

Stat 375: Inference in Graphical Models

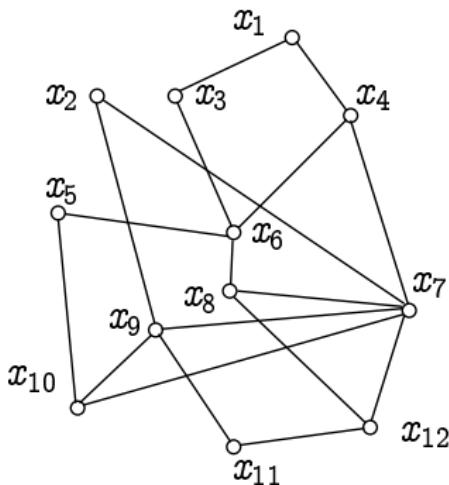
Lectures 3-4

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Undirected Pairwise Graphical Model



$G = (V, E)$, $V = [n]$, $x = (x_1, \dots, x_n)$, $x_i \in \mathcal{X}$, $|\mathcal{X}| < \infty$ (**small**)

$$\mu(x) = \frac{1}{Z} \prod_{(ij) \in E} \psi_{ij}(x_i, x_j).$$

$$\psi \equiv (\psi_{ij})_{(i,j) \in E}$$

Probabilistic inference tasks

Given (G, ψ)

- ▶ Compute the partition function $Z_G(\theta)$
- ▶ Compute expectations $\mathbb{E}_\mu\{f(x_i)\}$
- ▶ Sample from $\mu_{G,\psi}(\cdot)$

The above are roughly equivalent.

These two lectures

- ▶ Computational hardness and reductions.
- ▶ Graph structure and hardness: the case of one-dimensional models.

Computational hardness and reductions

Formalization: The class #P

$S = \{0, 1\}^* \equiv \{ \text{ binary strings } \},$

$|s| = \text{length of string } s.$

Formalization: The class #P

A p -relation:

$$R : S \times S \rightarrow \{0, 1\},$$

- ▶ $R(s, x) = 1 \Rightarrow |x| \leq \text{Poly}(|s|).$
- ▶ $R(s, x) = 1$ can be checked in polynomial time

Formalization: The class #P

Decision: Given s is there an x such that $R(s, x) = 1$?

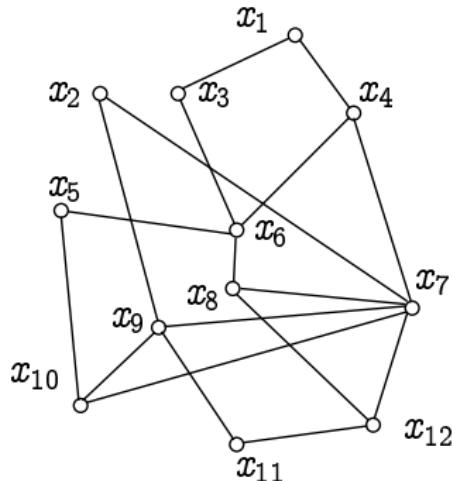
Construction: Given s construct an x such that $R(s, x) = 1$.

Counting: Given s how many x 's are such that $R(s, x) = 1$?
(Class #P)

Sampling: Given s , sample a uniformly random element in
 $\{x : R(s, x) = 1\}$.

Marginals: Need a bit more structure.

Example: Independent sets



$$s = G = (V, E)$$

$$x = (x_i)_{i \in V}, x_i \in \mathcal{X} = \{0, 1\}$$

$$R_{\text{IS}}(G, x) = \begin{cases} 0 & \text{if there exists } (i, j) \in E \text{ with } x_i = x_j = 1 \\ 1 & \text{otherwise.} \end{cases}$$

Associated graphical model

$$\mu(x) = \frac{1}{Z(G)} \prod_{(i,j) \in E} \psi(x_i, x_j), \quad \psi(x_i, x_j) = \begin{cases} 0 & \text{if } x_i = x_j = 1, \\ 1 & \text{otherwise.} \end{cases}$$

Counting is computing a partition function.

$$\begin{aligned} Z(G) &= |\text{Independent Sets}(G)| \\ &= |\{x : R_{\text{IS}}(G, x) = 1\}| \end{aligned}$$

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Does #P require $\psi_{ij}(x_i, x_j) \in \{0, 1\}$?

Can have $\psi_{ij} : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{N}$:

- ▶ Add factor node and auxiliary variable for edge (i, j) .

Can have $\psi_{ij} : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{Q}$:

- ▶ Normalize to make all weights integer.

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Reductions within #P

Counting R_1 reduces to counting R_2 if you have

- ▶ A poly-mapping $\varphi : S \rightarrow S$
- ▶ A poly-mapping $\gamma : \mathbb{N} \rightarrow \mathbb{N}$, s.t.

defining

$$Z_{1/2}(s) \equiv |\{x : R_{1/2}(s, x) = 1\}|,$$

we have

$$Z_1(s) = \gamma(Z_2(\varphi(s)))$$

(Usually one takes $\gamma(N) = N$.)

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Bad news

R is $\#P$ -complete if any problem in $\#P$ can be reduced to it.

Theorem (Valiant 1979)

- ▶ *Independent sets is $\#P$ complete.*
- ▶ *q -coloring ($q \geq 3$) is $\#P$ complete.*
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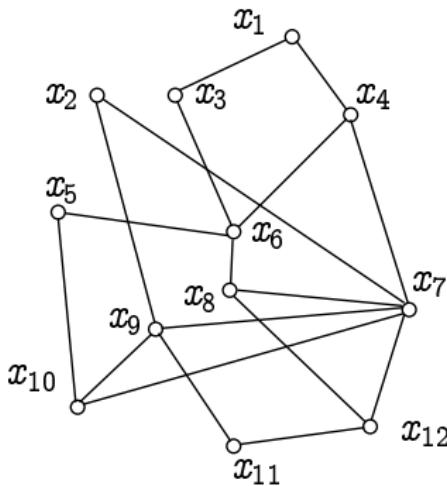
Reductions from other inference problems

Proposition

*If R is **self-reducible**, then sampling and marginals reduce to counting.*

Will not define ‘self-reducible’ formally

Independent sets: Marginals → Counting

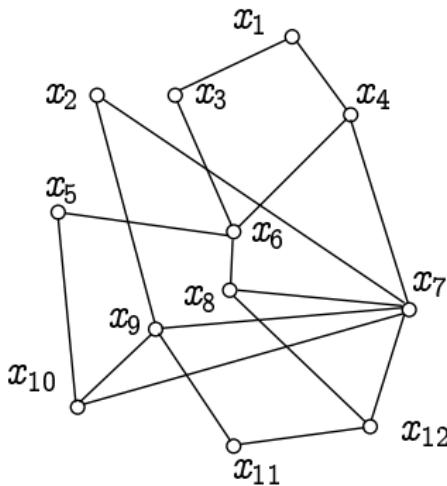


$$\mu(\{x_i = 1\}) = \frac{Z(G, x_i = 1)}{Z(G, x_i = 0) + Z(X_i = 1)},$$

$$Z(G, x_i = \xi) \equiv |\{\text{independent sets of } G \text{ with } x_i = \xi\}|,$$

$$Z(G, x_i = 0) = Z(G \setminus \{i\}), \quad Z(G, x_i = 1) = Z(G \setminus \{i\} \cup \partial i).$$

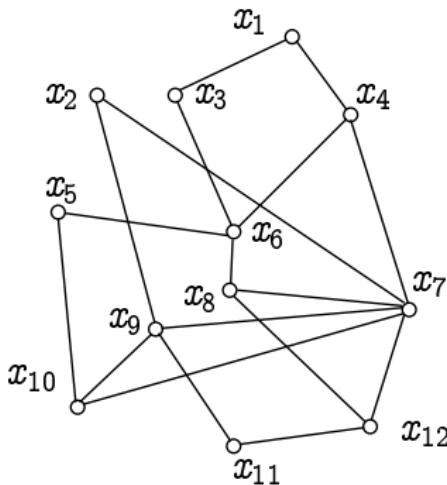
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Independent sets: Sampling → Marginals

SEQUENTIAL SAMPLING(graph G)

- 1: Set $G_0 = G$, $\ell = 0$;
 - 2: While G_ℓ is not empty:
 - 3: Choose $i \in V(G_\ell)$;
 - 4: Compute $p_i = \mu_{G_\ell}(\{x_i = 1\})$;
 - 5: Draw $x_i \sim \text{Bernoulli}(p_i)$;
 - 6: If $x_i = 1$, set $G_{\ell+1} = G_\ell \setminus \{i\} \cup \partial i$, $\ell \leftarrow \ell + 1$;
 - 7: If $x_i = 0$, set $G_{\ell+1} = G_\ell \setminus \{i\}$, $\ell \leftarrow \ell + 1$;
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Exercise

How should the above be modified for q -coloring? What about a general MRF?

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Independent sets: Counting → Marginals

$$\begin{aligned}\frac{1}{Z(G)} &= \frac{Z(G \setminus \{1\})}{Z(G)} \cdot \frac{Z(G \setminus \{1, 2\})}{Z(G \setminus \{1\})} \cdots \frac{Z(\emptyset)}{Z(G \setminus \{1, \dots, n-1\})} \\ &= \mu_G(x_1 = 0) \cdot \mu_{G \setminus \{1\}}(x_2 = 0) \cdots \mu_{G \setminus \{1, \dots, n-1\}}(x_n = 0)\end{aligned}$$

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Independent sets: Marginals → Sampling

Samples: $x^{(1)}, x^{(2)}, \dots x^{(m)} \sim_{i.i.d} \mu_G$

$$\hat{\mu}_G(x_i = 1) = \frac{1}{m} \sum_{\ell=1}^m x_i^{(\ell)}$$

Taking $m = \text{Poly}(n)$, can estimate marginals with relative error $1/\text{Poly}(n)$.

Not exact reduction!

Indeed there are problems for which sampling is easy and exact counting #P hard (e.g. #DNF).

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The next best thing

FPTAS (Fully Polynomial Time Approximation Scheme)

Given G, ε , outputs $\widehat{Z}(G, \varepsilon)$ in time $\text{Poly}(n, 1/\varepsilon)$ such that

$$(1 - \varepsilon)Z(G) \leq \widehat{Z}(G, \varepsilon) \leq (1 + \varepsilon)Z(G).$$

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Reductions between approximation problems

Counting: Computing partition function within multiplicative error ε .

Marginals: Computing marginals within multiplicative error ε .

Sampling: Generating $x \sim \tilde{\mu}$, with $\|\tilde{\mu} - \mu\|_{TV} \leq \varepsilon$.

These are polynomially equivalent.

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Bad news and good news

Theorem

No FPRAS exists for { INDEPENDENT SETS, #SAT, ISING, ... } unless $P=NP$.

[e.g. Luby, Vigoda, 1999, ..., Sly 2010]

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FPRAS/FPTAS exist for { INDEPENDENT SETS AT SMALL DEGREE, PERMANENT, FERROMAGNETIC ISING, ... }.

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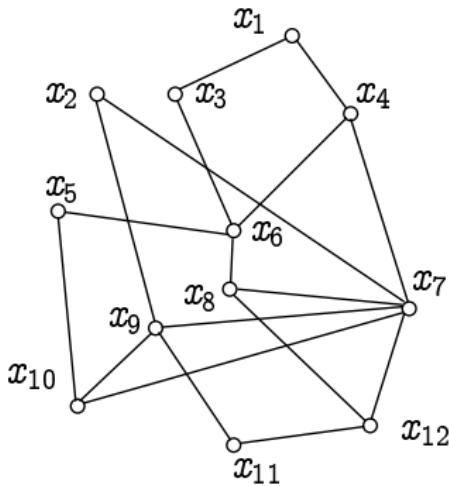
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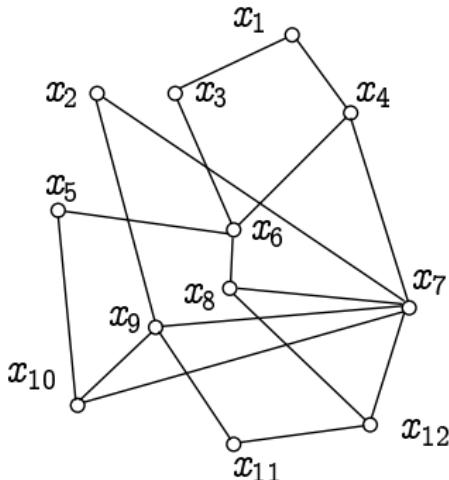
How do you prove inapproximability?



Antiferromagnetic Ising model: $x = (x_1, \dots, x_n) \in \{+1, -1\}^V$

$$\mu(x) = \frac{1}{Z(G, \beta)} \exp \left\{ -\beta \sum_{(i,j) \in E} x_i x_j \right\}$$

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Antiferromagnetic Ising model: $x = (x_1, \dots, x_n) \in \{+1, -1\}^V$

$$\text{cut}_G(x) \equiv |\{ \text{edges connecting } x_i = +1 \text{ and } x_j = -1 \} |,$$

$$\mu(x) = \frac{1}{Z(G, \beta)} \exp \left\{ 2\beta \text{cut}_G(x) \right\}$$

For β large enough

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concentrates on **MAXCUTS**

Exercise

If $\beta = n^2$ then, for all n large enough

$$\mathbb{P}_\mu \{ \text{cut}_G(x) = \text{MAXCUT}(G) \} \geq \frac{1}{2}.$$

But MAXCUT is NP-complete

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But you cheated taking β so large!!!

Keep $\beta = 1$ and

- ▶ Replicate each edge n^2 times.

Keep $\beta = 1$ and

- ▶ Replicate each vertex k times.
- ▶ Place a completely connected graph on the k copies.
- ▶ Replace each edge by a bipartite completely connected graph on k vertices.

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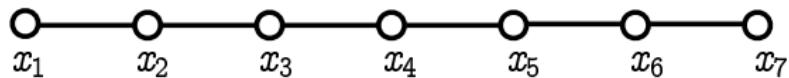
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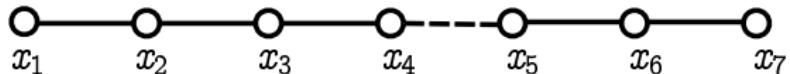
The case of one-dimensional models

Line graph



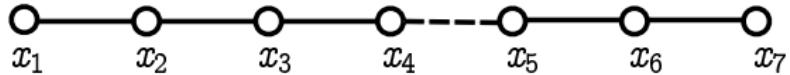
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Line graph: Computing marginals



$$\begin{aligned}\mu_{i \rightarrow} (x_1, \dots, x_i) &\equiv \frac{1}{Z_{i \rightarrow}} \prod_{j=1}^{i-1} \psi_j(x_j, x_{j+1}), \\ \mu_{(i+1) \leftarrow} (x_{i+1}, \dots, x_n) &\equiv \frac{1}{Z_{(i+1) \leftarrow}} \prod_{j=i+1}^{n-1} \psi_j(x_j, x_{j+1}), \\ \nu_{i \rightarrow} (x_i) &\equiv \sum_{x_1, \dots, x_{i-1}} \mu_{i \rightarrow} (x_1, \dots, x_i), \\ \nu_{(i+1) \leftarrow} (x_{i+1}) &\equiv \sum_{x_{i+2}, \dots, x_n} \mu_{(i+1) \leftarrow} (x_{i+1}, \dots, x_n).\end{aligned}$$

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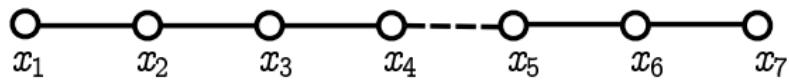
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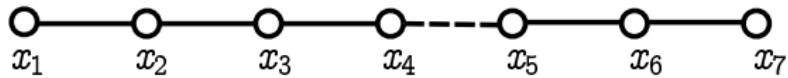


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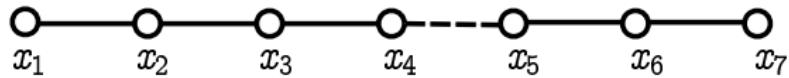


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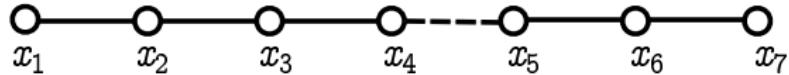


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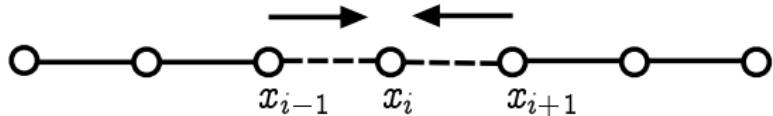


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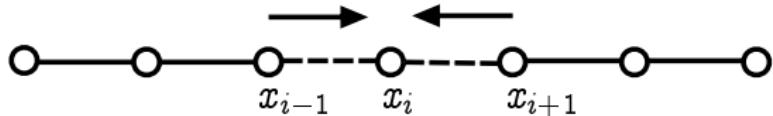


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Line graph: Computing marginals

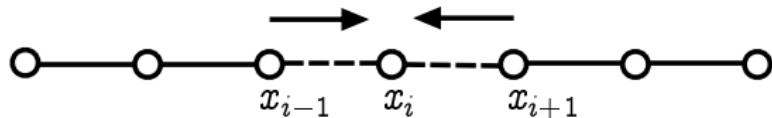


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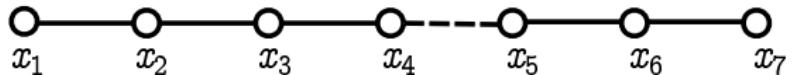
Line graph: Computing marginals



$$\mu(x_i) = \nu_i(x_i)$$

Marginals can be computed in $|\mathcal{X}|^2 n$ operations.

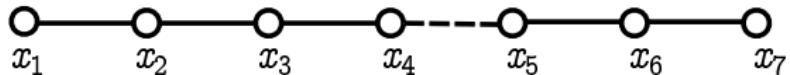
Line graph: Computing the partition function



$$\begin{aligned} Z_{i \rightarrow} &= Z_{(i-1) \rightarrow} \tilde{Z}_{i \rightarrow}, \\ \tilde{Z}_{i \rightarrow} &= \sum_{x_i, x_{i-1}} \nu_{(i-1) \rightarrow}(x_{i-1}) \psi_{i-1}(x_{i-1}, x_i), \\ Z_{1 \rightarrow} &= 1, \\ Z_{n \rightarrow} &= Z. \end{aligned}$$

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Line graph: Computing the partition function



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Example #1: Hidden Markov Models

Sequence of r.v.'s $\{(X_1, Y_1); (X_2, Y_2); \dots; (X_n, Y_n)\}$,

$\{X_i\}$ Markov Chain $\mathbb{P}\{x\} = \mathbb{P}\{x_1\} \prod_{i=1}^{n-1} \mathbb{P}\{x_{i+1}|x_i\},$

$\{Y_i\}$ noisy observations $\mathbb{P}\{y|x\} = \prod_{i=i}^n \mathbb{P}\{y_i|x_i\},$

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Example #1:

Time Homogeneous Hidden Markov Models

Sequence of r.v.'s $\{(X_1, Y_1); (X_2, Y_2); \dots; (X_n, Y_n)\}$,

$\{X_i\}$ Markov Chain $\mathbb{P}\{x\} = q_0(x_1) \prod_{i=1}^{n-1} q(x_i, x_{i+1}),$

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Parameters: $\{p(x, x') : x, x' \in \mathcal{X}\}, \{r(x, y) : x \in \mathcal{X}, y \in \mathcal{Y}\}$

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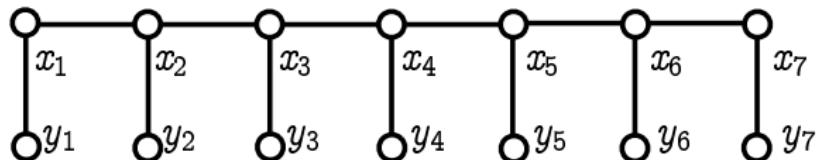
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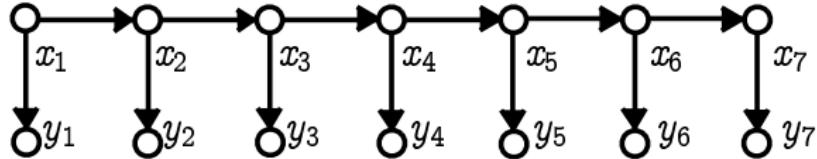
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Example #1: Unconditional Time Homogeneous Hidden Markov Models



$$\begin{aligned}\mu(x, y) &= \frac{1}{Z} \prod_{i=1}^{n-1} \psi_i(x_i, x_{i+1}) \prod_{i=1}^n \tilde{\psi}_i(x_i, y_i), \\ \psi_i(x_i, x_{i+1}) &= q(x_i, x_{i+1}), \quad \tilde{\psi}_i(x_i, y_i) = r(x_i, y_i).\end{aligned}$$

(equivalent directed form)



Question

Estimate $\hat{x}(y)$ from y as to minimize

$$\mathbb{E} N_{\text{err}} = \sum_{i=1}^n \mathbb{P}\{\hat{x}_i(y) \neq x_i\}.$$

Exercise: The optimal estimator is

$$\hat{x}_i(y) = \arg \max_{z \in \mathcal{X}} \mathbb{P}\{x_i = z | y\}.$$

Computational task.

Compute marginals of (q_0 uniform)

$$\mu_y(x) = \mathbb{P}\{x|y\} \stackrel{\text{Bayesthm}}{=} \frac{1}{Z(y)} \prod_{i=1}^{n-1} q(x_i, x_{i+1}) \prod_{i=1}^n r(x_i, y_i).$$

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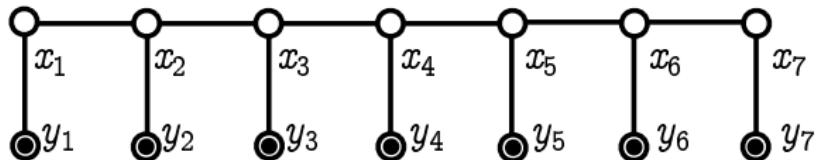
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Example #1: Conditional Hidden Markov Models

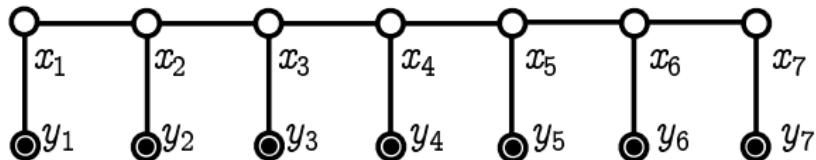


$$\begin{aligned}\mu_y(x) &= \frac{1}{Z(y)} \prod_{i=1}^{n-1} q(x_i, x_{i+1}) \prod_{i=1}^n r(x_i, y_i) \\ &= \frac{1}{Z(y)} \prod_{i=1}^{n-1} \psi_i(x_i, x_{i+1})\end{aligned}$$

$$\psi_i(x_i, x_{i+1}) = q(x_i, x_{i+1}) r(x_i, y_i) \quad (\text{for } i < n - 1)$$

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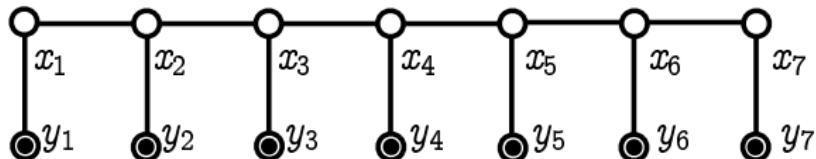


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Example #1: Computing marginals in HMMs

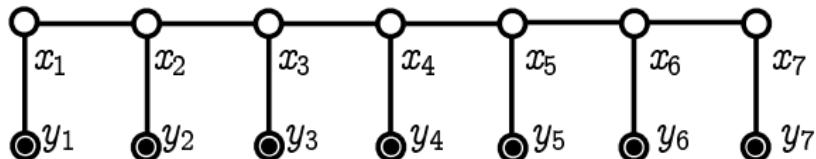


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(forward-backward algorithm)

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Example #1B: Linear system

HMM with $x_i \in \mathbb{R}^k$, $y_i \in \mathbb{R}^h$

Markov chain (parameters $F, Q \in \mathbb{R}^{k \times k}$, $Q \succeq 0$)

$$\begin{aligned} q(x_i, x_{i+1}) &\cong \exp \left\{ -\frac{1}{2}(x_{i+1} - Fx_i)^T Q^{-1}(x_{i+1} - Fx_i) \right\}, \\ x_i &\sim N(Fx_i, Q). \end{aligned}$$

Observations (parameters $H \in \mathbb{R}^{h \times k}$, $R \in \mathbb{R}^{h \times h}$, $R \succeq 0$)

$$\begin{aligned} r(x_i, y_i) &\cong \exp \left\{ -\frac{1}{2}(y_i - Hx_i)^T R^{-1}(y_i - Hx_i) \right\}, \\ y_i &\sim N(Hx_i, R). \end{aligned}$$

Forward iteration = Kalman filter

$$\nu_{i \rightarrow}(x_i) \cong \exp \left\{ -\frac{1}{2}(x_i - \hat{x}_i)^T P_i^{-1}(x_i - \hat{x}_i) \right\}$$

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Example #1C: Neuron firing patterns

Hypothesis

Assemblies of neurones activate in a coordinate way in correspondence to specific cognitive functions. Performing of the function corresponds sequence of these activity states.

Approach

Firing process \leftrightarrow Observed variables

Activity states \leftrightarrow Hidden variables

Example #1C: Neuron firing patterns

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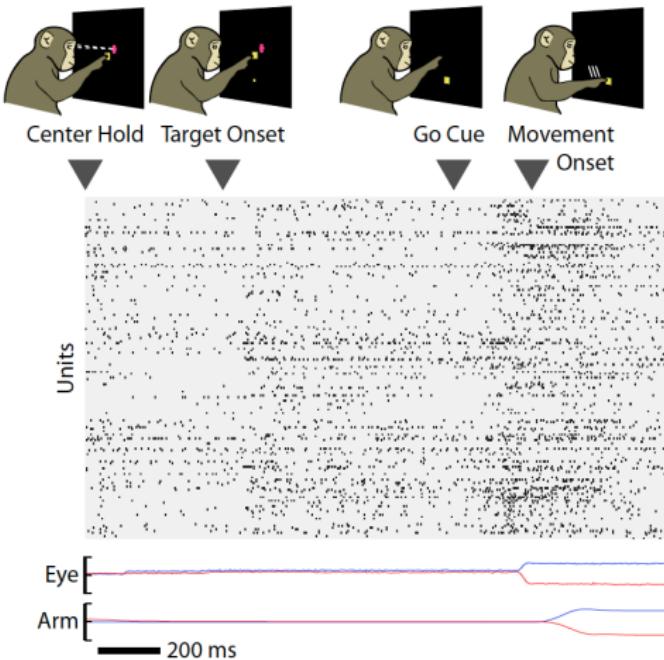
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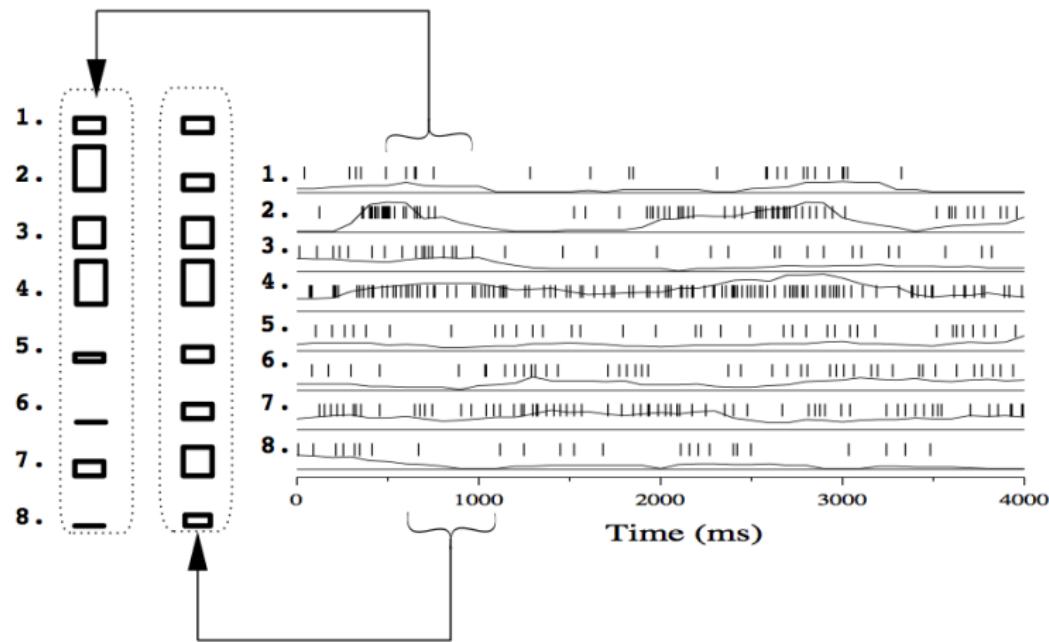
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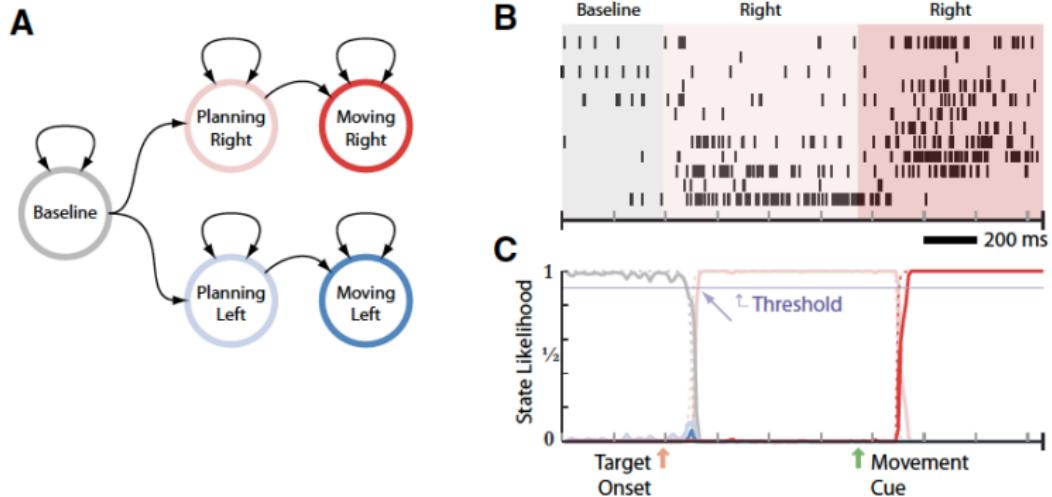
[C. Kemere, G. Santhanam, B. M. Yu, A. Afshar, S.I. Ryu, T. H. Meng and K.V. Shenoy, J. Neurophysiol. 100:2441-2452 (2008)]

Example #1C: Neuron firing patterns



[I. Gat, N. Tishby, and M. Abeles, Network: Computation in Neural Systems, 8:297-322 (1997)]

Example #1C: Neuron firing patterns



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