and have the anticipated signs. Identification in this regression comes from within country variation over time.

The coefficients on the instruments have the correct signs and are jointly significant. We evaluate joint significance with an F-test adjusted for country-year clustering. The F(4,58)-statistic is equal to 3.08, which is significant at the 0.023 level. Moreover, the corresponding F-statistic for the OLS regression with the interaction between Political Awareness and Elite Polarization as the dependent variable (unreported) is equal to 9.71, which is significant at less than the 0.001 level. This is clear evidence that the instruments are correlated with the endogenous regressors and thus satisfy the assumption that the selected instruments are relevant.¹⁶

Empirical Results

This section discusses the results of the estimation of Equation (1), testing our three hypotheses for the effect of elite communications on public support for European integration.

Table 2 reports the OLS estimation of this equation. The coefficient estimates are consistent with Hypothesis 1 that increasingly diverse and negative elite messages on Europe decreases support for integration. There is virtually no evidence for Hypothesis 2 that this effect is stronger among respondents with higher levels of political awareness. There is, however, some evidence consistent with Hypothesis 3 that political awareness has a positive effect on support for Europe when elites have a pro-integration consensus, that is, of mainstreaming effects.

To evaluate Hypotheses 1 and 2, it is necessary to calculate the marginal effect of Elite Polarization at various levels of Political Awareness. Figure 1 plots the magnitude of this effect along the vertical axis against different levels of Political Awareness along the horizontal axis. The solid line traces out the values of $\beta_1 + \beta_3 \cdot Political Awareness$ while the dashed lines indicate the bounds of the 95% confidence interval for this estimate.

Across all values of Political Awareness, a one-unit change in Elite Polarization reduces support for integration by roughly 0.2, and this result is significantly different from zero. This estimate is consistent with the argument that more diverse and negative elite messages about Europe have a negative effect on public opinion but inconsistent with the idea that this effect varies across individuals with different levels of political awareness.

To test Hypothesis 3 of mainstreaming effects, we want to evaluate the marginal effect of Political Awareness at different levels of elite consensus about European integration. Figure 2 plots the magnitude of this effect along the vertical axis against different levels of Elite Polarization along the horizontal axis. The solid line traces out regression of Europe on the instruments, Upper and LML* Ethnic, and the exogenous variables. These reduced form estimates are proportional to the coefficient of interest and are unbiased even if the instruments are weak and so indicate the sign of the coefficient of interest. The results of this regression show negative and statistically significant coefficients for the instruments Upper and LML* Ethnic, indicating that the coefficient for Elite Polarization is negative as we report in the IV results below. We also tried a number of other possible instruments for our full specification defined in Equation (1). For example, we added a variable set equal to the interaction between the logarithm of the median legislator’s district magnitude and the level of religious heterogeneity and this variable’s interaction with Political Awareness for a total of six instruments. For this specification, both F-tests were statistically significant at the 0.01 level, and the IV coefficient estimates of interest were qualitatively the same in magnitude and statistical significance as those reported below.

¹⁶Although significant, the correlation between the instruments and Elite Polarization is somewhat weak, which may bias the IV estimates toward the corresponding OLS estimates even if the exclusion restrictions hold. Concern about the strength of the instruments for Elite Polarization is mitigated by considering the reduced form
Table 2  Support for European Integration

<table>
<thead>
<tr>
<th>Regressor</th>
<th>OLS</th>
<th>IV</th>
</tr>
</thead>
<tbody>
<tr>
<td>Elite Polarization</td>
<td>-0.160</td>
<td>-0.566</td>
</tr>
<tr>
<td></td>
<td>(0.058)</td>
<td>(0.184)</td>
</tr>
<tr>
<td>Political Awareness</td>
<td>0.050</td>
<td>-0.084</td>
</tr>
<tr>
<td></td>
<td>(0.036)</td>
<td>(0.059)</td>
</tr>
<tr>
<td>Political Awareness*Elite Polarization</td>
<td>-0.019</td>
<td>0.084</td>
</tr>
<tr>
<td></td>
<td>(0.028)</td>
<td>(0.045)</td>
</tr>
<tr>
<td>Female</td>
<td>-0.071</td>
<td>-0.070</td>
</tr>
<tr>
<td></td>
<td>(0.012)</td>
<td>(0.012)</td>
</tr>
<tr>
<td>Education</td>
<td>0.026</td>
<td>0.025</td>
</tr>
<tr>
<td></td>
<td>(0.002)</td>
<td>(0.002)</td>
</tr>
<tr>
<td>Income</td>
<td>0.047</td>
<td>0.047</td>
</tr>
<tr>
<td></td>
<td>(0.003)</td>
<td>(0.003)</td>
</tr>
<tr>
<td>Unemployed</td>
<td>-0.087</td>
<td>-0.089</td>
</tr>
<tr>
<td></td>
<td>(0.015)</td>
<td>(0.015)</td>
</tr>
<tr>
<td>Manual Worker</td>
<td>-0.084</td>
<td>-0.084</td>
</tr>
<tr>
<td></td>
<td>(0.014)</td>
<td>(0.014)</td>
</tr>
<tr>
<td>Constant</td>
<td>2.475</td>
<td>3.125</td>
</tr>
<tr>
<td></td>
<td>(0.103)</td>
<td>(0.321)</td>
</tr>
<tr>
<td>Country Fixed Effects</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Year Fixed Effects</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Standard Error of Regression</td>
<td>0.695</td>
<td>0.697</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.127</td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>85,012</td>
<td>85,012</td>
</tr>
<tr>
<td>Hansen J-statistic</td>
<td></td>
<td></td>
</tr>
<tr>
<td>χ²(2) p-value</td>
<td>0.258</td>
<td>0.879</td>
</tr>
</tbody>
</table>

Notes: This table reports OLS and IV coefficient estimates and their robust standard errors clustered on country and year (in parentheses). The dependent variable is Europe. The variables Elite Polarization and Political Awareness*Elite Polarization are the endogenous regressors. The instruments are Upper and LML*Ethnic and their interactions with Political Awareness.

the values of \( \beta_2 + \beta_3 * \text{Elite Polarization} \) while the dashed lines indicate the bounds of the 95% confidence interval for this estimate.

At the highest level of consensus observed in our data—Elite Polarization between 0.5 and 1.0—political awareness is positively related to support for European integration. That is, as predicted by Zaller, more politically aware citizens are more likely to support Europe when the informational environment is characterized by consistent elite messages favoring the process of integration. This effect of political awareness, however, is sensitive to varying levels of elite consensus. As the elite informational environment polarizes, the difference in support for European integration between citizens with higher and lower levels of political awareness decreases and becomes statistically indistinguishable from no effect as elite polarization rises above its mean value of 1.3 in the sample. So, in short, political awareness is positively correlated with support for European integration when elite polarization is low (consensus high) and is uncorrelated with opinion when polarization is high (consensus low). These estimates are generally consistent with the mainstream hypothesis proposed by Zaller.

The results for the control variables included in the analysis are also consistent with theoretical expectations and prior research. Support increases with human and financial capital, and women express less support on average than men.

The main conclusions from the OLS analysis are that increasingly diverse and negative elite messages on Europe decrease support for integration among all citizens and that more politically aware citizens are more likely to take a mainstream view of Europe if elites are in agreement on the issue. These inferences, however, are based on the OLS analysis that assumes that elite messages about European integration are exogenous to mass opinions about this issue, that there is not unmeasured/unobserved stimuli simultaneously driving changes in elite messages and mass opinion, and that measurement error is not a significant problem. The discussion above suggests that these assumptions may not hold and that it is especially likely that elites respond to and/or anticipate changes in mass opinion. This latter concern is especially problematic because it suggests that the OLS estimates overestimate the effect of elites on mass opinion. Consequently, the OLS results do not provide convincing evidence of an elite effect at all, much less compelling evidence about how elite appeals interact with political awareness.

To address this issue, we reestimate the model employing the instrumental variables introduced in the previous section. Table 2 reports the results of the IV estimates of the model. These results are important because they provide a rigorous identification strategy for estimating the causal effect of elite messages on public opinion about European integration. Substantively, the results contrast to those based on the OLS analysis in two ways. First, the magnitude of the marginal effects of elite polarization on mass opinion is over twice as large in the IV framework. There is robust evidence consistent with Hypothesis 1 that increasingly diverse and negative elite messages on Europe decrease support for integration and, further, there is evidence that the OLS estimates are too small. This relationship between the OLS and IV estimates suggests that measurement error and/or omitted variable bias rather than endogeneity are quantitatively the most important issues for this empirical problem. This result is consistent with Bartels’s (1993) claim that measurement error accounts for why many media studies have failed to find large effects for media exposure on opinion. Second,
the IV estimates are inconsistent with Zaller’s mainstream model. The modest evidence consistent with Hypothesis 3 in the OLS framework is likely due to significant biases in the OLS estimator.

Again, we can evaluate Hypotheses 1 and 2 by calculating the marginal effect of Elite Polarization at various levels of Political Awareness. Figure 3 plots the magnitude of this effect using the IV estimates from Table 2. Across all values of Political Awareness, a one-unit change in Elite Polarization reduces support for integration by an average about 0.4, and this result is significantly different from zero. This estimate provides compelling evidence that more diverse and negative elite messages about Europe have a negative causal effect on public support for Europe. The estimate declines modestly in magnitude for individuals who are more politically aware, which is precisely the opposite effect suggested by Hypothesis 2 that elite polarization would have its biggest effect on individuals who were the most politically sophisticated.\footnote{We evaluated the robustness of this result by omitting the interaction between Elite Polarization and Political Awareness and employing only Upper and LML*Ethnic as instruments. The results from this specification were also consistent with Hypothesis 1, indicating a significant negative effect of increasingly diverse and negative elite messages about European integration.}

To assess Hypothesis 3 of a mainstreaming effect, we again calculate the marginal effect of Political Awareness at different levels of elite consensus about European integration. Figure 4 plots this effect ($\beta_2 + \beta_3 \times$ Elite Polarization) against Elite Polarization. At high levels of consensus (low Elite Polarization), the effect of Political Awareness is statistically insignificant. As the level of elite polarization increases, the marginal effect of political awareness increases and becomes significantly positive at high levels of polarization. These estimates that take into account the potential endogeneity and omitted variable problems with the OLS results are inconsistent with mainstream effects.

This lack of a mainstream effect may be due to the possibility of “cartelization” among party elites. One reason unrelated to voter opinion that parties converge on a policy position is that they collude with one another to keep particular issues off the agenda so as to maintain their cartel in legislative seats (Mair 1997). Some scholars contend that party consensus on the issue of membership in the European Union is an important component of these cartels (Blythe and Katz n.d.). Insofar as the politically aware are most cognizant of and dismayed by the formation of party cartels, we could imagine that the explicit efforts by parties to avoid competition on European integration actually lead to rejection of elite
Figure 2  Conditional Effect of Political Awareness on Support for European Integration: OLS Estimates

Notes: This figure plots the magnitude of the effect of a one-unit change in Political Awareness on the dependent variable Europe at different levels of Elite Polarization based on OLS estimates. The solid line traces out the estimated conditional effect while the dashed lines indicate the bounds of the 95% confidence interval for this estimate.

cues among the politically aware. This dynamic would work directly against any mainstreaming effect.

As in the OLS estimates of the model, the control variables have the expected effect on support for European integration and in the IV estimation. Support increases with education and income and is lower for manual workers, the unemployed, and women.

The two most obvious objections to this analysis are that the instruments may not be valid either because they are not correlated with the endogenous regressors or because they are not actually exogenous. In the previous section, we presented clear evidence on the first point. Although we cannot directly verify in the data that the instruments are exogenous and thus uncorrelated with the error term, Table 2 reports a test of the assumption that the model’s overidentifying restrictions are valid. Hansen’s J-statistic is equal to 0.258 and given that we have four instruments and two endogenous regressors, it is distributed $\chi^2$ with 2 degrees of freedom. The p-value for testing the null hypothesis that the overidentifying restrictions are valid is 0.879, indicating that there is not evidence for rejecting this assumption.

Conclusion

Whether and how elite communications influence public opinion is central to the study of voting behavior, parties and elections, public opinion, and representation. However, empirical estimation of elite influence is difficult because of endogeneity, omitted variables, and measurement error. In this article, we propose and test an estimation strategy to address these problems based on employing changes in political institutions as instrumental variables. The results indicate that the inferences about elite influence derived from estimates that ignore these concerns can be significantly biased. Specifically, we show that the IV estimates of several common hypotheses support substantively different empirical inferences about how the elite informational environment shapes public opinion.

Our results are important for understanding both the politics of European integration and more general pat-
FIGURE 3  Conditional Effect of Elite Polarization on Support for European Integration: IV Estimates

Notes: This figure plots the magnitude of the effect of a one-unit change in Elite Polarization on the dependent variable Europe at different levels of Political Awareness based on IV estimates. The solid line traces out the estimated conditional effect while the dashed lines indicate the bounds of the 95% confidence interval for this estimate.

FIGURE 4  Conditional Effect of Political Awareness on Support for European Integration: IV Estimates

Notes: This figure plots the magnitude of the effect of a one-unit change in Political Awareness on the dependent variable Europe at different levels of Elite Polarization based on IV estimates. The solid line traces out the estimated conditional effect while the dashed lines indicate the bounds of the 95% confidence interval for this estimate.
terns of representation and influence in contemporary democracies.

Studies of European integration have speculated a great deal about the role of elites in influencing public evaluations of the European project. Scholarly opinion has varied substantially over time from general acceptance of strong elite influence to considerable skepticism of such an effect. These contrasting views have substantial implications for our understanding of the determinants of European integration as well as for the quality of representation in European democracies. The evidence for each has been for the most part highly speculative with the balance of opinion seemingly changing from one referendum result to the next. Our results identify a causal effect of elite opinion on individual opinion formation suggesting that elites are indeed influential. Further, the substantive differences between our OLS and IV results indicate that future research attempting to identify how political elites and citizens influence each other in European politics requires careful attention to establishing a convincing identification strategy.

More generally, our findings, while not definitive for any particular theory of persuasion or opinion formation, raise important questions about interpreting past evidence supporting various theories of elite communication. For example, recent research has focused on how elite discourse influences U.S. mass attitudes toward military conflict but has not addressed the relevant identification issues that we raise in this article. Gelpi, Feaver, and Reiffer (2006, 38) conclude that the perceived level of elite consensus on the Iraq War has had a substantively small effect on public tolerance for U.S. casualties in the war in Iraq. In contrast, Berinsky (2006) concludes that elite discourse has a strong influence on public support for war (Iraq and World War II), but this elite effect is conditional on political awareness. Which, if any, of these accounts should inform how we understand the ability of elites to manage public support in times of war? Our analysis indicates that extant research, due to the endogeneity, omitted variable, and measurement problems that characterize the estimation of elite effects on mass opinion, does not provide a convincing answer.

Finally, our specific identification strategy is potentially useful to studies of representation and the effects of elite communication across contemporary democracies. Our approach to estimating the effect of elite consensus on mass opinions is, in principle, applicable to the broad set of policies where electoral institutions shape elite competition. European integration is simply one such policy. Furthermore, our general strategy of exploiting variation in political institutions to instrument for elite opinion could be adapted to address different aspects of elite communication or cue giving. Our study is but one example of how the literature on comparative democratic institutions could contribute to the study of mass behavior.

References


