Exact Simulation of Random Structures Depending on Infinite Future Information & Applications to Max-Stable Processes

Jose Blanchet (joint with Liu, Dieker, and Mikosch).

Columbia Departments of IEOR and Statistics

Present techniques to simulate processes that involve information from "infinite" future.

Application: Optimal Exact Simulation of Max-stable Fields & Density Estimation

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• Example 1: $M\left(\cdot\right)$ can be represented as

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- $A_n = n$ -th arrival of Poisson process with rate 1 (independent of $Y_n(\cdot)$).
- Brown-Resnick, de Haan, Engelke, Kabluchko, Schlather, Smith, Penrose...

• **Example 2:** Suppose that $S_n = \Delta_1 + ... + \Delta_n$ is a mean-zero multidimensional random walk

$$M_n = \max\{S_k - \mu\left(k
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where $\mu(k)$ grows faster than $O(k^{1/2+\epsilon})$ for some $\epsilon > 0$.

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- *Relevant in perfect simulation*: *M*₀ is steady-state workload in a class of stochastic networks.
- Propp & Wilson '96, Kendall '98,...

• **Example 3:** Stochastic Differential Equations, pick $\Delta_n = 2^{-n}$,

 $X_{n}\left(\left(k+1\right)\Delta_{n}\right)=X_{n}\left(k\Delta_{n}\right)+b\left(X_{n}\left(k\Delta_{n}\right)\right)\left(B\left(\left(k+1\right)\Delta_{n}\right)-B\left(k\Delta_{n}\right)\right)$

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- Let $X_n(\cdot)$ be continuous piecewise linear interpolation.
- Given ε , simulate $N(\varepsilon)$ such that with probability 1

$$\sup_{0 \le t \le 1} \left\| X_{N(\varepsilon)}(t) - X(t) \right\| \le \varepsilon.$$
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- Note that N (ε) is NOT a stopping time adapted to

 F_n = σ (B (kΔ_n) : 0 ≤ k ≤ 2ⁿ − 1).
- Finding a piecewise linear (or constant) process that satisfies (1) is *Tolerance Enforced Simulation* (TES).

Blanchet (Columbia)

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- Let $F\left(\cdot\right)$ be ANY sample path functional of $X\left(\cdot\right)$.
- Assume for simplicity that $F(\cdot) \ge 0$ is Lipschitz in uniform norm.
- Let T > 0 have density $g(\cdot)$ & independent of $X(\cdot)$ & $X_{N(\varepsilon_n)}(\cdot)$

$$E(F(X)) = E \int_0^\infty I(F(X) > t) dt$$

= $E \int_0^\infty \frac{I(F(X) > t)}{g(t)} g(t) dt$
= $E \left(\frac{I(F(X) > T)}{g(T)} \right).$

How is TES Related To Unbiased Simulation?

• Conclusion:

$$Z = \frac{I(F(X) > T)}{g(T)}$$

is unbiased estimator of E(F(X)).

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is unbiased estimator of E(F(X)).

- A lot sample I (F (X) > T) since this 1 or 0 variable (answer "YES" or "NO").
- Continue increasing $n \longleftarrow n+1$ until

$$egin{array}{lll} F\left(X_{N(arepsilon_n)}
ight) &> T+\kappaarepsilon_n$$
 or $F\left(X_{N(arepsilon_n)}
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where κ is the Lipschitz constant of $F(\cdot)$.

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- Records broken only finitely many times.
- Locate Record Breakers: YES or NO question ("will there be a next record?").
- Relevant future information encoded on finitely many YES or NO questions.

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- Today: B., Dieker, Liu, Mikosch (2015): Exact sampling and TES for Max-stable Processes.

• Wang & Stoev (2011) Conditional sampling for spectrally discrete max-stable random fields.

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- Oesting, Ribatet & Dombry (2014) Simulation of max-stable processes.
- Dieker & Mikosch (2014) Exact simulation of Brown Resnick random fields.

On Exact Simulation of Brown-Resnick Fields

• Dieker & Mikosch (2014): IF $Z_n(\cdot)$ stationary increments and $E \exp \left(Z_n\left(t
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$$e^{\mathcal{M}(t_i)} = \sup_{n\geq 1} \left\{ \frac{d}{A_n} \cdot \frac{\exp\left(Z_n\left(t_i - T_n\right)\right)}{\sum_{k=1}^d \exp\left(Z_n\left(t_k - T_n\right)\right)} \right\},$$

where $\{T_n\}_{n\geq 1}$ is i.i.d. uniform on $\{t_1, ..., t_d\} < -$ locations in advance.

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- Complexity O(d) points of $Y_n(\cdot)$ for each t_i .

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- Representation implies an exact sampler.
- Complexity O(d) points of $Y_n(\cdot)$ for each t_i .
- Complexity $O(d \times C(d))$ where C(d) = Complexity of sampling $(Y_n(t_1),...,Y_n(t_d))$.

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Intrinsic Complexity of Exact Sampling on Compact Domains

Is sampling
$$M(t_1)$$
, ..., $M(t_d)$ basically as "easy" as sampling $Z_1(t_1)$, ..., $Z_1(t_d)$?

Answer: Yes! This is what we mean by optimality (total complexity O(C(d)))

Our goal next is to explain how & use it for density estimation...

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• Given $\varepsilon > 0$ (deterministic & user defined) & K any given set

$$\sup_{t\in\mathcal{K}}\left|M_{\varepsilon}\left(t\right)-M\left(t\right)\right|\leq\varepsilon$$

with probability one (note that K can be uncountable).

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• Concept introduced in B. & Chen (2012).

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- Concept introduced in B. & Chen (2012).
- See also: ε-strong simulation G. Roberts, Beskos, Peluchetti, Murray, Pollock, Johansen...

Example of TES: Brownian Motion

• Consider Brownian Motion

$$Z_{n}(t) = \sum_{m=0}^{\infty} \lambda_{m} \Lambda_{m}(t) W_{m}(n).$$

where $W_m(n)$'s are i.i.d. N(0, 1) and $\lambda_m = 2^{-(j+1)/2}$ assuming $m = 2^{j-1} + k \ge 1$, $k = 0, 1, ..., 2^{j-1} - 1$ and $\lambda_0 = 1$.

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• Also, $\Lambda_{0}\left(t
ight)=t$

$$\Lambda_{1}\left(t
ight)=\left(1/2-\left|t-1/2
ight|
ight)$$
 / $\left(t\in\left[0,1
ight]
ight)$,

and $\Lambda_{n}\left(t
ight)$ are translations and dilations of $\Lambda_{1}\left(t
ight)$ along dyadic points...

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Wavelet Decomposition



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Wavelet Decomposition



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Wavelet Decomposition



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- m (ε) = 2^J ≥ N + 1/ε² ≥ 1/ε² <− after last record breaker.

$$\begin{split} \sum_{n=\mathbf{m}(\varepsilon)}^{\infty} \lambda_j \Lambda_j\left(t\right) |W_j| &\leq \sum_{n=\mathbf{m}(\varepsilon)}^{\infty} \lambda_j \Lambda_j\left(t\right) r\left(j\right) \\ &= \sum_{j=J}^{\infty} r\left(j\right) 2^{-(j+1)/2} \sum_{k=0}^{2^j-1} \Lambda_{2^j+k}\left(t\right) \\ &= \sum_{j=J}^{\infty} \rho\left(j+3\right)^{1/2} 2^{-(j+1)/2} \leq 5\rho \varepsilon \sqrt{\log\left(1/\varepsilon\right)}. \end{split}$$

• SIMULATE W_j 's jointly with times $R_m = \min\{n > R_{m-1} : |W_j| > r(j)\}; R_0 = -1$ (PENDING)

• Get $N = \max\{R_k : R_k < \infty\}$ (last record breaker).

We obtain $5\rho\epsilon\sqrt{\log(1/\epsilon)}$ guaranteed uniform error with $O(E(N+1/\epsilon^2)) = O(1/\epsilon^2)$ complexity (optimal).

Simulation of the Crucial Quantities

• Consider $R_1 = \min\{n \ge 1 : |W_j| > r(j)\}$ and let $p_1 = P(R_1 = \infty)$. How to sample $Ber(p_1)$?

$$p_1 \leq P(R_1 > m) := U(m) = \prod_{n=1}^m P(|W_j| \leq r(j))$$

$$p_{1} = U(m) \cdot \prod_{n=m+1}^{\infty} P(|W_{j}| \le r(j))$$

$$\ge D(m) = U(m) \times (1 - (m+1)^{1-\rho^{2}/2})$$

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• Let $V \sim U(0,1)$ and decide $V < p_1$ using "loop" $m \longleftarrow m+1$: Eventually finish when

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Since

$$P\left({{R_1} = m}
ight) = U\left({m - 1}
ight) - U\left(m
ight)$$
 ,

 $\text{ if }V>U\left(m\right)\text{ and }D\left(m\right)< V< U\left(m-1\right)\text{, then }R_{1}=m.$



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$$R_1 = 4$$

$$D(2)D(3)D(4)p U(3) U(2)$$

$$U(4)$$

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Algorithm: Output $\mathbf{m}(\varepsilon)$ jointly with W_n 's Step 0: Set, ε , $\rho = 4$, $G = 2 \lceil \varepsilon^{-2} \rceil$, $\mathcal{R} = []$. Step 1: Set U = 1, D = 0. Simulate $V \sim U(0, 1)$. Step 2: While U > V > D, set $G \leftarrow G + 1$, $U \leftarrow P(|W_1| \le \rho \sqrt{\log G}) * U$ and $D \leftarrow (1 - G^{1 - \rho^2/2}) U$. Step 3: If $V \le D$, $\mathcal{R} = [\mathcal{R}, G]$ and return to Step 1. Step 4: If V > U, $\mathbf{m}(\varepsilon) = G$, $\mathcal{R} = [\mathcal{R}, G]$. Step 5: If $j \in S$, W_j has law $(W \mid |W| > \rho \sqrt{\log(n)})$ (else given $|W| \le \rho \sqrt{\log(n)}$).

Theorem (B. & Chen '12)

The algorithm outputs a wavelet approximation with guaranteed ε error (with probability one) in uniform norm with complexity $O\left(\varepsilon^{-2}\log(1/\varepsilon)\right)$.

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• Technique generally applicable to

$$Z(t) = \sum_{n=1}^{\infty} Y_n(t)$$

if $Y_{n}\left(\cdot
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$$P\left(\max_{0\leq t\leq 1}\left|Y_{n}\left(t\right)\right|>r\left(n\right) \text{ i.o.}\right)=0.$$

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• Levy processes, fractional Brownian motion, ...

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• Consider random walk, say $\tau_i > 0$, are i.i.d.

$$S_n = \tau_1 + \dots + \tau_n - nv,$$

 $S_0 = 0, \& E(S_n) < 0.$

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• Goal: Sample

$$M_n = \max\{S_k : k \ge n\}.$$

• Record breakers = ascending ladder heights.

• $T_0 := 0$ and for $k \ge 1$

$$R_k = \inf\{n \ge T_k : S_n - S_{T_k} > 0\},\$$

$$T_k = \inf\{n \ge R_{k-1} : S_n - S_{R_{k-1}} \le 0\}.$$

• Suppose that Cramer root exists $heta^* > 0$

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• Standard change of measure trick:

$$\begin{array}{ll} P\left(R_1 < \infty\right) &=& E_{\theta^*} \exp\left(-\theta^* S_{R_1}\right) \\ &=& P_{\theta^*}\left(V < \exp\left(-\theta^* S_{R_1}\right)\right), \end{array}$$

where V is U(0, 1) independent of $S_{R_1} > 0$.

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$$E \exp(\theta^* S_n) = 1.$$

• Standard change of measure trick:

$$\begin{array}{ll} P\left(R_1 < \infty\right) &=& E_{\theta^*} \exp\left(-\theta^* S_{R_1}\right) \\ &=& P_{\theta^*}\left(V < \exp\left(-\theta^* S_{R_1}\right)\right), \end{array}$$

where V is U(0,1) independent of $S_{R_1} > 0$.

• Conclusion: We CAN answer YES or NO to "will there be a record breaker?" (Keep in mind that $E_{\theta^*}S_n > 0$).

• Moreover, for each $f(\cdot)$ bounded

$$= \frac{E(f(R_1, S_1, ..., S_{R_1}) | R_1 < \infty)}{\frac{E_{\theta^*}(f(R_1, S_1, ..., S_{R_1}) \exp(-\theta^* S_{R_1}))}{P(R_1 < \infty)}}$$

=
$$\frac{E_{\theta^*}(f(R_1, S_1, ..., S_{R_1}) I(V \le \exp(-\theta^* S_{R_1})))}{P(V \le \exp(-\theta^* S_{R_1}))}.$$

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• So, IF $V \leq \exp(-\theta^* S_{R_1})$ (i.e. $R_1 < \infty \implies$ YES there is Record Breaker) AND $S_1, ..., S_{R_1}$ from P_{θ_*} follows the law of $S_1, ..., S_{R_1}$ given $R_1 < \infty$.

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In both examples, answer to "Will there be a Record Breaker?" also gives the actual location of the Record Breaker (two birds with one stone!)

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• Split in two independent pieces: For any $k \ge 1$

$$\max_{n \le k} \{ -\log\left(\frac{A_n}{n}\right) - \log\left(n\right) + Z_n\left(t\right) \}.$$

$$\leq \max_{n \le k} \{ -\log\left(\frac{A_n}{n}\right) \} + \max_{n \le k} \{ -\log\left(n\right) + Z_n\left(t\right) \}.$$

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- Contribution of A_n easily handled (B. & Sigman (2011)).
- Contribution of Z_n (·) can be done using two approaches: TES & direct record breaking analysis.

• Exact simulation (& TES) for

$$\max_{n \le k} \{ -\log(n) + Z_n(t) \} \lor \max_{n \ge k} \{ -\log(n) + Z_n(t) \}$$

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can be done as we now explain.

• CRUCIAL: For compact C

$$\sup_{t\in\mathcal{C}}\max_{n\geq k}\left\{-\log\left(n\right)+Z_{n}\left(t\right)\right\}=O\left(-\log\left(k\right)\right)\rightarrow-\infty.$$

• Recall that
$$Z_n(t) = \sum_{m=0}^{\infty} \lambda_m \Lambda_m(t) W_m(n)$$

$$\sum_{m,n} P\left(|W_m(n)| > \rho \log^{1/2}(m+1) + \rho \log^{1/2}(n+1)\right)$$

$$\leq C \sum_{m,n} \exp\left(-\rho^2 \frac{\left(\log^{1/2}(m+1) + \log^{1/2}(n+1)\right)^2}{2}\right)$$

$$\leq C \sum_m \exp\left(-\rho^2 \left(\log(m+1)\right)\right) \sum_n \exp\left(-\rho^2 \left(\log(n+1)\right)\right).$$

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• Thus a pair (m, n) such that

$$\left| W_{m}\left(n\right) \right| >\rho \log ^{1/2}\left(m+1\right) +\rho \log ^{1/2}\left(n+1\right)$$

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is valid record breaker & finitely many of them.

• Use any convenient linear order of (m, n) (say lexicographic) to find them.

 \bullet We conclude if (\mathbf{m},\mathbf{n}) is last record breaker

$$\begin{aligned} |Z_{n}(t)| &\leq \sum_{m>m}^{\infty} \lambda_{m} \Lambda_{m}(t) |W_{m}(n)| \\ &\leq \sum_{m>m_{*}}^{\infty} \lambda_{m} \rho \log^{1/2}(m) + \log^{1/2}(n) \sum_{m>m_{*}}^{\infty} \lambda_{m} \rho. \end{aligned}$$

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• Simply choose *n* large enough so that

$$\log^{1/2}(n)\sum_{m>m_*}^{\infty}\lambda_m\rho<-\log(n).$$

Theorem (B., Dieker, Liu, Mikosch '15)

Algorithm outputs a wavelet approximation with guaranteed ε error (with probability one) in uniform norm with complexity $O\left(\varepsilon^{-2}\log(1/\varepsilon)\right)$.

• Applicable to fractional Brownian sheet (Dzhaparidze & van Zanten '05).

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- Applicable to fractional Brownian sheet (Dzhaparidze & van Zanten '05).
- But what if we don't have the wavelet expansion?

Theorem (B., Dieker, Liu, Mikosch '15)

Suppose the following:

1) Sampling $Z_n(t_1), ..., Z_n(t_d)$ with cost C(d). 2) $Z_n(\cdot)$ is Hölder continuous. 3) Can sample $Z_n(t_1), ..., Z_n(t_{d-1}) | Z_n(t_d) = z$ with cost C(d). 4) $\{t_1, ..., t_d\} \subset C$ compact. Then, can sample $M(t_1), ..., M(t_d)$

with complexity $O\left(C\left(d\right)^{1+\epsilon}\right)$ for any $\epsilon>0.$

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- Key ideas in proof of previous theorem are as follows.
- Define a record breaker at n if

$$\{\max_{i=1}^{d} Z_n(t_i) > \log(n)\}.$$

• Use algorithm by Adler, B., and Liu (2012) to optimally estimate

$$P(\max\{Z_n(t_1), ..., Z_n(t_d)\} > \log(n))$$

and sample

$$\{\max(Z_n(t_1), ..., Z_n(t_d)) > \log(n)\}$$

uniformly in d and n.

Combining with Malliavin Calculus for Max-Stable Processes

Theorem (B., Dieker, Liu, Mikosch '15)

Let N be the last record breaker for the Gaussian processes, then the density of $M := (M(t_1), ..., M(t_d))$ evaluated at $y = (y_1, ...y_d)$ satisfies

$$P(y_1, ..., y_d) = E\left(\sum_{i=1}^d G_i (y - M) \sum_{n=1}^N C_{i}^{-1} \bar{Z}_n\right),$$

where C is the covariance matrix of $Z_n := (Z_n(t_1), ..., Z_n(t_d))$ and $\overline{Z}_n(t_i) = Z_n(t_i) - E(Z_n(t_i))$, and

$$G_i(x_1,...,x_d) = \kappa_d \frac{x_i}{\|x\|_2^d}$$

for an explicit constant κ_d .

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- Presented unbiased Malliavin estimator for joint densities of max-stable processes.

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- Presented Optimal Exact Simulation & Tolerance Enforced Simulation (i.e. ε error in path space with probability 100% certainty) for max-stable fields.
- Presented unbiased Malliavin estimator for joint densities of max-stable processes.
- Key idea: define a sequence of finitely many record breakers & locate them with 0 1 questions (Bernoulli sampling).

Picture of a Max-Stable Gaussian Process

